



Social Well-Being at School: Development and Validation of a Scale for Primary Education Students

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This study presents the development and validation of a scale for Primary Education students that measures social well-being. A seven-factor structure was defined, with the factors being: achievement, cooperation, cohesion, coexistence, attitude towards school, attitude towards diversity and solidarity. 14 experts from independent European universities participated in the validation process of the scale. The 38-item scale showed considerable reliability (Cronbach's alpha =0.91). The confirmatory factor analysis confirmed the original seven-factor structure with consistent goodness and badness of fit indexes. The promising results in this study suggest that this scale may be suitable for an international audience.

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1 BACKGROUND

Students' social well-being may be defined as the extent to which they feel a sense of belonging and social inclusion in their academic environment (Pang, 2018). The importance of social well-being has been cited by different authors as one of the key factors in students' school success (Chen et al., 2018; Wrigley, 2019). They claim that it is a variable that influences students' physical and psychological health, academic performance and personal development, among others. Kokka (2019) and Alanko and Lund (2019) state that a connected student is a supported student at school. Hence, given its importance, it is essential to develop instruments that quantify and determine the degree of students' social well-being (Ryff, 2018). Moreover, authors such as Wrigley (2012), Niclasen et al. (2018) and Mowat (2019) state that Primary Education (ages 6–12) is arguably the most important educational stage in terms of students' social well-being because their experiences at this educational level may affect their development in future stages such as middle school, high school and college.

Several authors have examined the complex and multi-model structure of students' social wellbeing across different educational levels and countries. In this vein, a comparison between Turkish and California adolescents by Telef and Furlong (2017) concluded that factors associated with students' well-being may be difficult to determine and may vary across cultures or countries. In this sense, Maor and Mitchem (2018) believe that determining and measuring social well-being among schoolchildren is a difficult and complex task, as many different kinds of factors must be considered. One of the first factors linked to well-being at school that arises in the literature is achievement. Achievement may be defined as a student's competence in a given content area as a result of both intellectual and non-intellectual variables (Centeio et al., 2019). However, some authors state that while the association between social well-being and achievement is important, it is often wrongly ignored (Beard, 2018). That is not the case for other variables such as cooperation, cohesion and coexistence, as many authors in the field mention them as strongly linked to social well-being (Orkibi

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and Tuaf, 2017). Although several definitions of these variables may be found in the literature, cooperation may be defined as the presence of deliberate relations among autonomous students to jointly accomplish individual goals (Lu and Hallinger, 2018). One definition of coexistence is the way in which students live together with a series of shared rules and environments (Mayorga and Picower, 2018). Likewise, Mikulyuk and Braddock (2018) define cohesion as students' tendency to stick together and remain united as they pursue instrumental objectives and meet their affective needs. Cultural, gender, economic and physical differences among students must also be addressed when analysing their social well-being (Longobardi et al., 2019). Solidarity and attitudes towards diversity appear to be two of the most important variables (Dell'Anna et al., 2019; Piwoni, 2019). Solidarity can be defined as a concern for peers that implies making sacrifices based on a feeling of unity, while attitude towards diversity may be defined as the positive acceptance of other individuals regardless of their age, physical attributes, gender, culture, race, ethnicity, sexual orientation or economic or social class (Solhaug and Osler, 2018). Finally, institutional settings should also be considered when addressing students' well-being, with attitude towards school being a relevant factor (Albuquerque et al., 2019; Nikula et al., 2020). It can be defined as a student's feelings about their attendance at their school from an institutional perspective. Although many other variables linked to students' social well-being may be found in the literature, the ones mentioned above are the most common.

As stated by Lamb and Land (2013) and Bethell et al. (2017) assessing children well-being is many times very complex because not only school, but also family, health and other factors may influence this variable. Most recently, research on the effects of covid-19 on students' well-being has been documented (Cusinato et al., 2020; Morelli et al., 2020) and new directions have been taken towards the concept of well-being and the assessment of the factors that may define it. In this sense, a higher increase on the influence of screen exposure, which was a factor that was not considered as important some years ago, has now increased again given its capital importance (Bruggeman et al., 2019: Garcia-Hermoso et al., 2020). Hence, the dynamic nature of the concept of well-being makes it arguably one of the most difficult variables in educational contexts to be analysed. In this sense, differences among countries, cultures and even inside the same countries must also be considered (Unicef, 2012; Rees and Dinisman, 2015; Migliorini et al., 2019; Robayo-Abril and Millan, 2019).

Taking all of the above into consideration, two main objectives are defined for this study. The first is to develop and validate a scale for Primary Education students that measures their social well-being at school. Researchers, practitioners and other professionals in education and social sciences are provided with a simple tool which may be used to assess students' social well-being at their schools using. Through this assessment, specific needs in students' social well-being may be detected and interventions on teachers' training, students' behaviours or school policies may be arranged. In this sense, the instrument is intended to be suitable for an international audience so that comparisons between different contexts and countries may be possible. Hence, differences in results by gender, school and age will be analysed. The second objective is to determine the possible relationships among the factors in the scale, in order to contribute to a better understanding of the dimensions of students' social well-being at Primary Education. In this sense, understanding the relationship between these dimensions may also facilitate professionals' interventions on schools which are aimed at increasing students' social well-being.

2 METHODS

2.1 Participants

On the one hand, 14 volunteer experts from independent European universities participated in the validation process. All of them were full-time professors employed by their universities, and they participated on a voluntary basis with no conflicts of interests or monetary compensation. Nine came from public universities and 5 from private universities. Their ages ranged from 30 to 62, with an average age of 48.73 years old and a standard deviation of 15.3 years. Eight were male and 6 were female. More information on their selection process can be found in the Design and Procedure section of this manuscript.

On the other hand, 486 Primary Education students from five Spanish schools participated in the processes following the validation. 51.85% were female and 49.15% were male. All of them were enrolled in grades 4 to 6 (ages 10–12), with the average age being 10.81 years old and a standard deviation of 1.89 years. Of the total, 287 (59%) were Hispanic, 97 (20%) were Rumanian, 83 (17%) were African, 15 (3%) were Asian and the other 1% were from other ethnic groups. The socio-economic background of their families was average. Legal consent to perform this research was obtained from both the educational institutions and the students' parents, and the ethical standards provided by the Spanish Ministry of Education were followed.

2.2 Representativity of the Sample

According to the Spanish Statistical Institute, in 2018 there were about 2,000,000 students in grades 4 to 6 (ages 10–12) enrolled in public and private schools in Spain. Using Cochran's formula for large populations with a confidence coefficient of 95% (p < 0.05), at least 386 responses were needed in order to make that sample representative of the whole population. Hence, according to these authors, the sample was representative for Spanish Primary Education students from grades 4 to 6. The sampling procedure used in this study was stratified random sampling. Hence, Spanish schools were taken as strata for this purpose. The size of the sample, gender, ethnicity and ages in each stratum was intended to proportionately represent the size of the population.

2.3 Design and Procedure

Guidelines provided by the International Test Commission (International Test Commission, 2019) were followed regarding the creation, testing, translation and validation of the scale.

2.3.1 Phase 1: Scale Design

Buntins et al. (2017) state that there is no universally-accepted method to fully ensure the validation process of a scale.

Nevertheless, a series of common guidelines accepted by different educational organizations may be followed (Long, 2017). Hence, in accordance with SAGE model of social psychological research (Power et al., 2018), the validation process that was used in this research is described below.

First, a literature review was carried out on the existing scales which addressed the aforementioned variables. As indicated by Walliman (2017), broad knowledge of the area is needed before formulating the scale items. Then, an initial scale was designed following the recommendations of Wolf et al. (2016). Based on this initial design, researchers drafted the scale items. The Delphi method was used in order to bring validity to the scale (Belton et al., 2019). Dillman's Total Design Method (Axford et al., 1997) was used to develop the questions.

2.3.2 Phase 2: Validation

After the initial design of the instrument, independent experts judged it. As indicated by Steedle and Ferrara (2016), expert judgment allows the appropriateness of the items to be assessed and determines whether they belong to the defined factors or constructs. The scale was developed to be used by an international audience. To this end, the guidelines suggested by Arafat et al. (2016) were followed. A search for experts from different European universities was conducted. Their academic and professional background, their relationship to the field of study and their international experience were considered the selection criteria. As Jaeger (1991) indicates, the selection must be careful and the experts varied. Once the selection was finished, the experts were contacted via email. As Jann and Hinz (2016) state, conducting a survey using means other than email does not ensure greater validity or reliability of results. The guidelines by Whittaker and Worthington (2016) were followed to present the scale and its instructions to the experts. The experts' contributions were anonymous. According to Roberts and Allen (2015), anonymity gives the study rigor while also decreasing the possibilities that opinions may be affected by others. In order to carry out the validation, experts used Lawshe's method (Baghestani et al., 2019) to quantify the need to include each scale item. This method means that the experts had to grade each item depending on its suitability for inclusion in the scale. A grade of 1 indicates that the item is not necessary, a grade of 2 indicates that the item is useful but not essential and a grade of 3 indicates that the item is essential. Their answers were collected using Google Forms. Only items showing a significance level of p < 0.05 were included in the final version of the scale. The CVR (content validity ratio) for each item was reported. Furthermore, the inter-expert degree of agreement was analysed. As Kane (2016) indicate, addressing the individual validity of each item is not sufficient, and instead a good inter-expert degree of agreement is needed to ensure validity. Hence, taking into account the type of quantitative answers provided by the experts, Fleiss' Kappa coefficient and the intra-class correlation coefficient were used to measure the degree of agreement between experts. According to Vanbelle (2019), these are the most appropriate means when analysing ordinal responses. SPSS version 25 was used to calculate these coefficients.

2.3.3 Phase 3: Reliability analysis

The reliability of the scale was tested after the validation process. To this end, Cronbach's alpha, Spearman-Brown & Guttman split-half coefficient were used to assess the overall reliability of the scale, and if the item was deleted the Cronbach's alpha was also used (Raykov and Marcoulides, 2019). SPSS version 25 software was used to conduct these analyses.

2.3.4 Phase 4: Exploratory Factor Analysis

An EFA was conducted after the validation and reliability analyses (Zhang et al., 2019). Following guidelines given Watkins (2018), first, a testing for normality of distribution using Kolmogorv-Smirnov parameter was carried out. Once normality of the distribution was guaranteed, responses from 200 students were used to conduct the EFA. Previosly, its feasibility was analysed using Kaiser-Meyer-Olkin and Bartlett's sphericity tests (Beavers et al., 2013). After assuring the feasibility of carrying out a factor analysis, the Kaiser method with varimax rotation (Osborne, 2015) was used. According to this method, the final version of the scale should have the same number of factors as eigenvalues greater than 1, and confirmatory factor analysis trials should be based on these results. SPSS version 25 software was used to conduct the EFA and the aforementioned preliminary tests.

2.3.5 Phase 5: Confirmatory Factor Analysis

A CFA was carried out after considering the results from the EFA (Schreiber et al., 2006). According to Jorgensen et al. (2018) at least 250 subjects are necessary to conduct a CFA. Hence, responses from another 286 students were used in this phase. Taking into account the number of eigenvalues greater than 1 shown by the EFA, several trials for different dimensional structures were performed until the goodness-of-fit indices showed suitable values (Pendergast et al., 2017). For this purpose, Bentler's comparative fit index (CFI) and Joreskog-Sorbom's goodness of fit index (GFI) were used. Furthermore, chi-squared divided by degrees of freedom (χ^2/df) and root mean square residual (RMR) were used as badness-of-fit indices. Values for CFI and GFI should be 0.9 or greater (the higher the better), RMR should be 0.06 or lower and χ^2/df should be lower than 4 (the lower the better) in order to achieve a suitable factor structure. EQS 6.2 software was used to perform the CFA.

2.3.6 Phase 6: Descriptive and Inferential Statistical Analysis

After the CFA, quantitative descriptive and inferential statistical analyses were carried out. Descriptive results included calculation of average scores, standard deviations and Cronbach's alpha for each factor and the scale as a whole. Responses were analysed by gender, age and school overall and by factors. Student's t-tests were conducted for the gender and school variables (De Winter, 2013). Moreover, one-way ANOVA tests (Wetzels et al., 2012) were carried out for the age and factor variables. Finally, the Pearson correlation coefficient was used to measure the correlation among factors (De Winter et al., 2016) taking into account the cut-off value suggested by Dancey and Reidy (2007). According to these authors, in educational and psychological

Item	1	2	3	CVR	Item	1	2	3	CVR
1	0	1	13	0.86 ^a	29	0	6	8	0.14
2	0	3	11	0.57 ^a	30	3	4	7	0
3	0	3	11	0.57 ^a	31	1	4	9	0.29
4	0	1	13	0.86 ^a	32	1	2	11	0.57 ^a
5	0	3	11	0.57 ^a	33	3	5	6	-0.14
6	0	2	12	0.71 ^a	34	1	2	11	0.57 ^a
7	0	6	8	0.14	35	1	5	8	0.14
8	0	2	12	0.71 ^a	36	0	2	12	0.71 ^a
9	0	2	12	0.71 ^a	37	0	6	8	0.14
10	0	5	9	0.29	38	1	1	12	0.71 ^a
11	2	5	7	0	39	3	2	9	0.29
12	0	0	14	1 ^a	40	0	2	12	0.71 ^a
13	1	1	12	0.71 ^a	41	0	0	14	1 ^a
14	0	3	11	0.57 ^a	42	0	1	13	0.86 ^a
15	2	5	7	0	43	0	2	12	0.71 ^a
16	0	3	11	0.57 ^a	44	0	0	14	1 ^a
17	0	3	11	0.57 ^a	45	0	2	12	0.71 ^a
18	0	5	9	0.29	46	0	0	14	1 ^a
19	0	3	11	0.57 ^a	47	1	1	12	0.71 ^a
20	1	4	9	0.29	48	0	2	12	0.71 ^a
21	0	2	12	0.71 ^a	49	0	2	12	0.71 ^a
22	0	5	9	0.29	50	0	0	14	1 ^a
23	0	5	9	0.29	51	0	0	14	1 ^a
24	0	1	13	0.86 ^a	52	0	1	13	0.86 ^a
25	2	4	8	0.14	53	0	2	12	0.71 ^a
26	0	3	11	0.57 ^a	54	0	3	11	0.57 ^a
27	1	5	8	0.14	55	1	5	8	0.14
28	0	3	11	0.57 ^a	56	2	5	7	0

Note. CVR, is calculated using the following formula: CVR = $(N_E - N/2)/(N/2)$. N_E = number of experts indicating "essential item," N = total number of experts. ^ap < .05.

studies, Pearson correlation coefficients greater than 0.4 indicate strong correlation. SPSS version 25 software was used to conduct calculations in this phase.

3 RESULTS

3.1 Phase 1: Scale Design

After performing the literature review and following the procedures and guidelines indicated in the section above, the original version of the scale was designed. That original version had 56 items distributed across seven factors: achievement, cooperation, cohesion, coexistence, attitude towards school, attitude towards diversity and solidarity. Factors were determined after a thorough literature review (see Background section) and taking into account the latest studies in the field that addressed which factors moderate students' social well-being (Bücker et al., 2018; Storli and Hansen Sandseter, 2019).

3.2 Phase 2: Validation

A total of 14 responses were obtained from 14 different experts. In **Table 1**, expert responses are summarised, and the number of assessments received for each item is shown. Furthermore, CVR and its significance level for the two-tailed test were also included.

As shown in **Table 1**, items rated as "essential" by fewer than 10 of 14 experts were excluded, as they revealed a CVR that is not

significant. The rest of the items revealed a high significance level and were kept in the scale.

Upon analysing the degree of agreement between the experts, an intra-class correlation value for the coefficient of 0.83 was found. Furthermore, a Fleiss' kappa value of 0.74 was also found, showing substantial agreement among the experts (Vanbelle, 2019). The end result of this phase is a 38-item scale.

3.3 Phase 3: Reliability Analysis

For this phase, we relied on a sample of 221 students who responded to the scale. A Cronbach's alpha of 0.91, a Spearman-Brown coefficient of 0.70 and a Guttman split-half coefficient of 0.67 were obtained. According to Vaske et al. (2017), these values can be considered as consistent.

Apart from the aforementioned reliability analyses, the Cronbach's alpha was also calculated for each item according to whether the item is deleted. Based on the results, it was observed that eliminating items did not improve the overall reliability of the scale but instead worsened it. Therefore, all 38 items were kept for use in the following phase. The final version of the scale that was used for the EFA and CFA is provided as **Supplementary Material**.

3.4 Phase 4: Exploratory Factor Analysis

The Kaiser Meyer Olkin value (0.76) and Bartlett's sphericity test (1,239.05 p < 0.00) revealed satisfactory results, indicating that it was appropriate to conduct an EFA. The results of the Kaiser method are shown in **Table 2** (only components with eigenvalues greater than 1 are shown).

3.5 Phase 5: Confirmatory Factor Analysis

A CFA was carried out, taking the results of the EFA into consideration. Analysing the seven eigenvalues from Table 2 reveals that the distance between the highest (2.57) and the other eigenvalues is not considerable. Hence, according to Chatfield (2018), all eigenvalues in this study are equally important and none should be considered "small." Furthermore, as Auerswald and Moshagen (2019) state, only considering eigenvalues greater than one has a statistical basis, assuring that at least a major part of the variance is explained by these eigenvalues. Hence, a CFA for the original seven-factor structure was conducted. The analysis revealed the following values: CFI = 0.92; GFI = 0.90; χ^2/df = 2.13 and RMR = 0.05. As these consistent values of goodness and badness of fit indices belonged to the original seven factor structure, no more confirmatory factor analysis were performed. Factor loadings, which also seemed consistent for factors and items are provided in Table 3.

3.6. Phase 6: Descriptive and Inferential Statistical Analysis

The analysis by gender did not show statistically significant differences in either overall scores on the scale (t = 1.19, p = 0.24) or by factors. The analysis by grades did not show statistically significant differences either with the overall scores on the scale (F = 0.91, p = 0.44) or analysing the different factors

TABLE 2 | Global information from the Kaiser method.

Component		Initial eigenvalu	es	R	otation sums of square	ed loadings
	Total	% of variance	Cumulative %	Total	% of variance	Cumulative %
1	5.26	21.02	21.02	2.57	10.29	10.29
2	1.91	7.63	28.65	2.56	10.27	20.57
3	1.54	6.14	34.80	2.31	9.25	29.82
4	1.41	5.62	40.42	1.85	7.43	37.26
5	1.14	4.57	44.99	1.61	6.44	43.70
6	1.13	4.51	49.51	1.33	5.34	49.05
7	1.10	4.39	53.90	1.21	4.84	53.90

	Factor loading	Item	Item loading
		1	0.75
1. Cooperation	0.87	8	0.73
		12	0.83
		13	0.81
		24	0.85
2. Solidarity	0.89	25	0.88
		14	0.69
		15	0.85
3. Cohesion	0.81	21	0.76
		16	0.81
		19	0.82
		20	0.67
		22	0.73
4. Coexistence	0.77	23	0.65
		17	0.78
5. Attitudes towards school	0.75	18	0.71
		26	0.64
6. Attitude towards diversity	0.69	38	0.61
		9	0.65
7. Achievement	0.62	11	0.61

TABLE 4 | Overall results by factors

Ν	X	SD	α
8	2.94	0.41	0.95
4	2.89	0.53	0.86
3	2.39	0.74	0.83
5	2.63	0.57	0.89
2	2.45	1.01	0.76
13	3.60	0.31	0.93
3	2.59	0.61	0.81
38	3.02	0.56	0.91
	N 8 4 3 5 2 13 3 3 8	N X 8 2.94 4 2.89 3 2.39 5 2.63 2 2.45 13 3.60 3 2.59 38 3.02	N X SD 8 2.94 0.41 4 2.89 0.53 3 2.39 0.74 5 2.63 0.57 2 2.45 1.01 13 3.60 0.31 3 2.59 0.61 38 3.02 0.56

separately. Likewise, no differences were detected among schools (F = 1.42, p = 0.23) using either the overall scores or factors. The descriptive results are shown in **Table 4** with N, X, SD and α representing the number of factors, average score, standard deviation and Cronbach's alpha for each factor.

Table 5 shows the analysis of correlations between factors. In cases in which the correlation is statistically significant (p < 0.01) and the Pearson r-value is greater than 0.4, the box is shaded in grey and there is an asterisk (*) next to the value.

4 DISCUSSION

In this study, strong relationships between the factors were detected, as well as a single factor that could not be linked to any other. Thus, it is justified that the factors of cooperation and solidarity showed a high correlation given that they are strongly related from a theoretical standpoint. It should be borne in mind that the concept of cooperation is an attitudinal manifestation of solidarity towards solving common needs. Cooperation and participation in the classroom foster the development of positive attitudes such as solidarity. Students help each other to learn, share ideas and resources. Thus, the best way to educate in solidarity is practising it through cooperation (Koomen et al., 2020). The same holds true with the factors of solidarity and coexistence, with the research by authors such as Markham (2019) and Valero et al. (2018) showing a strong interaction between the two. Solidarity is fundamental to achieve coexistence; it is an attitude that is necessary not only in emergencies caused by war, hunger, drought, etc., but also when dealing with the different problems that affect other people: the next ones to us, those of our street, neighbourhood or city, school and also many others whom we do not know but who need our help, so we can help them according to our capability of solving these problems (Migliorini et al., 2019). In the same way, the strong similarities and correlations between the factors of coexistence, cohesion, attitude towards school and academic achievement have previously been documented by authors such as Putwain et al. (2019) and Burman and Miles (2018). One of the cornerstones of peaceful coexistence is group cohesion (Twenge, 2019). For a group to function well, it must be cohesive, and its members must feel part of it and proud to belong to it. This scenario enables the conditions for students to feel satisfaction attending the school, feel integrated into the group and behave properly because of the climate of coexistence that is generated.

On the contrary, the fact that attitude towards diversity does not show statistical similarities or correlations with any other factor is due to the uniqueness of this factor. Recent studies carried out by Juvonen, Kogachi and Graham (2018) and Azorín and Ainscow (2020) indicate that it is difficult to relate attitude towards diversity with other attitudinal, emotional or academic factors. Like any attitude, attitude toward diversity is a concept that is based on three elements (Heintzelman and Diener, 2019): feelings or affections, what we like or dislike in terms of the feeling it provokes in us; beliefs and knowledge, the belief that something is good or bad

	F1	F2	F3	F4	F5	F6	F7
F1—cooperation	1	0.49 ^a	0.38	0.49 ^a	0.30	0.20	0.50 ^a
F2-solidarity	0.49 ^a	1	0.35	0.52 ^a	0.43 ^a	0.21	0.38
F3—cohesion	0.38	0.35	1	0.63 ^a	0.32	0.17	0.39
F4-coexistence	0.49 ^a	0.52	0.63	1	0.47 ^a	0.28	0.42 ^a
F5-attitude towards school	0.30	0.43	0.32	0.47	1	0.10	0.28
F6-attitude towards diversity	0.20	0.21	0.17	0.28	0.10	1	0.11
F7-achievement	0.50 ^a	0.38	0.39	0.42	0.28	0.11	1

TABLE 5 | Analysis of correlations.

^ap < 0.05

according to the degree of knowledge, opinion or beliefs we have about it; and behaviours or actions, the idea that something is favourable or unfavourable for us based on the behaviour that we manifest when confronted with it. Attitude is not behaviour itself; it is a precursor (Tilleczek et al., 2019), and given its enormous complexity, it is an entity unto itself and, as mentioned above, is complex to relate to other factors analysed here.

The fact that no differences were detected by gender, age or schools, as well as the participation of independent international experts, are promising results for this scale in that it is meant to be used by an international audience and different cultures (Volk et al., 2018; Wu et al., 2019).

4.1 Practical and Research Implications

The main practical implication of the developed scale in this research is the possibility of assessing students' well-being in primary education students. Moreover, it gives researchers the chance to assess some factors such as cohesion, coexistence or attitude towards school that have not been so broadly studied in the literature. In this sense, it is quite difficult to find validate scales that assess one or several of these variables, so the potentiality of this scale must also be considered regarding this fact. From the research implication, researchers in the field may use it as a base to create and validate their specific instruments, questionnaires or scales adapting it for their specific contexts, that is, regarding the cultural, economic, academic, religious or contextual factors in which they want to carry out their research. In this regard, it can also serve as a guide for researchers in the field for other educational levels such as middle-school, high school or college, since many of the factors defined in this scale should also be considered in these educational levels when analysing students' well-being.

4.2 Limitations

Although the results and conclusions of this study are very promising, there are a number of limitations that must be taken into account when analysing them, which necessitate a cautious interpretation. First, the number of experts involved in the validation, though not limited, is far from ideal (Baghestani et al., 2019), which could compromise the universal validity of the scale (Sahin, 2017). The fact that all of them came from European countries must also be taken into account. Similarly, although the degree of agreement among the experts is adequate, it cannot be considered ideal (Penfield and Miller, 2004). Likewise, the reliability of scales, though not negligible, is considered questionable by various authors in the field of education (Vaske et al., 2017; Lisawadi et al., 2018). The goodness and badness of fit indices should also be considered. Although some authors (Reise et al., 1993) consider them suitable, they may seem insufficient for the some of the more demanding authors in this field (Heene et al., 2011). The sample size and its limited representativeness must also be considered (Krejcie and Morgan, 1970).

Regarding future directions, the scale has not yet been tested with 1st, 2nd, and 3rd grade students. In this sense, administering scales to students under the age of 9 poses many technical difficulties in terms of both their reading comprehension levels and their limited cognitive development (Simonds et al., 2007). It is not clear that the scale is suitable for students of these courses. In fact, the 4th grade students already experienced certain difficulties and required more help and time than the higher grades. A study with students of these ages or an adaptation of the scale would be interesting further research. Finally, although this study aspires to produce a scale that can be used by an international audience, it must be borne in mind that it has only been tested with a representative sample in Spain. Its use in different cultures and countries will determine whether it is indeed a valid scale to be used universally (Wrigley, 2018; Sin et al., 2019). In this sense, studies using this scale in different countries and contexts may be considered worthwhile future research.

5 CONCLUSION

The results from the scale show a strong relationship between different factors. Thus, it can be concluded that the factors of cooperation and solidarity, as well as the variables of solidarity and coexistence, are directly related. In the same way, these conclusions can be extended to the variables of cohesion, attitude toward school, coexistence and academic achievement, which are strongly related. On the contrary, attitude toward diversity is unique in that it is not directly related to and does not depend on any of the aforementioned factors.

This study presents a validated scale that allows Primary Education students' social well-being at school to be measured. Its considerable reliability and its confirmed factor structure also allow it to quantify the factors of cooperation, cohesion, coexistence, attitude towards school, attitude towards diversity, solidarity and academic achievement. It was designed and validated so it can be used by an international audience, and the promising results from this research suggest that it could be administered regardless of the students' ages or gender or the school that they attend. It is an instrument that provides a quick diagnosis of students' social well-being with a low number of items and a survey time of between 20 and 25 min. Based on the results, educational actions can be proposed by researchers or practitioners in the field to try to address those factors in which students show lower results.

DATA AVAILABILITY STATEMENT

The original contributions presented in the study are included in the article/**Supplementary Materials**, further inquiries can be directed to the corresponding author.

ETHICS STATEMENT

The studies involving human participants were reviewed and approved by the Conselleria de Educación de la Comunidad Valenciana. Written informed consent to participate in this study was provided by the participants' legal guardian/next of kin.

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AUTHOR CONTRIBUTIONS

LM: Literature review and discussion; FA: Methodology and statistical analizes; AC-M and OC: Discussion and Conclusion.

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SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/feduc.2021.800248/full#supplementary-material

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