



# Probabilistic Seismic and Tsunami Hazard Analysis Conditioned on a Megathrust Rupture of the Cascadia Subduction Zone

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This paper presents a methodology for probabilistic hazard assessment for the multi-hazard seismic and tsunami phenomena [probabilistic seismic and tsunami hazard analysis (PSTHA)]. For this work, a full-rupture event along the Cascadia subduction zone is considered and the methodology is applied to a study area of Seaside, Oregon, which is located on the US Pacific Northwest coast. In this work, the annual exceedance probabilities (AEPs) of the tsunami intensity measures (*IMs*) are shown to be qualitatively dissimilar to the *IMs* of the seismic ground motion in the study area. Specifically, the spatial gradients for the tsunami *IM* are much stronger across the length scale of the study area owing to the physical differences of wave propagation and energy dissipation of the two mechanisms. Example results of probabilistic seismic hazard analysis and probabilistic tsunami hazard analysis are shown for three observation points in the study area of Seaside. For the seismic hazard, the joint mean annual rate of exceedance of *IMs* shows similar trends for the three observation points, even though for a given observation point there is a large scatter between two ground-motion *IMs* analyzed, which were peak ground acceleration (PGA) and spectral acceleration at a period of vibration of 0.3 s, i.e., PGA and  $S_a$  ( $T_1 = 0.3$  s). For the tsunami hazard, the joint AEP of maximum flow depth ( $h_{\max}$ ) and maximum momentum flux ( $(M_F)_{\max}$ ) shows a high correlation between the two *IMs* in the study area. The joint AEP at each of the three observation points follows a particular Froude number ( $Fr$ ) due to the local site-specific conditions rather than the distributions of fault slip distributions used to generate the scenarios that are the basis of the AEP maps developed. The joint probability distribution of  $h_{\max}$  and  $(M_F)_{\max}$  throughout the study region falls between  $0.1 \leq Fr < 1.0$  (i.e., the flow is subcritical), regardless of return interval (500-, 1,000-, and 2,500-year). However, the peak of the joint probability distribution with respect to  $h_{\max}$  and  $(M_F)_{\max}$  varies with the return interval, and the largest values of  $h_{\max}$  and  $(M_F)_{\max}$  were observed with the highest return intervals (2,500 years) as would be expected. The results of the PSTHA can be the basis for a probabilistic multi-hazard damage and loss assessment and help to evaluate the uncertainties of the multi-hazard assessments.

**Keywords:** seismic hazard analysis, tsunami hazard analysis, multi-hazard risk, Cascadia subduction zone, community resilience

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## INTRODUCTION

Over the past decade and a half, megathrust earthquakes accompanied by near-field tsunamis have devastated coastal regions throughout the world, including the 2004 Indian Ocean tsunami (e.g., Jaffe et al., 2006; Rossetto et al., 2007), events in Chile in 2010 (e.g., Mas et al., 2012), and the 2011 Tohoku tsunami (e.g., Mori et al., 2013). These events remind us that when assessing for life safety, it is often desirable to plan for a “worst case” or “most credible” scenario. However, when considering damage to the built environment, it is often more practical to employ risk-informed decision-making to help minimize the overall damage, annualized financial loss, and increase the rate of recovery. In other words, risk-based decisions can increase the overall resilience of a community to earthquake and tsunami events. However, risk-based methods require a probabilistic understanding of the hazard. The case of subduction zone (SZ) earthquakes and tsunamis is particularly challenging because of the multi-hazard phenomena. A large building located in the megathrust earthquake region and in a potentially tsunami prone zone, for example, will first experience intense ground shaking followed by the subsequent hydrodynamic demands imposed by the tsunami. Liquefaction, local scour, landslides, debris, and other cascading consequences further exacerbate the problem. This paper marks one of the first attempts to provide a methodology for conducting a joint hazard analysis, which is termed as probabilistic seismic and tsunami hazard analysis (PSTHA), by combining probabilistic seismic hazard analysis (PSHA) with probabilistic tsunami hazard analysis (PTHA) based on a consistent process for concurrent earthquake occurrence and tsunami generation. As an illustrative example, the PSTHA is applied to a coastal community based on a conditional rupture of the Cascadia subduction zone (CSZ) along the northwest coast of North America.

## Background and Literature Review

Probabilistic seismic hazard analysis provides the evaluation of annual frequencies of exceedance of ground motion intensity measures (*IMs*) [typically designated by peak ground acceleration (PGA) or by linear elastic damped response spectral ordinates at specific periods of vibration] at a site. The result of a PSHA is a seismic hazard curve [annual frequency of exceedance versus ground-motion intensity measure (*IM*) amplitude], a uniform hazard spectrum (spectral amplitude versus structural period, for a fixed annual frequency of exceedance), or conditional mean spectrum (Baker, 2011; Lin et al., 2013). First reports of PSHA date back to the 1960s. Since then, PSHA has become the basis for seismic assessment and design of new and existing engineered facilities ranging from civil structures, such as buildings and bridges, to critical facilities, such as nuclear power plants. In PSHA, all possible earthquake fault sources contributing to the hazard need to be characterized first. Second, ground-motion prediction equations (GMPEs) are used to relate ground-motion *IMs* to variables describing earthquake source, path, and site effects. Extensive research has been performed on GMPEs for use in PSHA. Douglas (2003, 2011, 2016) summarized over 400 GMPEs that were developed since 1964–2016 for estimation of PGA and over 250 GMPEs for estimation of spectral ordinates at a site. Douglas and Edwards (2016) provides a recent discussion of

current and future trends in ground-motion prediction. Stewart et al. (2015) provides a discussion of the selection of GMPEs for hazard assessments for the three principal tectonic regimes: active crustal regions, SZs, and stable continental regions for a global earthquake model. Of interest to this paper, Stewart et al. (2015) recommended the use of three models for SZ ground-motion predictions, “BC Hydro” model of Abrahamson et al. (2016), the global earthquake model described in Atkinson and Boore (2003), and the model in the study by Zhao et al. (2006). The BC Hydro model was developed using different data sets of SZ strong-motion recordings (e.g., Crouse et al., 1988; Crouse, 1991; Youngs et al., 1997; Atkinson and Boore, 2003, 2008; Zhao et al., 2006; Lin and Lee, 2008). While GMPEs are most often used for PSHA, it is worth noting that other methods for generating the *IMs* that involve the generation of synthetic ground motions have also been recently proposed, however, all involving extremely computational intensive methods. These methods that could serve as alternatives to the GMPEs include kinematic earthquake models (e.g., Olsen et al., 2008; Frankel et al., 2014; Pulido et al., 2015; Iwaki et al., 2016), stochastic finite-fault ground-motion methods (e.g., Atkinson et al., 2009), or hybrid broadband ground-motion methods (e.g., Atkinson et al., 2011; Skarlatoudis et al., 2015).

Relative to PSHA, only recently has there been much work on PTHA. Mori et al. (submitted)<sup>1</sup> summarized 30 PTHA studies conducted worldwide, all but one of which were conducted after the 2004 Indian Ocean tsunami. Prior to the 2004 event, tsunami hazards were generally characterized by “worst credible” or “worst case” scenarios. Rikitake and Aida (1988) were the first to use historical records of past run-up events to characterize the probability of the tsunami run-up hazard in Japan. Subsequent to the 2004 event and with the more recent SZ tsunami events in Chile in 2010 and in Japan in 2011, there has been an increasing interest in developing PTHA. Studies have focused on regions throughout the Pacific Rim, including Japan (Burroughs and Tebbens, 2005; Annaka et al., 2007; Yanagisawa et al., 2007; Fukutani et al., 2015; and Goda and Song, 2016), the US Pacific Coast and Canada (Geist and Parsons, 2006; González et al., 2009; Thio and Somerville, 2009; Priest et al., 2010; Witter et al., 2013; Leonard et al., 2014; and Park and Cox, 2016), South China Sea (Liu et al., 2007; Li et al., 2016), New Zealand, and Australia (Power et al., 2007, 2013; Burbidge et al., 2008; and Mueller et al., 2015), as well as places in Europe (Tinti et al., 2005; Grezio et al., 2010; Anita et al., 2012) and the Northwestern Indian Ocean (Thio et al., 2007; Heidarzadeh and Kijko, 2011). As explained in the study by see text footnote 1, the PTHA generally uses one of three approaches for the tsunami generation: (1) historical record approach, (2) logic-tree approach, and (3) random phase approach. Generally, the historical record of tsunamis is not sufficient to build a credible probabilistic model. Therefore, the second two approaches are favored. The logic-tree approach (e.g., Park and Cox, 2016) is based on combinations of slip conditions (e.g., magnitude, peak slip location, and slip distribution), and these combinations are given a weighting based on expert opinion, historical record, or

<sup>1</sup>Mori, N., Goda, K., and Cox, D. T. (submitted). “Recent progress in probabilistic tsunami hazard analysis (PTHA) for mega thrust subduction earthquakes,” in *Reconstruction and Restoration after the 2011 Japan Earthquake and Tsunami* (Springer).

equal weighting in some cases. For the random phase approach (e.g., Goda et al., 2014), the slip distributions are created by an assumed slip wavenumber spectrum using random phases. For this paper, the logic-tree approach is used since it is straightforward to combine PTHA with the PSHA. In general, the output of the PTHA focuses on the flow depth at the shoreline, the maximum extent of inundation for planning, the evaluation of annual exceedance probabilities (AEPs), or the estimation of the extent of damage through a probabilistic tsunami damage analysis (e.g., Wiebe and Cox, 2014; Park et al., 2017).

The authors are aware of only one study by De Risi and Goda, 2016 (DG16 hereafter), which considers the combined PSTHA. In that paper, the authors present a consistent method to account for a common rupture process. They use GMPEs based on magnitude and rupture distance to quantify the shaking intensity at a particular location subjected to a SZ earthquake, and then solve the shallow water wave equations for tsunami propagation and inundation from the source to the site of interest. The main output of their work is seismic hazard curves and tsunami hazard curves that represents the mean annual rate (MAR) of exceedance of a given *IM*. Their generalized framework was applied for assessing combined earthquake and tsunami hazard at a single location on the coast line of Sendai City based on the subduction fault plane in the Tohoku region of Japan.

## Study Objectives and Outline

Similar to the work of DG16, the objective of this work is to develop a consistent framework for a multi-hazard analysis considering large magnitude SZ earthquakes and a subsequent near-field tsunami. However, the study performed herein is developed for a different site with a different methodology to define earthquake and tsunami sources when compared to DG16. The long-term objective is to be able to use the PSTHA as the basis for a probabilistic multi-hazard damage assessment to quantify the separate contributions of seismicity and tsunami hazards in the estimation of damage to the built environment at a community scale or for design of specific infrastructure elements such as a critical facility. A general methodology is first presented in Section “Methodology,” with a general review of the combined probabilistic hazard analysis (PHA) methodology proposed (see General PHA Methodology), the earthquake fault source models and their characteristics (see Earthquake Fault Source Models and Their Characteristics), the GMPEs used (see Earthquake Simulation), and the tsunami generation, propagation, and inundation (see Tsunami Generation, Propagation, and Inundation). Then, the multi-hazard logic-tree model is presented (see Multi-Hazard Logic-Tree Model), and the methods to estimate the AEP (see Estimates of the AEP) are presented next. In Section “Application: Full-Rupture Event of the CSZ Impacting the City of Seaside, OR,” the methodology is applied to an application example in which the combined seismic and tsunami hazard are quantified for the City of Seaside on the Oregon Coast of the US Pacific Northwest, conditional on a full rupture of the CSZ. In this application, the CSZ earthquake source model, and the geological, geographic, and morphological features of the City of Seaside, Oregon, are characterized first (see Characterization of the Source) and then details of the CSZ fault model and tsunami model (see CSZ Fault

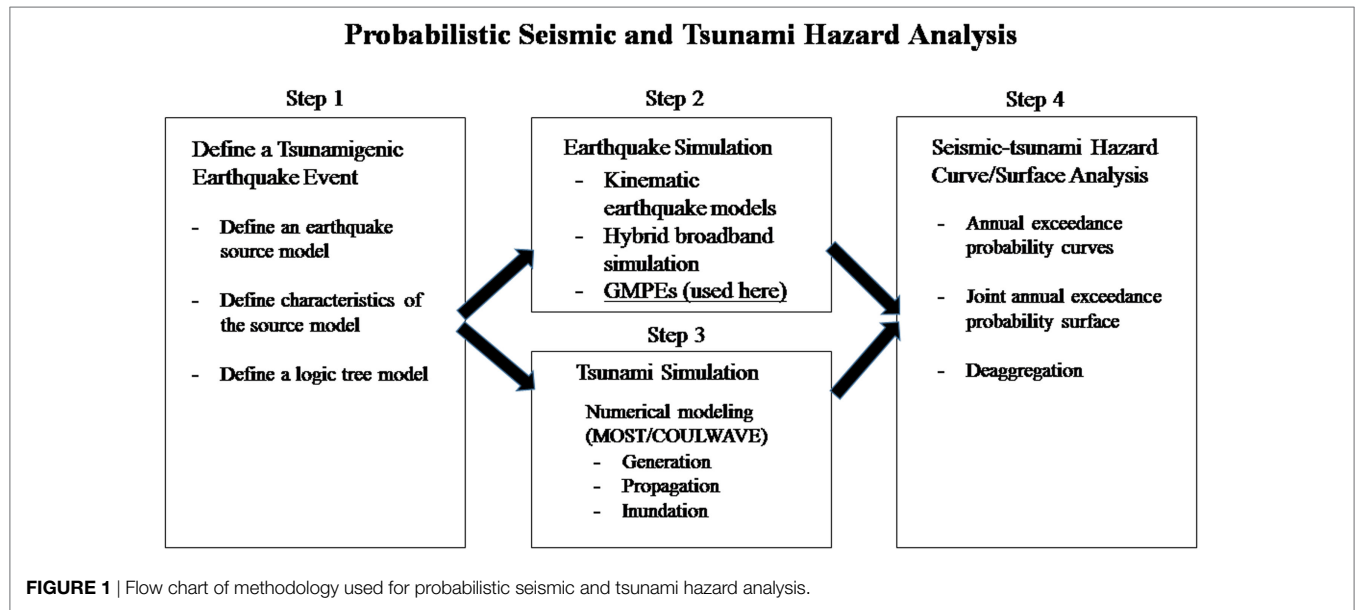
Modeling) are provided. In Section “Results,” results are presented in terms of the seismicity (see Seismicity) and tsunami intensity (see Tsunami Intensity) estimated for the area of interest. A spatial representation of both the seismic and tsunami intensity is presented, including the study of the granularity needed to characterize the hazards and joint distribution for different vector-valued *IMs* (see Spatial Representations Seismic and Tsunami Hazard). Finally, in Section “Summary, Conclusion, and Future Work,” discussion and summary of the results are provided, followed by the conclusion and recommendations for future research.

## METHODOLOGY

The methodology for consistent PSTHA is presented in this section. The methodology behind PSHA is well known (e.g., McGuire, 2004). However, a few adaptations are required for performing the combined PSTHA. The statistical earthquake model behind PSTHA is the same for PSHA and PTHA, although in PSHA the earthquakes generated solely inland do not contribute to the tsunami hazard. Even though the statistical earthquake model may be the same, the methods used in PSHA and PTHA for propagation of the effects of a fault slip to a specific site of interest vary. For PSHA, GMPEs are most often used. The GMPEs relate a certain moment release on an earthquake source (a line source or an area source) and the source-to-site rupture distance (or in a few instances, the hypocentral distance) to the ground-motion *IMs*. For the tsunami hazards, and due to the way tsunamis are generated, instead of using equations analogous to GMPEs, the conscientious decision is made herein of using the waveform excitation and propagation approach by solving non-linear shallow water equations [e.g., the “Method of Splitting Tsunami” (MOST) model, Titov et al., 2011] in which an earthquake, transoceanic propagation, and inundation of dry land are modeled. This type of detailed modeling of the tsunami propagation is analogous to computationally intensive recent trends in PSHA, which involves the use of methods for generating ensembles of synthetic ground-motion time histories as described above.

## General PHA Methodology

The general formulation presented here is aimed at developing the probabilistic earthquake and tsunami hazard at a site. In this formulation, the first topic that has to be addressed is the definition of an *IM* for the hazard. For example, with respect to earthquake ground shaking, typical measures of interest are spectral acceleration ( $S_a$ ), while common *IMs* used for the tsunami hazard are the inundation flow depth ( $h$ ), flow velocity ( $V$ ), and specific momentum flux ( $M_F = hV^2$ ). Since there is great uncertainty in the quantification of the earthquake that induces ground shaking and possible tsunamis, be it in the focal mechanism including definition of the location (e.g., location of the epicenter, extension of fault rupture), size (magnitude), and resulting intensity of a future earthquake at a specific site of interest, PSHA (e.g., Cornell, 1968; McGuire, 1995; Kramer, 1996) was developed as an analytical tool to characterize the seismic hazard probabilistically. PSHA has become the most widely used method for assessing seismic hazard at a specific site. More recently, PTHA has also been developed (PTHA—e.g., Geist and Parsons, 2006; Annaka et al., 2007; Power



et al., 2007) as reviewed by see text footnote 1 and summarized in Section “Background and Literature Review.”

The general PHA provides the  $MAR(\lambda)$  of  $IM$  exceeding an intensity measure value  $im$  that is computed using the Total Probability Theorem (Benjamin and Cornell, 1970), by integrating the contributions of all possible tsunami-seismogenic sources, and for each of the sources, all possible values of earthquake magnitude as

$$\lambda_{IM>im}(im) = \sum_{i=1}^{N_{\text{sources}}} \lambda_i(M \geq m_{\min}) \times \int_{\Theta} \int_{m_{\min}}^{m_{\max}} P(IM > im | \Theta, M) s_{\Theta_i}(\Theta | M) f_{M_i}(m) dm d\Theta \quad (1)$$

where  $N_{\text{sources}}$  denotes the total number of seismic sources contributing to the hazard at the site,  $\lambda_i(M \geq m_{\min})$  is the MAR of occurrence of earthquakes with magnitude greater than a lower bound threshold value,  $m_{\min}$ , of seismic source  $i$ ,  $P(IM > im | \Theta, M)$  represents the probability that intensity measure  $IM$  will exceed a given intensity value  $im$  at the site conditional on a given magnitude  $M$  and source parameters  $\Theta$ ,  $s_{\Theta_i}(\Theta | M)$  represents the characteristic probability density function (PDF) of earthquake source parameters  $\Theta$  obtained from a hazard parameter prediction model conditional on the magnitude of the earthquake, and the function  $f_{M_i}(m)$  denotes the PDF of the magnitude  $M$  given the occurrence of an earthquake on seismic source  $i$ .

In Eq. 1, it is assumed that earthquake occurrences at different seismic sources are statistically independent (in terms of occurrence time, scaling relationship,  $M$ , etc.), implying that earthquake occurrences from all possible sources can be assumed to follow a Poisson process. It is also assumed that within each seismic source  $i$ , the magnitude  $M_i$  earthquake events are statistically independent. Thus, the summation in Eq. 1 considers the contributions from all seismic sources while the integrations over  $M_i$  and source parameters account for earthquakes of all possible magnitudes

and source parameters conditional on the magnitude for each seismic source, respectively. In Eq. 1, the conditional probability  $P(IM > im | \Theta, M)$ , with  $im > 0$ , represents the complementary cumulative distribution function (CCDF) of the  $IM$  conditional on  $M$  and  $\Theta$ . **Figure 1** illustrates the steps that can be used to compute the probabilistic seismic and tsunami hazard relationships. In Step 1, the earthquake fault source models and characteristics of the earthquake source models are defined. In Step 2, the ground-motion  $IM$ s at a given site are obtained through the earthquake simulation, performed through either explicit source-to-site wave propagation models and synthetic ground-motion generation or, most commonly, with GMPEs. In Step 3, the tsunami is simulated, including tsunami generation, propagation and inundation modeling. In Step 4, the seismic-tsunami hazard curves and surfaces are determined. It is notable that Step 2 (seismic simulations) and Step 3 (tsunami simulations) are produced from a consistent earthquake event.

## Earthquake Fault Source Models and Their Characteristics

To describe the distribution of earthquake magnitudes in a given region of interest, the Gutenberg–Richter relationship (GR) is widely adopted. The GR relationship is given by

$$\log \lambda_m = a - bm \quad (2)$$

where  $\lambda_m$  is the MAR of exceedance of an earthquake of magnitude  $m$ ,  $a$  represents the overall rate of earthquakes in a region of interest, and  $b$  represents the relative ratio of small and large magnitudes. The parameters  $a$  and  $b$  are estimated based on statistical analysis of the database of seismicity for the seismic source zone of interest. The GR relationship is developed from a regional dataset of seismicity accounting for many different source zones and has been found to be inadequate to represent the earthquake recurrence relationship for the tail-end of the magnitude–frequency distribution representing large magnitude earthquakes (Schwartz

and Coppersmith, 1984; Wesnousky, 1994). Several models have been proposed to address the shortcomings of the GR recurrence law (Kagan, 1997, 2002a; Bird and Kagan, 2004). For SZs, such as the CSZ used as an example in this paper, a Tapered Gutenberg–Richter (TGR) distribution has been shown to be a robust model (Rong et al., 2014). The TGR is expressed as a function of the seismic moment  $M_0$ , instead of magnitude  $m$  and an exponential taper is applied to the number of events with very large seismic moment. The TGR CCDF is given by Kagan (2002a)

$$F(M_0) = \left(\frac{M_{0t}}{M_0}\right)^\beta \exp\left(-\frac{M_{0t} - M_0}{M_{0c}}\right), \text{ for } M_{0t} \leq M_0 < \infty \quad (3)$$

where  $\beta$  is the index parameter of the distribution, and  $\beta = (2/3)b$ ,  $M_{0c}$  is the corner moment, and  $M_{0t}$  is the threshold moment above which the earthquake catalog is assumed to be complete. The conversion between seismic moment  $M_0$  and moment magnitude  $m$  is given by

$$M_0 = 10^{1.5m+C} \quad (4)$$

where  $C = 9 - 9.1$ . The corner moment  $M_{0c}$  can be estimated using the seismic moment conversion principle (Kagan, 2002b) and given approximately by

$$M_{0c} = \left[ \frac{\chi \dot{M}_T (1 - \beta)}{\alpha_t M_{0t}^\beta \Gamma(2 - \beta)} \right]^{1/(1-\beta)} \quad (5)$$

where  $\chi$  is the seismic coupling coefficient,  $\dot{M}_T$  is the tectonic moment rate,  $\alpha_t$  is the recurrence rate for earthquake with moment  $M_{0t}$  and greater, and  $\Gamma$  is the gamma function. Equation 5 can be used to derive the recurrence interval as

$$T(M_{0t}) = \left[ \frac{1}{1 - \beta} \right] \frac{M_{0t}^\beta M_{0c}^{1-\beta}}{\dot{M}_T} \Gamma(2 - \beta) \exp\left(-\frac{M_{0t}}{M_{0c}}\right) \quad (6)$$

The TGR CCDF can be rewritten in terms of magnitude  $m$  and given as

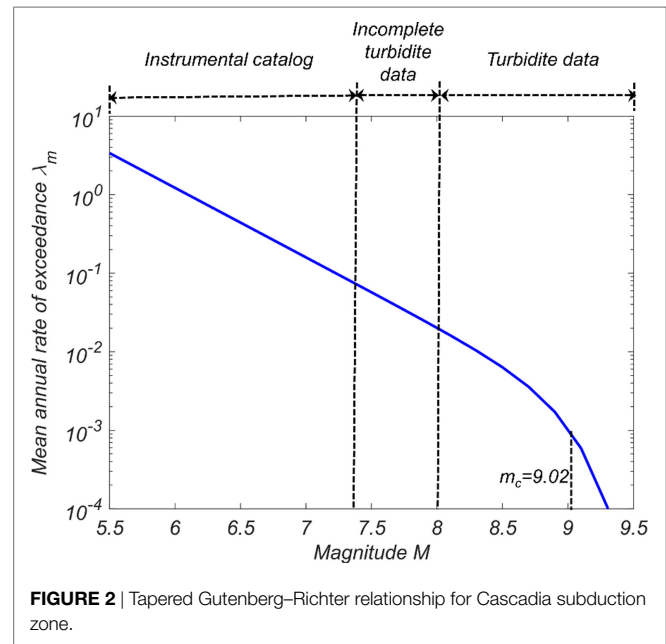
$$F(m) = \left[ 10^{1.5(m_t - m)} \right]^\beta \exp \left[ 10^{1.5(m_t - m_c)} - 10^{1.5(m - m_c)} \right] \quad (7)$$

Rong et al. (2014) estimated the probable maximum earthquake that is likely to occur within a given time for circum-Pacific SZs using TGR distributions. For the CSZ, using maximum likelihood estimation method, the values of  $\beta = 0.59$  and  $m_c = 9.02$  were estimated considering the 10,000-year paleoseismic record based on the turbidite studies by Goldfinger et al. (2012), coupled with the limited number of instrumental earthquake data. In the implementation performed, it is convenient to convert the continuous distribution of magnitudes into a discrete set of possible magnitudes  $m_j$ , which are given by

$$P[M = m_j] = G(m_j + 0.5\Delta m) - G(m_j - 0.5\Delta m) \quad (8)$$

where the  $G(m) = 1 - F(m)$  is the cumulative density function and  $\Delta m$  is the adopted discretization interval.

Figure 2 shows the median earthquake recurrence relationship developed for the CSZ using a TGR distribution (Eqs 3–7) and



with  $\beta = 0.59$  and  $m_c = 9.02$ , as recommended by Rong et al. (2014) for the CSZ. According to the TGR distribution shown in Figure 2,  $m \geq 8.8$  earthquakes are expected with a return period of 500 years ( $\lambda_m = 0.002$ ), while  $m \geq 9.0$  earthquakes are expected with a return period of 1,000 years ( $\lambda_m = 0.001$ ). Goldfinger et al. (2012) reconstructed the large earthquake history of the CSZ for approximately 10,000 years based on strong shaking-induced turbidite deposits in marine sediments and onshore paleoseismic records. The study suggested four types of earthquake rupture along the CSZ based on the interpretation of the turbidite data during the past 10,000 years: (1) 19–20 full-margin or nearly full-margin ruptures, (2) 3–4 ruptures along the 50–70% of the southern margins, (3) 10–12 southern ruptures from central Oregon southward, and (4) 7–8 southern Oregon/northern California ruptures. Though the turbidite data do not provide direct indication of the probable earthquake magnitudes, Goldfinger et al. (2012) estimated the earthquake magnitudes of different rupture events based on the relations observed among the rupture length (distance between offshore core sites containing turbidites from same events), turbidite thickness, and turbidite mass and estimated that full-rupture events constituted  $m = 8.7\text{--}9.3$ . Considering 20 full rupture over the past 10,000 years, the MAR of full-rupture events  $\lambda_{full-rupture} \approx \frac{20}{10,000} = 0.002$ , which is consistent with the  $\lambda_{m \geq 8.8} = 0.002$  estimated by Rong et al. (2014) using TGR distribution shown in Figure 2.

### Earthquake Simulation

Near-field ground motions are strongly affected by the heterogeneity of earthquake rupture processes, such as slip distribution, rupture directivity, and the acceleration and deceleration of the rupture front. To estimate the ground motion quantitatively for a seismic hazard assessment, characterization of this heterogeneity is essential, and usually accounted for in ground-motion hybrid broadband simulation procedures (e.g., Somerville et al., 2012) or 3D simulations of earthquake event scenarios (e.g., Olsen et al.,

2008; Delorey et al., 2014). While these explicit source-to-site wave propagation models and synthetic ground-motion generation models are extremely useful, especially in generating synthetic waveforms, these typically involve very large number of computations in large parallel computing centers, which is currently still not feasible when performing probabilistic seismic hazard analyses. Instead, when performing PSHA, GMPEs are still typically used today.

In this study, GMPEs, which are dependent on the local and regional site conditions, are used to simulate the ground-motion *IMs*. The recently developed Abrahamson et al. (2016) GMPE, which is based on the global datasets on SZ earthquakes described in Section “Background and Literature Review” is used. The functional form of this GMPE for interface SZ earthquakes is given by:

$$\ln(Sa_{\text{interface}}) = \theta_1 + \theta_4 \Delta C_1 + (\theta_2 + \theta_3 (M - 7.8)) \ln(R_{\text{rup}}) + C_4 \exp(\theta_9 (M - 6)) + \theta_6 R_{\text{rup}} + f_{\text{MAG}}(M) + f_{\text{FABA}}(R_{\text{rup}}) + f_{\text{site}}(\text{PGA}_{1,000}, V_{S30}) + \sigma \epsilon \quad (9)$$

where  $\theta_i$  are regression parameters,  $R_{\text{rup}}$  is the closest distance to the rupture area from site,  $V_{S30}$  is the shear wave velocity of the uppermost 30 m of soil,  $\text{PGA}_{1,000}$  is the median PGA corresponding to  $V_{S30} = 1,000$  m/s,  $\sigma$  is the total SD, which is obtained combining intra-event uncertainty  $\phi$  and inter-event uncertainty  $\tau$ , and  $\epsilon$  is the standard normal error term. In this study, only the intra-event uncertainty  $\phi$  is considered since only the CSZ is used for hazard analysis. The magnitude scaling term is given by

$$f_{\text{MAG}}(M) = \begin{cases} \theta_4 (M - (C_1 + \Delta C_1)) + \theta_{13} (10 - M)^2, & \text{for } M \leq C_1 + \Delta C_1 \\ \theta_5 (M - (C_1 + \Delta C_1)) + \theta_{13} (10 - M)^2, & \text{for } M > C_1 + \Delta C_1 \end{cases} \quad (10)$$

where  $C_1 = 7.8$ , and the  $\Delta C_1$  term represents the epistemic uncertainty around the distinct break in magnitude scaling between frequent smaller magnitudes events and rare large interface events. The forearc and backarc scaling term in Eq. 9 is given by

$$f_{\text{FABA}}(M) = \left[ \theta_{15} + \theta_{16} \ln \left( \frac{\max(R_{\text{rup}}, 100)}{40} \right) \right] F_{\text{FABA}}; \quad F_{\text{FABA}} = \begin{cases} 0, & \text{for forearc or unknow sites;} \\ 1, & \text{for backarc sites.} \end{cases} \quad (11)$$

The model for site response scaling is given by

$$f_{\text{site}}(\text{PGA}_{1,000}, V_{S30}) = \begin{cases} \left[ \theta_{12} \ln \left( \frac{V_S^*}{V_{\text{lin}}} \right) - b \ln(\text{PGA}_{1,000} + c) \right] & \text{for } V_{S30} < V_{\text{lin}} \\ + b \ln \left( \text{PGA}_{1,000} + c \left( \frac{V_S^*}{V_{\text{lin}}} \right)^n \right), & \\ \left[ \theta_{12} \ln \left( \frac{V_S^*}{V_{\text{lin}}} \right) + b \ln \left( \frac{V_S^*}{V_{\text{lin}}} \right) \right], & \text{for } V_{S30} \geq V_{\text{lin}} \end{cases} \quad (12)$$

$$V_S^* = \begin{cases} 1,000, & \text{for } V_{S30} > 1,000; \\ V_{S30}, & \text{for } V_{S30} \leq 1,000. \end{cases}$$

All other model coefficients in Eqs 9–12 are listed in the study by Abrahamson et al. (2016).

## Tsunami Generation, Propagation, and Inundation

In general terms, the tsunami hazard at a particular site requires the three steps of (1) tsunami generation, (2) propagation, and (3) inundation. For most modeling efforts, tsunami generation is given as an initial surface water displacement along the fault. The Okada (1985) model, which is based on the linear co-seismic dislocation of fault slips, is often used for simplicity, and the initial displacement is assumed to occur simultaneously along the fault.

Tsunami propagation is generally considered a solved problem in that the equations are well defined for long wave propagation in the open ocean. Of course, the propagation phase requires accurate knowledge of the underlying bathymetry which affects the wave through refraction, diffraction, and shoaling. The third phase, tsunami inundation, considers the flow of water over dry land. This is considered a difficult problem to solve because of the complex interaction of the flow with the built and natural environment that is also changing due to the destructive nature of the flow. However, most inundation models assume “bare earth” conditions, that is a digital elevation model (DEM) in which the natural and built environments are removed. Typically, the effects of the vegetation and structures are replaced by a suitable friction factor, although there are additional uncertainties in this step (e.g., Park et al., 2013; Bricker et al., 2015).

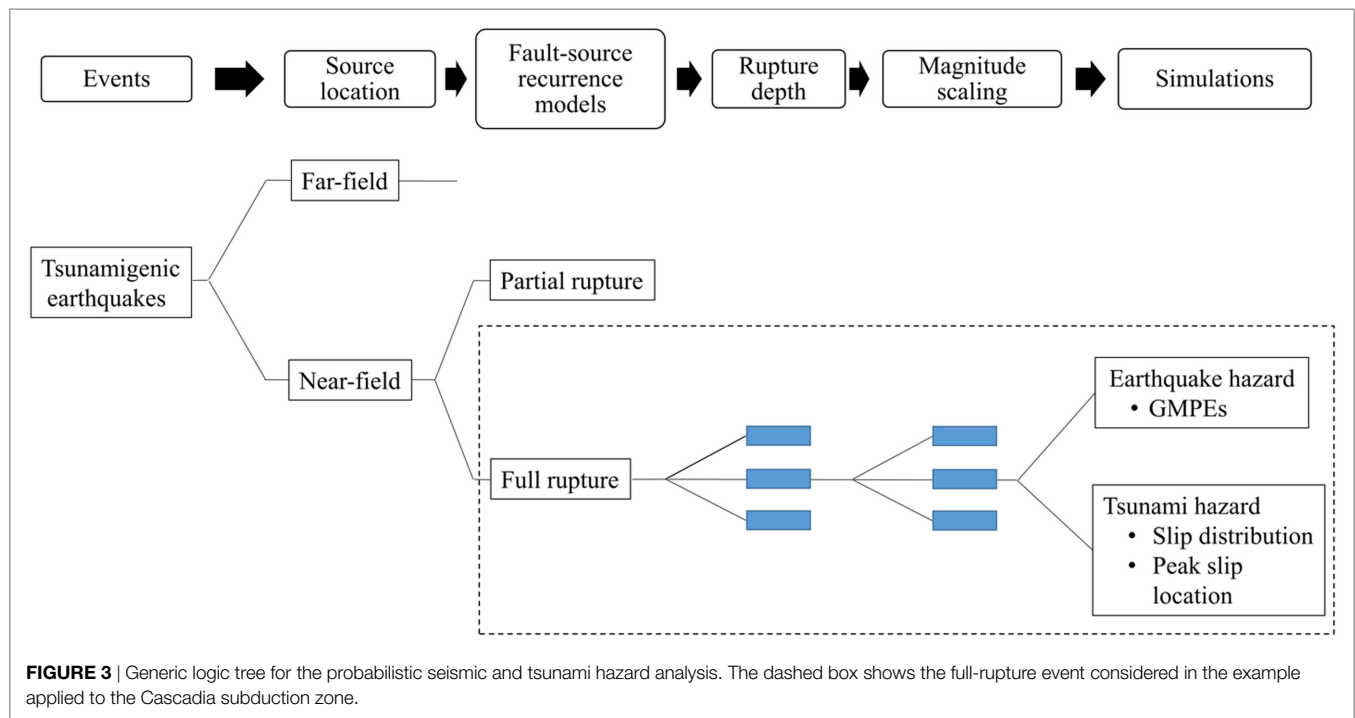
In this study, the logic-tree model by Park and Cox (2016) is applied for tsunami generation to characterize the fault slip. The slip model from the study by Park and Cox (2016) characterizes the randomness of fault slip distribution at the CSZ as a Gaussian shape, parameterized in terms of the moment magnitude, peak slip location, and a fault slip shape as

$$f(Y'/dL|\alpha, \beta) = \frac{1}{\beta\sqrt{2\pi}} \exp\left(-\frac{(Y'/dL - \alpha)^2}{2\beta^2}\right) \quad (13)$$

where  $\alpha$  and  $\beta$  are the slip distribution parameters along a rupture strike direction ( $Y'$ ), and  $dL$  is the unit length of sub-fault utilized in the slip model. Each parameter  $\alpha$  and  $\beta$  controls the location of the peak slip and the shape of slips. A total of 72 scenarios from the three seismic moments, three slip shapes, and eight peak slip locations are proposed for the full-length rupture CSZ event. The occurrence rate of each seismic moment estimated from the paleoseismic data at CSZ (Goldfinger et al., 2012). Each of the fault slip distributions are determined from the slip model and applied to the ComMIT/MOST model (Titov et al., 2011) as an input to simulate the tsunami generation and propagation parts. Although MOST can also be used for the inundation phase, the software COULWAVE (Lynett et al., 2002) is used to model this final step (Park and Cox, 2016).

## Multi-Hazard Logic-Tree Model

The generic multi-hazard logic tree is presented in Figure 3. The first step is to identify all tsunamigenic earthquakes that could potentially affect the site of interest. These earthquakes are then classified as near-field or far-field earthquakes depending on the source location with respect to the site. For each of the potential fault sources, the next step is to use an appropriate recurrence



model. As mentioned in Section “Earthquake Fault Source Models and Their Characteristics,” the widely used modified-GR underpredicts the recurrence rate for large magnitude tsunamigenic earthquakes and hence TGR (Kagan, 2002a) and other characteristic moment–frequency distributions (e.g., Wesnousky, 1994) are preferred. Since most of the GMPEs use a distance metric of closest distance from the site to the rupture plane ( $R_{rup}$ ), the depth of rupture which determines the position of the rupture edge close to the site of interest can have significant effect on the ground-motion intensity observed. Several geophysical models are available to determine the rupture depth and can be used for a specific fault of interest. As for example, for 2014 National Seismic Hazard Map of United States, three geophysical models were used to constrain the eastern edge of the CSZ rupture zone (Frankel et al., 2015).

Reliable magnitude scaling relationships for subduction earthquakes is a prerequisite for accurate estimation of earthquake and tsunami hazard intensities. There are several magnitude scaling relationships available in the literature that are derived from global observation of earthquakes on the plate interface of SZ (Papazachos et al., 2004; Strasser et al., 2010; Murotani et al., 2013; Goda et al., 2016; Skarlatoudis et al., 2016). Of these, Skarlatoudis et al. (2016) compiled an updated database of interface earthquakes that occurred worldwide in last decade in the major SZs (e.g., 2004 M 9.1 Sumatra, 2010 M 8.8 Chile, 2011 M 9.0 Japan earthquake) and proposed new and improved source scaling laws with reduced uncertainty compared to other currently available source scaling laws for subduction earthquakes. Once the appropriate source magnitude and scaling laws are identified, the final step is to perform the earthquake and tsunami simulations for a multiple logic-tree branch. In the case of earthquakes, GMPEs are utilized, which typically make use of the  $M$  and  $R_{rup}$  as the input parameter

to compute the ground-motion  $IM$ s at the site. On the other hand, for tsunamis, the details of the slip distribution (e.g., slip shape, peak slip location within rupture zones) and seismic moment for a given rupture are critical input parameters needed to compute the tsunami hazard intensity at a given site.

### Estimates of the AEP

Four  $IM$ s were selected for the PSTHA, including the PGA, the 5% damped linear elastic spectral acceleration at a fundamental period of vibration of 0.3 s [ $S_a(T_1) = 0.3$  s], the maximum flow depth ( $h_{max}$ ), and the specific maximum momentum flux ( $M_F$ )<sub>max</sub> because these  $IM$ s are often linked with structural damage at a structure scale (e.g., Faggella et al., 2013; Gidaris et al., 2016). Here, the flow depth is the net elevation of the free surface elevation above the local land elevation, and the specific momentum flux is given by the product of the flow depth with the square of the flow velocity ( $M_F = hV^2$ ).

The Poisson process (Cornell, 1968) is used to estimate the AEP of both earthquake and tsunami hazard curves at a specific location. The probability of  $IM$ s exceeding a certain level of that hazard conditional on a given time  $t$  is equal to the probability of at least one event occurring in time  $t$ , and is given by

$$P[IM > im|t] = 1 - e^{-\lambda t} \quad (14)$$

where  $\lambda$  is the mean occurrence rate at which the  $IM$  will exceed a specific  $im$  at a given location during the time  $t$ . The MAR exceedance of each  $IM$  is computed using Eq. 1.

The AEP surfaces for the joint seismic-tsunami hazard can be computed using the formulation presented next, designated here as vector-valued probabilistic seismic and tsunami hazard analysis (VPSTHA). The presentation starts from the definition

of the joint mean rate density (MRD) of the hazard (e.g., Bazurro, 1998; Barbosa, 2011), here expanded to account for  $IM$ s related to ground-motion shaking intensity as well as tsunami  $IM$ s. For a vector of ground-motion and/or tsunami intensity parameters  $\mathbf{IM} = \{IM_1, IM_2\}$ , the joint MRD of the hazard is given by

$$\begin{aligned} \text{MRD}_{IM_1, IM_2}(im_1, im_2) &= \sum_{i=1}^{N_{\text{sources}}} \lambda_i(M \geq m_{\text{min}}) \\ &\times \int_{\Theta} \int_M f_{IM_1, IM_2}(im_1, im_2 | \Theta, M) s_{\Theta_i}(\Theta | M) f_{M_i}(m) dm d\Theta \end{aligned} \quad (15)$$

where  $f_{IM_1, IM_2}(im_1, im_2 | \Theta, M)$  is the joint PDF of  $IM_1$  and  $IM_2$  conditional on earthquake of magnitude  $M$  and source parameters  $\Theta$ . Once the MRD is defined, the MAR of events at the site with  $IM_1$  and  $IM_2$  being between  $im_{1,1} < IM_1 \leq im_{1,2}$  and  $im_{2,1} < IM_2 \leq im_{2,2}$ , respectively, is given by:

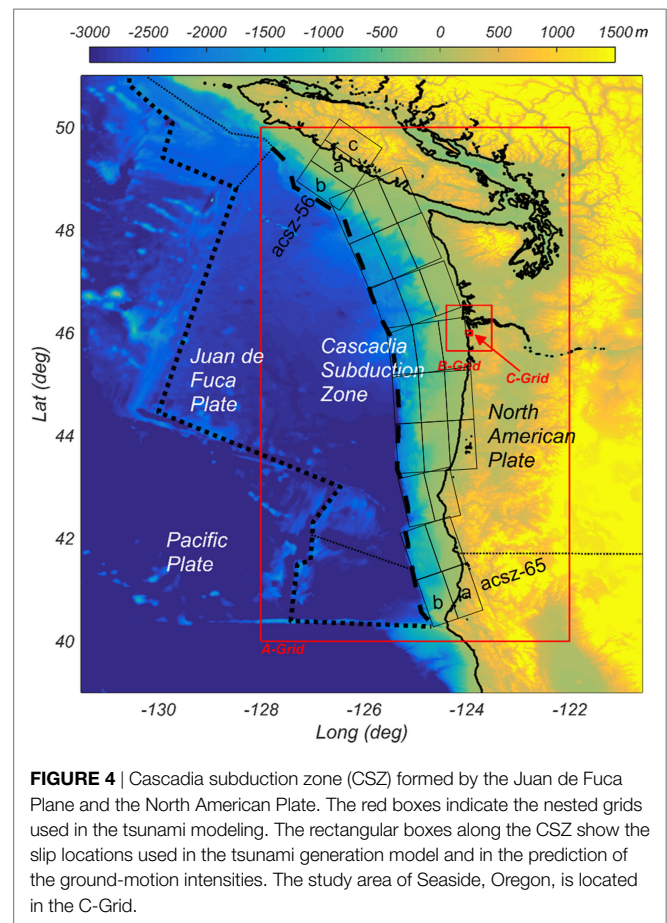
$$\lambda_{IM_1 \in [im_{1,1}, im_{1,2}], IM_2 \in [im_{2,1}, im_{2,2}]} = \int_{im_{1,1}}^{im_{1,2}} \int_{im_{2,1}}^{im_{2,2}} \text{MRD}(im_1, im_2) dim_2 dim_1 \quad (16)$$

It is worth highlighting that the vector of intensity parameters  $\mathbf{IM} = \{IM_1, IM_2\}$  may contain ground-motion scalar  $IM$ s and/or tsunami scalar  $IM$ s. Examples are vector-valued  $IM$ s such as  $\mathbf{IM} = \{S_a(T_1), h_{\text{max}}\}$ ,  $\mathbf{IM} = \{\text{PGA}, S_a(T_1)\}$ ,  $\mathbf{IM} = \{h_{\text{max}}, (M_F)_{\text{max}}\}$ . The formulation in Eqs 15 and 16 is generic, but its implementation requires the computation of joint MRDs, which can be computationally expensive.

## APPLICATION: FULL-RAPTURE EVENT OF THE CSZ IMPACTING THE CITY OF SEASIDE, OR

### Characterization of the Source

The northern coast of the North American continent from Northern California in the United States to Vancouver Island in Canada is facing the threat of a megathrust earthquake event with near-field tsunami from the CSZ along the converging plate boundary between Juan de Fuca Plate and North American Plate. The Juan de Fuca Plate is sinking beneath the North American Plate with the mean rate of 0.04 m/year (Heaton and Hartzell, 1987) to the northeast direction (Figure 4). The accumulated potential energy between two plates is released in the megathrust earthquake events and can cause ground shaking and rapid displacement of the seafloor which generates the initial deformation of surface water. Each megathrust rupture of the converging plate boundary triggers the earthquake and tsunami event in both offshore and onshore directions. The last full-rupture event occurred on January 26, 1700, with moment magnitude estimated between 8.7 and 9.2 (Satake et al., 2003). It was also reported that there were smaller but more frequent partial rupture events on the north or south margins of the CSZ (Atwater and Griggs, 2012; Goldfinger et al., 2012). However, in this application example of

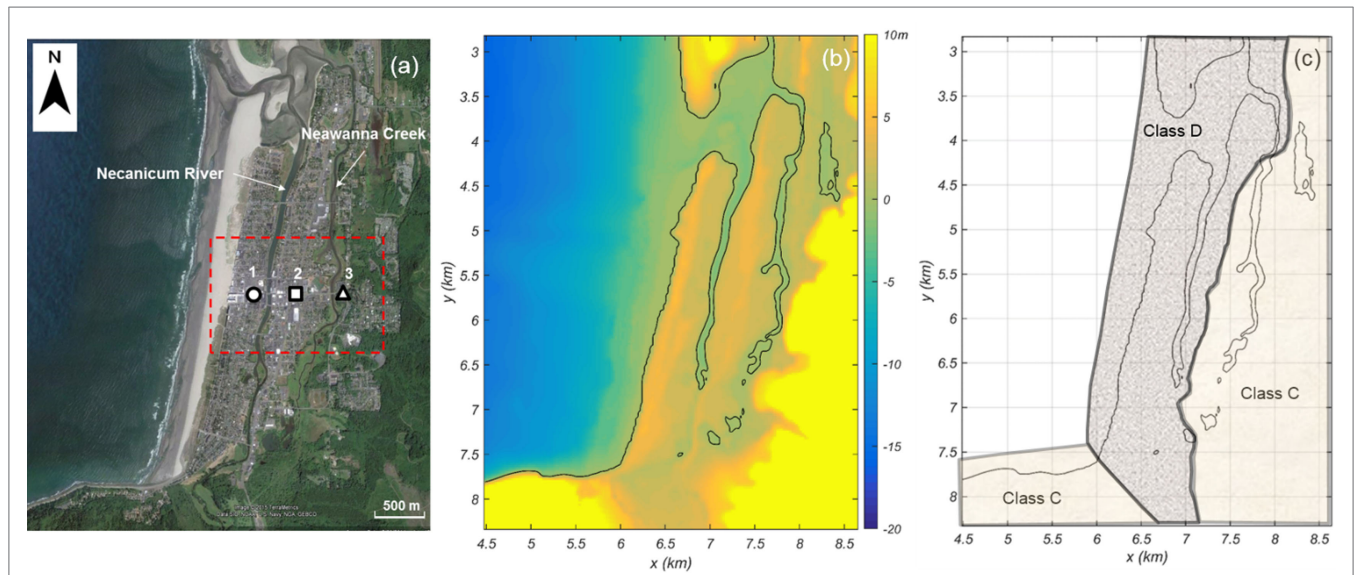


**FIGURE 4** | Cascadia subduction zone (CSZ) formed by the Juan de Fuca Plate and the North American Plate. The red boxes indicate the nested grids used in the tsunami modeling. The rectangular boxes along the CSZ show the slip locations used in the tsunami generation model and in the prediction of the ground-motion intensities. The study area of Seaside, Oregon, is located in the C-Grid.

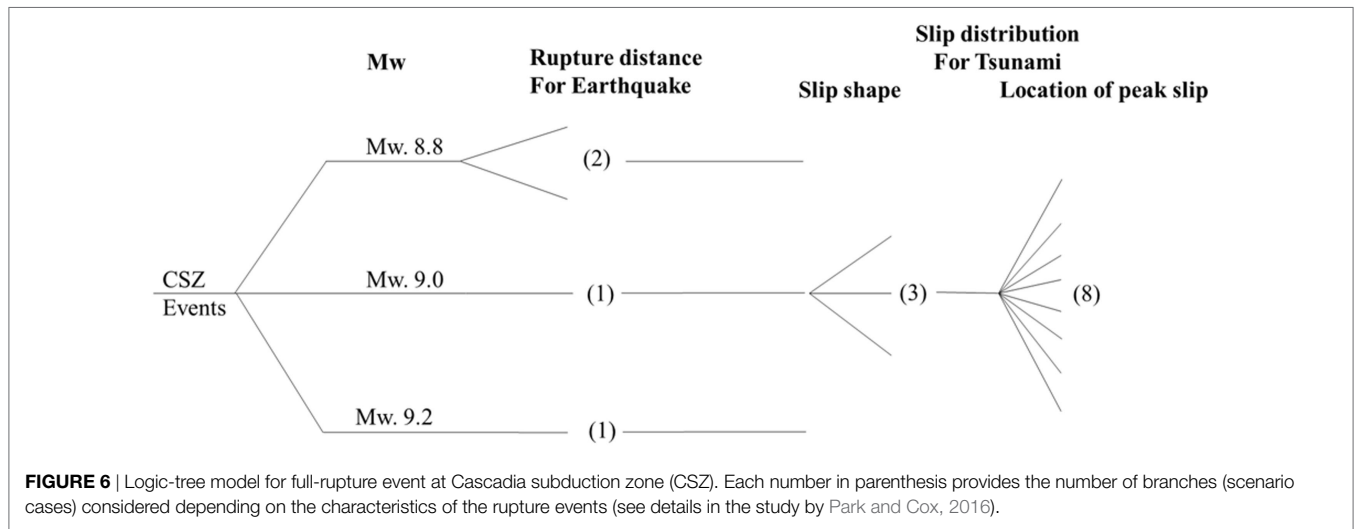
the proposed multi-hazard assessment, the partial rupture events are not considered to reduce the number of computations needed for the tsunami inundation.

Seaside, Oregon, is the study area chosen to demonstrate the proposed PSTHA methodology. Figure 5A shows an aerial view of Seaside with approximately 4 km of shoreline facing the Pacific Ocean and two small rivers, the Necanicum River and Neawanna Creek, that run parallel to the shoreline. The city has a population of approximately 6,500 residents, and the number of people in the city can increase to over 20,000 during the summer tourist season. There are over 5,700 buildings in Seaside with the larger hotels constructed out of steel and reinforced concrete in the center of the city (highlighted by the red box in Figure 5A) and surrounded by mostly wooden residential structures to the north and south of the city center. Figure 5B shows the bathymetry and topography of the study area. The remnant coastal dune on which the city was built can be seen running parallel to the shoreline with a secondary rise between the Necanicum River and Neawanna Creek. To the east of the Neawanna Creek, there is a steep gradient leading to the foot hills and a large headland to the south west. Figure 5C shows the distribution of the soil classes for this region. The steeper mountain areas are considered as Class C, and a majority of the city area is Class D (ASCE, 2010). The soil class distribution is also aligned approximately in the shore parallel direction.





**FIGURE 5** | Study area of Seaside, Oregon, at the landward side of the C-Grid shown in **Figure 4**. **(A)** Satellite image of the City of Seaside area with Necanicum River and Neawanna Creek bisecting the city. Areas marked by 1 (circle), 2 (square), and 3 (triangle) are example locations used in later figures. Red box highlights detailed area in later figures. **(B)** Bathymetry/topography for study area. Note that waterfront area of the city is built on a remnant dune and has higher elevation than river areas. **(C)** Soil classes assumed based on DOGAMI Oregon Department of Geology and Mineral Industries (2017) maps.



**FIGURE 6** | Logic-tree model for full-rupture event at Cascadia subduction zone (CSZ). Each number in parenthesis provides the number of branches (scenario cases) considered depending on the characteristics of the rupture events (see details in the study by Park and Cox, 2016).

Seaside has been the subject of multiple tsunami studies in the past because it is considered to be one of the most vulnerable locations to a CSZ event. As reported by Wood et al. (2010), the low lying city has approximately 87% of its land within the inundation zone of a CSZ full-rupture event, and 89% of the employees work within this zone. Because of the high vulnerability of Seaside to a CSZ event, there have been several recent studies in this area as reviewed by Park and Cox (2016) and Park et al. (2017), including the work by González et al. (2009) that conducted a PTHA using 14 historic tsunami far-field scenarios and 12 scenarios of the CSZ event.

### CSZ Fault Modeling

**Figure 6** shows the logic-tree model utilized for both earthquake and tsunami for CSZ full-rupture event (Park and Cox, 2016).

This logic-tree model characterizes the randomness of the fault slip in terms of the moment magnitude, peak slip location, and a fault slip shape. To simplify the scenarios, 27 sub-faults distributed along the CSZ were considered, which are the default sub-fault setup for the ComMIT model (Titov et al., 2011). The sub-faults are distributed from the northern Vancouver Island to Northern California. Each sub-fault is 100 km long by 50 km wide, as shown in **Figure 4**. The detailed geologic information of the location and slip information is summarized in **Table 1**. Each of the 27 sub-faults has a constant rake of 90° and has varied strike, dip, and depth conditions along the entire fault. The x and y coordinates indicate the middle point on the right edge boundary of each sub-faults in **Figure 4**.

The full-rupture event at the CSZ is discretized as three moment magnitude scenarios (*M* 8.8, 9.0, and 9.2), and the corresponding

**TABLE 1** | Geologic sub-faults information at Cascadia subduction zone.

Sub-faults	Longitude (°E)	Latitude (°N)	Strike (°)	Dip (°)	Depth (m)
acsz-56a	234.0588	49.1702	315	11	12.82
acsz-56b	233.5849	48.8584	315	9	5.00
acsz-56c	234.5230	49.4727	315	18	22.36
acsz-57a	234.9041	48.2596	341	11	12.82
acsz-57b	234.2797	48.1161	341	9	5.00
acsz-57c	235.5249	48.4018	341	14	22.36
acsz-58a	235.3021	47.3812	344	11	12.82
acsz-58b	234.6776	47.2597	344	9	5.00
acsz-58c	235.9298	47.5029	344	11	22.36
acsz-59a	235.6432	46.5082	345	11	12.82
acsz-59b	235.0257	46.3941	345	9	5.00
acsz-59c	236.2591	46.6216	345	13	22.36
acsz-60a	235.8640	45.5429	356	11	12.82
acsz-60b	235.2363	45.5121	356	9	5.00
acsz-60c	236.4860	45.5734	356	14	22.36
acsz-61a	235.9106	44.6227	359	11	12.82
acsz-61b	235.2913	44.6150	359	9	5.00
acsz-61c	236.5211	44.6303	359	15	22.36
acsz-62a	235.9229	43.7245	359	11	12.82
acsz-62b	235.3130	43.7168	359	9	5.00
acsz-62c	236.5213	43.7321	359	16	22.36
acsz-63a	236.0220	42.9020	350	11	12.82
acsz-63b	235.4300	42.8254	350	9	5.00
acsz-64a	235.9638	41.9818	345	11	12.82
acsz-64b	235.3919	41.8677	345	9	5.00
acsz-65a	236.2643	41.1141	345	11	12.82
acsz-65b	235.7000	41.0000	345	9	5.00

rupture areas ( $S$ ) were estimated using the relationship between seismic moment  $M_0$  and rupture area provided by Murotani et al. (2013) and given by

$$S = 1.34 \times 10^{-10} M_0^{2/3} \quad (17)$$

Park and Cox (2016) added upper and lower limits to the surface area moment magnitude relationships of Eq. 17 by multiplying Eq. 17 with 2.0 and 0.5, respectively, and these limits bounded most of historic recent tsunami events at SZ around the Pacific Ocean which are generated by earthquake scenarios with  $M > 8.5$ . Including those two limits, a total of three slip shapes along the strike direction were utilized as possible scenarios per moment magnitude condition. Additionally, 8 possible peak locations are considered along the rupture length, giving rise to 24 scenarios applied for each moment magnitude condition. A total of 72 scenarios (3 moment magnitudes  $\times$  3 slip shapes  $\times$  8 peak slip locations) are proposed here to characterize the full-rupture CSZ event. Park and Cox (2016) applied 72 scenarios as inputs to the ComMIT/MOST (Titov et al., 2011) model, which considers the elastic dislocation model (Okada, 1985) for tsunami generation and solves the non-linear shallow water equations implemented with a finite difference scheme. The ComMIT/MOST model consists of three nested grids (A, B, and C-Grid) shown in Figure 4. For the inundation modeling, the results of COULWAVE (Lynett et al., 2002) were used, which solves a set of Boussinesq equations with a high-order finite-volume method by using the output of ComMIT/MOST modeling at B-Grid as the input of COULWAVE at C-Grid. All bathymetry data were originated from NOAA's National Geophysical Data Center and the DEM for the Seaside

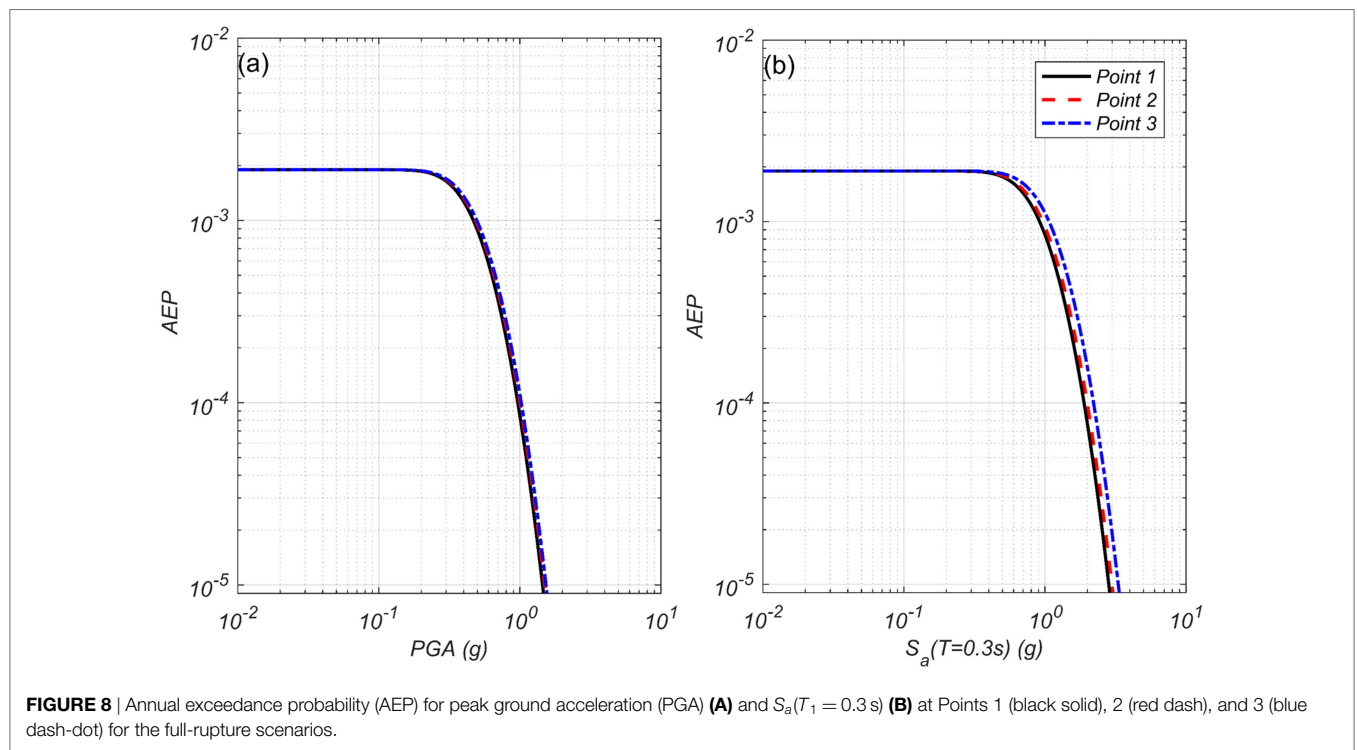
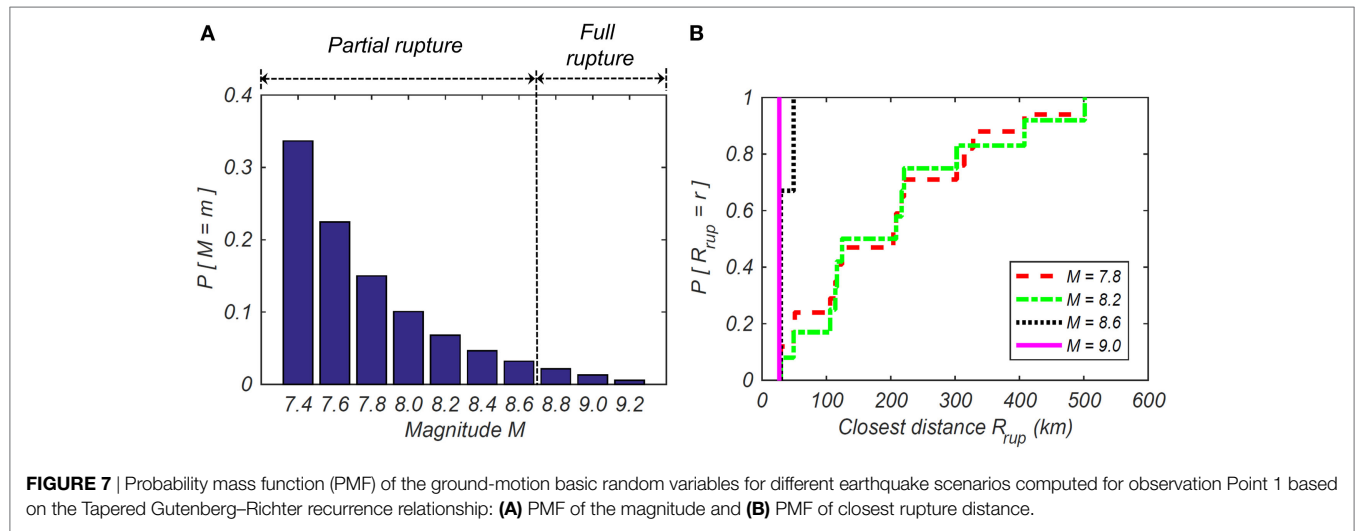
area whose resolution is 1/3 arc second was used for the C-Grid. Each A, B, and C-Grids has 1 min ( $400 \times 400$ ), 3 s ( $800 \times 800$ ), and 24 m ( $416 \times 390$ ) resolutions. As a tide condition, the mean high water level was fixed as conservative tsunami hazards estimation in this study. The model results provide surface elevation and velocity time series over the entire study area at the 24 m grid resolution. More details of fault slip distributions and tsunami simulations are available in the study by Park and Cox (2016). In case of earthquakes, the same fault model is used. However, the BC Hydro GMPE is utilized in this study, which is based on the closest rupture distance,  $R_{rup}$ , between the site of interest and rupture surface for the interface earthquakes considered.

## RESULTS

### Seismicity

Though larger magnitude earthquakes generally have higher damage potential compared to smaller magnitude earthquakes, smaller and frequent events can be dominant contributors to the structural damage risk measured in terms of mean annual frequency of exceeding a damage state such as the collapse damage state (e.g., Zareian and Krawinkler, 2007; Eads et al., 2013). Figure 7A shows the probability mass function (PMF) of the earthquakes in CSZ deemed capable of generating tsunamis. Tsunamigenic earthquakes with lowest magnitude of  $m_{min} = 7.3$  is assumed and largest magnitude is assumed to be  $m_{max} = 9.3$ , consistent with the earthquake expected over a 10,000-year period in CSZ (Rong et al., 2014). A discretization interval of  $\Delta m = 0.2$  is adopted to develop the PMF of Figure 7, resulting in 10 central magnitude values  $m = 7.4$  to  $m = 9.2$ . The PMF for different magnitudes are computed based on the TGR distribution described in Section "Earthquake Fault Source Models and Their Characteristics" and using Eq. 8. Of these 10 magnitudes,  $m \geq 8.8$  is assumed to produce the full rupture along the CSZ following the recommendations by Goldfinger et al. (2012), which is also consistent with Park and Cox (2016). These full-rupture events ( $m = 8.8, 9.0, 9.2$ ) are only considered for earthquake-tsunami hazard analysis for Seaside. According to Figure 7A,  $\lambda_{m \geq 8.8} = 0.002$  results in 500-year return period for  $m \geq 8.8$  representing full-rupture scenarios. To be consistent with 526-year return period used in Park and Cox (2016) for full-rupture CSZ events,  $\lambda_{m \geq 8.8} = 0.0019$  is used for the PSHA computations for the study area of City of Seaside.

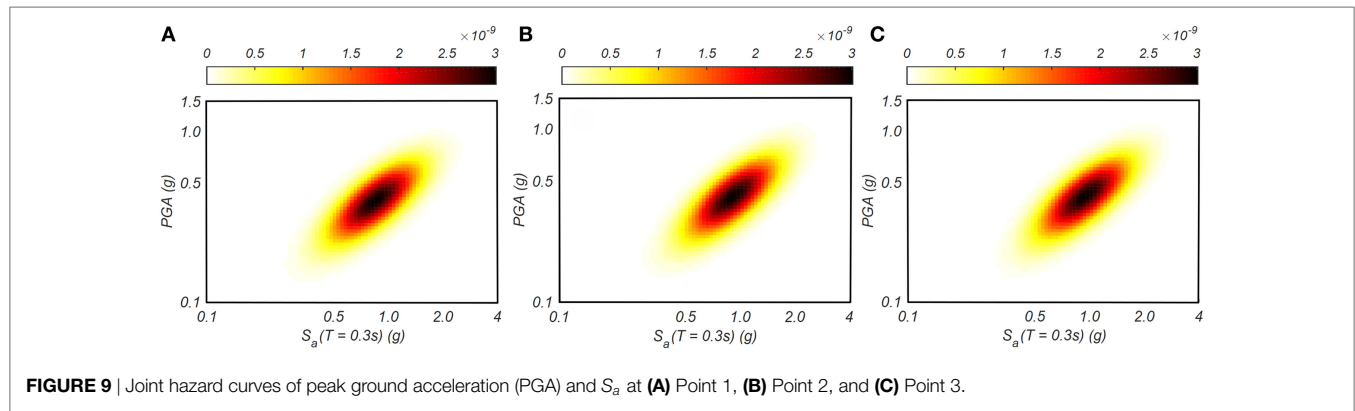
One of the inputs required in the GMPEs for computing ground-motion intensities at a given site is the source-to-site distance. The BC Hydro GMPE used in this study is based on the closest rupture distance  $R_{rup}$  between the site of interest and the rupture surface for interface SZ earthquakes. To compute the  $R_{rup}$  for different locations in Seaside for different magnitude scenarios, the  $R_{rup}$  for all the C-grid locations are precomputed from the 27 sub-fault areas shown in Figure 4. The rupture surface for different magnitude scenarios of Figure 7A are then computed based on the magnitude scaling law by Murotani et al. (2013), which are randomly positioned along the 27 sub-faults corresponding to the approximately 1,000 km long by 150 m width fault plane. The randomly positioned rupture surfaces corresponding to different scenarios are subsequently used to interpolate the  $R_{rup}$  for each



scenario from the precomputed closest distances. **Figure 7B** shows the PMF of  $R_{rup}$  computed for observation point 1 in **Figure 5A**. The closest distance computed for all the scenarios is 27.5 km, which is the distance computed from the fault area immediately below the C-grid. It can be observed in **Figure 7B** that smaller variability of  $R_{rup}$  is observed for larger magnitudes. For example, for  $m = 9.0$  that is a full-rupture event results in a single source-to-site distance and the resulting  $R_{rup}$  can be considered as a deterministic event with  $P[R_{rup} = 27.5 \text{ km}] = 1$ . For the partial rupture events, the rupture surface is randomly positioned within the fault plane, which results in variability in the closest rupture distance computed as shown in **Figure 7B**.

AEP for two earthquake IMs, PGA and  $S_a(T_1 = 0.3 \text{ s})$ , are computed for three observation points (shown in **Figure 5A**) in Seaside. These points are aligned shore-normal along the urban center of Seaside with approximately 400 m distance apart from each other. The observation Points 1, 2, and 3 are located between shoreline and the Necanicum River, Necanicum River and the Neawanna Creek, and Neawanna Creek and the edge of inundation zone, respectively.

**Figure 8** shows the AEP of PGA and  $S_a(T_1 = 0.3 \text{ s})$  at three observation points computed for the full-rupture scenarios ( $M = 8.8$ ,  $M = 9.0$ ,  $M = 9.2$ ) using the BC Hydro GMPE. As per the soil site class map of **Figure 5C**, point 1 and point 2 fall in



**FIGURE 9** | Joint hazard curves of peak ground acceleration (PGA) and  $S_a$  at (A) Point 1, (B) Point 2, and (C) Point 3.

site class D where as point 3 is in site class C. The  $V_{s30}$  values assigned to Point 1, Point 2, and Point 3 are 300, 360, and 420 m/s, respectively. A value of  $\lambda_m = 0.0019$  corresponding to  $M \geq 8.8$  is used for the hazard curves and subsequent AEP computations. In AEP computations, the PMF values considered for the three magnitudes is consistent with those used in the study by Park and Cox (2016), which were obtained from the study by Goldfinger et al. (2012). Since the weights assigned to the M 8.8, M 9.0, and M 9.2 are 5/19, 13/19, and 1/19, respectively, the M 9.0 is the dominant contributor to the hazard curves presented in **Figure 8** for both PGA and  $S_a(T_1 = 0.3 \text{ s})$ , followed by M 8.8, and M 9.2, respectively. As shown in **Figure 8**, AEP of PGA for three points are almost identical whereas slight variation in AEP for  $S_a(T_1 = 0.3 \text{ s})$  is observed for those three points. The slight variation in AEP for  $S_a(T_1 = 0.3 \text{ s})$  is due to the different  $V_{s30}$  values assigned to those points. Overall AEP of both *IMs* are insensitive to their geographical conditions at this 800 m distance between the three points, which are all approximately 27.5 km from the fault.

The joint MAR of exceedance of *IMs* contour plot of PGA and  $S_a(T_1 = 0.3 \text{ s})$  is shown in **Figure 9** for the three points indicated in **Figure 5** for the 1,000-year event. The joint MAR of exceedance of *IMs* is computed based on the formulation presented in Eqs 15 and 16, in which the joint probability function of the two *IMs* conditional on the magnitude at a given site is computed assuming that the *IMs* follow a joint lognormal distribution. For the joint MAR of exceedance computation of PGA and  $S_a(T_1 = 0.3 \text{ s})$ , correlation between the two *IMs* are considered and computed based on the study by Baker and Jayaram (2008). It can be seen from **Figure 9** that for all three observation points, the most likely joint intensities occur for values of PGA = 0.5 g and  $S_a(T_1 = 0.3 \text{ s}) = 1.0 \text{ g}$ . Moreover, there is no discernable difference in the joint hazard intensities among the observation points considered as can be observed in **Figure 9**, as expected.

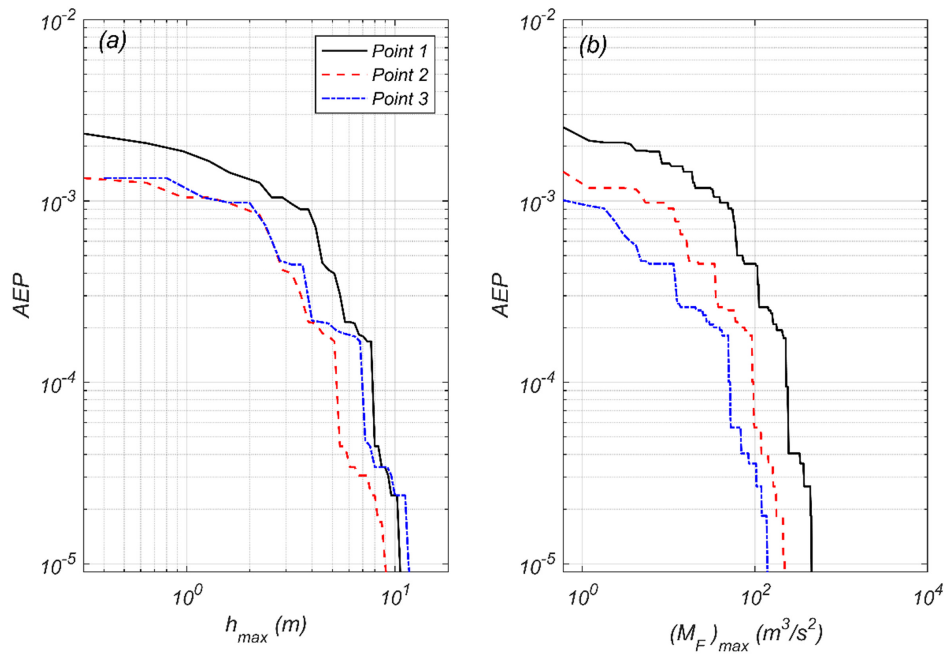
## Tsunami Intensity

**Figure 10** shows the AEP of two tsunami *IMs*  $h_{\max}$  and  $(M_F)_{\max}$  at the three observation points similar to **Figure 8**. Generally, the tsunami *IMs* decrease as distance from the shoreline increases because of the dissipation of energy during the inundation process as shown for the AEP of  $h_{\max}$  (**Figure 10A**) and  $(M_F)_{\max}$  (**Figure 10B**). For example, the AEP of both  $h_{\max}$  and  $(M_F)_{\max}$  at Point 1 are higher than Points 2 and 3. However, for lower

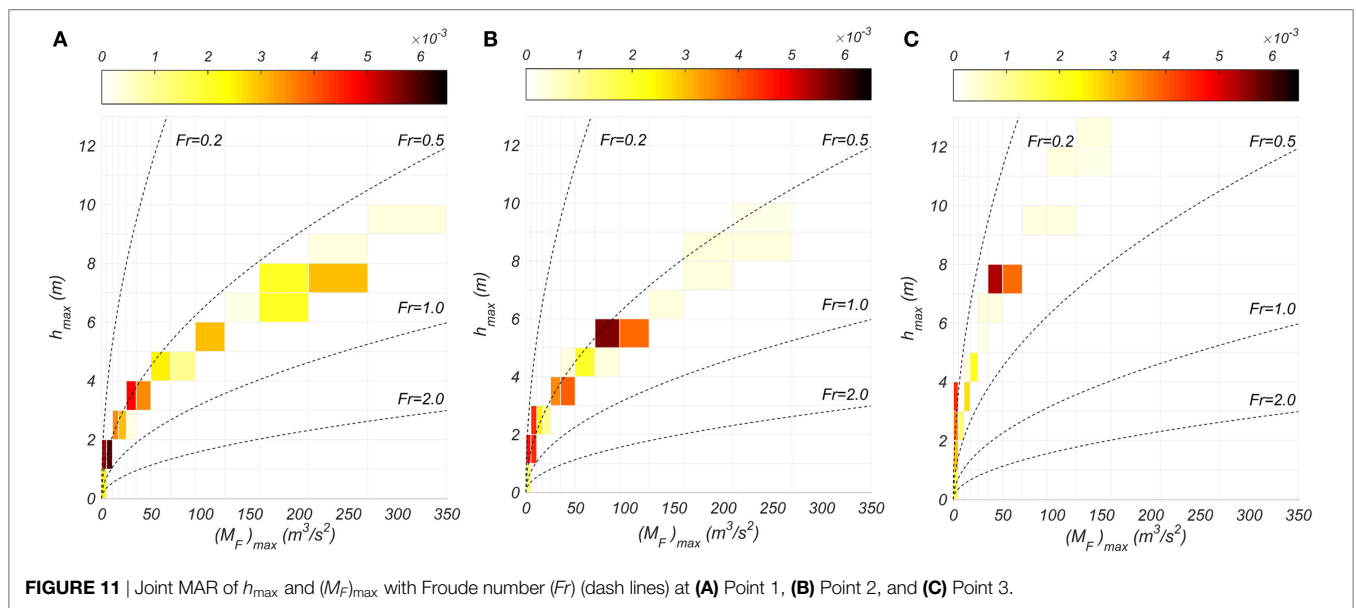
probability events in the range of  $0 < h_{\max} \leq 3 \text{ m}$ , Points 2 and 3 have nearly the same AEP. Point 3 has higher  $h_{\max}$  at the higher probability events in the range of  $3 < h_{\max} \leq 10 \text{ m}$  and also has the highest  $h_{\max}$  due to the local topographic effects, including effects of the river and the creek running parallel to the shoreline (**Figure 5A**), highlighting the sensitivity of the tsunami *IM* to site-specific conditions. On the other hand,  $(M_F)_{\max}$  shown in **Figure 10B** has three distinct curves with a generally decreasing trend from the shoreline to inundation limits. Comparing **Figures 8** and **10**, it is noted that the variation of AEP of tsunami *IMs* among three observation points are qualitatively dissimilar to the results of earthquakes *IMs* in that there are strong spatial gradients for the tsunami *IMs* across the length scale of the city. This is somewhat expected because the AEP of both PGA and  $S_a(T_1 = 0.3 \text{ s})$  depend primarily on the soil types and  $R_{\text{rup}}$  that have relatively small variations over the study region at Seaside, especially for the full-rupture scenarios considered. Moreover, this difference underscores the differences in the fundamental physics of the propagation of the seismic energy through the subsurface and the propagation of the hydrodynamic tsunami energy.

**Figures 11A–C** shows the joint AEP of  $h_{\max}$  and  $(M_F)_{\max}$  at Point 1, 2, and 3, respectively, where the same number of bins for both *IMs* are utilized to calculate the vector-valued *IMs* of the joint  $h_{\max}$  and  $(M_F)_{\max}$  from Eqs 15 and 16. **Figure 11** also includes the isolines for four Froude numbers ( $Fr = 0.2, 0.5, 1.0, \text{ and } 2.0$ ) to show the possible range of distributions of velocity fields at the given flow depth. In general, the results show the high correlation between  $h_{\max}$  and  $(M_F)_{\max}$  at all three observation points, and each joint AEP follows a particular  $Fr$  for that area. For example, at Point 1 when  $h_{\max}$  is in the range  $6 < h_{\max} < 7 \text{ m}$ , the corresponding range of  $(M_F)_{\max}$  is  $150 < (M_F)_{\max} < 200 \text{ m}^3/\text{s}^2$  and corresponds to a  $Fr$  range of approximately  $0.6 < Fr < 0.7$ . The  $Fr$  generally decreases shoreward, and seems to be everywhere subcritical ( $Fr < 1.0$ ) which is physically realistic based on observed inundations following the 2011 Tohoku tsunami. This is not to say that the flow is always subcritical, just that at the point of maximum flow depth or momentum flux, that the flow can be expected to be subcritical.

**Figure 11** reveals that the different distributions of the joint surface of the AEP following the specific  $Fr$  at three observation points originate from the local bathymetric/topographic conditions and not from the fault slip distributions. These results



**FIGURE 10** | Annual exceedance probability (AEP) of  $h_{max}$  (A) and  $(M_F)_{max}$  (B) at Points 1 (black solid), 2 (red dash), and 3 (blue dash-dot).

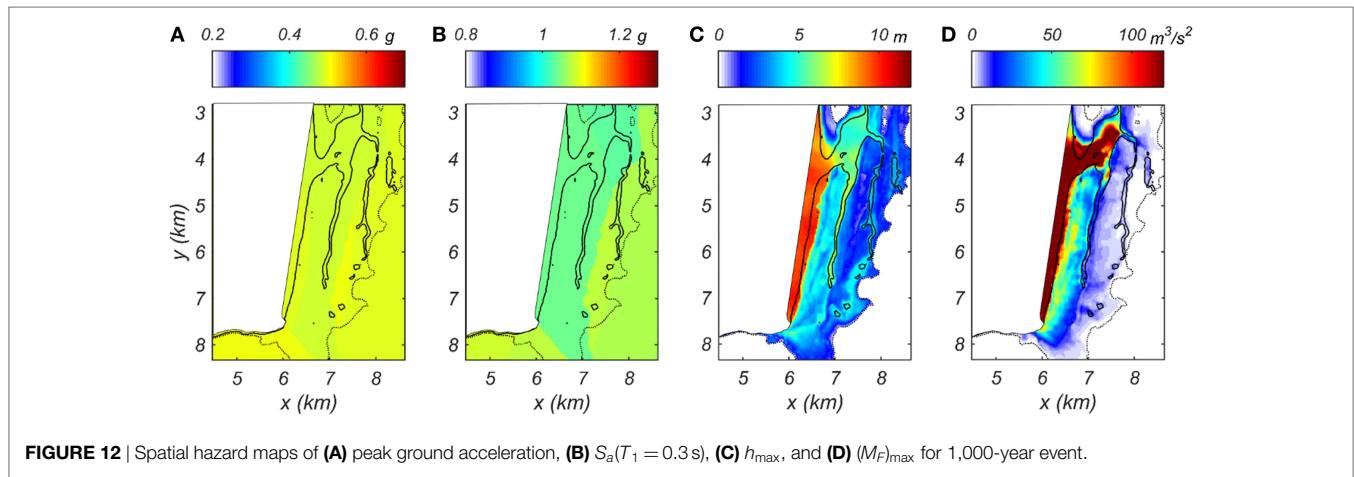


**FIGURE 11** | Joint MAR of  $h_{max}$  and  $(M_F)_{max}$  with Froude number ( $Fr$ ) (dash lines) at (A) Point 1, (B) Point 2, and (C) Point 3.

are helpful to understand realistic input conditions of flow depth and velocity fields for different  $Fr$  regime that can be used to develop the tsunami fragility curves, which utilize the random combinations of  $h_{max}$  and flow velocity in the generation of fragility curves (Attary et al., 2017a; Attary et al., 2017b; Alam et al., submitted<sup>2</sup>). In addition, it is worth noting that

the maximum flow depth and momentum flux typically do not occur at the same time (Park et al., 2013). Therefore,  $Fr$  at the maximum momentum flux and at the maximum flow depth can be expected to be slightly different. However, plots similar to **Figure 11** (not shown in the interest of brevity) in which the flow depth  $h$  at the instant of the maximum momentum flux  $(M_F)_{max}$  or the instantaneous momentum flux is plotted against the corresponding maximum inundation flow depth lead to similar conclusions that the flow is generally subcritical for the maximum IMs.

<sup>2</sup>Alam, M. S., Barbosa, A. R., Scott, M. H., Cox, D., and van de Lindt, J. W. (submitted). Development of physics-based tsunami fragility functions considering structural member failures. *ASCE J. Struct. Eng.*



## Spatial Representations Seismic and Tsunami Hazard

**Figure 12** shows the spatial distributions of the earthquake and tsunami  $IM$ s at Seaside, Oregon, for the AEP = 0.001 (often referred to as the “1,000-year event”) for the full-rupture scenario at CSZ. **Figures 12A–D** presents the spatial distribution of PGA,  $S_a$  ( $T_1 = 0.3 \text{ s}$ ),  $h_{\max}$ , and  $(M_F)_{\max}$  respectively, and the dotted contour lines in each panel show the maximum inundation limits ( $h_{\max} = 0.3 \text{ m}$ ). Two distinct regions of earthquake  $IM$ s (PGA and  $S_a$ ) are observed for the study area depending on the soil site class assumed in **Figure 5C** and within each region distributions of  $IM$ s are generally uniform. In the case of tsunami  $IM$ s ( $h_{\max}$  and  $(M_F)_{\max}$ ), irregular distributions are observed depending on the bathymetry conditions. However, both the  $IM$ s generally decreases from the shore toward inland (i.e., in the positive  $x$ -direction).

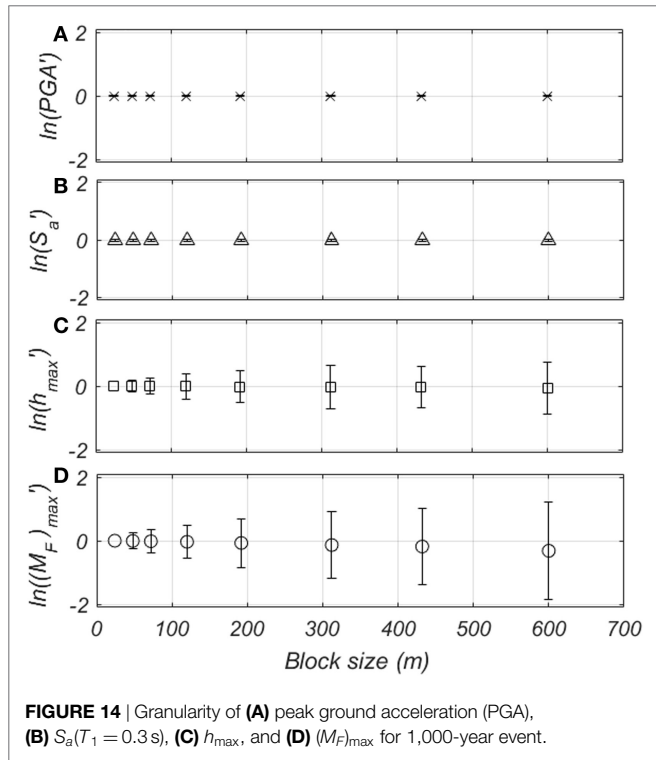
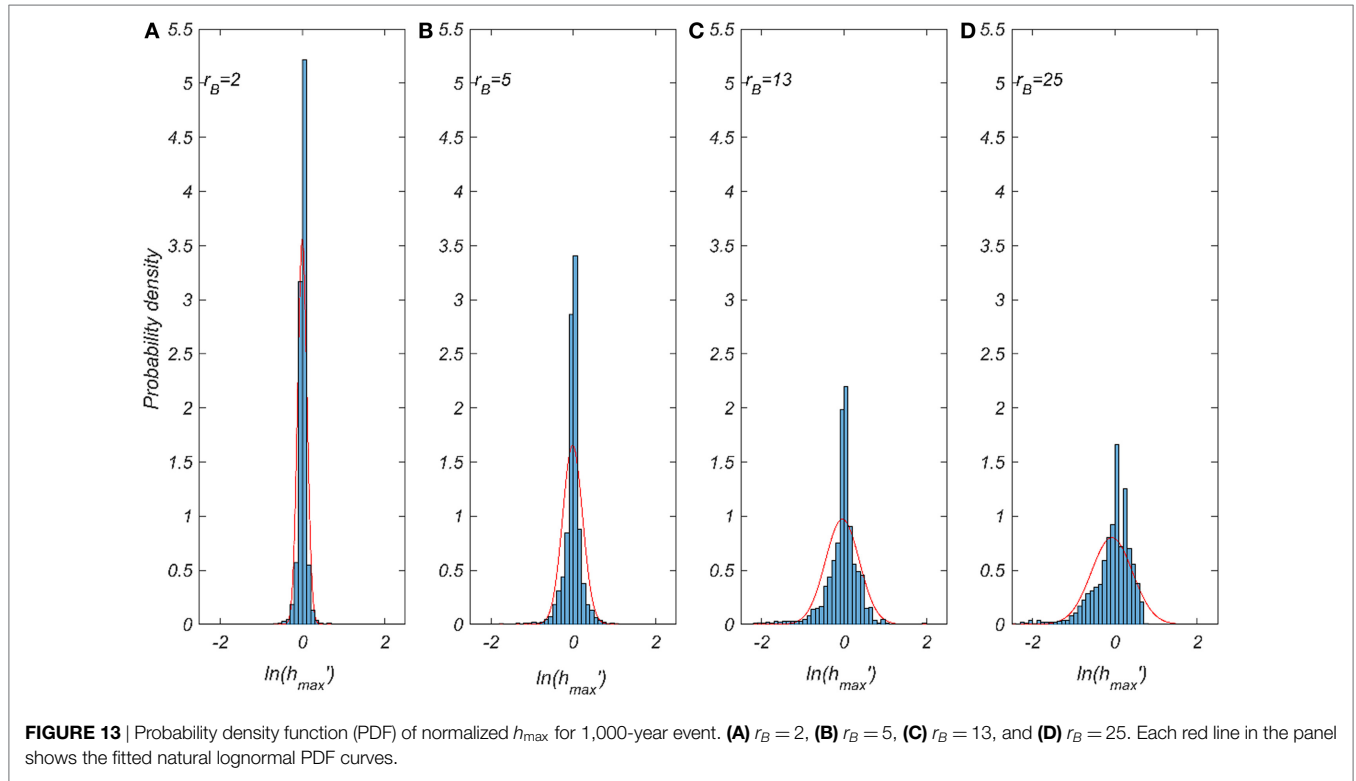
The range of PGA and  $S_a(T_1 = 0.3 \text{ s})$  over the study region is 0.47–0.50 g and 1.03–1.08 g, respectively, while the  $h_{\max}$  and  $(M_F)_{\max}$  range from 0 to 12 m and 0 to 120  $\text{m}^3/\text{s}^2$ , respectively for the 1,000-year event. The values of PGA and  $S_a$  are uniformly distributed over the entire study area, while the area affected by  $h_{\max}$  and  $(M_F)_{\max}$  are limited to the maximum inundation limits. To understand the granularity of both  $IM$ s of the earthquake and tsunami at the same event, the spatial mean and deviation of  $IM$ s conditional on varying unit block size is computed. This analysis is performed on the rectangle region, shown in **Figure 5A**, which is a 1,990 m length and 3,000 m wide rectangle near the center of the City of Seaside. The smallest unit block size, designated as reference or unit block size here forth, is equal to the Cartesian mesh grid size ( $dx = dy = 24 \text{ m}$ ) considered for tsunami inundation modeling in the C-grid. The ratio of block size ( $r_B = \frac{\text{Block size}}{\text{reference mesh size}}$ ) is increased systematically, keeping the same block shape and the mean and SD of the  $IM$ s are estimated as a function of the changes in the block size. The mean of normalized  $IM$  is given by

$$IM' = \frac{\sum_{j=1}^{N_b} \left( \sum_{i=1}^{N_g} im'_{i,j} \right)}{N_b N_g} \quad (18)$$

where  $N_b$  is the number of blocks depending on  $r_B$  conditions, and  $N_g$  is the number of grid points in a  $j$ th block ( $N_g = r_B^2$ ). The  $im'_{ij}$  is the normalized  $IM$  of grid points of the  $j$ th block. The  $IM$ s in the  $j$ th block are normalized by the mean of the  $IM$ s for the  $j$ th blocks, and this calculation is performed for eight  $r_B$  conditions:  $r_B = 1, 2, 3, 5, 8, 13, 18,$  and  $25$ . The corresponding block size and number of grid points per each block ( $N_g$ ) are 24 m (1), 48 m (4), 72 m (9), 120 m (25), 192 m (64), 312 m (169), 432 m (324), and 600 m (625). For example, the total number of grid points in the study region is 10,625. For the case of  $r_B = 5$ , there are 425 block subsets over the study region, and each block is composed of  $N_g = 25$  mesh grid points. The mean of  $IM$  at each of the 425 blocks is first computed, and then, each  $IM$  is normalized by the calculated mean.

**Figure 13** shows the PDF of  $\ln(h_{\max}')$  over the study region for AEP = 0.001 (1,000-year event). **Figures 13A–D** show the PDFs at  $r_B = 2, 5, 13,$  and  $25$ , respectively. The red line in each panel shows the natural log-normal fitting curve. When  $\ln(h_{\max}')$  is equal to zero, the mean of block subsets match exactly with the  $h_{\max}$  of the unit grids, and each negative or positive value indicates an overestimation or underestimation of the block means. In the case of  $h_{\max}$ , the PDF shape becomes wider and shows a larger deviation as the block size increases. This process was extended to the three other  $IM$ s: PGA and  $S_a$  and  $(M_F)_{\max}$ .

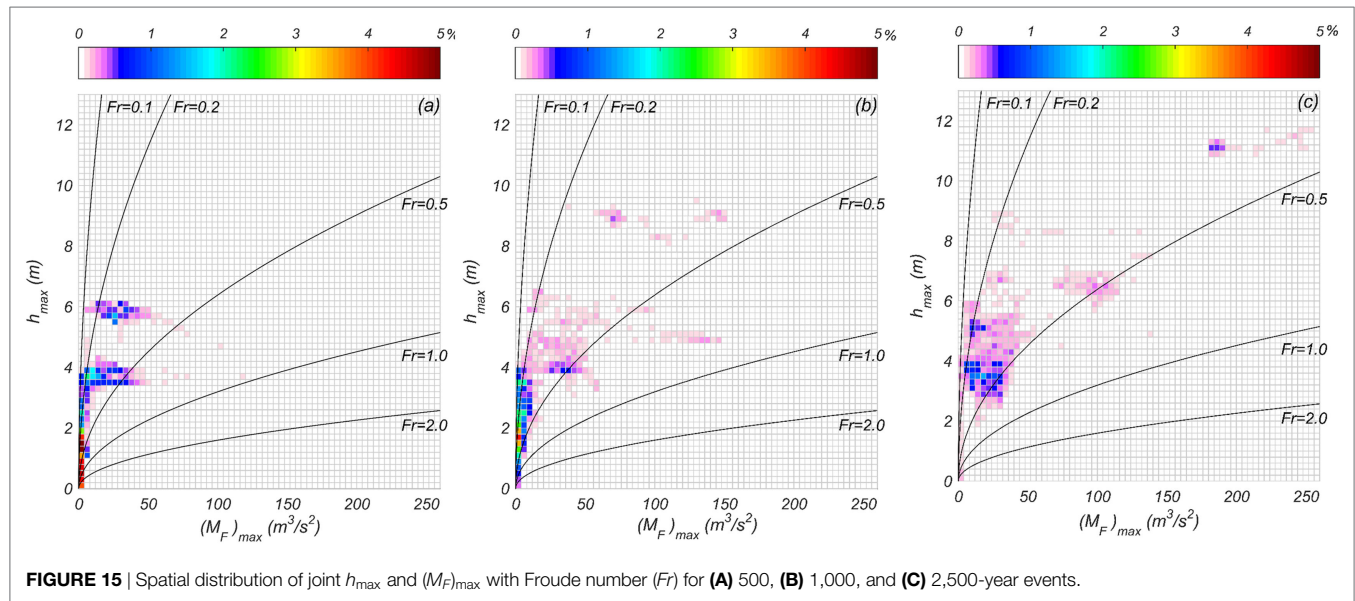
**Figure 14** summarizes the results of the four granularity tests for AEP = 0.001 (1,000-year event) by plotting the mean of the four normalized  $IM$ s with 90% confidence intervals for each unit block size. Both PGA and  $S_a$  ( $T_1 = 0.3 \text{ s}$ ) show almost insensitivity to the block sizes. All of the normalized means are essentially zero and the confidence intervals are small for all block sizes. However, both  $h_{\max}$  and  $(M_F)_{\max}$  results show significant variation with the 90% confidence intervals increasing as the block size increases. The largest deviation is found at  $(M_F)_{\max}$ , and relatively smaller deviation if found  $h_{\max}$ . The range of confidence interval of both  $h_{\max}$  and  $(M_F)_{\max}$  increases sharply with  $r_B = 13$  (block size = 312 m). These results highlight the different sensitivity between two earthquake and tsunami  $IM$ s to the block size used to aggregate the  $IM$  results. In case of the PGA and  $S_a$ , the information of the site and location are less significant for determining each  $IM$  while  $h_{\max}$  and  $(M_F)_{\max}$  require detailed information of the site to minimize the uncertainty.



Lastly, the spatial distribution of joint IMs is analyzed next. Only  $h_{\max}$  and  $(M_F)_{\max}$  are considered herein since significant variations of tsunami IMs across the study area is observed in **Figure 12**. For the joint distribution computations uniform bins at

0.2 m interval are used for  $h_{\max}$ , while a  $4\text{ m}^3/\text{s}^2$  interval is used for  $(M_F)_{\max}$ . The probability (%) of the joint distribution is computed by counting number of grids which involved in the joint  $h_{\max}$  and  $(M_F)_{\max}$  bin, for the whole study region, at the three AEP conditions, i.e., the 500, 1,000, and 2,500-year events, respectively. The results of spatial distributions of joint probability of  $h_{\max}$  and  $(M_F)_{\max}$  are shown in **Figures 15A–C** for the 500, 1,000, and 2,500-year events, respectively, with four Froude numbers plotted on each figure in a manner similar to **Figure 11**.

In the case of the 500-year event (**Figure 15A**), more than 45% of the joint  $h_{\max}$  and  $(M_F)_{\max}$  is distributed within  $0.1 < Fr \leq 0.5$ . The larger  $(M_F)_{\max}$  is observed within  $0.5 < Fr \leq 1.0$ . The joint peak of  $h_{\max}$  and  $(M_F)_{\max}$  is located at  $1.2 < h_{\max} \leq 1.4\text{ m}$ , and  $0.0 < (M_F)_{\max} \leq 4.0\text{ m}^3/\text{s}^2$ . Two distinct narrow banded regions of  $h_{\max}$  are observed near  $h_{\max} = 3.8$  and  $h_{\max} = 5.8\text{ m}$ , which correspond to different ranges of  $(M_F)_{\max}$  even at the similar flow depth conditions. In the case of the 1,000-year event (**Figure 15B**), more than 60% of the joint  $h_{\max}$  and  $(M_F)_{\max}$  is distributed within  $0.1 < Fr \leq 0.5$ . Large  $(M_F)_{\max}$  values are observed at both  $0.1 < Fr \leq 0.5$  and  $0.5 < Fr \leq 1.0$ . The joint peak is located at  $1.6 \leq h_{\max} < 1.8\text{ m}$ , and  $0.0 < (M_F)_{\max} \leq 4.0\text{ m}^3/\text{s}^2$  which is a slightly higher  $h_{\max}$  condition than the 500-year event. Several isolated islands of joint distributions are observed that have relatively higher  $(M_F)_{\max}$  values. In the case of the 2,500-year event (**Figure 15C**), more than 68% of the joint  $h_{\max}$  and  $(M_F)_{\max}$  is located within  $0.1 < Fr \leq 0.5$ . Large  $(M_F)_{\max}$  values are located primarily within  $0.1 < Fr \leq 0.5$ . The joint peak is located at  $3.4 < h_{\max} \leq 3.6\text{ m}$ , and  $16.0 < (M_F)_{\max} \leq 20.0\text{ m}^3/\text{s}^2$ , which are significantly larger than the 500 and 1,000-year events. A few isolated points are observed for 2500-year event similar to that



**FIGURE 15** | Spatial distribution of joint  $h_{\max}$  and  $(M_F)_{\max}$  with Froude number ( $Fr$ ) for **(A)** 500, **(B)** 1,000, and **(C)** 2,500-year events.

observed in 1,000-year event. Overall, the joint  $h_{\max}$  and  $(M_F)_{\max}$  is distributed within  $0.1 \leq Fr < 1.0$  and are typically in the range  $0.1 < Fr \leq 0.5$ . It is also observed that each event has a different joint peak location. Further research is needed to understand the extent to which these are site-specific results, how this work can be generalized for other coastal archetypes (e.g., embayments), and how these results can be parameterized based on conditions offshore of the study area.

## SUMMARY, CONCLUSION, AND FUTURE WORK

This paper presents a framework for a consistent probabilistic hazard assessment for the multi-hazard seismic and tsunami phenomena (PSTHA). For this work, full-rupture event along the CSZ is considered and the PSTHA methodology is applied to the study area of Seaside, Oregon, along the US Pacific Northwest coast. In this work, it is shown that:

1. The AEPs of the tsunami  $IMs$  are qualitatively dissimilar to the  $IMs$  of the seismic ground motion in the study area. Specifically, the spatial gradients for the tsunami  $IMs$  are much stronger across the length scale of the city owing to the physical differences of energy dissipation of the two mechanisms.
2. For the seismic hazard, similar trends were observed for the joint MAR of exceedance of the PGA and  $S_a(T_1 = 0.3 \text{ s})$  at the three observation points, which may be attributed to the proximity of the observation points with respect to each other as well as from the source of the megathrust full rupture of CSZ.
3. For the tsunami hazard, the joint AEP of  $h_{\max}$  and  $(M_F)_{\max}$  show a high correlation between  $h_{\max}$  and  $(M_F)_{\max}$  in the study area. The joint AEP at each of the three observation points follows a particular  $Fr$  due to the local site-specific conditions at each location rather than the distributions of fault slips.
4. The joint probability distribution of  $h_{\max}$  and  $(M_F)_{\max}$  throughout the study region falls between  $0.1 \leq Fr < 1.0$  (i.e., the flow

is subcritical), regardless of return interval (500-, 1,000-, and 2,500-year). However, the peak of the joint probability distribution with respect to  $h_{\max}$  and  $(M_F)_{\max}$  varies with the return interval, and the largest values of  $h_{\max}$  and  $(M_F)_{\max}$  were observed with the highest return intervals (2,500 year) as would be expected.

The tsunami inundation simulations were conducted with a bare earth DEM, and the effects of the natural and built environment were simply modeled using a single friction factor. It is known that the tsunami inundation velocity is sensitive to the choice of friction factor (e.g., Park et al., 2013) and that the friction factor can vary significantly for the built environment (e.g., Bricker et al., 2015). Therefore, future research should consider the effect of bottom friction uncertainty in modeling the probabilistic tsunami hazard assessment. Moreover, the difference in the types of models being used with respect to the approximation of the governing equations (e.g., non-linear shallow water equations or Boussinesq equations), solution techniques (finite element, finite difference), grid resolution, and so on introduce model source uncertainty and should be evaluated in a similar manner as bottom friction.

In terms of multi-hazards, future PSTHA frameworks should include additional sources such as other earthquake faults (on land), which would not produce tsunamis. Nonetheless, these would contribute to the seismic hazard in the region. Conversely, distance sources tsunamis should also be included, which may not produce ground shaking.

The results of the PSTHA can be the basis for a probabilistic multi-hazard damage assessment (PTSDA) to quantify the separate contributions of seismicity and tsunami hazards in the estimation of damage to the built environment over the community scale. In addition, it will be necessary to understand the propagation of uncertainties of the hazard assessments combined with the uncertainties of the damage estimates to evaluate the overall community risk to the multi-hazards. Moreover, the PTSDA should be evaluated considering the spatial gradients of the building damages at a



community scale due to the site information (e.g., building types) or fragility functions applied to each building for both earthquake and tsunami *IMs* across the length scale.

## AUTHOR CONTRIBUTIONS

Each author contributed equally to the manuscript. HP and DC contributed to the tsunami hazard modeling. MA and AB contributed to the seismic hazard modeling.

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**Conflict of Interest Statement:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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