



RETRACTED: Assessing the Role of Environmental Expenditures and Green Transport in Emissions Released by Transport: An Application of ARDL Approach

Zahid Hussain¹, Cuifen Miao², Weitu Zhang¹, Muhammad Kaleem Khan³ and Zhiqing Xia^{1*}

¹School of Finance, Qilu University of Technology (Shandong Academy of Sciences), Jinan, China, ²National Institute of International Strategy, Chinese Academy of Social Sciences, Beijing, China, ³School of Economics and Management, East China Jiao tong University, Nanchang, China

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*Correspondence:

Zhiqing Xia
xiazhiqing@hotmail.com

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This study investigates the effects of transport and environmental factors on transport carbon dioxide emissions (TCO₂). It employs cross-sectional autoregressive distributed lags for the estimation in the short and long runs and examines the panel time-series data from 2000 to 2020 in the OECD countries. This method allows heterogeneity in the dependencies and slope parameters across the countries. The results demonstrate that road and railway traffic movements increase the amount of TCO₂ in the short and long runs. In addition, transport energy consumption is the driving factor in releasing TCO₂ in the long run. Moreover, the joint effect of locomotives and transport energy consumption significantly reduces TCO₂ in the short run. By contrast, the findings support the argument that environmental expenditures and green transport mitigate TCO₂ in the long run. The findings also show an inverted u-shaped relationship between TCO₂ and transport energy consumption. With the empirical findings as a basis, we suggest that the OECD countries should reduce traffic movements and enhance the environmental expenditures so that they may produce green transport vehicles to combat environmental issues.

Keywords: traffic, transport energy consumption, environmental expenditures, green transport, OECD

INTRODUCTION

The transport sector has gained great attention owing to concern over environment quality. Environment quality is deteriorating because of carbon dioxide (CO₂) emissions, the most critical challenge for countries (Ahmad et al., 2021). The transport sector is a major contributor to CO₂ emissions, accounting for approximately 23% (Seum et al., 2020; Churchill et al., 2021; Sohail et al., 2021). Globally, CO₂ emissions from the transport sector will increase to nearly 60% by 2050 in the absence of effective mitigation measures (ITF, 2019). IEA (2019) also reports that approximately one-third of global CO₂ emissions are from the transport sector. Several economic activities such as wide-range development of transport infrastructure, traffic movement of vehicles, population growth, and economic growth have enhanced demand for transport vehicles, which indicate critical threats to sustainable development.

The transport energy consumption in the OECD countries is projected to reach an average rate of 1.2% from 2012 to 2040. Nevertheless, the patterns of transport energy consumption in the OECD countries are well established with vehicle efficiency improvements (Conti et al., 2016). Energy is the

most important factor for economic development, particularly in the transport sector (movement of passengers and freights), and the prime reason for environment deterioration. The association between transport, energy use, environment and economic development has been considered a debatable topic in the current century (Mehmood 2021; Habib et al., 2021; Tan et al., 2021).

Traffic movements are also creating challenges for the transport sector and the environment. The passenger- and freight-intensive movements resulted in trillion trips per annum, which increase the demand for locomotives (vehicles). These movements and locomotives substantially contribute to CO₂ emissions (Hadavi et al., 2020). Verlinghieri. (2020) argues that traffic movements indirectly affect CO₂ emissions through the usage of road and railway locomotives. Likewise, Heinold. (2020) asserts that road and railway locomotives produce a substantial amount of CO₂ emissions through energy consumption.

Considerable literature examines the different scenarios of transport modes with regard to CO₂ emissions. We seek to answer the following questions: What is the effect of traffic movements and transport energy consumption on CO₂ emissions? Do green transport and environmental expenditures affect CO₂ emissions?

This study concentrates on the OECD countries for a number of several reasons. First, the OECD countries are high-income economies that significantly contribute to the global economy, with approximately 42.8% of GDP at purchasing power parity. Second, the OECD countries are responsible for producing CO₂ emissions, and their transport energy consumption increases at an average rate of 1.4% per annum (from 104 quadrillion Btu to 155 quadrillion Btu) from 2012 to 2040. Furthermore, 55% of global transport energy consumption is by the OECD countries, whereas non-OECD countries account for around 45% (IEA, 2019). Third, the OECD countries are knowledge-based economies OECD. (2018). Billion trips of passengers and freights are made per year, which lead to traffic movements. Usually, traffic movements for international economic activities are toward the OECD countries. Fourth, the OECD countries are serious in controlling the emissions and improving environment quality. They are spending a larger amount of GDP for mitigating CO₂ emissions (Petrović and Lobanov, 2020). Fifth, the OECD countries are shifting their transport resources from carbon transport to green transport (e.g. electric vehicles). Therefore, regardless of the determinants of CO₂ emissions, whether and how transport energy consumption and green transport affect the association between CO₂ emissions and traffic movements, locomotives and environmental government budget remains a gap in the literature. Acknowledging the influence of traffic movements and green transport on CO₂ emissions in OECD countries is imperative to identify policy implications for sustainable transport policies (Rafique et al., 2022; Shahzad et al., 2021; OECD).

This study motivates by making numerous contributions to the literature. First, this study does not only investigate the linear effect of transport energy consumption on transport CO₂ emissions (TCO₂) but also analyzes the nonlinear effect. Second, this study examines the joint effect of transport energy consumption and locomotives on

TCO₂. Third, this study provides insights into the relationship between traffic movements and TCO₂. Fourth, we explore the effect of environmental research and development expenditures (ERDE) on TCO₂ and the joint effect of ERDE and environmental taxes on TCO₂. In addition, this study examines whether green transport significantly affects TCO₂. Fifth, this study makes methodological contribution by applying the cross-sectional autoregressive distributed lags (CS-ARDL) model to examine the relationship between traffic movements, transport energy consumption, ERDE, and green transport on TCO₂. In the presence of cross-dependency, heterogeneity, endogeneity problem, nonstationarity, and misspecification bias, the CS-ARDL is a robust method (Zeqiraj et al., 2020). For robustness check, we use a common correlated effects mean group (CCEMG) approach.

The structure of this article is described as follows. Previous studies are presented in **section 2**. **Section 3** provides the theoretical framework, data source, and methodology. The empirical results and discussions are given in **section 4**. Last, **section 5** summarizes the conclusions and policy implications.

LITERATURE REVIEW

Several decades take into account the attention on climate change and accelerating the degradation of environmental quality, which are the most critical challenges and big threat for the world. Under such circumstances, the governments need to acknowledge the importance of environmental issues and put into practice to counteract them. Therefore, the reasons behind those human activities are being executed e.g., economic growth, trade, energy use, urbanization, and so on. These activities cause environmental issues pertaining to global warming and climate change. For instance, enormous productivity of greenhouse gases (especially CO₂ emissions) is increasing the global temperature, pollution, and degrading the natural resources. Consequently, environment quality is being drastically deteriorated over the time period (Shahzad et al., 2020; Polloni-Silva et al., 2021; Polloni-Silva et al., 2021). Thus, we move to debate on association between transport sector and environment as follows.

Transportation has a significant effect on CO₂ emissions around the world. Transport and the environment can be studied in different ways. Generally, the growth of the transport sector has resulted in environment quality costs. Studies show that the environmental effects from the transport sector vary depending on transportation methods and regulations.

Recently, substantial research on the association between environment and transport has gained much attention. Churchill et al. (2021) research the effect of transport infrastructure on CO₂ emissions using parametric and nonparametric approaches for a panel of OECD countries. Their findings confirm that a 1% increase in transport infrastructure is associated with a 0.4% increase in CO₂ emissions. Furthermore, nonparametric estimation suggests a time-varying relationship between transport infrastructure and CO₂ emissions, which is positive throughout World War II and

up until now. From a critique perspective, transport infrastructure (e.g., construction of road, railway, and airport) is not the only one responsible for CO₂ emissions. Traffic activities also play a significant role in the emissions. Thus, this study includes traffic activities (e.g. movement and locomotives) in analyzing the effect on emissions released by the transport sector.

Subsequently, Ángel et al. (2021) focus on the relationship between road transport and CO₂ emissions in 22 European countries. Their results unveil that the transport sector is releasing CO₂ emissions that account for approximately 27% of the total emissions. Furthermore, road TCO₂ are almost entirely determined by (fossil) fuel consumption. Hence, energy use is the main determinant in CO₂ emissions, particularly the transport sector. Another explanation by Pani et al. (2021) also reports that freight transport upsurges the greenhouse emissions in the largest countries. The reason is that truck vehicles for forwarding the shipments stimulate the energy demand. Consequently, CO₂ emissions are released and deteriorate the environment quality. Similarly, Cardenete and López-Cabaco. (2021) document that transport of cargo is the most effective factor and that more than 30% of all modes of transport contribute to CO₂ emissions in Spain. Likewise, Arvin et al. (2021) analyze the correlation between and energy consumption in Germany. The fuel (e.g. gasoline and diesel) used in the transport vehicles upsurges CO₂ emissions through traffic locomotives.

Another evidence documented by Umar et al. (2021) highlights the effects of biomass energy consumption and fossil fuel energy consumption on CO₂ emissions in the transport sector in the United States. They find that fossil fuel energy has a positive and significant effect on CO₂ emissions released by the transport sector. Biomass energy consumption is negatively associated with CO₂ emissions though. Furthermore, they note an inverted U-shaped relationship between energy consumption and CO₂ emissions.

By contrast, Hussain et al. (2020) document that climate change potential (CO₂ emissions) has a negative association with transport infrastructure. Nonetheless, development infrastructure is also a driving factor of CO₂ emissions. Interestingly, extreme climate change potential reduces transport activities through critical infrastructure. Transport emissions are also investigated by Ahmed et al. (2020). Their results confirm that energy consumption in economic growth and the road sector increases emissions.

In support of mitigating CO₂ emissions, Sohail et al. (2021) emphasize the association between green transport and environment. They find that green transport is the better strategy to reduce CO₂ emissions. Electric vehicles reduce the demand for fossil energy consumption whilst increasing electricity demand. However, electricity shortage issues occur in the market. In this situation, resources are shifted from fossil energy to electric vehicles. Consequently, CO₂ emissions by the transport sector tend to decrease.

Another study supports green transport aimed at reducing the CO₂ emissions. Oryani et al. (2021) argue that renewable electric vehicles are supported to reduce CO₂ emissions per capita. A

substantial reduction in CO₂ emissions is possible by shifting internal combustible engine vehicles to alternative fuel vehicles.

Moreover, some studies provide evidence for reducing CO₂ emissions through environmental tax channels. For instance, Bergantino et al. (2021) analyze the effect of taxes on CO₂ emissions to improve environment quality. They argue that taxes on cars decrease CO₂ emissions because demand for and supply of cars decline simultaneously in the market. Consequently, an increase in sharing of cars reduces CO₂ emissions. Khastar et al. (2020) also point out that an adequate environmental tax level poses a mitigating effect on CO₂ emissions in the European Union countries. Furthermore, Reaños (2020) argue that carbon tax has a significant effect on CO₂ emissions. To reduce CO₂ emissions, carbon taxes need to be imposed (at least 30 Euros per tonne of CO₂ emissions) on vehicle owners. Price elasticities suggest that additional carbon taxes may stimulate vehicle owners toward intense energy consumption.

Mariano et al. (2016) evaluated the efficiency analysis on the transport logistics performance. They used a nonparametric approach e.g., slacks-based measure of the data envelopment analysis (DEA) with CO₂ emissions (treated as an input) and seven outputs e.g., GDP and six LPI components (treated as an output). Furthermore, window analysis and Malmquist index are also employed to evaluate the efficiency levels over the time. In addition, the DEA technique is based on mathematical foundation and has no specific assumption for analysis. It does not estimate the short-run and long-run relationships between the variables. In contrast, this study focuses on econometric approaches that have specific assumptions for analysis, and examines the short-run and long-run relationships between the variables e.g., transport carbon emissions, traffic, environmental expenditures, and green transport, used in this model.

Extant literature also considers the effect of environmental R&D on CO₂ emissions released from the transport sector. Substantial research proves that R&D supports the reduction of CO₂ emissions and improvement of environment quality. In this context, Petrović and Lobanov et al. (2020) report that the average effect of R&D is negatively associated with CO₂ emissions in the OECD countries. On average, a 0.15% decrease in CO₂ emissions is due to a 1% increase in R&D expenditures. Their results confirm that a higher level of R&D expenditures can reduce CO₂ emissions, but it does not apply to 40% of countries owing to scarce resources. Wang and Zhang. (2020) also find that a 1% increase in R&D expenditures reduces CO₂ emissions by 0.8122% in the BRICS countries.

The research on the association between traffic movements and locomotives must be considered in panoramic aspects. Besides, the joint effect of transport energy consumption and traffic locomotives remains unexplored. The nonlinear effect of transport energy consumption is ignored in the previous literature as well. In addition, the nexus between TCO₂ and traffic, transport energy consumption, ERDE, and green transport frequently neglects the potential heterogeneity and cross-sectional dependence. Consequently, a substantial gap exists in the prevailing literature. Therefore, the wide

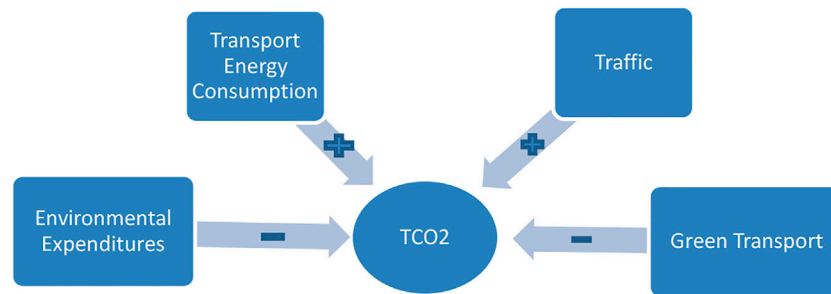


FIGURE 1 | Theoretical framework.

indicators of traffic, environmental expenditures, and green transport should be explored with different methods.

METHODOLOGY

Theoretical Framework

After reviewing the literature, we develop the theoretical framework to investigate the relationship between the concerned variables. **Figure 1** illuminates that traffic is intensely associated with CO₂ emissions. Traffic affects CO₂ emissions through movements and locomotives. Specifically, road and railway movements (e.g., passengers and freight) are indirectly associated with CO₂ emissions. Therefore, road and railway movements happen through locomotives (which require energy and transport-related resources). Consequently, CO₂ emissions are released. In addition, transport energy consumption is investigated in this analysis. Transport energy consumption directly influences CO₂ emissions.

The empirical model can be estimated as follows:

$$\begin{aligned} \text{TCO}_{2it} = & \beta_0 + \beta_1 (\text{ROADMOV}_{it}) + \beta_2 (\text{RAILMOV}_{it}) \\ & + \beta_3 (\text{ROADLOCO}_{it}) + \beta_4 (\text{RAILLOCO}_{it}) \\ & + \beta_5 (\text{ECT}_{it}) + \varepsilon_{it} \end{aligned} \quad (1)$$

Eq. 1 indicates that TCO₂ is a function of road movement, railway movement, road locomotive, railway locomotive, transport energy consumption, interaction of locomotive and transport energy consumption, and square of transport energy consumption. In addition, cross-sections are denoted by “*i*” (e.g. 35 OECD countries), whereas “*t*” represents the period from 2000 to 2020. The term “ β ’s” indicates the intercept and parameters, whereas the error term is denoted by ‘ ε ’. TCO₂ represents the TCO₂ measured in tons per year. ROADMOV indicates road movement that is defined as passenger and freight traveled distance per year. Likewise, RAILMOV represents the railway movement of the passengers and freight traveled distance per year. ROADLOCO indicates that locomotive (i.e. truck per year). RAILLOCO also represents the locomotive related to railway (i.e. freight coaches). ECT shows that transport energy consumption is in terms of tons. Specifically, it is the total energy (i.e. fuel, petrol, and diesel) consumed by the road and railway vehicles and locomotives in terms of tons per year. The quadratic term of transport energy consumption is used to examine

whether CO₂ emissions reduce once transport energy consumption achieves a threshold level. The joint effect of locomotive and transport energy consumption is denoted by the interaction term (LOCO*ECT).

The previous argument shows that road movement is expected to have a positive effect on TCO₂ ($\beta_1 = \frac{\partial \text{TCO}_2}{\partial \text{ROADMOV}} > 0$). Likewise, railway movement plays a crucial role in releasing TCO₂. A larger volume of freight (million kilogram) is moved to different locations. Hence, railway movement is predicted to have a positive effect on TCO₂ ($\beta_2 = \frac{\partial \text{TCO}_2}{\partial \text{RAILMOV}} > 0$). The road locomotives (e.g., traction engine and diesel engine) extensively move due to a larger amount of passengers and freights. Thus, it is also predicted to have a positive effect on TCO₂ ($\beta_3 = \frac{\partial \text{TCO}_2}{\partial \text{ROADLOCO}} > 0$). The rail locomotives (e.g., steam engine and diesel engine) substantially contribute to CO₂ emissions released by the transport sector (Cipek et al., 2021). Therefore, rail locomotive is anticipated to have a negative effect on TCO₂ ($\beta_4 = \frac{\partial \text{TCO}_2}{\partial \text{RAILLOCO}} > 0$). Mehmood. (2021) argues that energy consumption in the transport sector is a crucial factor in releasing CO₂ emissions, and it is also anticipated to have a positive effect on CO₂ emissions ($\beta_5 = \frac{\partial \text{TCO}_2}{\partial \text{ECT}} > 0$).

$$\begin{aligned} \text{TCO}_{2it} = & \beta_0 + \beta_1 (\text{ROADMOV}_{it}) + \beta_2 (\text{RAILMOV}_{it}) \\ & + \beta_3 (\text{ROADLOCO}_{it}) + \beta_4 (\text{RAILLOCO}_{it}) \\ & + \beta_5 (\text{ECT}_{it}) + \beta_6 (\text{LOCO*ECT}_{it}) + \varepsilon_{it} \end{aligned} \quad (2)$$

This study also investigates the joint effect of locomotive and transport energy consumption on TCO₂. Thus, we add the interaction term in the empirical model (**Eq. 1**) to examine the effect of locomotive on the association between TCO₂ and transport energy consumption. Hence, ($\beta_6 = \frac{\partial \text{TCO}_2}{\partial \text{LOCO*ECT}} < 0$).

Subsequently, we include the quadratic term of transport energy consumption in **Eq. 3** to estimate the U-shaped or inverted U-shaped between TCO₂ and ECT.

$$\begin{aligned} \text{TCO}_{2it} = & \beta_0 + \beta_1 (\text{ROADMOV}_{it}) + \beta_2 (\text{RAILMOV}_{it}) \\ & + \beta_3 (\text{ROADLOCO}_{it}) + \beta_4 (\text{RAILLOCO}_{it}) \\ & + \beta_5 (\text{ECT}_{it}) + \beta_6 (\text{LOCO*TEC}_{it}) + \beta_7 (\text{ECT}_{it}^2) + \varepsilon_{it} \end{aligned} \quad (3)$$

Furthermore, it is supposed to have a positive effect on TCO₂ ($\beta_7 = \frac{\partial \text{TCO}_2}{\partial \text{ECT}^2} > 0$).

Interestingly, we add environmental government budget, ERDE, and green transport factors to analyze their effect on TCO_2 .

$$TCO_{2it} = \beta_0 + \beta_1 (ROADMOV_{it}) + \beta_2 (RAILMOV_{it}) + \beta_3 (ROADLOCO_{it}) + \beta_4 (RAILOCO_{it}) + \beta_5 (EGB_{it}) + \beta_6 (ERDE_{it}) + \beta_7 (ETAX_{it}) + \beta_8 (ROEV_{it}) + \beta_9 (RAEV_{it}) + \varepsilon_{it} \quad (4)$$

Eq. 4 shows that TCO_2 is also a function of environmental government budget (EGB), ERDE, environmental taxes (ETAX), road electric (ROE), and railway electric (RAE). Environmental government budget is another important factor to protect environment quality. Hussain et al. (2020) argue that the governments allocate the national fund for different environmental projects that enable the government bodies to control the greenhouse emissions released from several sectors, particularly the transport sector and related industries. Thus, environmental government budget is anticipated to have a negative effect on transport energy consumption ($\beta_5 = \frac{\partial TCO_2}{\partial EGB} < 0$). Likewise, ERDE is an important factor to control TCO_2 . Wang et al. (2021) and Chishti et al. (2021) argue that ERDE can reduce TCO_2 through advancement in technology. ERDE is expected to have a negative effect on TCO_2 ($\beta_6 = \frac{\partial TCO_2}{\partial ERDE} < 0$). ETAX is a significant factor as well. Li et al. (2021) argue that environmental taxes discourage the production entities or companies in releasing a certain amount of CO_2 emissions. Thus, ETAX is predicted to have a negative effect on TCO_2 ($\beta_7 = \frac{\partial TCO_2}{\partial ETAX} < 0$). In addition, road electric vehicles have a significant effect on TCO_2 . Hu et al. (2021) argue that road electric vehicles play a crucial role in reducing TCO_2 . These vehicles are considered a green transport strategy to protect the environment. Thus, it is anticipated to have a negative effect on TCO_2 ($\beta_8 = \frac{\partial TCO_2}{\partial ROEV} < 0$). Railway electric vehicles also effect TCO_2 as an alternative approach. Kejun et al. (2021) debate that railway electric vehicles impede TCO_2 pathways. These vehicles are regarded as an important transport measure to control CO_2 emissions released by the transport sector. RAE is expected to have a negative effect on TCO_2 ($\beta_9 = \frac{\partial TCO_2}{\partial RAEV} < 0$).

$$TCO_{2it} = \beta_0 + \beta_1 (ROADMOV_{it}) + \beta_2 (RAILMOV_{it}) + \beta_3 (ROADLOCO_{it}) + \beta_4 (RAILOCO_{it}) + \beta_5 (EGB_{it}) + \beta_6 (ERDE_{it}) + \beta_7 (ETAX_{it}) + \beta_8 (ROEV_{it}) + \beta_9 (RAEV_{it}) + \beta_{10} (ERDE_{it}^2) + \beta_{11} (ROEV_{it}^2) + \varepsilon_{it} \quad (5)$$

We extend the empirical model (Eq. 4) to analyze the nonlinear effect of ERDE and ROE, specifically whether the nexus between TCO_2 , ERDE and ROE is U-shaped or inverted U-shaped. The coefficients of square of ERDE and ROE are predicted to have a negative effect on TCO_2 ($\beta_{10} = \frac{\partial TCO_2}{\partial ERDE^2} < 0$), ($\beta_{11} = \frac{\partial TCO_2}{\partial ROE^2} < 0$).

We also include the interaction term (EGB*ETAX) in the empirical model (Eq. 6) to analyze the joint effect of EGB and

ETAX on TCO_2 . Environmental government budget is interacted with environmental taxes to estimate the joint effect on TCO_2 emissions. More precisely, EGB may mitigate the effect on TCO_2 emissions through the environmental taxes. Yuelan et al. (2021) also support that environmental government budget has a significant impact on emissions. The reason is that environmental taxes are also main sources of government budget related to environment (Mirović et al., 2021; Rafique et al., 2022).

$$TCO_{2it} = \beta_0 + \beta_1 (ROADMOV_{it}) + \beta_2 (RAILMOV_{it}) + \beta_3 (ROADLOCO_{it}) + \beta_4 (RAILOCO_{it}) + \beta_5 (EGB_{it}) + \beta_6 (ERDE_{it}) + \beta_7 (ETAX_{it}) + \beta_8 (ROEV_{it}) + \beta_9 (RAEV_{it}) + \beta_{10} (ERDE_{it}^2) + \beta_{11} (ROEV_{it}^2) + \beta_{12} (EGB_{it} * ETAX_{it}) + \varepsilon_{it} \quad (6)$$

To analyze the joint effect of EGB and ETAX, we predict that the coefficient of the interaction term has a negative effect on TCO_2 ($\beta_{12} = \frac{\partial TCO_2}{\partial EGB * ETAX} < 0$).

Data and Source

This study uses balanced panel data set from 2000 to 2020 for 35 OECD countries, namely, Australia, Austria, Belgium, Canada, Chile, Czech Republic, Denmark, Estonia, Finland, France, Germany, Greece, Hungary, Iceland, Ireland, Italy, Japan, Korea, Latvia, Lithuania, Luxembourg, Mexico, The Netherlands, New Zealand, Norway, Poland, Portugal, Slovak Republic, Slovenia, Spain, Sweden, Switzerland, Turkey, the United Kingdom, and the United States. We employ 11 variables in the empirical models. Table 1 reveals the operational variables and data sources.

Cross-Sectional Dependence Test

We employ a cross-sectional dependence (CD) test on the variables. The reason is that countries are interconnected through multiple aspects such as economic, cultural, political, and social. As a result, dependency may exist. We follow Pesaran's (2004) CD and scaled LM test. The equation of CD test is given as

$$CD = \sqrt{\frac{2T}{N(N-1)}} \left(\sum_{i=1}^{N-1} \sum_{j=i+1}^N \tilde{\rho}_{ij} \right), \quad (7)$$

where $\tilde{\rho}_{ij}$ indicates pairwise correlation of the cross-sectional residuals that are obtained from augmented Dickey-Fuller (ADF). Therefore, "T" and "N" are indicators of time and cross-section dimensions, respectively.

Unit Root Test

This study employs cross-sectionally augmented ADF and cross-sectionally augmented IPS test (Pesaran, 2007) to examine the stationarity characteristics. These also can be known as CADF and CIPS, respectively. Furthermore, arguments by Pesaran. (2007), Moon and Perron. (2004), and Bai and Ng (2004)

TABLE 1 | Variables and data sources.

Variable	Code	Measurement	Source
Transport-carbon emission	TCO ₂	Million ton	OECD
Road Movement	ROADMOV	Passenger and freight Kilo meter (million)	OECD
Rail Movement	RAILMOV	IV-07 (1000)	OECD
Road Locomotive	ROADLOCO	Number at (31.12) total traction engine (million)	OECD
Rail Locomotive	RAILLOCO	Number at (31.2) total steam engine, diesel (million)	OECD
Transport-energy consumption	ECT	Percentage of total energy consumption	OECD
Road-electric vehicle	ROEV	Million	OECD
Rail-electric vehicle	RAEV	million	OECD
Environmental government budget	EGB	Share of percentage of total EGB	WDI
Environmental R&D expenditures	ERDE	Percentage of GDP	WDI
Environmental Tax	ETAX	Percentage of GDP	WDI

Source: author's calculation.

draw a picture. That is, CIPS approach is a second-generation panel unit root test, and it is an efficient method for CD and heterogeneity. The estimated equation is as follows:

$$\Delta CA_{it} = \varnothing_i + \varnothing_i Z_{i,t-1} + \varnothing_i \overline{CSA}_{t-1} + \sum_{i=0}^p \varnothing_{it} \Delta \overline{CSA}_{t-1} + \sum_{i=0}^p \varnothing_{it} \Delta CA_{i,t-1} + \mu_{it} \quad (8)$$

where cross-section averages are denoted by \overline{CSA}_{t-1} and $\Delta \overline{CSA}_{t-1}$. Therefore, the test statistics of CIPS is given as

$$\widehat{CIPS} = 1/N \sum_{i=1}^n CDF_i \quad (9)$$

In Eq. 9, CDF represents the cross-sectional ADF.

Cointegration Test

To investigate the long-run association between the model parameters, we employ the ECM-based cointegration method (Westerlund, 2007). This method is more appropriate than conventional methods such as Pedroni and Kao. The reason is that the Westerlund cointegration method provides unbiased results in the presence of cross-dependency and heterogeneity. The test statistics can be stated as

$$\alpha_i(L) \Delta y_{it} = \delta_{1i} + \delta_{2it} + \alpha_i (y_{it-1} - \beta_i x_{it-1} + \lambda_i(L) v_{it} + \varepsilon_{it}), \quad (10)$$

where $\delta_{1i} = \alpha_i(1)\phi_{2i} - \alpha_i\phi_{1i} + \alpha_i\phi_{2i}$, but $\delta_{2i} = -\alpha_i\phi_{2i}$. α_i represents the error correction term. The test statistics is described as

$$G_t = 1/N \sum_{i=1}^N \alpha_i' / SE(\alpha_i') \quad (10.1)$$

$$G_\alpha = 1/N \sum_{i=1}^N T_i' / (\alpha_i') 1 \quad (10.2)$$

$$P_t = \alpha_i' / SE(\alpha_i') \quad (10.3)$$

$$\alpha_i' = P_\alpha / T \quad (10.4)$$

The error correction parameters (α_i') in Eq. 10 are computed by replacing the value of $P_\alpha = T_\alpha'$ (Eq. 10.4). Hence, the error correction parameter is specified as (α_i') = P_α/T that indicates the

ratio of error to be corrected each year in the short run for the disequilibrium case.

Cross-Sectionally Augmented Autoregressive Distributed Lags

We analyze the short- and long-run relationship among the variables (used in the current empirical models) by employing the CS-ARDL model suggested by Chudik and Pesaran. (2015). The framework of this model includes long-run parameters, short-run ones with error correction and cross-sectional mean. It overcomes the issues such as cross-sectional dependence, heterogeneity, and non-stationarity as robust and produces reliable outcomes (Zeqiraj et al., 2020; Ahmad et al., 2021). The model for CS-ARDL is proposed as follows:

$$\Delta TCO_{2,it} = \vartheta_i + \sum_{j=1}^p \vartheta_{it} \Delta TCO_{2,i,t-1} + \sum_{j=1}^p \vartheta_{ij} AV_{i,t-1} + \sum_{j=1}^p \vartheta_{it} \overline{Z}_{t-j} + \varepsilon_{it}, \quad (11)$$

where $\Delta TCO_{2,it}$ is the dependent variable, $AV_{i,t-1}$ and \overline{Z}_t are the independent variables and the averages for cross-sections, respectively.

Robustness and Causality Tests

After estimation on CS-ARDL, we check the robustness by using the CCEMG of Pesaran. (2006). This test allows parameters to be heterogeneous in the long run. CS-ARDL is criticized due to imposition of homogeneity restriction in the long run, but economies are diverse with regard to economic and social structures. Subsequently, despite the reliable outcomes obtained by the CS-ARDL and CCEMG, the direction of association between observed variables is not analyzed (which is important for policy implications). We then apply the Dumitrescu and Hurlin (D&H, 2012) method to investigate the casual relationship among the variables (used in the current models). This method provides two statistics that are test average (\bar{W}) and standard normal distribution (\bar{Z}). The model can be described as

TABLE 2 | Descriptive statistics.

Variable	Obs	Mean	Std.Dev	Min	Max
TCO ₂	735	92.65	72.998	11.2	861
ROAD-MOV	735	84581.4	548000	1	5210000
RAIL—MOV	735	59434.79	182000	1	1140000
RAIL-LOCO	735	59434.79	182000	1	1140000
ROAD-LOCO	735	666000	2950000	1	2.26e+07
ECT	735	253.327	1759.159	1	27615
EGB	735	25.695	11.399	1	58.94
ERDE	735	2.49	2.094	0.05	17.66
ETAX	735	0.104	0.119	0.01	1
ROAD-ELECT	735	20.32	0.123	0.60	652143
RAIL-ELECT	735	10.23	0.432	0.30	32154

Source: author's calculations.

$$Z_{i,t} = \alpha_i + \sum_{j=1}^p \beta_i^j Z_{i,t-1} + \sum_{j=1}^p \gamma_i^j T_{i,t-j}, \quad (12)$$

where β^j (j) and j indicate the autoregressive parameters and lag length, respectively.

EMPIRICAL RESULTS AND DISCUSSIONS

Descriptive Statistics

Table 2 reveals the descriptive statistics. The mean values and standard deviations of road movement (ROAD-MOV), rail movement (RAIL-MOV), rail locomotive (RAIL-LOCO), road locomotive (ROAD-LOCO), and transport energy consumption (ECT) are very high, indicating high heterogeneity in the variables across the OECD countries. Moreover, the standard deviations of EGB and TCO₂ are higher (11.39 and 72.99, respectively), implying that observations follow a skewed distribution across the sample countries. However, the mean value and standard deviation of environmental tax (ETAX) are smaller, indicating that observations vary within a narrow range over the period.

We investigate the cross-dependency on the observed variables. To analyze the models, detecting the presence of CD is important. Pesaran. (2004) explains that ordinary econometric methods often cannot overcome the bias in the panels because of presence of CD. Table 3 reports the estimate of Pesaran's CD and scaled LM test. The results reveal that CD is supported by the value of absolute mean (ranging from 0.696 to 107.619). Therefore, the outcomes of Pesaran's CD and Pesaran scaled LM test are highly significant for the entire observed variable, indicating that variables have CD. It is worth noticing because with globalization, emerging economies are interconnected. Consequently, the outcomes of the CD test are significantly projected in the model. Some possible changes in the observed variables of emerging economies may affect those of the other economies.

To investigate the order of integration, we employ a second-generation panel unit root test (CIPS and CADF) by Pesaran. (2007), which shows that cross-sectionally unbiased is a primary

feature of CIPS. Investigation on the order of integration is a unique factor in the estimation technique. The outcomes of the CIPS and CADF are summarized in Table 4. The findings exhibit that all variables are stationary at 1 and follow a mixed order of integration. The presence of CD and mixed order of integration require the usage of CS-ARDL framework. Subsequently, we employ a Westerlund cointegration approach to investigate the long-run relationship in the models. The outcomes are given in Table 5, which shows that a long-run relationship exists in the models. Furthermore, error correction (EC) can be calculated by $P\alpha$ value in the models. Hence, the parameter of EC is ($\alpha' = \frac{P\alpha}{T}$) = $-3.966/18 = -0.220$ for model 1, $-5.893/18 = -0.327$ for model 2, and $-3.854/18 = -0.214$ for model 3. The errors around 25.37% between TCO₂ and its determinants are corrected each year, so disequilibrium in the short run becomes stable in the long run.

Role of Traffic and Transport Energy Consumption

After evaluation of cointegration, we employ a CS-ARDL method to gauge the dynamic effect of traffic and transport energy consumption in the short and long runs. Table 6 displays the estimation using CS-ARDL. The magnitude of each coefficient indicates a significant relationship between explanatory variables and TCO₂ in the short and long runs. A 29.5% increase in TCO₂ is due to a 1% increase in road movement in the short run, whereas a 4.3% augmentation in TCO₂ is affected by a similar rise in rail movement in the short run. Conversely, road and rail movement variables have relatively less influence on TCO₂ in the long run. Payus et al. (2019), Chen et al. (2018), and Grote et al. (2018) argue that movement of passengers upsurges the use of road vehicles. Consequently, CO₂ emissions released in the environment are due to traffic activities.

Besides, the positive and significant coefficient of ECT indicates an increase in TCO₂. Numerous studies (e.g., Adams et al., 2020; Figliozzi 2020; Peng et al., 2020) find that the use of

TABLE 3 | Cross-Sectional Dependence.

Variable	Pesaran CD		Pesaran Scaled LM
	CD-test	abs (corr)	CD-test
TCO ₂	101.153***	0.90	132.333***
ROAD-MOV	22.129***	0.20	231.433***
RAIL—MOV	44.54***	0.40	154.298***
RAIL-LOCO	28.464***	0.26	120.938***
ROAD-LOCO	17.447***	0.16	121.433***
ECT	107.619 ***	0.96	201.432***
EGB	9.149***	0.30	103.322***
ERDE	4.186***	0.08	98.322***
ETAX	54.142***	0.55	123.543***
ECT2	100.157***	0.90	190.432***
EGB2	5.315***	0.28	134.329***
ERDE2	3.953***	0.08	102.433***

Source: author's calculations.

Note: Table 2 reveals the estimate of cross-dependency (CD) test of Pesaran. CD and Pesaran Scaled LM of observed variables of 35 OECD countries.

TABLE 4 | Panel Unit Root.

Variable	Cross-sectionally Augmented IPS (CIPS)		Cross-sectionally Augmented Dicky-Fuller (CADF)	
	Level	First-difference	Level	First-difference
TCO ₂	-2.498	-1.232***	-4.479***	-11.086***
ROAD-MOV	-0.871	-0.432*	10.961	9.678***
RAIL-MOV	-1.698	-1.327***	2.823	-0.799***
RAIL-LOCO	-0.190 ***	—	12.269	6.966***
ROAD-LOCO	-0.252***	—	13.296	12.567***
ECT	-2.617	-1.432***	-1.404	-9.930***
EGB	-2.614*	-1.232***	-1.576 *	-9.157***
ERDE	-0.530 *	-0.432*	14.714	7.925**
ETAX	-1.754	-1.422*	3.558	-4.340***
ECT2	-2.610	0.323*	-1.609*	-10.079***
EGB2	-2.869***	—	-2.222 ***	-9.548***
ERDE2	-0.516*	-0.431*	14.714	7.925**

Source: author's calculations.

Note: CIPS, CADF.

TABLE 5 | Cointegration Test.

Variable	Model 1	Model 2	Model 3
Gt	-2.557*** (-5.133)	-1.540 (1.566)	-1.635*** (0.938)
Ga	-2.636*** (4.897)	-3.354 (4.117)	-3.983* (3.433)
Pt	-17.093*** (-8.578)	-16.099*** (-7.578)	-9.259*** (-0.699)
Pa	-3.966*** (0.352)	-5.893*** (-2.216)	-3.854*** (0.501)

Source: author's calculations.

Note: Westerlund cointegration test.

energy is the most influencing factor of TCO₂, and a positive relationship exists between energy consumption and CO₂ emissions in the transport sector. A higher volume of economic activities particularly in the transport sector enhance energy consumption, which then increases CO₂ emissions.

The effect of road and rail locomotive is positive and significant with respect to TCO₂, which implies that extensive use of locomotive is causing TCO₂. Entities buy more fossil fuel vehicles (e.g., truck, bus, car, and motorcycle), so TCO₂ increase. Several other studies (e.g., Rietmann et al., 2020; Arvin et al., 2021; Li et al., 2021; Xu et al., 2020; Yan et al., 2021) endorse this finding, arguing that locomotives are the main driving factors of CO₂ emissions in the transport sector. The demand for trip/traveling within a specific time stimulates the use of fuel vehicles by individuals/consumers. Therefore, road and rail locomotive augments TCO₂. The coefficient of ECT square indicates that a 1% increase in ECT² decreases TCO₂ by 43.2 and 54.3% in the short and long runs, respectively. An inverted U-shaped relationship occurs, indicating that a particular amount of energy consumption in the transport sector can reduce TCO₂ by adopting environmental strategies (e.g. efficient energy use, electric vehicles, and sustainable locomotive production and consumption patterns). Furthermore, the outcome reflects that OECD countries' environmental strategies are in the right

direction as their economies are progressively adopting alternative uses of transport vehicles to protect their environment (Lin, (2020); He et al., 2021; Rietmann et al., 2020; Lin, (2020); Sharma and Chandel. 2020; Harvey 2020). Subsequently, the coefficient of interaction term (ECT*LOCOM) indicates that a joint effect of transport energy consumption and locomotives increases TCO₂, implying that locomotives increase TCO₂ through ECT.

Role of Environmental Expenditures and Green Transport

After assessing the effect of traffic and transport energy consumption on TCO₂, we can move forward to investigate the role of environmental expenditure and green transport to resolve the issue of CO₂ emissions released by the transport sector. The coefficients of environmental government budget (EGB) are negative and significant in **Table 7**, which indicate that a 1% increase in EGB reduces TCO₂ by approximately 146% (model 1), 13.2% (model 2), and 2.1% (model) in the short run. On average, 28.6, 1.7, and 22.1% decreases in TCO₂ in all the cases are due to EGB in the long run.

The results reveal that the current environmental policies of the OECD countries are in the right direction for protecting the environment, particularly regarding the release of TCO₂. Some related studies (e.g. Fan et al., 2020; Yang et al., 2020) argue that the government budget related to the environment has a remarkable effect on CO₂ emissions. It is considered a driving force to tackle environmental issues. Lack of funding can impede the identification of environmental issues though. Thus, the government budget supports the relevant departments or entities to tackle the problems. However, OECD countries can also constrain the governmental budget related to the environment.

ERDE is another important factor to mitigate CO₂ emissions. The coefficient of ERDE is negative and significant. On average, a 1% upsurge in ERDE reduces TCO₂ by approximately 3.1% (model 1), 32.1% (model 2),

TABLE 6 | The role of traffic and transport energy consumption.

Variable	Short-run			Long-run		
	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3
ROAD-MOV	0.183*** (0.232)	0.295*** (0.028)	0.254*** (0.034)	0.091*** (0.023)	0.181*** (0.24)	0.127*** (0.032)
RAIL-MOV	0.110*** (0.034)	0.043*** (0.057)	0.098*** (0.032)	0.543** (0.432)	0.087*** (1.216)	0.432** (0.321)
ECT	0.403*** (0.332)	0.314*** (0.847)	0.323*** (0.023)	0.220*** (0.329)	12.6*** (0.498)	0.201** (0.232)
ROAD-LOCO	0.782*** (0.143)	0.890*** (0.096)	0.891*** (0.213)	0.792** (0.321)	9.576*** (9.221)	0.702*** (0.485)
RAIL-LOCO	0.348*** (0.432)	0.085*** (0.078)	0.232*** (0.043)	0.038*** (0.043)	1.431*** (1.048)	0.243*** (0.093)
ECT*LOCOMOTIVE	—	0.015* (0.470)	0.002* (0.123)	—	0.0613* (1.123)	0.006* (0.083)
ECT ²	—	—	−0.432*** (0.432)	—	—	−0.543*** (0.093)
SR Error Correction	−0.325*** (0.065)	−0.234*** (0.098)	−0.308*** (0.088)	—	—	—
Observations	630	630	630	630	630	630
R-squared	0.499	0.639	0.604	0.575	0.475	0.363
Number of groups	35	35	35	35	35	35

Note: *, **, and *** denote 10, 5 and 1% significance level. The standard errors are in parentheses.

TABLE 7 | The role of environmental expenditures and green transport.

Variable	Short-run			Long-run		
	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3
ROAD-MOV	0.021*** (0.032)	0.021** (0.032)	−0.048 (0.094)	0.032** (0.038)	0.048* (0.032)	0.043*** (0.052)
RAIL-MOV	0.023*** (0.032)	0.039*** (0.003)	0.321* (0.054)	0.021*** (0.043)	0.003** (0.085)	0.031** (0.076)
ROAD-LOCO	−0.383 (0.438)	0.038*** (0.083)	−0.003** (0.084)	0.048** (0.043)	−0.039 (0.998)	0.010** (0.076)
RAIL-LOCO	0.002*** (0.053)	0.004* (0.021)	−0.033 (0.098)	0.323*** (0.048)	0.021* (0.443)	0.001** (0.028)
EGB	−0.021*** (0.212)	−0.132*** (0.021)	−1.461*** (0.691)	−0.221*** (0.021)	−0.017*** (0.012)	−0.286*** (0.211)
ERDE	−0.217*** (0.021)	−0.321*** (0.032)	−0.031*** (0.031)	−0.028*** (0.032)	−0.018*** (0.028)	−0.581*** (3.178)
ETAX	−0.021** (0.081)	−0.028** (0.212)	−0.55** (0.11)	−0.021*** (0.028)	−0.038*** (0.321)	−0.468*** (0.791)
RAIL-ELECTRIC	−0.432*** (0.023)	−0.124*** (0.082)	−0.103*** (0.089)	−0.765*** (0.033)	−0.543*** (0.032)	−0.155*** (0.005)
ROAD-ELECTRIC	−0.472*** (0.032)	−0.328*** (0.024)	−0.276*** (0.231)	−0.573*** (0.076)	−0.382*** (0.063)	−0.335*** (0.155)
ERDE ²	—	−0.128** (0.022)	−0.323** (0.082)	—	−0.053*** (0.032)	−0.281*** (0.321)
ROAD-ELECT-SQU	—	−0.543*** (0.032)	−0.707*** (0.887)	—	−0.654*** (0.203)	−0.179*** (0.318)
EGB*ETAX	—	—	−0.007*** (0.841)	—	—	−0.004*** (0.045)
SR Error Correction	−0.398*** (0.083)	−0.403*** (0.032)	−0.432*** (0.093)	—	—	—
Observations	630	630	630	630	630	630
R-squared	0.513	0.404	0.287	0.482	0.288	0.295
Number of groups	35	35	35	35	35	35

Note: *, **, and *** denote 10, 5 and 1% significance level. The standard errors are in parentheses.

and 21.7% (model 3) in the short run. In the long run, a 58.1% decrease in TCO_2 is due to ERDE. These outcomes partially coincide with related studies by Wang et al. (2021) and Chishti et al. (2021) who explain the remarkable effect of ERDE on TCO_2 because the former drives invention of new technology for tackling the environmental issues.

Additionally, the coefficient of ERDE square explains that a 1% increase in ERDE square decreases TCO_2 by approximately 32.3% (model 1), 12.8% (model 2), and 43.2% (model 3) in the short run. The negative relationship between ERDE square and TCO_2 is also apparent in the long run in OECD countries. This indicates that TCO_2 can be reduced with technological innovation and environmental economic policies after reaching a specific level of R&D. It also shows that OECD countries are giving greater attention on environment by adopting the alternative approaches, such as green transportation through green financing (Kong et al., 2021; Song et al., 2021).

Subsequently, our investigation focuses on the effect of environmental taxes on TCO_2 . The coefficients of ETAX are negative and significant in all cases, which shows that 55% (model 1), 2.8% (model 2), and 2.1% (model 3) decrease in TCO_2 is due to environmental taxes in the short run. A slight change in the coefficients of ETAX is noted, and the negative correlation among the variables is validated in the long run. This finding is also consistent with those of other studies, for instance, Hao et al. (2021), Ma et al. (2021), and Zhai et al. (2021) who document that environmental taxes have a remarkable effect on CO_2 emissions. The reason is that taxes are sources of government bodies which are used for mega projects, including sustainable development, agriculture and industrial growth, and infrastructure (Arvin et al., 2021).

Now our concentration moves to the effect of green transport. Electric vehicles are used as a proxy to analyze the effect on TCO_2 . The coefficients of railway electric vehicles are negative and significant for TCO_2 in all the cases. A 1% increase in railway electric vehicles decreases TCO_2 by approximately 43.2 and 76.5% in the short and long runs, respectively (see Table 6). On the contrary, the coefficients of road vehicles also validate the negative relationship with TCO_2 . On average, in the short run, 27.6% decrease in TCO_2 for model 1, 32.8% for model 2, and 47.2% for model 3 are due to road vehicles. Furthermore, a 1% increase in road vehicles decreases TCO_2 by approximately 33.3% (model 1), 38.2% (model 2), and 57.3% (model 3) in the long run (see Table 6).

An inverted U-shaped relationship between TCO_2 and road vehicles shows that after reaching a specific level of road vehicles, TCO_2 can be reduced with better transport and environmental policies. OECD countries are progressively shifting their resources from diesel to electric vehicles. Our findings are consistent with those of related studies by Xu et al. (2021) and Zhang and Hanaoka (2021) who conclude that electric vehicles have a remarkable effect on TCO_2 as an alternative transport strategy.

The interaction term of EGB and ETAX is negative and significant. The joint effect of EGB and ETAX reduces TCO_2 ,

TABLE 8 | Robustness Check.

	Model 1	Model 2	Model 3
ROAD-MOV	0.0295*** (0.002)	0.040*** (0.009)	3.245** (0.026)
RAIL-MOV	0.0438** (0.057)	0.012*** (0.002)	0.009*** (0.002)
ROAD-LOCO	0.085*** (0.000)	0.096*** (0.000)	-0.112 (0.091)
RAIL-LOCO	0.089*** (0.009)	2.210 (0.505)	3.041*** (0.000)
ECT	0.0382** (0.012)	0.124*** (0.132)	0.132*** (0.023)
EGB	—	0.425*** (0.127)	—
ERDE	—	-2.1705* (0.470)	—
ETAX	—	0.314*** (0.847)	—
RAIL-ELECTRIC	—	—	0.546*** (0.127)
ROAD-ELECTRIC	—	—	0.031*** (0.082)
ROAD-ELECT-SQU	—	—	-0.001*** (0.142)
EGB*ETAX	—	-0.002** (0.032)	—
Observations	630	630	630
R-squared	0.639	0.316	0.355
Number of groups	35	35	35

Note: *, **, and *** denote 10, 5 and 1% significance level. The standard errors are in parentheses.

implying that governments simultaneously allocate the budget and impose taxes related to environment to reduce CO_2 emissions.

Endogeneity and Robustness Check

Endogeneity is the concept of econometric that explains about the correlation between explanatory variables and error terms. Distinguish between endogenous and exogenous variables created in simultaneous equations models where a separate variable is determined by the model that is predetermined. Usually, endogeneity problem arises due to correlation between the explanatory variables and errors term from unobserved or omitted variables is confound both independent variable and dependent variable.

Thus, we attempt endogeneity and robustness of the modeling techniques to present the potential reverse causality problem. We use the best strategy, namely, CCEMG, to estimate the correlation of unobserved variables with explanatory variables and error term. This model can correlate with unobserved variables, such as investment in transport, total energy consumption, freight and passenger volume, and CO_2 emission release from other sectors. Table 8 reports the estimate, which indicates that the outcomes are significant. The results portray that the outcomes produced by CS-ARDL are valid, indicating consistent findings. In addition, Table 8 (model 1) shows that traffic factors (road movements, rail movements, road locomotives, rail locomotives, and transport energy

TABLE 9 | Panel Causality Test Results.

Null Hypothesis	W-statistics	Zbar-statistics	Prob	Conclusion
TCO ₂ ↔ ROADMOV	1.323***	0.343	0.000	ROADMOV → TCO ₂
ROADMOV ↔ TCO ₂	1.237***	0.432	0.000	
TCO ₂ ↔ RAILMOV	1.543	0.733	0.234	
RAILMOV ↔ TCO ₂	1.454***	0.873	0.000	RAILMOV → TCO ₂
TCO ₂ ↔ ECT	2.1230	0.3638	0.7160	ECT → TCO ₂
ECT ↔ TCO ₂	1.4327***	-1.6781	0.0033	
TCO ₂ ↔ ROLOC	1.554	0.432	0.432	
ROLOC ↔ TCO ₂	1.654***	0.232	0.000	ROLOC → TCO ₂
TCO ₂ ↔ EGB	2.006	0.0177	0.9859	EGB → TCO ₂
EGB ↔ TCO ₂	3.3591***	4.0204	0.0001	
TCO ₂ ↔ ERDE	1.343	0.393	0.0898	
ERDE ↔ TCO ₂	1.543***	0.493	0.000	ETAX ↔ TCO ₂
TCO ₂ ↔ ETAX	3.4939***	4.4189	0.0000	
ETAX ↔ TCO ₂	6.9714***	14.7057	0.0000	

Source: author's calculations.

Note: "→" indicates one-way causality, while "↔" two-way causality between the variables.

consumption) are statistically significant at 1% level. It indicates that the relationship is consistent with the main approach. On the other hand, model 2 also reveals that there is no drawback in findings with the CS-ARDL estimation. Likewise, the coefficients of green transport in model 3 also indicate that there is no robust effect in the model. Generally, unobserved or omitted variables are not correlated with explanatory and dependent variable (Hussain et al., 2020; Mehmood et al., 2021; Isik et al., 2021).

We then carry out the D&H panel causality test to estimate the causal association (Dumitrescu and Hurlin, 2012). The outcomes are given in **Table 9**. Road movement and TCO₂ and environmental taxes and TCO₂ have bidirectional causality. Therefore, any policy shock in road movement and environmental taxes can substantially affect TCO₂ and vice versa. For instance, an increase in CO₂ emissions requires higher amount of environmental taxes (that are source of environmental revenue) (Tao et al., 2021). Conversely, one-way causality to TCO₂ cannot be reversed among variables such as transport energy consumption, environmental government budget, and road locomotive. Tao and Wu. (2021) also support that any policy change in transport energy consumption cannot affect TCO₂ emissions indicating that the transport energy consumption independently does work in its own mechanism without being affected by external factors.

CONCLUSION

We investigate the role of traffic and transport energy consumption in the release of TCO₂ and examine the effect of environmental expenditures and green transport of OECD countries from 2000 to 2020. We employ the second-generation empirical tools by Pesaran. (2004) and Pesaran and Yamagata (2008) to check the CD and heterogeneity, respectively. We also employ the unit root tests of CIPS and CADF by Pesaran. (2007). To examine the long-run equilibrium association among

the variables, we adopt the Westerlund (2007) cointegration technique.

We find a CD issue presence in the dataset, and the model is suffering from slope heterogeneity. Consequently, the long-run relationship persists between the variables suggested by the cointegration technique. The findings from the CS-ARDL test show that traffic movement (i.e. road and railway) escalates TCO₂. Extensive use of road and railway locomotives augments CO₂ emissions. The findings further reveal that transport energy consumption has a remarkable effect on CO₂ emissions, which means that energy consumption in the transport sector is a driving factor of CO₂ emissions. The joint effect of traffic locomotives (road and railway) and transport energy consumption substantially effects TCO₂ as well.

Additionally, environmental government budget greatly diminishes TCO₂ through the R&D channel, which stimulates the innovations related to environmental mitigation. R&D expenditures generate innovations or new technology. On the contrary, road and railway vehicles (green transport) substantially degrade TCO₂. More precisely, electric vehicles are a better alternative strategy to mitigate TCO₂ in OECD countries.

Lastly, the causality results (D&H) portray that any policy that targets road movement and environmental taxes significantly changes TCO₂ and vice versa. Any policy related to transport energy consumption, road locomotives, and environmental government budget remarkably change TCO₂. On the contrary, any supporting policy to TCO₂ does not affect road locomotives and environmental government budget.

POLICY IMPLICATIONS

After concluding the findings, this study recommends the following policy implications for transport experts/economists, urban planners, and transport modelers. First, the results reveal that policymakers in OECD countries should take radical steps to moderate the traffic movement, particularly road movements, to

curb the deterioration of the environment. The movement of vehicles on the road must be controlled by transport institutes. The movement strategy can reduce CO₂ emissions approximately 18.3 and 11% for road and rail movement respectively (ref. **Table 6**). Second, government bodies or transport experts/policymakers should reduce 78.3 and 19.3% road and rail locomotives to improve better environment. Third, the governments of the OECD countries should increase their budget by 12.1% on average related to the environment to control CO₂ emissions (ref. **Table 7**). A better expenditure strategy related to the environment can be devised to preserve nonrenewable resources. Fourth, green transport (railway and road electric vehicles) should be produced around 59.85 and 52.25% by the OECD countries respectively (ref. **Table 7**). Fifth, the policy related to environmental taxes and R&D expenditures must be planned for the long run to achieve desirable consequences. Sixth, urban planners must design the transport infrastructure at city level, where traffic movements could be to avoid the releasing of CO₂ emissions. Seventh, transport modelers can formulate a policy by considering the demand for and supply of products of green transport e.g., electric vehicles. Transport experts (producers) should reduce the production of rail and road locomotives, as these consume a greater amount of energy during movements. Furthermore, they should control traffic movements within the urban areas that cause the transport carbon emissions.

The limitations of this study are as follows: This study is limited to OECD countries. Only road and railway traffic movements, energy use, green transport, and environmental expenditures are considered to estimate the effect on TCO₂.

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More specifically, transport sector has other factors such as transport household expenditure, transport infrastructure stock, transport logistics performance. It can be reasoned that these factors have been entrenched in/by traffic that directly or indirectly cause environmental degradation. In addition, the current study focuses on the OECD countries and it would be appropriate to enlarge the investigation to other countries e.g., BRICS, OBOR, or emerging economies and non-OECD as well. This research can be extended in the future by including environmental technology, agglomeration of population, and institutional quality. In particular, investigating the effects of green financing, green transportation, and green technology including emerging economies can provide policy makers/practitioners with more specific insights.

DATA AVAILABILITY STATEMENT

Publicly available datasets were analyzed in this study. This data can be found here: <https://stats.oecd.org>.

AUTHOR CONTRIBUTIONS

ZH conceived, carried out the literature review, and drafted the manuscript. MK, XZ, CM and WZ conceived, designed, and coordinated the study, and contributed to and finalized the draft. All authors read and approved the final manuscript.

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