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Machine learning and synthetic outcome estimation for individualised antimicrobial cessation

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The decision on when it is appropriate to stop antimicrobial treatment in an individual patient is complex and under-researched. Ceasing too early can drive treatment failure, while excessive treatment risks adverse events. Underand over-treatment can promote the development of antimicrobial resistance (AMR). We extracted routinely collected electronic health record data from the MIMIC-IV database for 18,988 patients (22,845 unique stays) who received intravenous antibiotic treatment during an intensive care unit (ICU) admission. A model was developed that utilises a recurrent neural network autoencoder and a synthetic control-based approach to estimate patients' ICU length of stay (LOS) and mortality outcomes for any given day, under the alternative scenarios of if they were to stop vs. continue antibiotic treatment. Control days where our model should reproduce labels demonstrated minimal difference for both stopping and continuing scenarios indicating estimations are reliable (LOS results of 0.24 and 0.42 days mean delta, 1.93 and 3.76 root mean squared error, respectively). Meanwhile, impact days where we assess the potential effect of the unobserved scenario showed that stopping antibiotic therapy earlier had a statistically significant shorter LOS (mean reduction 2.71 days, p-value <0.01). No impact on mortality was observed. In summary, we have developed a model to reliably estimate patient outcomes under the contrasting scenarios of stopping or continuing antibiotic treatment. Retrospective results are in line with previous clinical studies that demonstrate shorter antibiotic treatment durations are often non-inferior. With additional development into a clinical decision support system, this could be used to support individualised antimicrobial cessation decision-making, reduce the excessive use of antibiotics, and address the problem of AMR.

KEYWORDS

antimicrobial resistance, artificial intelligence, clinical decision support systems, decision-making, individualised antimicrobial prescribing, precision prescribing, antibiotic cessation, outcome estimation

Introduction

Bacterial antimicrobial resistance (AMR) is a global threat (1, 2), which resulted in an estimated 1.27 million deaths in 2019 (3). One key strategy to tackle AMR is to optimise antimicrobial use and prolong current antimicrobials' therapeutic life. Clinical decision support systems (CDSSs) are software designed to provide information to healthcare professionals, patients, or other individuals in order to make informed clinical decisions. With the advent of artificial intelligence (AI) and the ever increasing prevalence of electronic health records (EHRs), numerous CDSSs utilising machine learning (ML) trained on historical patient data have been developed to assist with managing infections (4). Recent research has focused on the diagnoses of bacterial infections (5–7), resistance prediction (8), and antimicrobial therapy selection (9, 10).

One challenge when treating a patient who has a bacterial infection is determining when it is appropriate to stop antibiotic treatment (11). The decision to cease antibiotics too early can result in the patient's condition worsening, while unnecessary exposure increases the risk of toxicity (12) and drives the evolution of AMR (13). Even over-treating for a short duration can have a significant impact on a population level and enhances the development of resistance (14). Furthermore, excessive treatment is responsible for most avoidable antibiotic adverse events including gastrointestinal distress and allergic reactions (15, 16). Numerous studies have shown that on a population level, shorter treatment durations are often non-inferior to longer ones (17-21). The challenge is that the resulting recommendations do not take into account the individual patient's characteristics or specific scenarios. It is difficult for clinicians to have confidence in individualised treatment decisions for their patient, when there is a poor understanding of the factors that facilitate or inhibit an individual from receiving a short duration of antibiotic therapy. Therefore, durations are often unnecessarily extended (22) and decided by habit or arbitrarily based on population evidence. Antibiotic cessation should be a collective, data-driven decision, given choices are made in a more favourable environment once time has passed from presentation and significant amounts of information have been gathered. Despite this, systems to help support individualised antibiotic duration and cessation decisionmaking are often neglected and under-researched with little innovation in this area (23, 24).

Given the current standard of care uses clinical factors to determine if a patient should stop antibiotics or not, we hypothesise that an AI-based CDSS using routinely collected EHR data may be able to support individualised antibiotic cessation decision making and overcome prescriber concerns of poor patient outcomes that is likely a major driver of over treatment (25, 26). We approach this problem by estimating clinical outcomes under alternative scenarios with the aim of showing non-inferiority or a direct benefit of antibiotic cessation. More specifically, a machine learning and synthetic control-based approach was developed to estimate patients' LOS and mortality outcomes for any given day, if they were to stop vs. continue antibiotic treatment. Figure 1 shows a graphical abstract of the approach and methodology employed in this retrospective research study.

Methods

Dataset

MIMIC-IV is a large de-identified real-world clinical dataset that is publicly available for clinical research (27, 28). It contains EHR information for over 40,000 patients admitted to the Beth Israel Deaconess Medical Center (BIDMC) in Boston, MA, Unites States, between 2008 and 2019. The patient population was filtered to those who received intravenous antibiotic treatment for a duration between 1 and 21 days during an ICU stay. Input features were extracted, analysed, and selected based on prevalence, correlation, as well as infectious disease doctors and critical care consultants advice. Length of stay (LOS) (continuous value) and mortality (binary) labels were extracted for each patient stay; however, it should be noted that these are not temporally dynamic. An overview of statistics for each dataset is shown in Table 1

Some features were calculated based on other variables. Cumulative overall antibiotic treatment length was determined for each day of each ICU stay that considered consecutive treatment days irrespective of the antibiotic given. In addition, whether the patient had received re-treatment for antibiotics or not and their age at the time of ICU admission were also computed. Standard pre-processing was applied to features including outliers being removed and values normalised, as well as missing values forward filled or highlighted. Features were aggregated by day for each unique stay to create a regular temporal dataset. In general, there was a high degree of missingness, and so patients with greater than 50% of values missing each day were removed. The resulting dataset contained 43 input features (supplementary Table S1) including lab test results, clinical parameters, ventilation settings, and demographics.

Model architecture

The objective of our model is to estimate the patients' LOS and mortality outcomes for any given day, if they were to stop

1 DATASET	Filter relevant patient stays from MIMIC-IV, extract and aggregate features including lab test results, clinical parameters, ventilation settings and demographics to create a regular temporal dataset for estimation of their length of stay and mortality outcomes. MIMIC-IV >40,000 ICU patients ICU intravenous antibiotic treatment (1 <days<21) OUR DATASET 18,988 ICU patients 22,845 unique stays</days<21)
2 AUTO ENCODER	Train autoencoders to create an embedding that is representative of the patients' temporal features and a linear predictor of their outcomes. $ \begin{array}{c} $
3 SYNTHETIC OUTCOME ESTIMATION	Apply the adapted synthetic control methodology to estimate patient outcomes if they were to stop vs continue treatment on each antibiotic day within their ICU stay.
4 EVALUATION AND VALIDATION	Evaluate stop and continue estimations through 'impact' and 'control' days. Validate the model through numerous tests and application to pneumonia and UTI datasets. STOP STOP IMPACT CONTROL CONTINUE CONTINUE CONTROL - Mean delta - RMSE - AUROC

vs. continue antibiotic treatment. It uses a bi-directional long short-term memory (LSTM) autoencoder, which takes in a sequence of patient input features $(x^{(1)}, x^{(2)} \dots x^{(T)})$, creates an embedding representation, and outputs a sequence of reconstructed features $(\tilde{x}^{(T)} \dots \tilde{x}^{(2)}, \tilde{x}^{(1)})$. This autoencoder is trained through two loss functions (29), which are summed together to create a combined loss for backpropagation. First, the reconstruction loss L_r is calculated by the root mean squared error (RMSE) between outputs that are trying to reproduce the inputs and the real input data. Second, a supervised learning loss L_s is calculated by doing a linear transformation of the embedding representation (\tilde{Y}) to try and predict the real label (Y) and taking either the RMSE loss for the LOS outcome or the binary cross-entropy loss for mortality classification. L_s ensures that the embedding space created by the autoencoder is a good linear predictor of the outcome of interest, which is important for the subsequent adapted synthetic control method. Overall, an embedding representation is created that considers a patient's past and is representative of their state on that day.

TABLE 1 Datasets statistics.

	Dataset							
Statistic	Overall	Train	Validation	Test	Pneumonia	UTI		
Number of stays	22,845	15,991	3,427	3,427	2,473	923		
Mortality rate	18.60	18.47	18.30	19.52	24.02	18.96		
Mean LOS	5.63	5.62	5.74	5.55	9.05	5.50		
LOS standard deviation	4.23	4.24	4.31	4.15	5.19	4.32		
Mean length of treatment	4.38	4.38	4.48	4.30	6.95	4.77		
Length of treatment standard deviation	3.32	3.32	3.48	3.18	4.28	3.55		
Spearman's correlation between LOS and treatment length	0.72	0.72	0.72	0.73	0.73	0.74		
Percentage of patients that stopped treatment during their ICU stay	41.56	41.64	41.17	41.55	31.95	26.54		

LOS, length of stay; ICU, intensive care unit; UTI, urinary tract infection.

Once the antoencoder is trained and an embedding representation for each antibiotic day in all patient stays have been created, an adapted synthetic control approach (30) is utilised, where the act of stopping or continuing treatment on a particular day is considered an intervention and each patient acts as a singular unit. This method is useful when evaluating an intervention using randomised controlled trials is challenging, as is the case with antibiotic cessation, and hence retrospective observational data are assessed. Synthetic controls have frequently been applied to understand public health interventions (31, 32), but their use within digital health research is limited. In this study, we want to know what are the predicted outcomes if a given patient was to stop vs. continue antibiotics on a given day within their ICU stay. To this extent, two synthetic controls are created, one can be labelled the "stop synthetic control," which is based on subjects who stopped antibiotics on that particular day, and the second labelled the "continue synthetic control," which is created from subjects who continue antibiotic treatment on that particular day. To achieve this for each day (t), two separate donor pools are created based on subjects associated embedding representation and antibiotic treatment status. In other words, those who continue antibiotics on day t are partitioned into the "continue" embedding space while those who stop antibiotics are placed in the "stop" embedding space. In this way, the estimated outcomes for stopping and continuing on day t are driven by representative donors who experienced analogous treatment. To create the stop and continue synthetic controls for a particular patient i, the kmost closely related to embedding representations from each relevant donor pool are selected based on a distance metric (in this study k = 10 and Euclidean distance were used for stop and continue estimations). Given both that embeddings are representative of the patients' state, those selected donors will be similar, giving a considered into potential alternative outcomes insight under antibiotic temporality. A ridge regression function

 $\begin{aligned} (Loss_{i}^{S,t} &= \sum_{d=1}^{D} [z_{i,d}^{t} - \sum_{j=1}^{k} x_{j,d}^{S,t} w_{i,j}^{S,t}]^{2} + \sum_{j=1}^{k} w_{i,j}^{S,t^{2}} & \text{for stop} \\ \text{estimations} & \text{and} & Loss_{i}^{C,t} &= \sum_{d=1}^{D} [z_{i,d}^{t} - \sum_{j=1}^{k} x_{j,d}^{C,t} w_{i,j}^{C,t}]^{2} + \\ \sum_{j=1}^{k} w_{i,j}^{C,t^{2}} & \text{for continue estimations, where } d \text{ are the} \end{aligned}$ embedding dimensions, j are the donors, and z represents the particular patient i's embedding for a given dimension and time) is then applied to the subject and their respective stop and continue donor embeddings. This returns two sets of weights $(w_{i,j}^{S,t}$ for "stop" and $w_{i,j}^{C,t}$ for "continue") that minimise the square difference between the subject of interest and the selected units in the donor pools $(Y_{i,j}^{S,t}$ for "stop" and $Y_{i,i}^{C,t}$ for "continue"). The objective of this L2 regularisation is to fairly distribute weights across the donors for stop and continue estimations. Finally, the stop and continue synthetic control outcomes $(\tilde{Y}_i^{S,t}$ and $\tilde{Y}_i^{C,t}$, respectively) for the particular patient *i* are computed from the weighted average of donor labels. To this extent during outcome estimation for a given patient i, we assume that we know the outcomes for all other patients within the dataset. Overall outcomes are estimated for each patient on each relevant antibiotic day of their stay if they were to stop vs. continue antibiotic treatment. An overview of the model's architecture and this process for stop and continue outcome estimation is shown in Figure 2.

Model development and software

The model was applied on the MIMIC-IV EHR dataset, which was randomly split based on patients' "stay_id" into training, validation, and testing sets (70%, 15%, and 15%, respectively). PyTorch (33) was used to create a bi-directional LSTM recurrent neural network (RNN) with a custom dataset class to extract labels and features. In order to address the mortality class imbalance (Table 1), over-sampling was used during training. To be specific, those cases with positive mortality were replicated three times within the custom dataset class to achieve a more balanced mortality rate of



51.90% within the train dataset. The Adam optimiser (34) was used with binary cross-entropy loss for classification, mean squared error loss for regression, and Ray Tune for hyperparameter optimisation (35). Training utilised 50 epochs, during which the model with the best performance on the validation dataset (RMSE or area under the receiver operating characteristic curve for LOS and mortality prediction, respectively) was selected as the final model. Two separate LSTM autoencoder models were trained on the whole training dataset to create embedding representations relevant to patients' LOS and mortality outcomes. Models were evaluated using functions and metrics from the TorchMetrics, Scikit-learn, and SciPy libraries. Further details of the two models' hyperparameters and their optimisation are shown in the supplementary material (supplementary Figure S1 and Table S2).

Model evaluation and metrics

Commonly with the synthetic control method, the delta difference between the single unit and the counterfactual in the pre-intervention period is minimised and the treatment effect is then observed in the post-intervention period. For our research question, this is not possible due to the nature of stopping antibiotics being the final event at one point in time, after which the patient is not applicable to our research population or question. An analogue can be applied for this study where we define "control" and "impact" days that are equivalent to the pre- and post- intervention periods. For estimating outcomes when continuing antibiotics, all the days the patient actually continues antibiotics are "control" days where we expect minimal difference between the true and estimated outcomes. On the other hand, on the single day the patient stops antibiotics, we can assess the "impact" if they were to instead continue. When estimating outcomes upon stopping antibiotics, the reverse is true, whereby each day antibiotics were continued the "impact" of stopping can be assessed and the final day where the patient stops treatment acts as a "control." Note that it is not possible to define this for every patient, given not every individual will stop antibiotics during their ICU stay. The percentage of patients who stopped antibiotic treatment during their ICU stay is shown in **Table 1**. Outcomes are estimated in the same way for impact and control days as discussed in the "Model architecture" subsection. However, for control days, we know the real outcome and so can compare our estimations, while for impact days, the real outcome is unknown. Each day, therefore, acts as both a "control" and "impact" across the two "stop" and "continue" scenario outcome estimations. An outline of this is shown in **Figure 3** and the number of continue and stop donors for each day in the test dataset is illustrated in **supplementary Figure S2**.

For outcome estimation, the mean delta is calculated to evaluate the difference between the real labels and the estimations, through the following formula: $\mu \Delta^{S} = (1/n) \sum_{i=1}^{n} \left[(1/T_{i}) \sum_{t=1}^{T_{i}} \left[Y_{i}^{S,t} - \tilde{Y}_{i}^{S,t} \right] \right]$ for stop estimations and $\mu \Delta^{C} = (1/n) \sum_{i=1}^{n} [(1/T_{i}) \sum_{t=1}^{T_{i}} [Y_{i}^{C,t} - \tilde{Y}_{i}^{C,t}]]$ for continue estimations, where T_i is the number of days that the patient receives antibiotics. Minimal difference should be seen on control days where our model aims to reproduce labels, while on impact days you can assess the effect of the unobserved scenario. Statistical analysis can be used to determine if the difference between the true LOS labels and the estimated outcomes are statistically significant. Given the non-normal data distribution, the non-parametric Wilcoxon rank-sum (Mann-Whitney U) test was used with the alpha set at 0.05. Furthermore, the mean absolute percentage error (MAPE) and mean absolute error (MAE) can be calculated through the following notations: can be calculated through the following hotatoris. $MAPE^{S} = (1/n) \sum_{i=1}^{n} [(1/T_{i}) \sum_{t=1}^{T_{i}} |Y_{i}^{S,t} - \tilde{Y}_{i}^{S,t}|/Y_{i}^{S,t}] \text{ and } MAE^{S} = (1/n) \sum_{i=1}^{n} [(1/T_{i}) \sum_{t=1}^{T_{i}} |Y_{i}^{S,t} - \tilde{Y}_{i}^{S,t}|], \text{ respectively,} for stop estimations and MAPE^{C} = (1/n) \sum_{i=1}^{n} [(1/T_{i}) \sum_{t=1}^{T_{i}} |Y_{i}^{C,t} - \tilde{Y}_{i}^{C,t}|/Y_{i}^{C,t}] \text{ and } MAE^{C} = (1/n) \sum_{i=1}^{n} [(1/T_{i}) \sum_{t=1}^{T_{i}} |Y_{i}^{C,t} - \tilde{Y}_{i}^{C,t}|]$ $Y_i^{C,t} - \tilde{Y}_i^{C,t}$], respectively, for continue estimations. Standard ML metrics can also be used to evaluate model prediction



performance. For LOS regression estimation, the RMSE is used, while for the mortality classification task, Area Under the Receiver Operating Characteristic curve (AUROC) is most appropriate given the class imbalance (**Table 1**), but accuracy, precision, recall, sensitivity, F1 score, and Area Under the Precision Recall curve (AUPRC) can also be calculated. Metrics were calculated as global averages, across all samples, meaning every day of antibiotic treatment within each patients stay is considered equally. 95% confidence intervals were calculated through 1,000 bootstrapped samples on the test set with n = 1, 000 for mortality metrics and the sum of the squared errors method for LOS RMSE.

To validate our findings beyond the hold out test set, we applied our model to patients who were diagnosed with pneumonia or a urinary tract infection (UTI). The effects of short vs. longer antibiotic treatment regimes have been extensively studied in pneumonia and UTIs. In general, research supports the notion that shorter antibiotic treatments durations are non-inferior to longer ones in these infections, especially for non-complicated cases (19, 36-40). Based on this evidence and the latest antimicrobial prescribing guidelines (41, 42), we defined a long treatment duration as any patient receiving antibiotics for longer than 7 days, and applied our model to estimate their outcomes if they were to instead stop treatment after 7 days. In addition, there is increasing evidence that even shorter courses of antibiotics can be used in such infections, in particular, pneumonia (19, 41). Hence, we investigated the estimated outcomes of those patients who received the standard of care 7 days treatment, for slightly shorter treatment durations (5 or 6 days).

Results

Autoencoder

In total, 18,988 patients, associated with 22,845 unique ICU stays, were included across datasets. Through a linear transformation of a given patient day embedding, outcome estimations could be made on the unseen test set (3427 unique ICU stays). The LOS model achieved an RMSE of 3.88 (95% CI 3.84–3.92), while the mortality estimation model obtained an AUROC of 0.77 (95% CI 0.73–0.80) [accuracy 0.73 (95% CI 0.71–0.75), precision 0.44 (95% CI 0.36–0.46), recall 0.67 (95% CI 0.61–0.72), specificity 0.75 (95% CI 0.72–0.78), F1 0.53 (95% CI 0.46–0.56), and AUPRC 0.55 (95% CI 0.42–0.56)] (Figure 4), indicating that the model was relatively effective at balancing false-positive and false-negative mortality predictions.

Synthetic outcome estimation

LOS and mortality estimation results on the unseen test set are shown in **Table 2**. For LOS estimation on control days, the mean delta under both stopping and continuing scenarios was 0.24 and 0.42 days, respectively, showing a minimal difference Bolton et al.



between predictions and the ground truth labels. Furthermore, a MAPE of 0.26, MAE of 1.32, and RMSE of 1.93 for stop control days show that the corresponding impact estimations are more reliable. On impact days, stopping earlier had a statistically significant shorter LOS (mean difference 2.71 days, *p*-value <0.01). This indicates that on average LOS estimations for stopping antibiotics earlier are shorter in duration than those when the patient continues antibiotics. For mortality, no impact was observed by stopping or extending antibiotic treatment. Estimations had modest performance with an average AUROC of 0.67 and accuracy of 0.82; however, the model clearly struggled with falsenegative predictions.

Estimations were made for each day of each patient's stay within all the extracted data (i.e., train, validation, and testing sets combined) to understand if results would deviate by dataset size. For LOS, reliable estimations were once again obtained (mean stop control difference of 0.33 days and mean continue control difference of 0.42 days). Continuing showed no given impact (mean difference of -0.30 days), while stopping once again showed a significant impact with a mean reduction of 1.87 days. Little difference in mortality estimations was seen between stop and continue controls and impacts (stop impact-0.03, stop control-0.03, continue control-0.05, continue impact-0.05). Mortality predictions were relatively reliable with a mean AUROC of 0.72.

To show the importance of the temporality in our predictions, we created estimations for each antibiotic day of each patients stay, without segregating the embedding space (by time or by antibiotic treatment given they are mutually dependent). The resulting estimations had a mean LOS difference of 2.60 days from the true labels, an RMSE of 5.05, and a statistically significant difference in medians (*p*-value <0.01).

The performance of the model on subjects towards the edges of the distribution in terms of the correlation between LOS and overall antibiotic treatment length was investigated. Subjects in the 10th and 90th percentiles were selected leading to a smaller Spearman's correlation of 0.35. As expected, given the dataset size (n=686) and donor distribution, results were quite poor with a mean stop control difference of 2.92 days and a mean continue control difference of 2.13 days. The impact of stopping early though was still much greater than the control at 4.36 days mean difference.

Pneumonia and UTIs

A total of 2,473 stays where patients were diagnosed with pneumonia were identified, with a mean LOS of 9.05 days and a mean antibiotic treatment length of 6.95 days. Overall estimation of the results on this whole pneumonia population reflected the wider dataset and are shown in **Table 3**. When focusing on those with long treatment durations and the question of what if they stopped after 7 days of treatment, statistically significant results show that average LOS were 2.82 days shorter when stopping earlier. No difference in estimated mortality was observed; however, estimations were consistent across groups with an average AUROC of 0.75. No significant difference in LOS or mortality was estimated for pneumonia patients who received the standard of care 7 days treatment, if they had slightly shorter treatment durations of 5 or 6 days.

For UTIs, 923 patient stays were selected having a mean LOS and antibiotic treatment length of 5.50 and 4.77 days respectively. Once again, overall estimation results (Table 3) were similar to previous findings with trustworthy controls, stopping early being associated with a shorter LOS and no

		LOS				Mortality			
Scenario	Day(s)	Mean delta (days, <i>p</i> -value)	MAPE	MAE	RMSE	Mean delta	MAE	AUROC	
Stop	Impact	2.71*, <0.01	0.36	3.30	4.80	0.06	0.25	0.66	
	Control	0.24, 0.60	0.26	1.32	1.93	0.05	0.15	0.72	
	Impact	-2.09*, <0.01	0.77	2.85	3.16	0.05	0.18	0.67	
	Control	0.42*, 0.01	0.48	2.72	3.76	0.07	0.24	0.64	

TABLE 2 Outcome estimation results for patients in the unseen test set.

*Statistical significance with alpha set at 0.05.

LOS, length of stay; MAPE, mean absolute percentage error; MAE, mean absolute error; RMSE, root mean squared error; AUROC, Area Under the Receiver Operating Characteristic curve.

TABLE 3 Outcome estimation results for patients with pneumonia and UTIs.

Infection		Scenario	Day (s)	LOS	Mortality		
	Analysis			Mean delta (days, <i>p</i> -value)	RMSE	Mean delta	AUROC
Pneumonia	Whole dataset	Stop	Impact	3.72*, <0.01	5.87	0.00	0.71
			Control	0.26, 0.47	2.14	0.07	0.76
		Continue	Impact	-2.79*, <0.01	3.65	0.10	0.69
			Control	0.49*, <0.01	4.01	0.05	0.68
	Long treatment durations stopping after 7 days	Stop	Impact	2.82*, <0.01	4.65	-0.03	0.74
			Control	0.43, 0.08	2.11	0.05	0.80
		Continue	Impact	_	_	_	_
			Control	0.41, 0.21	3.47	0.05	0.73
UTI	Whole dataset	Stop	Impact	2.36*, <0.01	4.70	0.14	0.63
			Control	0.36, 0.89	2.04	0.07	0.87
		Continue	Impact	-1.91*, <0.01	3.26	0.03	0.79
			Control	0.38, 0.05	3.82	0.04	0.71
	Long treatment durations stopping after 7 days	Stop	Impact	2.08*, <0.01	4.35	0.30	0.52
			Control	1.04, 0.23	2.42	0.17	0.93
		Continue	Impact	_	_	_	_
			Control	0.26, 0.05	3.48	0.05	0.76

Results are shown for both the whole population and analysis of what if those who received long treatment durations stopped after day 7.

*Statistical significance with alpha set at 0.05.

LOS, length of stay; RMSE, root mean squared error; AUROC, Area Under the Receiver Operating Characteristic curve; UTI, urinary tract infection.

difference in mortality but reliable estimations (AUROC ranging from 0.63 to 0.87). Estimations for stopping after 7 days for those with long treatment durations did show a positive impact in terms of reduced LOS (mean difference 2.08 days, p-value <0.01). The stop control where we expect to see minimal difference showed a larger mean deviation of 1.04 days, but statistical analysis showed the medians between control estimations and labels were not significantly different. Mortality estimations here were for the most part dependable; a high predictive performance on stop and continue controls was achieved with an AUROC of 0.93 and 0.78, respectively, but a lower score for the stop impact of 0.52. When analysing those patients who received the standard of care 7 days treatment, for slightly shorter treatment durations (5 or 6 days). A statistically significant result was observed where estimated LOS outcomes were on average longer by 1.45 days if the patients stopped antibiotics slightly earlier (p-value <0.01, RMSE 2.72).

Discussion

We demonstrate that our RNN autoencoder and synthetic control-based approach trained on a large ICU EHR dataset can estimate patient outcomes under the alternative scenarios of stopping vs. continuing antibiotic treatment. Results across experiments were consistent, with stop control days often showing the greatest performance indicating our stop impact estimations, which occur on days where the true outcome upon stopping is unknown, are more reliable. The stop impact results from this retrospective study show that stopping antibiotics earlier can be associated with a statistically significant average LOS reduction of 2.71 days. Overall minimal impact on mortality was observed, which is to be expected given death can be caused by a large number of factors beyond those included as model features. Figure 5 shows some specific illustrative examples of patient LOS and mortality estimations. The pneumonia dataset demonstrated



particularly positive results with overall and stopping on day 7 analysis indicating antibiotic cessation can have a significant impact on LOS in this population (mean difference 3.72 and 2.82 days, respectively). This reflects current clinical thinking that shorter treatments are optimal for this infection (19, 36, 37, 41). However, there is a balance to be made with antibiotic treatment durations. The UTI analysis indicated courses shorter than 7 days may be detrimental to the patient and that the current standard of care treatment duration is likely appropriate. As such, care must be taken to consider the patients and the public's best interests with respect to current infections and the threat of AMR.

Our methodological approach to the problem of antibiotic cessation is novel. This model can in principal assist with individualised antibiotic cessation decisions as it takes into account numerous patient characteristics and the specific treatment scenario with regards to patient outcomes, factors that previously could not be considered together in their entirety. This study has approached the problem of antibiotic cessation from the perspective of making a clinically useful tool designed to support decision-making by estimating direct measures that may influence clinical decision-making under

alternative scenarios. We believe it could be useful for prescribing physicians during their daily clinical round to compare between stop and continue estimated outcomes and understand when it is appropriate to cease antibiotic treatment. In particular, this system should help show shorter treatment durations can be safe and support individualised antimicrobial decision-making through hard outcome estimation. From a behaviour change perspective, this approach may provide reassurance to support early cessation of therapy, while promoting improved knowledge and understanding on the issue of antimicrobial optimisation and stewardship (43). It should be noted though that too short a course of antibiotics can cause harm and have negative knock-on effects. As such, the aim of this research is to optimise antimicrobial use and determine the most appropriate antibiotic treatment duration for each individual patient. One significant outstanding question is how clinicians treating a patient would adopt recommendations provided by such a system and if it would influence antimicrobial clinical decision-making. Holistically, we believe antibiotic cessation is a collective, data-driven decision, meaning a CDSS in this area can have a larger influence and acceptance by end users.

However, the degree to which this tool would be accepted and work alongside clinical decision-making behaviour requires investigation.

We have shown that our model is able to reliably estimate alternative patient outcomes depending on their antibiotic treatment status. Based on our results, the size and consistency of the dataset used and, hence, the number of available donors are strongly related to the reliability of outputs. Experiments utilising small datasets often led to poor results given there were not enough suitable patients within a given embedding space to create an appropriate synthetic estimation. On the other hand, there does seem to be a ceiling above which more instances are not necessary. For example, similar results were obtained across the pneumonia, test, and whole datasets even though they had sizes of 2,476, 3,427, and 22,845 patient stays, respectively. As such, we can infer that this method is likely to produce suitable estimations if several thousand patient examples are available. Although this should be reasonable for most clinical scenarios, it does act as a dataset constraint when evaluating less common infections, where potentially more interesting nuanced findings could be made.

The quality of the initial autoencoder model is another significant implication that determines performance. The standard autoencoder model without the synthetic control methods applied achieved higher performance on the LOS prediction task than estimations generated without segregating the embedding space (RMSE of 3.88 and 5.05, respectively). This confirms first that the model has been trained to appropriately represent the patient in the embedding space with respect to their outcome. Second, the temporal aspect of the embeddings assists with synthetic outcome estimations and finally the subsequent synthetic outcome estimation methodology applied ensures that outputs can be clinically applicable with regards to antibiotic treatment. As such, the autoencoder is critical for appropriate temporal representations and subsequent estimations.

It is important to note that there is a high degree of correlation between LOS and overall treatment length in the datasets (Table 1, supplementary Figure S3). This is to be expected given those patients who are less sick will likely receive fewer antibiotics and leave the ICU sooner. Although the model architecture is designed to account for this, through representative and segregated embeddings, it is still likely that the model "learned" this association causing some confounding. Results on outliers when there is reduced correlation still illustrate that stopping can impact LOS outcomes, even if the predictions themselves are not reliable in this situation given the skewed dataset analysed. Numerous factors influence ICU LOS; hence, even if the model predicts that stopping antibiotics could be neutral or beneficial, other random factors may make this an impossibility. Nevertheless, our results and the strong correlation observed between

antibiotic treatment length and LOS in this dataset mean this model can act as a proxy with the ultimate aim of reducing the unnecessary use of antibiotics.

This study has several limitations. We focused on addressing what would happen if antibiotic cessation occurred earlier during a patient's ICU stay. The synthetic control methodology was chosen and adapted as it allows us to address this problem while more traditional causal discovery seems intractable. MAPE and MAE LOS estimation results are in the region of days which could limit clinical utility but are comparable to that of recent research (44). Unlike most synthetic control applications, we do not have an extensive pre-intervention period making confidence in results more challenging. Furthermore, one of our analogues stop "control" days would not be available on a patient-specific level during clinical use due to the nature of cessation occurring after treatment. Other types of interpretability such as being able to investigate selected donors to see if they are clinically meaningful could counteract this. Second, the use of historical EHR data to estimate the synthetic outcome means all our estimations are biased based on past antibiotic prescribing policies. These methodological approaches were required to answer our question of interest but mean that historical approaches towards antimicrobial stewardship govern our model's outputs. The analysis of such a large dataset along with estimations being the weighted average of donors does, however, mitigate this to some extent. In conjunction with this, the analysis presented here is of a macro-scale; however, to realise the potential of this approach for true antimicrobial optimisation, more nuanced, relative, and individualised studies will be required, which we plan to conduct in future. Finally, given the high degree of missingness in the dataset, a number of clinically important features have to be excluded. In particular, research shows that procalcitonin (PCT) and C-reactive protein (CRP) are useful biomarkers for determining when it is safe and appropriate to stop antibiotic therapy (45-48). Neither of these were included as features due to insufficient data. As such, this approach and the subsequent results could potentially be more powerful if applied to a complete dataset focused on a narrow type of infection with defined variables of interest.

In conclusion, we have developed an AI-driven model to estimate patient outcomes if they were to stop or continue antibiotic treatment in the ICU. With further development into a CDSS, we envisage that this can assist clinicians with antimicrobial optimisation and reduce the excessive use of antibiotics to tackle AMR. Future research will investigate which variables promote or hinder cessation and discern the ability of this tool to influence antimicrobial decisionmaking.

Data availability statement

Publicly available datasets were analysed in this study. These data can be found here: https://physionet.org/content/mimiciv/ 1.0/.

Author contributions

WB, TR, BH, RW, and DA contributed to study concept and design. WB and BH contributed to data acquisition. WB, BH, and TR contributed to data analysis and accessed and verified the underlying data. WB, TR, and BH contributed to the initial manuscript drafting, discussion of the results, and review of the data. All authors contributed to data interpretation and final revisions of the manuscript. DA, PG, and AH contributed to study supervision. All authors contributed to the article and approved the submitted version.

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References

1. Nations U. Political declaration of the high level meeting of the general assembly on antimicrobial resistance: draft resolution/submitted by the president of the general assembly New York: UN (2016) 6 p.

2. World Health Organization. Global action plan on antimicrobial resistance. World Health Organization (2015) 28 p.

3. Murray CJ, Ikuta KS, Sharara F, Swetschinski L, Aguilar GR, Gray A, et al. Global burden of bacterial antimicrobial resistance in 2019: a systematic analysis. *Lancet.* (2022) 399:629–55. doi: 10.1016/S0140-6736(21)02724-0

4. Rawson TM, Moore LSP, Hernandez B, Charani E, Castro-Sanchez E, Herrero P, et al. A systematic review of clinical decision support systems for antimicrobial management: are we failing to investigate these interventions appropriately? *Clin Microbiol Infect.* (2017) 23:524–32. doi: 10.1016/j.cmi.2017.02.028

5. Hernandez B, Herrero P, Rawson TM, Moore LSP, Evans B, Toumazou C, et al. Supervised learning for infection risk inference using pathology data. *BMC Med Inform Decis Mak.* (2017) 17:168. doi: 10.1186/s12911-017-0550-1

6. Rawson TM, Hernandez B, Moore LSP, Blandy O, Herrero P, Gilchrist M, et al. Supervised machine learning for the prediction of infection on admission to hospital: a prospective observational cohort study. *J Antimicrob Chemother*. (2019) 74:1108–15. doi: 10.1093/jac/dky514

7. Rawson TM, Hernandez B, Wilson RC, Ming D, Herrero P, Ranganathan N, et al. Supervised machine learning to support the diagnosis of bacterial infection

for Antimicrobial Optimisation (CAMO) at Imperial College London. The views expressed in this publication are those of the authors and not necessarily those of the NHS, the National Institute for Health Research or the UK Department of Health.

Conflict of interest

TR was employed by Sandoz (2020), Roche Diagnostics Ltd (2021), and bioMerieux (2021–2022). These commercial entities were not involved in the study design, collection, analysis, interpretation of data, the writing of this article, or the decision to submit it for publication. All authors declare no other competing interest.

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Supplementary material

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fdgth. 2022.997219/full#supplementary-material.

in the context of COVID-19. JAC-Antimicrob Resist. (2021) 3:dlab002. doi: 10. 1093/jacamr/dlab002

8. Hernandez B, Herrero-Viñas P, Rawson TM, Moore LSP, Holmes AH, Georgiou P. Resistance trend estimation using regression analysis to enhance antimicrobial surveillance: a multi-centre study in London 2009–2016. *Antibiotics*. (2021) 10:1267. doi: 10.3390/antibiotics10101267

9. Hernandez B, Herrero P, Rawson T, Moore L, Charani E, Holmes A, et al. Data-driven web-based intelligent decision support system for infection management at point-of-care: case-based reasoning benefits and limitations. In Proceedings of the 10th International Joint Conference on Biomedical Engineering Systems and Technologies – HEALTHINF, (BIOSTEC 2017). (2017). p. 119-27.

10. Rawson TM, Hernandez B, Moore LSP, Herrero P, Charani E, Ming D, et al. A real-world evaluation of a case-based reasoning algorithm to support antimicrobial prescribing decisions in acute care. *Clin Infect Dis.* (2021) 72:2103–11. doi: 10.1093/cid/ciaa383

11. Tamma PD, Miller MA, Cosgrove SE. Rethinking how antibiotics are prescribed: incorporating the 4 moments of antibiotic decision making into clinical practice. *JAMA*. (2019) 321:139–40. doi: 10.1001/jama.2018.19509

12. Langford BJ, Morris AM. Is it time to stop counselling patients to "finish the course of antibiotics"? *Can Pharm J.* (2017) 150:349–50. doi: 10.1177/1715163517735549

13. Holmes AH, Moore LSP, Sundsfjord A, Steinbakk M, Regmi S, Karkey A, et al. Understanding the mechanisms, drivers of antimicrobial resistance. *Lancet.* (2016) 387:176–87. doi: 10.1016/S0140-6736(15)00473-0

14. Spellberg B. The new antibiotic mantra—"shorter is better". JAMA Intern Med. (2016) 176:1254–5. doi: 10.1001/jamainternmed.2016.3646

15. Curran J, Lo J, Leung V, Brown K, Schwartz KL, Daneman N, et al. Estimating daily antibiotic harms: an umbrella review with individual study meta-analysis. *Clin Microbiol Infect.* (2022) 28:479–90. doi: 10.1016/j.cmi.2021.10.022

16. Vaughn VM, Flanders SA, Snyder A, Conlon A, Rogers MA, Malani AN, et al. Excess antibiotic treatment duration and adverse events in patients hospitalized with pneumonia. *Ann Intern Med.* (2019) 171:153–63. doi: 10. 7326/M18-3640

17. Spellberg B, Rice LB. Duration of antibiotic therapy: shorter is better. Ann Intern Med. (2019) 171:210–1. doi: 10.7326/M19-1509

18. Yahav D, Franceschini E, Koppel F, Turjeman A, Babich T, Bitterman R, et al. Seven versus 14 days of antibiotic therapy for uncomplicated gramnegative bacteremia: a noninferiority randomized controlled trial. *Clin Infect Dis.* (2019) 69:1091–8. doi: 10.1093/cid/ciy1054

19. Royer S, DeMerle KM, Dickson RP, Prescott HC. Shorter versus longer courses of antibiotics for infection in hospitalized patients: a systematic review and meta-analysis. J Hosp Med. (2018) 13:336–42. doi: 10.12788/jhm.2905

20. Wald-Dickler N, Spellberg B. Short-course antibiotic therapy-replacing Constantine units with "shorter is better". *Clin Infect Dis.* (2019) 69:1476-9. doi: 10.1093/cid/ciy1134

21. Hanretty AM, Gallagher JC. Shortened courses of antibiotics for bacterial infections: a systematic review of randomized controlled trials. *Pharmacotherapy*. (2018) 38:674–87. doi: 10.1002/phar.2118

22. Janssen RME, Oerlemans AJM, Van Der Hoeven JG, Ten Oever J, Schouten JA, Hulscher MEJL. Why we prescribe antibiotics for too long in the hospital setting: a systematic scoping review. *J Antimicrob Chemother*. (2022) 77(8): dkac162. doi: 10.1093/jac/dkac162

23. Charani E, McKee M, Ahmad R, Balasegaram M, Bonaconsa C, Merrett GB, et al. Optimising antimicrobial use in humans: review of current evidence and an interdisciplinary consensus on key priorities for research. *Lancet Reg Health Eur.* (2021) 7:100161. doi: 10.1016/j.lanepe.2021.100161

24. Peiffer-Smadja N, Rawson TM, Ahmad R, Buchard A, Georgiou P, Lescure FX, et al. Machine learning for clinical decision support in infectious diseases: a narrative review of current applications. *Clin Microbiol Infect.* (2020) 26:584–95. doi: 10.1016/j.cmi.2019.09.009

25. Pandolfo AM, Horne R, Jani Y, Reader TW, Bidad N, Brealey D, et al. Understanding decisions about antibiotic prescribing in ICU: an application of the Necessity Concerns Framework. *BMJ Qual Saf.* (2022) 31:199–210. doi: 10. 1136/bmjqs-2020-012479

26. Rawson TM, Charani E, Moore LSP, Hernandez B, Castro-Sánchez E, Herrero P, et al. Mapping the decision pathways of acute infection management in secondary care among UK medical physicians: a qualitative study. *BMC Med.* (2016) 14:208. doi: 10.1186/s12916-016-0751-y

27. Johnson A, Bulgarelli L, Pollard T, Horng S, Celi LA, Mark R. MIMIC-IV (2021)[Dataset]. doi: 10.13026/s6n6-xd98

28. Goldberger AL, Amaral LAN, Glass L, Hausdorff JM, Ivanov PC, Mark RG, et al. PhysioBank, PhysioToolkit, and PhysioNet. *Circulation*. (2000) 101:e215–20. doi: 10.1161/01.CIR.101.23.e215

29. Qian Z, Zhang Y, Bica I, Wood A, van der Schaar M. SyncTwin: treatment effect estimation with longitudinal outcomes. In: *Advances in Neural Nnformation Processing Systems 34 (NeurIPS 2021).* Vol. 34. Vancouver Canada: Curran Associates, Inc. (2021). p. 3178–3190.

30. Abadie A, Gardeazabal J. The economic costs of conflict: a case study of the Basque country. Am Econ Rev. (2003) 93:113–32. doi: 10.1257/000282803321455188

31. Bouttell J, Craig P, Lewsey J, Robinson M, Popham F. Synthetic control methodology as a tool for evaluating population-level health interventions. *J Epidemiol Community Health.* (2018) 72:673–8. doi: 10.1136/jech-2017-210106

32. Kreif N, Grieve R, Hangartner D, Turner AJ, Nikolova S, Sutton M. Examination of the synthetic control method for evaluating health policies with multiple treated units. *Health Econ.* (2016) 25:1514–28. doi: 10.1002/hec.3258

33. Paszke A, Gross S, Massa F, Lerer A, Bradbury J, Chanan G, et al. Pytorch: an imperative style, high-performance deep learning library. In: *Advances in Neural Information Processing Systems 32 (NeurIPS 2019)*. Vol. 32. Vancouver Canada: Curran Associates, Inc. (2019). p. 8024–8035.

34. Kingma DP, Ba J. Adam: a method for stochastic optimization (2014). Available from: https://arxiv.org/abs/1412.6980

35. Liaw R, Liang E, Nishihara R, Moritz P, Gonzalez JE, Stoica I. Tune: A research platform for distributed model selection and training [Preprint] (2018). Available at: http://arxiv.org/1807.05118.

36. Dimopoulos G, Poulakou G, Pneumatikos IA, Armaganidis A, Kollef MH, Matthaiou DK. Short- vs long-duration antibiotic regimens for ventilatorassociated pneumonia: a systematic review and meta-analysis. *Chest.* (2013) 144:1759–67. doi: 10.1378/chest.13-0076

37. Pugh R, Grant C, Cooke RPD, Dempsey G. Short-course versus prolongedcourse antibiotic therapy for hospital-acquired pneumonia in critically ill adults. *Cochrane Database Syst Rev.* (2015) (8):CD007577. doi: 10.1002/14651858. CD007577.pub3

38. Drekonja DM, Trautner B, Amundson C, Kuskowski M, Johnson JR. Effect of 7 vs 14 days of antibiotic therapy on resolution of symptoms among afebrile men with urinary tract infection: a randomized clinical trial. *JAMA*. (2021) 326:324–31. doi: 10.1001/jama.2021.9899

39. de Gier R, Karperien A, Bouter K, Zwinkels M, Verhoef J, Knol W, et al. A sequential study of intravenous and oral fleroxacin for 7 or 14 days in the treatment of complicated urinary tract infections. *Int J Antimicrob Agents.* (1995) 6:27–30. doi: 10.1016/0924-8579(95)00011-V

40. Peterson J, Kaul S, Khashab M, Fisher AC, Kahn JB. A double-blind, randomized comparison of levofloxacin 750 mg once-daily for five days with ciprofloxacin 400/500 mg twice-daily for 10 days for the treatment of complicated urinary tract infections and acute pyelonephritis. *Urology*. (2008) 71:17–22. doi: 10.1016/j.urology.2007.09.002

41. National Institute for Health and Care Excellence. *Pneumonia (hospital-acquired): antimicrobial prescribing NICE guideline [NG139].* (2019). Available from: https://www.nice.org.uk/guidance/ng139

42. National Institute for Health and Care Excellence. Urinary tract infection (lower): antimicrobial prescribing NICE guideline [NG109]. (2018). Available from: https://www.nice.org.uk/guidance/ng109

43. Pauwels I, Versporten A, Vermeulen H, Vlieghe E, Goossens H. Assessing the impact of the Global Point Prevalence Survey of Antimicrobial Consumption and Resistance (Global-PPS) on hospital antimicrobial stewardship programmes: results of a worldwide survey. *Antimicrob Resist Infect Control.* (2021) 10:138. doi: 10.1186/s13756-021-01010-w

44. Rocheteau E, Liò P, Hyland S. Temporal pointwise convolutional networks for length of stay prediction in the intensive care unit. *Proceedings of the Conference on Health, Inference, and Learning, CHIL '21.* 2021 April 8 – 10; New York, NY: Association for Computing Machinery (2021). p. 58–68. Available at: https://doi.org/10.1145/3450439.3451860

45. Schuetz P, Wirz Y, Sager R, Christ-Crain M, Stolz D, Tamm M, et al. Procalcitonin to initiate or discontinue antibiotics in acute respiratory tract infections. *Cochrane Database Syst Rev.* (2017) 2017:CD007498. doi: 10.1002/ 14651858.CD007498.pub3

46. Rhee C. Using procalcitonin to guide antibiotic therapy. Open Forum Infect Dis. (2016) 4:ofw249. doi: 10.1093/ofid/ofw249

47. Oliveira CF, Botoni FA, Oliveira CRA, Silva CB, Pereira HA, Serufo JC, et al. Procalcitonin versus C-reactive protein for guiding antibiotic therapy in sepsis: a randomized trial. *Crit Care Med.* (2013) 41:2336–43. doi: 10.1097/CCM. 0b013e31828e969f

48. Coelho L, Póvoa P, Almeida E, Fernandes A, Mealha R, Moreira P, et al. Usefulness of C-reactive protein in monitoring the severe community-acquired pneumonia clinical course. *Crit Care.* (2007) 11:R92. doi: 10.1186/cc6105