



# RETRACTED: Revealing the Impacts of Banking Sector Development on Renewable Energy Consumption, Green Growth, and Environmental Quality in China: Does Financial Inclusion Matter?

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China aims to reduce its carbon dioxide emissions and promote green growth. This study aims to examine the effect of banking sector performance indicators (banks assets and return on asset) and financial inclusion on renewable energy consumption, green growth, and carbon emissions for China from 1995 to 2020 using the ARDL approach. The long-run results suggest that bank assets increase renewable energy consumption and green growth. While return on assets also increases green growth and decreases carbon emission in the long run. Financial inclusion enhances renewable energy consumption and green growth, curbing CO<sub>2</sub> emissions. Banking sector performance and financial inclusion have short-run effects on renewable energy consumption, green growth, and carbon dioxide emissions. The findings thus point to the need for policies that promote banking sector performance and financial inclusion to boost green growth and alleviate CO<sub>2</sub> emissions.

**Keywords:** banking sector development, renewable energy consumption, green growth, CO<sub>2</sub> emissions, Pakistan

## INTRODUCTION

Indeed, economic growth is crucial for reducing poverty, raising the people's living standards, and developing the necessary infrastructure to provide health and educational facilities to most people. Energy is the key element in the growth of the economy (Ahmad et al., 2019). Economic development, being the potential of an economy, is the most viable factor observed by the world's nations (Khan et al., 2022). Economies grow over time, whereas the initial state of growth usually remains slow (Sun et al., 2020). Economic development is the potential of any economy. Nevertheless, economic development has some social costs attached to it in the form of environmental degradation due to the overutilization of natural resources (Ullah et al., 2021a). Although China has achieved unprecedented economic development in the past few decades, Chinese people have witnessed an enormous increase in living standards. Still, China has become the world's leading carbon emitter (Sun et al., 2020). The banking sector is necessarily required to upgrade and reduce the CO<sub>2</sub> emission as the banking performance affects heterogeneously to the banking performance which categories the measure of renewable energy consumption (Zhang et al., 2021).

To increase the energy efficiency, it is necessarily to improve the size of the bank, the quality of the asset, and the efficient management of the HI nations (Gyimah et al., 2022). A detailed description of

the term “inclusive development” is hypothetical (Mohsin et al., 2019). As an effect, the idea is correspondently used with associated thoughts such as “inclusive growth.” this leads to significant uncertainty in the literature and policy agendas from several organizations, administrations, and the country’s government (Pasimeni & Pasimeni, 2016). In each field, the environmental characteristic of any plan and strategy is a vital topic to address. Climate change has brought serious consequences for the Earth’s ecological balance, such as rising temperature, melting glaciers, droughts, floods, extreme weather, food insecurity, and deterioration of the worker’s efficiency (Solarin et al., 2017; Wei & Ullah, 2022).

Study of Zakarya et al. (2015) investigated the dynamic factors affecting CO<sub>2</sub> emissions in the 12 Middle East countries using panel data from 1990 to 2009. CO<sub>2</sub> used, and trade is used affect the financial up-gradation, and there is a bidirectional relationship both in the short and long run. Therefore, the issue of climate change and its detrimental impact on humanity has got paramount importance in the eyes of empirics, policymakers, and civil society. Due to the significance of the issue of climate change, environmental economists and policymakers have shifted their focus from a traditional growth concept to a more sophisticated, modern, sustainable one. The environmentally friendly notion of growth is green growth (Li et al., 2022). From 1990 to 2018, inflation instability is linked to pollution and carbon emissions in 40 countries. An increase in the price variety affects investment projects and consumption (Mohsin et al., 2019). It has resulted in economic growth, increasing the breakdown and causing the environment lower condition. To improve the environment, financial stability is essential (Al Asbahi et al., 2019).

The continuous degradation of the environment due to rising temperature and its harmful impact on human lives has made the goal of sustainable development a top focus of the international community. As a result, the policymakers in many countries have tried to implement various environmental strategies and policies that can address the issue of climate change. The Paris Pact and Sustainable Development Goals (SDGs) can be termed practical steps toward improving the environment (OECD, 2020). According to the OECD (2020), the SDGs have become a crucial objective for policymakers in developed and developing economies (Ahmad et al., 2019; Mohsin et al., 2019). However, the steps taken by the various countries in this regard are quite different so as their outcomes. For instance, due to various policy measures, Australia and Belgium are enjoying high green growth compared to Portugal and Turkey, facing stagnant green growth (Ahmad et al., 2021).

In this age of increasing globalization and modernization, when the environmental problem is multi-dimensional, cannot deal with this problem with one size fit all strategy; hence, it required to examine different factors that can help decouple economic growth from CO<sub>2</sub> emissions. Study suggested that green growth is a type of growth that can be achieved through efficient and sustainable use of natural resources and capital. Green growth represents the path towards a more green and sustainable economy (Iram et al., 2020). Green growth is a process of utilizing natural resources sustainably (Dercon,

2014). The ultimate goal is to raise people living standards, reduce inequality and poverty, develop infrastructure without exerting an extra burden on the environment, and protect the environment for future generations.

On one side, green economic growth has become the norm of the new environmental strategy adopted by almost every nation (Dercon, 2014; Linster & Yang, 2018). On the other side, traditional economic activities are considered the primary drivers of CO<sub>2</sub> emissions. Therefore, literature is enriching as far as the determinants of environmental quality are concerned. Most studies have confirmed that economic growth adversely affects the environmental quality at the early stages and improves ecological quality (Stern, 2004). Although the literature on the nexus between economic growth and environmental quality is still flourishing, the inclusion of new variables has cleared many questions in the mind of policymakers about this relationship. One such variable is financial development which has become a primary driver of economic growth (Li et al., 2022). However, its impact on environmental quality is underexplored. Development of the financial sector, represented mainly by banking sector development, may influence energy-related emissions positively or negatively.

On the one hand, the development of the banking sector may promote technological innovations in the energy sector and help to reduce emissions. However, the banking sector also fosters industrial and manufacturing activities that increase energy-driven emissions (Amuakwa-Mensah & Näsström, 2022). Moreover, the research and development activities also flourish due to progress in the financial sector which spurs economic activities and consequently impacts environmental quality (Frankel and Romer, 1999). Lie et al. (2021) used data from China from 1970 to 2015 to estimate the linkage between CO<sub>2</sub> emission per capita, energy power, real economic growth (per capita), industrial sectors, urban rate of population, and renewable energy consumption. Analysis of this study (ARDL) was used for both short-run and long-run (VECM) to find causality. Renewable energy and GDP had positive and unidirectional causality relations in the short-run, while other variables had bidirectional causality. In this study, the researcher gave policy suggestions to improve the environment, reduce carbon emissions, and improve energy use and banking efficiency.

Zhang et al. (2019) used panel data from 1996 to 2017 to estimate the carbon dioxide emission in BRICS countries. This study estimated short termed and long termed data was analysed and shows that using variables like; per capita GDP, energy structure (ratio of renewable energy consumption to total energy consumption), and economic structure (value-added service industry to GDP and contribution service industry to economic growth). The CO<sub>2</sub> correlated effect mean group (CCEMG) was used to find the relationship between economic development and carbon emissions. The results according to EKC are inverted U-shape relation in GDP with CO<sub>2</sub> and CCEMG method; reduction in carbon dioxide by an increase in renewable energy consumption. The banking performance affects heterogeneously the banking performance, which categories the measure of renewable energy consumption. To increase

the energy efficiency it is necessarily to increase the size of bank, the quality of asset, efficiency management of the HI nations.

Although banking sector development has become a crucial variable in determining the environmental quality and economic growth (Yang et al., 2020). Financial inclusion has become an integral part of the development of the financial sector (Ullah et al., 2021b). Financial inclusion implies the easy availability of a wide variety of financial products and services to a majority of people (Ullah et al., 2020).

An inclusive and vibrant financial sector provides easy credit facilities for investment in green energy technologies that can significantly improve environmental quality (IPA, 2017). Conversely, an increase in financial inclusiveness helps the economy to achieve scale effect through increased production and manufacturing activities that contribute to environmental degradation. Similarly, an inclusive financial system allows individuals to consume more energy-intensive products, infusing carbon emissions into the atmosphere (Ahmad et al., 2021).

In the light of the above discussion, it is imperative to include the variables of financial development and financial inclusion as determinants of environmental quality, green growth, and renewable energy consumption. This is the first-ever study that has tried to capture the effect of banking sector development and financial inclusion on China's environmental quality, green growth, and renewable energy consumption. Although previous studies have included one or the other variable related to the financial sector, none have observed the collective impact of banking sector development and financial inclusion on the CO<sub>2</sub> emissions, green growth, and renewable energy consumption in the context of China. The choice of China as a country is not by chance but instead based on the fact that China is the largest consumer of energy and the largest emitter of carbon in the world.

## Financial Inclusion Review

Financial inclusion development has increased in the early 1990s with development and technological revolution on policy-making for economically deprived people (Leyshon and Thrift 1995; Sanderson et al., 2018). Financial inclusion is the easy accumulation of capital, security, and a safe way to save, and it is a significant driver of economic growth (Claessens and Perotti 2007). Women will play a larger role in banking if families and companies have simple access. Such actions boost economic growth (Sahay et al., 2015). Financial inclusion can increase economic growth through small business value creation, which improves human capital through improving health and education, which is vital for economic development, as well as reducing inequality and poverty (Agnello et al., 2012). Sarma (2008) created an index based on several characteristics to calculate financial inclusion for a group of countries. Arora (2010) used a similar method for developed and developing economies. The researchers looked at economic growth and financial inclusion from multiple aspects. Sethi & Acharya, (2018) stated dynamic, resilient, and efficient financial sector drives economic growth. If private enterprises have easy access to credit and financial services, business would boost per capita income (Kiplimo et al., 2015).

Mohan (2006) noted that improving the role of financial intermediaries boosts economic growth and financial development. Estrada et al. (2010) found a relationship between poverty and financial access. Donou-Adonsou and Sylwester (2016) found microfinance hurts poverty in emerging countries. The banking sector decreases poverty in the same economies. Research on the effects of financial inclusion on economic growth and other variables is extensive, but research on the environment is in its infancy. Li et al. (2022) collected data from 31 countries between 2004 and 2014 to study the association between financial inclusion and CO<sub>2</sub> emissions.

Renzhi and Baek (2020) found a negative link between financial inclusion and CO<sub>2</sub> emissions. Financial inclusion may improve environmental quality by lowering economic expansion's negative effects. Zhang et al. (2019) examined the impact of financial inclusion on CO<sub>2</sub> emissions in OECD nations from 2004 to 2014 and found that it reduces CO<sub>2</sub> emissions in some economies in the short and long term. Increasing financial inclusion in China lowered CO<sub>2</sub> emissions. Therefore this study attempted to capture the effect of banking sector development and financial inclusion on China's environmental quality, green growth, and renewable energy consumption. Although previous studies have included one or the other variable related to the financial sector, none have observed the collective impact of banking sector development and financial inclusion on the CO<sub>2</sub> emissions, green growth, and renewable energy consumption in the context of China. The choice of China as a country is not by chance but instead based on the fact that China is the largest consumer of energy and the largest emitter of carbon in the world.

## Model and Methods

Theoretically, banking sector development works as a channel of green credit and financial services. Renewable energy consumption impacts green growth and environmental quality (Amuakwa-Mensah & Näsström 2022). Empirically, banking sector development has increased green growth and environmental quality by increasing renewable energy consumption (Cao et al., 2021; Amuakwa-Mensah & Näsström, 2022). Following the standard earlier studies (Samour et al., 2019; Zafar et al., 2019; Amuakwa-Mensah & Näsström, 2022), renewable energy consumption, green growth, CO<sub>2</sub> emissions are given as a function of banking sector development and financial inclusion as follows in Eq. 1.

$$REC_t = Y_{10} + Y_{11}BSP_t + Y_{12}FI_t + Y_{13}ICT_t + Y_{14}FDI_t + Y_{15}Trade_t + \mu_{1t} \quad (1)$$

$$GG_t = Y_{20} + Y_{21}BSP_t + Y_{22}FI_t + Y_{23}ICT_t + Y_{24}FDI_t + Y_{25}Trade_t + \mu_{2t} \quad (2)$$

$$CO_{2,t} = Y_{30} + Y_{31}BSP_t + Y_{32}FI_t + Y_{33}ICT_t + Y_{34}FDI_t + Y_{35}Trade_t + \mu_{3t} \quad (3)$$

In the above specification Eqs 1–3, renewable energy consumption (REC), green growth (GG), and CO<sub>2</sub> emissions (CO<sub>2</sub>) are determined by the banking sector performance (BSP),

financial inclusion (FI), information and communication technologies (ICTs), foreign direct investment, inflow (FDI), trade openness (Trade), and random error term ( $\epsilon_t$ ). Based on the standard literature, we expect an estimate of  $\Upsilon_{11}$  and  $\Upsilon_{21}$  to be positive if the banking sector may enable customers and firms to provide low rate funds for renewable energy consumption and green technologies. While we expect the estimate of banking sector development elasticity's to be negative in CO<sub>2</sub> emissions. The analysis is not limited to the long-run estimates only; hence, the above long-run equation is respecified in the form of an error correction format given below:

$$\begin{aligned} \Delta REC_t = & Y_{10} + \sum_{k=1}^n \beta_{11k} \Delta REC_{t-k} + \sum_{k=0}^n \beta_{12k} \Delta BSP_{t-k} \\ & + \sum_{k=1}^n \beta_{13k} \Delta FI_{t-k} + \sum_{k=0}^n \beta_{14k} \Delta ICT_{t-k} \\ & + \sum_{k=1}^n \beta_{15k} \Delta FDI_{t-k} + \sum_{k=1}^n \beta_{16k} \Delta Trade_{t-k} \\ & + Y_{11} REC_{t-1} + Y_{12} BSP_{t-1} + Y_{13} FI_{t-1} + Y_{14} ICT_{t-1} \\ & + Y_{15} FDI_{t-1} + Y_{16} Trade_{t-1} + \lambda ECM_{t-1} + \mu_{1t} \end{aligned} \quad (4)$$

$$\begin{aligned} \Delta GG_t = & Y_{20} + \sum_{k=1}^n \beta_{21k} \Delta GG_{t-k} + \sum_{k=0}^n \beta_{22k} \Delta BSP_{t-k} \\ & + \sum_{k=1}^n \beta_{23k} \Delta FI_{t-k} + \sum_{k=0}^n \beta_{24k} \Delta ICT_{t-k} \\ & + \sum_{k=1}^n \beta_{25k} \Delta FDI_{t-k} + \sum_{k=1}^n \beta_{26k} \Delta Trade_{t-k} \\ & + Y_{21} GG_{t-1} + Y_{22} BSP_{t-1} + Y_{23} FI_{t-1} + Y_{24} ICT_{t-1} \\ & + Y_{25} FDI_{t-1} + Y_{26} Trade_{t-1} + \lambda ECM_{t-1} + \mu_{2t} \end{aligned} \quad (5)$$

$$\begin{aligned} \Delta CO_{2,t} = & Y_{30} + \sum_{k=1}^n \beta_{31k} \Delta CO_{2,t-k} + \sum_{k=0}^n \beta_{32k} \Delta BSP_{t-k} \\ & + \sum_{k=1}^n \beta_{33k} \Delta FI_{t-k} + \sum_{k=0}^n \beta_{34k} \Delta ICT_{t-k} \\ & + \sum_{k=1}^n \beta_{35k} \Delta FDI_{t-k} + \sum_{k=1}^n \beta_{36k} \Delta Trade_{t-k} \\ & + Y_{31} CO_{2,t-1} + Y_{32} BSP_{t-1} + Y_{33} FI_{t-1} + Y_{34} ICT_{t-1} \\ & + Y_{35} FDI_{t-1} + Y_{36} Trade_{t-1} + \lambda ECM_{t-1} + \mu_{3t} \end{aligned} \quad (6)$$

**Equations 4–6** can simultaneously produce short and long-run estimates, known as the ARDL model of Pesaran et al. (2001). Multiple advantages characterize the PMG/ARDL approach. One of the main advantages of ARDL is that the estimation is possible even when the explanatory variable is endogenous. Furthermore, as long as the ARDL model is free of the residual correlation, endogeneity is a minor problem.

In recent years, the Johnson test was used in the short term and long term relationship among variables for the empirical analysis (Spann, 1974). Still, these tests are problem-creating and have some disadvantages. After a few times, in 1998, Pesaran and Shin introduced the model to find long-run and short-run results. The Autoregressive Distributed Lag Model (ARDL) is moderately used as an econometrician approach (Shahbaz et al., 2020; Zheng et al., 2021). If the explanatory variables refer to one or more lagged values of the dependent variable, it is identified as an autoregressive or dynamic model. If any structural break exists in the series, the Auto-Regressive Distributed Lag model gives us information. The ARDL bound test can automatically resolve the serial correlation of residuals

(Pesaran & Shin, 1995). Another advantage of using the ARDL model is that it is unnecessary to specify whether variables have integrated order one or a zero. Hence, there is no need to check the stationarity of variables. ARDL can be used to check the long-run association between variables regardless of the level of stationarity. However, according to Shehbaz et al. (2011) if variables have integrated order two I(2), then the value of the F-statistic will not remain valid because of violation of the assumption that the integrated order of variables should be zero or one, it is necessary to ensure that no variable is stationary at integrated order two I(2).

Pesaran et al. (1999) the suitable lags in the ARDL model may be used to account for both residual correlation and endogeneity. ARDL has an advantage over single equation cointegration analysis such as having the problem of endogeneity, but ARDL can tell which variables are dependent and which are explanatory, which is very important. To estimate the long-run parameters that are linearly normal, the ARDL technique is utilized, regardless of whether the variables are in the range of zero to one **Eq. 1**. Furthermore, it calculates the long-run and short-run outcomes and the error correction for each parameter (Zhu & Peng, 2012).

It is possible to get the short-run estimates from the first-differenced variables by solving **Eqs 4–6**. According to Pesaran et al. (2001), the limits F-test is used to determine whether or not long-run variables are cointegrated. In contrast to all other time series approaches such as Johansen (1988) has used the ARDL model may perform effectively even if specific variables are stationary at the level and other variables are fixed at the first difference. Another feature of the ARDL approach is that it performs well even when a small sample size is used (Bahmani-Oskooee et al., 2020; Usman et al., 2021). It incorporates a short-run dynamic process, it is a better approach in terms of its ability to cope with endogeneity and serial correlation, among other things (Pesaran et al., 2001). Finally, diagnostic tests such as the Lagrange Multiplier (LM) and the Breusch-Pagan (BP) tests have been used to discover autocorrelation and heteroscedasticity issues in the data. Ramsey's RESET command is used to correct model misspecification. We use the cumulative sum (CUSUM) and the cumulative sum-square (CUSUM-square) to assess the stability of short- and long-run characteristics.

## Data (1995–2020)

The study intends to investigate the impact of banking sector expansion on renewable energy consumption, green growth, and environmental quality in China from 1995 to 2020.

**Table 1** presents descriptive data, definitions, and variable symbols. There are three dependent variables in the research; three models regress it. These dependent variables include renewable energy consumption (total energy consumption in quad Btu), green growth (environmentally adjusted multifactor productivity), and CO<sub>2</sub> emissions (in kilotons). Three characteristics are used to measure the development of the banking sector: bank assets, return on assets, and ATMs. Bank



**TABLE 1 |** Definitions and descriptive statistics.

| Variables         | Definitions   | Mean  | Median | Maximum | Minimum | Std. Dev. | Skewness | Kurtosis |
|-------------------|---|-------|--------|---------|---------|-----------|----------|----------|
| REC               | Total energy consumption from nuclear, renewables, and other (quad Btu) | 1.614 | 1.604  | 2.797   | 0.523   | 0.798     | 0.006    | 1.519    |
| Green growth (GG) | Environmentally adjusted multifactor productivity                       | 8.864 | 8.615  | 13.134  | 7.103   | 1.298     | 1.825    | 6.620    |
| CO <sub>2</sub>   | CO <sub>2</sub> emissions (kilotons)                                    | 15.64 | 15.77  | 16.26   | 14.93   | 0.476     | -0.355   | 1.525    |
| BA                | Deposit money banks assets to GDP (%)                                   | 124.9 | 120.3  | 174.5   | 82.39   | 23.64     | 0.485    | 2.606    |
| ROA               | Bank return on assets (% , after tax)                                   | 0.940 | 0.978  | 1.545   | 0.431   | 0.332     | 0.110    | 2.307    |
| ATMS              | ATMs per 100,000 adults   | 30.26 | 13.70  | 91.36   | 1.808   | 31.78     | 0.922    | 2.266    |
| ICT               | ICT index   | 32.80 | 30.39  | 73.56   | 1.304   | 23.97     | 0.150    | 1.611    |
| FDI               | Foreign direct investment, net inflows (% of GDP)                       | 3.297 | 3.500  | 4.725   | 1.311   | 1.077     | -0.507   | 2.146    |
| TRADE             | Trade (% GDP)   | 45.28 | 43.82  | 64.47   | 32.42   | 10.12     | 0.539    | 2.026    |

**TABLE 2 |** Unit root tests.

|                 | DF-GLS    |           |          | ADF     |           |          |
|-----------------|-----------|-----------|----------|---------|-----------|----------|
|                 | I(0)      | I(1)      | Decision | I(0)    | I(1)      | Decision |
| REC             | -2.875*** |           | I(0)     | -2.785* |           | I(0)     |
| GG              | -2.785*** |           | I(0)     | -2.658* |           | I(0)     |
| CO <sub>2</sub> | -0.287    | -2.987*** | I(1)     | -0.589  | -3.898*** | I(1)     |
| BA              | -0.485    | -3.289*** | I(1)     | -1.785  | -3.248**  | I(1)     |
| ROA             | -1.855*   |           | I(0)     | -2.654* |           | I(0)     |
| ATMS            | -0.145    | -2.754*** | I(1)     | 2.012   | -2.756*   | I(1)     |
| ICT             | -0.423    | -1.986**  | I(1)     | 0.325   | -3.658**  | I(1)     |
| FDI             | -0.415    | -5.365*** | I(1)     | -0.921  | -5.965*** | I(1)     |
| TRADE           | -1.452    | -3.248*** | I(1)     | -1.564  | -3.556**  | I(1)     |

Note: \*\*\* $p < 0.01$ ; \*\* $p < 0.05$ ; \* $p < 0.1$ .

assets are calculated as a percentage of GDP, returns on assets are calculated after-tax, and ATM data is calculated per 100,000 individuals. The ICT index, foreign direct investment, and trade role are included as control variables. The IMF, OECD, and World Bank provided the data for the study.

## RESULTS AND DISCUSSION

The unit root features of the data were checked before undertaking regression analysis in this research. According to the findings the DF-GLS and ADF unit root tests were used to complete this work. The coefficient estimates for both unit root tests are shown in **Table 2**. DF-GLS and ADF coefficient estimates show that renewable energy consumption, green growth, and asset returns are level stationary variables. In contrast, according to the tests, CO<sub>2</sub> emissions, bank assets, ATMs, information and communications technology, foreign direct investment (FDI), and trade are the first difference static variables. As a result of these results, the ARDL technique was used for the regression analysis in the research. **Table 3** summarizes the models' conclusions for renewable energy consumption, energy growth, and CO<sub>2</sub> emissions throughout the long and short-term timeframes.

The long-run results of the renewable energy consumption model show that bank assets report a statistically significant and positive rise in renewable energy consumption over the long run. A one percent increase in bank assets results in a 0.011 percent rise in renewable energy usage over the long term. ATMs have been shown to have a significant and considerable impact on the use of renewable energy over the long term. It has been discovered that a one percent increase in the number of ATMs has the potential to increase renewable energy usage by 0.003 percent. Results indicated that there is a favorable relationship between the expansion of the financial industry and the use of renewable energy. Amiakwa-Mensah and Näsström (2022) provide evidence to support this good relationship. According to the literature, there are two avenues *via* which the expansion of the banking industry might impact the use of renewable energy.

Firstly, the expansion of the financial sector encourages the use of renewable energy sources by encouraging technological innovation and the accumulation of wealth. Second, allocating financing to lucrative projects helps increase the use of renewable energy by encouraging more people to do so. These results imply that a well-organized banking sector can minimize liquidity risk and increase financial resources for green investments in renewable energy resources, as shown by the findings. For example, the banking industry in China, Germany, and Sweden has launched green auto loans and mortgages based on the usage of environmentally friendly cars and renewable energy sources. Aside from that, large-scale investments are necessary for advancing technological innovation and the development of renewable energy resources, which might be made feasible by the availability of financial assistance (Zheng et al., 2021). In the long term, information and communications technology (ICT) and trade factors contribute to positive and significant growth in renewable energy consumption. However, foreign direct investment (FDI) has a negligible impact on renewable energy use in the long term. The short-run findings of the renewable energy consumption model show that bank assets and ATMs tend to increase renewable energy consumption in the near term. ICT, trade, and foreign direct investment (FDI) all indicate a beneficial impact on renewable energy use in the near term.

Concerning the green growth model, the data over the long term suggest that bank assets and ATMs significantly promote

**TABLE 3** | Short and long-run estimates of renewable energy consumption, green growth, and CO<sub>2</sub> emissions.

| Variable  | REC         |        | GG          |        | CO <sub>2</sub> |        |
|-----------|-------------|--------|-------------|--------|-----------------|--------|
|           | Coefficient | t-Stat | Coefficient | t-Stat | Coefficient     | t-Stat |
| Short-run |             |        |             |        |                 |        |
| BA        | 0.009***    | 2.279  | 0.157**     | 2.029  | 0.006***        | 2.852  |
| BA(-1)    | 0.024***    | 6.046  | -0.284***   | 2.648  | -0.018***       | 3.305  |
| BA(-2)    | -0.009***   | 2.731  |             |        | 0.004           | 1.535  |
| ATMS      | 0.007       | 1.644  | 0.071       | 0.839  | -0.007*         | 1.894  |
| ATMS(-1)  | 0.014***    | 3.063  | 0.082       | 0.931  | -0.008**        | 2.298  |
| ATMS(-2)  |             |        | 0.312***    | 2.894  | 0.011***        | 2.904  |
| ICT       | 0.057***    | 5.688  | 1.194***    | 2.684  | -0.027***       | 3.375  |
| ICT(-1)   |             |        | -2.326***   | 3.009  | -0.087***       | 2.911  |
| ICT(-2)   |             |        | 1.095***    | 2.593  | 0.071**         | 2.485  |
| FDI       | 0.099**     | 2.013  | 0.079       | 0.117  | -0.035          | 1.096  |
| FDI(-1)   | 0.092*      | 1.817  |             |        | -0.101***       | 2.632  |
| FDI(-2)   | -0.087*     | 1.723  |             |        | -0.060***       | 2.722  |
| TRADE     | 0.029***    | 4.100  | 0.012       | 0.123  | 0.011***        | 3.357  |
| TRADE(-1) | 0.007       | 1.037  | 0.225**     | 2.156  | 0.002           | 0.811  |
| TRADE(-2) | 0.031***    | 4.080  | 0.100       | 1.625  | -0.010***       | 2.679  |
| Long-run  |             |        |             |        |                 |        |
| BA        | 0.011***    | 3.565  | 0.046**     | 2.203  | -0.002          | 1.267  |
| ATMS      | 0.003*      | 1.679  | 0.057***    | 2.882  | -0.001          | 0.798  |
| ICT       | 0.027***    | 22.65  | 0.013       | 0.497  | -0.021***       | 10.83  |
| FDI       | 0.050       | 1.200  | 0.028       | 0.117  | 0.063***        | 2.662  |
| TRADE     | 0.004**     | 2.322  | 0.114***    | 8.084  | 0.011***        | 6.100  |
| C         | 1.040**     | 2.472  | 8.266***    | 3.033  | 1.035***        | 6.540  |

Note: \*\*\* $p < 0.01$ ; \*\* $p < 0.05$ ; \* $p < 0.1$ .

green growth. A one percent increase in the number of bank branches and ATMs increases green growth by 0.046 and 0.057 percent, respectively, according to the data collected by the researchers. A good link is observed between banking sector development and green growth in our research, and this relationship is reported in the studies done by (Aslam et al., 2021; Cao et al., 2021). According to research issued by the Asian Expansion Bank (2012), cited Abo (2013) the development of the banking sector is essential to produce new technologies and industrial goods that take environmental quality, green growth, and energy efficiency into mind. Two routes are involved in developing the banking sector's contribution to green development.

The expansion of the banking sector may help lessen the impact of environmental fluctuations and damages on human

society and the economic system. Second, the development of the banking sector facilitates access to financial resources and contributes to promoting green growth. The literature has stated that low-cost loans are essential to achieve the objectives of greenhouse gas emission reduction. In the case of control variables, the results indicate that information and communications technology (ICT) and foreign direct investment (FDI) have a negligible connection with green growth. Still, trade is strongly and positively connected with green growth in the long term. The short-run results of the green growth model suggest that bank assets have a statistically significant and positive relationship with green growth. Still, ATMs have a statistically insignificant relationship with green growth. The control variables suggest that information and communications technology (ICT) is strongly and positively

**TABLE 4** | Short and long-run estimates of renewable energy consumption, green growth, and CO<sub>2</sub> emissions.

| Variable    | REC         |        | GG          |        | CO <sub>2</sub> |        |
|-------------|-------------|--------|-------------|--------|-----------------|--------|
|             | Coefficient | t-Stat | Coefficient | t-Stat | Coefficient     | t-Stat |
| Short-run   |             |        |             |        |                 |        |
| Diagnostics |             |        |             |        |                 |        |
| F-test      | 12.55***    |        | 5.987***    |        | 8.655***        |        |
| ECM(-1)*    | -0.406***   | 12.59  | -0.462***   | 8.785  | -0.470***       | 9.049  |
| LM          | 0.542       |        | 0.785       |        | 2.012           |        |
| BP          | 0.654       |        | 0.452       |        | 1.021           |        |
| RRESET      | 1.254       |        | 1.325       |        | 0.365           |        |
| CUSUM       | S           |        | S           |        | S               |        |
| CUSUM-sq    | S           |        | S           |        | S               |        |

Note: \*\*\* $p < 0.01$ ; \*\* $p < 0.05$ ; \* $p < 0.1$ .

**TABLE 5** | Short and long-run estimates of renewable energy consumption, green growth, and CO<sub>2</sub> emissions (Robustness).

| Variable    | REC         |        | GG          |        | CO <sub>2</sub> |        |
|-------------|-------------|--------|-------------|--------|-----------------|--------|
|             | Coefficient | t-Stat | Coefficient | t-Stat | Coefficient     | t-Stat |
| Short-run   |             |        |             |        |                 |        |
| ROA         | 0.073       | 0.382  | 2.869***    | 3.181  | 0.161**         | 2.261  |
| ROA(-1)     |             |        | 1.594       | 1.239  | 0.053           | 1.245  |
| ROA(-2)     |             |        | 4.662***    | 2.666  | 0.133***        | 2.739  |
| ATMS        | 0.014       | 1.095  | 0.011       | 0.086  | 0.012           | 0.029  |
| ATMS(-1)    | 0.018       | 1.645  | 0.338***    | 2.770  | 0.010**         | 2.169  |
| ATMS(-2)    | 0.011       | 1.362  | 0.174       | 1.587  | 0.010***        | 2.581  |
| ICT         | 0.039       | 1.592  | 1.038**     | 2.469  | 0.016*          | 1.664  |
| ICT(-1)     | 0.126***    | 3.019  | 0.949**     | 2.536  | 0.013           | 0.923  |
| ICT(-2)     | 0.165***    | 3.569  | 0.440       | 1.066  | 0.027           | 1.355  |
| FDI         | 0.069       | 0.956  | 0.669       | 1.030  | 0.002           | 0.074  |
| FDI(-1)     | 0.117*      | 1.910  | 0.845       | 1.346  | -0.017          | 0.834  |
| FDI(-2)     | 0.128       | 1.407  | 0.790       | 1.434  | -0.050**        | 2.425  |
| TRADE       | 0.010       | 1.449  | 0.052*      | 0.738  | 0.007***        | 3.086  |
| TRADE(-1)   | 0.001       | 0.143  | 0.194**     | 1.960  | 0.004           | 1.389  |
| TRADE(-2)   | 0.016       | 1.544  | 0.188**     | 2.266  | -0.004          | 1.018  |
| Long-run    |             |        |             |        |                 |        |
| ROA         | 0.044       | 0.370  | 1.008***    | 4.252  | -0.233***       | 3.488  |
| ATMS        | 0.013**     | 2.564  | 0.052**     | 2.243  | -0.005***       | 3.400  |
| ICT         | 0.047***    | 7.565  | 0.118***    | 3.596  | -0.021**        | 2.396  |
| FDI         | 0.188***    | 2.710  | 0.207       | 0.610  | 0.060           | 1.355  |
| TRADE       | 0.005*      | 1.781  | 0.111***    | 8.124  | 0.011***        | 8.036  |
| C           | 1.284***    | 4.232  | 2.252**     | 2.480  | 4.587***        | 6.383  |
| Diagnostics |             |        |             |        |                 |        |
| F-test      | 4.325*      |        | 3.987*      |        | 6.354***        |        |
| ECM(-1)*    | -0.670***   | 7.416  | -0.571***   | 8.187  | -0.470***       | 17.82  |
| LM          | 1.988       |        | 1.365       |        | 1.355           |        |
| BP          | 0.365       |        | 1.355       |        | 0.325           |        |
| RESET       | 1.587       |        | 0.254/      |        | 1.302           |        |
| CUSUM       | S           |        | S           |        | S               |        |
| CUSUM-sq    | S           |        | S           |        | S               |        |

Note: \*\*\* $p < 0.01$ ; \*\* $p < 0.05$ ; \* $p < 0.1$ .

connected with green growth. Still, that trade and foreign direct investment (FDI) reports have no meaningful influence on green development in the near term.

In the CO<sub>2</sub> emissions model, the long-term data indicate that bank branches and ATMs have a negligible impact on CO<sub>2</sub> emissions in the long run. ICT has both substantial and adverse effects on CO<sub>2</sub> emissions, with evidence indicating that increased usage of ICT leads to a reduction in CO<sub>2</sub> emissions and, consequently, an improvement in environmental quality. Thus, higher energy consumption reduces ecological quality by increasing carbon emissions due to the expansion of the industrial setup. The growth of the banking industry has resulted in an increase in CO<sub>2</sub> emissions, which has been increasing. In contrast, after reaching a certain level of development, the banking sector may see a surge in energy innovations, resulting in cleaner energy from other sources (Shahbaz et al., 2013; Smulders et al., 2014). The amount of CO<sub>2</sub> emissions might begin to decline. The studies conducted by Shahbaz et al. (2013) and Zafar et al. (2019) both found a similar association between the two variables.

In contrast, foreign direct investment (FDI) and international commerce (trade) have a significant and beneficial influence on CO<sub>2</sub> emissions in the long term, suggesting that growth in FDI

and international trade increases CO<sub>2</sub> emissions, hence deteriorating environmental quality. Bank branches have considerable and growing impacts on CO<sub>2</sub> emissions in the near term, but ATMs have significant and decreasing effects on CO<sub>2</sub> emissions in the short term. In the near term, information and communications technology (ICT) results in a significant reduction in CO<sub>2</sub> emissions. In the short term, trade, on the other hand, has the effect of increasing CO<sub>2</sub> emissions. The relationship between FDI and CO<sub>2</sub> emissions is statistically negligible in the short run.

The diagnostic testing results demonstrated in Table 4 indicated that the coefficient estimates derived from ARDL models are accurate and reliable. LM and BP tests revealed no evidence of heteroskedasticity or autocorrelation, which was verified by the outcomes of the study. According to the Ramsey RESET test results, all of the models have been appropriately defined. The CUSUM and CUSUM-sq tests demonstrate that all of the models are stable. It is possible to have long-run cointegration among variables, as shown by the coefficient estimates of the ECM test and the F-statistics.

The coefficient estimates of robustness models for the long-run and short-run are shown in Table 5. The renewable energy consumption model results demonstrate that returns on assets

have a negligible impact on renewable energy consumption in the long term. In contrast, automated teller machines (ATMs) positively impact renewable energy consumption. The return on assets and ATMs both state that the effect on renewable energy usage is modest in the near term. Long-term, the green growth model findings demonstrate that the return on assets and the use of ATMs contribute to a significant and positive increase in green growth, respectively. Increasing return on investments has improved green growth in the short run; however, ATMs have had a minor influence on green development in the near term. The results of the CO<sub>2</sub> model demonstrate that return on assets and automated teller machines (ATMs) have the potential to cut CO<sub>2</sub> emissions over the long term considerably. However, return on investments has a considerable and favorable influence on CO<sub>2</sub> emissions in a short time, but the return on equity does not (Mohsin et al., 2019). The results of control variables in robust models are consistent with the findings of fundamental models in both the short- and long-run.

## DISCUSSION

The study examines the impact of financial inclusion on environmental quality and economic performance from 1995 to 2019. Financial inclusion is measured by ATM, ICT, FDI and premiums. Empirical research uses panels. The study's variables show long-run panel. Long-term, bank branches positively affect economic efficiency, implying that as bank branches expand, economic performance rises. Long-term bank loan increases economic growth and pollution. The influence of insurance premiums on economic growth is statistically insignificant, indicating no long-term correlation. Insurance premiums have a significant adverse impact on carbon emissions, confirming that rate increases reduce emissions over time.

Long-term energy use favorably affects economic growth and pollutant emissions (Ullah and Ozturk, 2021). Long-term, trade's influence on CO<sub>2</sub> emission is mainly negative in the insurance rate regression model. Long-term bank branch regression is the only economic-environmental consequence of population growth. Short-term, banking institutions increase carbon dioxide emissions, therefore more branches means more emissions. Short-term bank lending doesn't affect the economic performance (Mohan, 2006). In the near run, higher insurance premiums reduce economic growth and carbon emissions. The findings reveal policy recommendations. Financial inclusion fosters economic growth, according to our research. Our study suggests that long-term economic development requires financial sector changes. To reduce the effects of economic inequality by extending access to financial services, it is important to understand the relationship between financial inclusion, economic growth, and poverty.

Authorities and policymakers in these economies should adopt and implement digital financial inclusion. Financial inclusion should support Asia's growing markets. Through financial inclusion, financial institutions may support green energy initiatives (Iram et al., 2020).

The monies should be allocated to green firms, enterprises, and individuals. The government should also set strong criteria for financial institutions to lend for renewable energy and environmentally-friendly initiatives. Digital financial inclusion can help with this. Governments should reduce digital financial inclusion hurdles including cost, paperwork, and trust. Policymakers should expand climate finance to help economically disadvantaged groups cope with rising CO<sub>2</sub> emissions. Individual and SMB climate finance can help reduce CO<sub>2</sub> emissions. Sohail et al. (2021) agree that green investment in public and commercial sectors, through the financial sector, is vital if the economy wishes to transform from a high emission to low carbon green economy.

## CONCLUSION AND IMPLICATIONS

The growth of the financial sector, particularly the development of the banking sector, has played a critical role in the economic success of a country. Over the last several years, financial inclusion has emerged as a vital component of any modern monetary system. Financial inclusion refers to the availability and accessibility of economic goods and services to as wide a range of individuals as feasible. A financially inclusive system may aid in promoting environmentally friendly productive activities in the economy. Still, it can also help in decoupling economic growth from environmental damage. Increased inclusion in banking sector development offers cash for investment in clean and green technology and renewable energy projects, which contribute to the promotion of both green growth and environmental quality at the same time (De Gregorio and Guidotti, 1995). Our primary goal in this research is to determine the influence of the expansion of the banking sector and financial inclusion on renewable energy consumption, green growth, and the quality of the environment.

Checking if the variables are stable at the level or initial difference is step one in doing an empirical study of a data set. ADF and DF-GLS are two unit root tests used to demonstrate that our variables combine I(0) and I(1). Therefore, the ARDL model is the most suitable model to use to get estimates of the variables. Banking sector growth and financial inclusion are represented by three separate variables: assets held by banks, return on assets earned by banks, and the number of ATMs available. The findings of the baseline ARDL model imply that bank assets have a beneficial effect on green growth and renewable energy consumption in the long term while having a negative but minor impact on CO<sub>2</sub> emissions. These results suggest that the improvement in the banking sector leads to an increase in renewable energy consumption and green growth over the long term.

Similarly, the increase in financial inclusion leads to a rise in renewable energy use and green growth while having no discernible influence on CO<sub>2</sub> emissions. In the robust model, we have utilized return on assets rather than bank assets to measure risk. Only in the robust green growth model the long-run estimates of return on the investment are positively significant. Still, they are negatively substantial in the CO<sub>2</sub>



emissions model alone. On the other hand, ATMs have a favorable impact on the estimates of renewable energy consumption and green growth while having a negative impact on the estimates of CO<sub>2</sub>. The F-test and the ECMt-1 tests are also used to validate the robustness of the long-run findings obtained. On the other hand, the short-term outcomes are equivocal and inconsistent. The findings are significant in the development of specific policies. According to one point of view, the banking sector is critical for decoupling economic growth from CO<sub>2</sub> emissions and increasing renewable energy consumption; as a result, policymakers should strengthen its position in every area of the economy, especially in the energy and environmental sectors. As a second point, the policies of growth of the financial industry must be coordinated with those of the environment and energy. Third, authorities should emphasize the importance of inclusion in the financial sector. This will give finances and credit for investment in green production activities and renewable energy projects, which will benefit all of society. Last but not least, governments should make an effort to employ information resources rather than physical resources since this would aid in decoupling economic growth from CO<sub>2</sub> emissions and making economies more weightless.

## REFERENCES

- Abe, M. (2013). "Expansion of Global Value Chains in Asian Developing Countries," in *Global Value Chains in a Changing World*, 385–409. WTO iLibrary. doi:10.30875/82bf4caf-en
- Agnello, L., Mallick, S. K., and Sousa, R. M. (2012). Financial Reforms and Income Inequality. *Econ. Lett.* 116 (3), 583–587. doi:10.1016/j.econlet.2012.06.005
- Ahmad, M., Beddu, S., binti Itam, Z., and Alanimi, F. B. I. (2019). State of the Art Compendium of Macro and Micro Energies. *Adv. Sci. Technol. Res. J.* 13 (1), 88–109. doi:10.12913/22998624/103425
- Ahmad, W., Ullah, S., Ozturk, I., and Majeed, M. T. (2021). Does Inflation Instability Affect Environmental Pollution? Fresh Evidence from Asian Economies. *Energy & Environ.* 32 (7), 1275–1291. doi:10.1177/0958305x20971804
- Amuakwa-Mensah, F., and Nässström, E. (2022). Role of Banking Sector Performance in Renewable Energy Consumption. *Appl. Energy* 306, 118023. doi:10.1016/j.apenergy.2021.118023
- Arora, R. U. (2010). Measuring Financial Access. *Griffith Bus. Sch. Discuss. Pap. Econ.* 1 (7), 1–21. Available at: <https://citeseerx.ist.psu.edu/viewdoc/download?doi=10.1.1.1083.3903&rep=rep1&type=pdf> (Accessed July 31, 2022).
- Asbahi, A. A. M. H. A., Gang, F. Z., Iqbal, W., Abass, Q., Mohsin, M., and Iram, R. (2019). Novel Approach of Principal Component Analysis Method to Assess the National Energy Performance via Energy Trilemma Index. *Energy Rep.* 5, 704–713. doi:10.1016/j.egyr.2019.06.009
- Asghar Khan, M., Rehan, R., Umer Chhapra, I., and Bai, A. (2022). Inspecting Energy Consumption, Capital Formation and Economic Growth Nexus in Pakistan. *Sustain. Energy Technol. Assessments* 50, 101845. doi:10.1016/j.seta.2021.101845
- Aslam, B., Hu, J., Majeed, M. T., Andlib, Z., and Ullah, S. (2021). Asymmetric Macroeconomic Determinants of CO<sub>2</sub> Emission in China and Policy Approaches. *Environ. Sci. Pollut. Res.* 28 (31), 41923–41936. doi:10.1007/s11356-021-13743-7
- Bahmani-Oskooee, M., Usman, A., and Ullah, S. (2020). Asymmetric J-Curve in the Commodity Trade between Pakistan and United States: Evidence from 41 Industries. *Eurasian Econ. Rev.* 10 (2), 163–188. doi:10.1007/s40822-019-00137-x
- Cao, J., Law, S. H., Samad, A. R. B. A., Mohamad, W. N. B. W., Wang, J., and Yang, X. (2021). Impact of Financial Development and Technological Innovation on the Volatility of Green Growth-Evidence from China. *Environ. Sci. Pollut. Res.* 28 (35), 48053–48069. doi:10.1007/s11356-021-13828-3
- Claessens, S., and Perotti, E. (2007). Finance and Inequality: Channels and Evidence. *J. Comp. Econ.* 35 (4), 748–773. doi:10.1016/j.jce.2007.07.002
- De Gregorio, J., and Guidotti, P. E. (1995). Financial Development and Economic Growth. *World Dev.* 23 (3), 433–448. doi:10.1016/0305-750x(94)00132-i
- Dercon, S. (2014). Is Green Growth Good for the Poor? *World Bank Res. Observer* 29 (2), 163–185. doi:10.1093/wbro/lku007
- Donou-Adansou, F., and Sylwester, K. (2016). Financial Development and Poverty Reduction in Developing Countries: New Evidence from Banks and Microfinance Institutions. *Rev. Dev. Finance* 6 (1), 82–90. doi:10.1016/j.rdf.2016.06.002
- Estrada, G. B., Park, D., and Ramayandi, A. (2010). *Financial Development and Economic Growth in Developing Asia*. Manila, Philippines: Asian Development Bank Economics Working Paper, 233. doi:10.2139/ssrn.1751833
- Frankel, J. A., and Romer, D. (1999). Does Trade Cause Growth? *Am. Econ. Rev.* 89 (3), 379–399. doi:10.1257/aer.89.3.379
- Gyimah, J., Yao, X., Tachega, M. A., Sam Hayford, I., and Opoku-Mensah, E. (2022). Renewable Energy Consumption and Economic Growth: New Evidence from Ghana. *Energy* 248, 123559. doi:10.1016/j.energy.2022.123559
- IPA (2017). Climate change and financial inclusion. Available at: [https://www.povertyaction.org/sites/default/files/publications/Climate-Change-Financial-Inclusion\\_Final.pdf](https://www.povertyaction.org/sites/default/files/publications/Climate-Change-Financial-Inclusion_Final.pdf) (Accessed July 31, 2022).
- Iram, R., Zhang, J., Erdogan, S., Abbas, Q., and Mohsin, M. (2020). Economics of Energy and Environmental Efficiency: Evidence from OECD Countries. *Environ. Sci. Pollut. Res.* 27 (4), 3858–3870. doi:10.1007/s11356-019-07020-x
- Johansen, S. (1988). Statistical Analysis of Cointegration Vectors. *J. Econ. Dyn. control* 12 (2–3), 231–254. doi:10.1016/0165-1889(88)90041-3
- Kiplimo, J. C., Ngenoh, E., Koech, W., and Bett, J. K. (2015). Determinants of Access to Credit Financial Services by Smallholder Farmers in Kenya. *J. Dev. Agric. Econ.* 7 (9), 303–313. doi:10.5897/JDAE2014.0591
- Lei, W., Ozturk, I., Muhammad, H., and Ullah, S. (2021). On the Asymmetric Effects of Financial Deepening on Renewable and Non-renewable Energy Consumption: Insights from China. *Econ. Research-Ekonomska Istraživanja*, 1–18. doi:10.1080/1331677x.2021.2007413
- Leyshon, A., and Thrift, N. (1995). Geographies of Financial Exclusion: Financial Abandonment in Britain and the United States. *Trans. Inst. Br. Geogr.* 20, 312–341. doi:10.2307/622654
- Li, X., Ozturk, I., Majeed, M. T., Hafeez, M., and Ullah, S. (2022). Considering the Asymmetric Effect of Financial Deepening on Environmental Quality in BRICS Economies: Policy Options for the Green Economy. *J. Clean. Prod.* 331, 129909. doi:10.1016/j.jclepro.2021.129909

## DATA AVAILABILITY STATEMENT

The original contributions presented in the study are included in the article/Supplementary Material, further inquiries can be directed to the corresponding author.

## AUTHOR CONTRIBUTIONS

This idea was given by MK. MK and RR analyzed the data and wrote the complete paper. While RR read and approved the final version.

- Linster, M., and Yang, C. (2018). China's Progress towards Green Growth: An International Perspective (No. 2018/05). Paris, France: OECD Publishing. doi:10.1787/22260935
- Mohan, R. (2006). *Economic Growth, Financial Deepening and Financial Inclusion*. Mumbai, India: Reserve Bank of India Bulletin, 1305. (Accessed September 26, 2006).
- Mohsin, M., Abbas, Q., Zhang, J., Ikram, M., and Iqbal, N. (2019). Integrated Effect of Energy Consumption, Economic Development, and Population Growth on CO<sub>2</sub> Based Environmental Degradation: a Case of Transport Sector. *Environ. Sci. Pollut. Res.* 26 (32), 32824–32835. doi:10.1007/s11356-019-06372-8
- OECD (2020a). A territorial approach to the Sustainable Development Goals: synthesis report. OECD Urban Policy Reviews. Paris: OECD Publishing. Available at: <https://www.oecd-ilibrary.org/sites/0a305c62-en/index.html?itemId=/content/component/0a305c62-en> (Accessed July 30, 2022).
- Pasimeni, F., and Pasimeni, P. (2016). An Institutional Analysis of the Europe 2020 Strategy. *Soc. Indic. Res.* 127 (3), 1021–1038. doi:10.1007/s11205-015-1013-7
- Pesaran, M. H., and Shin, Y. (1995). *An Autoregressive Distributed Lag Modelling Approach to Cointegration Analysis*. (Oslo: Norwegian Academy of Science and Letters).
- Pesaran, M. H., Shin, Y., and Smith, R. J. (2001). Bounds Testing Approaches to the Analysis of Level Relationships. *J. Appl. Econ.* 16 (3), 289–326. doi:10.1002/jae.616
- Pesaran, M. H., Shin, Y., and Smith, R. P. (1999). Pooled Mean Group Estimation of Dynamic Heterogeneous Panels. *J. Am. Stat. Assoc.* 94 (446), 621–634. doi:10.1080/01621459.1999.10474156
- Renzhi, N., and Baek, Y. J. (2020). Can Financial Inclusion Be an Effective Mitigation Measure? Evidence from Panel Data Analysis of the Environmental Kuznets Curve. *Finance Res. Lett.* 37, 101725. doi:10.1016/j.frl.2020.101725
- Sahay, R., Čihák, M., N'Diaye, P., and Barajas, A. (2015). Rethinking Financial Deepening: Stability and Growth in Emerging Markets. *Rev. Econ. Inst.* 17 (33), 73–107. doi:10.5089/9781498312615.006
- Samour, A., Isiksal, A. Z., and Resatoglu, N. G. (2019). Testing the Impact of Banking Sector Development on Turkey's CO<sub>2</sub> Emissions. *Appl. Ecol. Env. Res.* 17 (3), 6497–6513. doi:10.15666/aer/1703\_64976513
- Sanderson, A., Mutandwa, L., and Le Roux, P. (2018). A Review of Determinants of Financial Inclusion. *Int. J. Econ. Financial Issues* 8 (3), 1. Available at: [https://media.proquest.com/media/hms/PFT/1/ymKs5\\_s\\_rT3LD4bW%2FjV1CXwINZW5yIcYwlg%3D](https://media.proquest.com/media/hms/PFT/1/ymKs5_s_rT3LD4bW%2FjV1CXwINZW5yIcYwlg%3D) (Accessed July 31, 2022).
- Sarma, M. (2008). Financial Inclusion and Development: A Cross Country Analysis. Available at: <http://hdl.handle.net/10419/176233>.
- Sethi, D., and Acharya, D. (2018). Financial Inclusion and Economic Growth Linkage: Some Cross Country Evidence. *J. Financial Econ. Policy* 10 (2). doi:10.1108/jfep-11-2016-0073
- Shahbaz, M., Raghutla, C., Song, M., Zameer, H., and Jiao, Z. (2020). Public-private Partnerships Investment in Energy as New Determinant of CO<sub>2</sub> Emissions: the Role of Technological Innovations in China. *Energy Econ.* 86, 104664. doi:10.1016/j.eneco.2020.104664
- Shahbaz, M., Solarin, S. A., Mahmood, H., and Arouri, M. (2013). Does Financial Development Reduce CO<sub>2</sub> Emissions in Malaysian Economy? A Time Series Analysis. *Econ. Model.* 35, 145–152. doi:10.1016/j.econmod.2013.06.037
- Shahbaz, M., Tang, C. F., and Shahbaz Shabbir, M. (2011). Electricity Consumption and Economic Growth Nexus in Portugal Using Cointegration and Causality Approaches. *Energy policy* 39 (6), 3529–3536. doi:10.1016/j.enpol.2011.03.052
- Smulders, S., Toman, M., and Withagen, C. (2014). Growth Theory and 'green Growth'. *Oxf. Rev. Econ. policy* 30 (3), 423–446. doi:10.1093/oxrep/gru027
- Sohail, M. T., Xiuyuan, Y., Usman, A., Majeed, M. T., and Ullah, S. (2021). Renewable Energy and Non-renewable Energy Consumption: Assessing the Asymmetric Role of Monetary Policy Uncertainty in Energy Consumption. *Environ. Sci. Pollut. Res.* 28 (24), 31575–31584. doi:10.1007/s11356-021-12867-0
- Solarin, S. A., Al-Mulali, U., Musah, I., and Ozturk, I. (2017). Investigating the Pollution Haven Hypothesis in Ghana: an Empirical Investigation. *Energy* 124, 706–719. doi:10.1016/j.energy.2017.02.089
- Spann, R. M. (1974). Rate of Return Regulation and Efficiency in Production: An Empirical Test of the Averch-Johnson Thesis. *Bell J. Econ. Manag. Sci.* 5, 38–52. doi:10.2307/3003091
- Stern, D. I. (2004). The Rise and Fall of the Environmental Kuznets Curve. *World Dev.* 32 (8), 1419–1439. doi:10.1016/j.worlddev.2004.03.004
- Sun, L., Cao, X., Alharthi, M., Zhang, J., Taghizadeh-Hesary, F., and Mohsin, M. (2020). Carbon Emission Transfer Strategies in Supply Chain with Lag Time of Emission Reduction Technologies and Low-Carbon Preference of Consumers. *J. Clean. Prod.* 264, 121664. doi:10.1016/j.jclepro.2020.121664
- Ullah, S., Guoqiang, H., Khan, U., and Niazi, K. (2020). State, Religion and the Marginalisation of Traditional Healing in Gwadar, Pakistan. *Med. Humanit.* 46 (4), 444–453. doi:10.1136/medhum-2019-011747
- Ullah, S., Khan, U., Rahman, K. U., and Ullah, A. (2021b). Problems and Benefits of the China-Pakistan Economic Corridor (CPEC) for Local People in Pakistan: A Critical Review. *Asian Perspect.* 45 (4), 861–876. doi:10.1353/apr.2021.0036
- Ullah, S., and Ozturk, I. (2020). Examining the Asymmetric Effects of Stock Markets and Exchange Rate Volatility on Pakistan's Environmental Pollution. *Environ. Sci. Pollut. Res.* 27 (25), 31211–31220. doi:10.1007/s11356-020-09240-y
- Ullah, S., Ozturk, I., Majeed, M. T., and Ahmad, W. (2021a). Do technological Innovations Have Symmetric or Asymmetric Effects on Environmental Quality? Evidence from Pakistan. *J. Clean. Prod.* 316, 128239. doi:10.1016/j.jclepro.2021.128239
- Usman, A., Bahmani-Oskoei, M., Anwar, S., and Ullah, S. (2021). Is There J-Curve Effect in the Trade between Pakistan and United Kingdom? Asymmetric Evidence from Industry Level Data. *Singap. Econ. Rev.*, 1–21. doi:10.1142/s0217590821500089
- Wei, L., and Ullah, S. (2022). International Tourism, Digital Infrastructure, and CO<sub>2</sub> Emissions: Fresh Evidence from Panel Quantile Regression Approach. *Environ. Sci. Pollut. Res.* 29, 36273–36280. doi:10.1007/s11356-021-18138-2
- Yang, L., Hui, P., Yasmeen, R., Ullah, S., and Hafeez, M. (2020). Energy Consumption and Financial Development Indicators Nexuses in Asian Economies: a Dynamic Seemingly Unrelated Regression Approach. *Environ. Sci. Pollut. Res.* 27 (14), 16472–16483. doi:10.1007/s11356-020-08123-6
- Zafar, M. W., Zaidi, S. A. H., Sinha, A., Gedikli, A., and Hou, F. (2019). The Role of Stock Market and Banking Sector Development, and Renewable Energy Consumption in Carbon Emissions: Insights from G-7 and N-11 Countries. *Resour. Policy* 62, 427–436. doi:10.1016/j.resourpol.2019.05.003
- Zakarya, G. Y., Mostefa, B., Abbes, S. M., and Seghir, G. M. (2015). Factors Affecting CO<sub>2</sub> Emissions in the BRICS Countries: a Panel Data Analysis. *Procedia Econ. Finance* 26, 114–125. doi:10.1016/S2212-5671(15)00890-4
- Zhang, D., Mohsin, M., Rasheed, A. K., Chang, Y., and Taghizadeh-Hesary, F. (2021). Public Spending and Green Economic Growth in BRI Region: Mediating Role of Green Finance. *Energy Policy* 153, 112256. doi:10.1016/j.enpol.2021.112256
- Zhang, M., Liu, X., Ding, Y., and Wang, W. (2019). How Does Environmental Regulation Affect Haze Pollution Governance?—An Empirical Test Based on Chinese Provincial Panel Data. *Sci. Total Environ.* 695, 133905. doi:10.1016/j.scitotenv.2019.133905
- Zheng, S., Yang, J., and Yu, S. (2021). How Renewable Energy Technological Innovation Promotes Renewable Power Generation: Evidence from China's Provincial Panel Data. *Renew. Energy* 177, 1394–1407. doi:10.1016/j.renene.2021.06.023
- Zhu, Q., and Peng, X. (2012). The Impacts of Population Change on Carbon Emissions in China during 1978–2008. *Environ. Impact Assess. Rev.* 36, 1–8. doi:10.1016/j.eiar.2012.03.003

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