



Translocation and Aquaculture Impact on Genetic Diversity and Composition of Wild Self-Sustainable *Ostrea edulis* Populations in the Adriatic Sea

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The European flat oyster, *Ostrea edulis*, is a keystone species suffering major population declines due to overfishing, habitat loss and parasite diseases. Knowledge of its fine-scale population genetic structure and connectivity, needed for effective conservation, restoration and management, is largely lacking. Along the eastern Adriatic Sea, genotyping of 1178 *O. edulis* individuals at 12 microsatellite loci was conducted, grouping the sampled populations by geographical origin (North, Middle, South Adriatic), shell-farm association (farmed, farm-impacted, wild oysters) and sampling year (2017, 2018), in order to explore spatio-temporal genetic variation and potential footprint of known human-mediated spat translocation events for aquaculture purpose. Short-term temporal genetic structuring of *O. edulis* populations was less pronounced compared to their spatial variability, which showed genetic discontinuity between *O. edulis* populations from different geographical regions, with the main boundary separating the North from the Middle and South Adriatic, and the weaker one limiting the flow between the Middle and South Adriatic. While the present culture practise and ongoing spat translocation promotes genetic heterogeneity in the investigated farms, reduced genetic diversity and smallest effective populations size of impacted, i.e., farm-associated *O. edulis* was consistently recorded in all geographical regions. Taken together, the results reflect regional oceanographic features, ongoing spat translocation and intensive harvesting, which might have reduced the wild *O. edulis* densities below the critical threshold for reproductive success, compromising settlement and favoring unidirectional gene-flow toward higher density farmed *O. edulis*. Genetic structure of Adriatic *O. edulis* populations revealed some concerning demographic changes and farm-wild oyster interactions and hence further investigation and management recommendations are given.

Keywords: aquaculture, bivalve, European flat oyster, spat translocation, genetic diversity, effective size, sPCA

INTRODUCTION

Population connectivity and spatial structure information provide a basis for understanding marine species population dynamics, and play a key role in the conservation and management of fisheries (Reiss et al., 2009). Gene flow is assumed to occur over large marine geographical scale due to the lack of obvious barriers to dispersal and to the existence of pelagic larvae phases in many species (Sa-Pinto et al., 2012). It appears that a different degree of connectivity among populations is greatly dependent on early life history traits, like pelagic larval dispersal as the presumed mechanism of primary connectivity (Huserbraten et al., 2013). Additionally, in bivalve species with relatively short pelagic larval duration, such as *Ostrea edulis*, larval settlement plays an important role in connectivity and successful stock restoration, as larvae are unable to metamorphose unless they are attached to a suitable substrate (Wieczorek and Todd, 1998).

The European flat oyster (*O. edulis*) is one of the bivalve species with the longest tradition of harvesting and aquaculture (e.g., Caceres-Martinez and Figueras, 1997; Edwards, 1997; Gouilletquer and Heral, 1997). It is a sessile, filter-feeding bivalve mollusc with a distribution ranging from Norway to Morocco in the Atlantic Ocean, and in the Mediterranean Sea and extending into the Black Sea. In the wild, *O. edulis* lives from the intertidal to 90 m depth and on different types of bottoms. Due its aquaculture potential, it has also been introduced into other parts of the world, including the United States and Canada (Carnegie and Barber, 2001; Vercaemer et al., 2006).

Although bivalve aquaculture has generally been showing a steady increase in recent decades, production of *O. edulis* in European aquaculture has decreased from an average production of 9152 tonnes per year in the period 1980–1989 to an average of 3305 tonnes per year in the period 2006–2015¹. Overharvesting, loss of habitat and the successive occurrence of two parasitic diseases, Marteiliopsis (*Marteilia refringens*) and the more serious bonamiosis (*Bonamia ostreae*) have been identified as the main causes of this drastic decline in the *O. edulis* production (Airoldi and Beck, 2007). Northern European countries are investing significant efforts in oyster restoration activities and programs (Laing et al., 2005; Shelmerdine and Leslie, 2009; Woolmer et al., 2011; Gravestock et al., 2014; Smaal et al., 2015; Harding et al., 2016), suggesting that *O. edulis* is an integral component of a biologically healthy functional benthic environment and, as such, the restoration of wild stocks is a matter of urgency (Smyth et al., 2018).

In Croatia, European flat oyster aquaculture has a long tradition, with the first organized oyster farming in the eastern Adriatic dating back to the 16th century (Horváth et al., 2013). Over the last 5 years, production limited to 49 tonnes of oysters per year was sold exclusively on the local market. Despite the strong resistance of Adriatic flat oyster populations to

both parasitic diseases over time, thanks mainly to a ban on spat imports spat from Western Europe (Zrnčić et al., 2007; Horváth et al., 2013), several other factors have contributed to declining oyster production, i.e., small domestic market, intense fisheries, great variability in larval dispersion and settlement, fish predation and alien species (Pineda et al., 2009; Šegvić-Bubić et al., 2011, 2016).

In line with most bivalve aquaculture production, oyster farming still relies on the collection of wild spat and its cultivation in long-line systems to a length of 3–4 mm. Due to the great spatial and temporal variability of spat settlement and its collection among years and at farming sites, spat translocation for farming purposes is commonplace at the regional scale, i.e., from the Mali Ston Bay in the south to the Lim Bay in the north of eastern Adriatic coast; however, such transfers are poorly documented. It is well known that the impact of translocation of individuals from wild populations into other genetically distinct populations is an important issue for the management of exploited or endangered species (Johnson, 2000), since the link between the risk associated with translocation and impacts on genetic integrity and diversity of native stock is well established (Brenner et al., 2014; Bromley et al., 2016). Despite the high importance of the European flat oyster as an aquaculture species, the implications that might arise from such practices on the receiving and farm-surrounding populations in the eastern Adriatic have not yet been assessed. Furthermore, there are no genetic descriptions of wild and farmed oyster populations to support the present population composition.

Several previous studies identified the importance of understanding genetic diversity of wild and cultured *O. edulis* populations in Europe (e.g., Naciri-Graven et al., 2000; Launey et al., 2002; Culloty et al., 2004; Diaz-Almela et al., 2004; Vercaemer et al., 2006; Taris et al., 2008; Lallias et al., 2010a). This is especially relevant for understanding the genetic basis for resistance to parasitic diseases (Culloty et al., 2004), differences in growth rates (Naciri-Graven et al., 2000) and reproductive success (Lallias et al., 2010a), and for conducting oyster restoration programs (Lallias et al., 2010a; Vera et al., 2016). High genetic diversity has been reported in oyster populations and the available data suggest that the genetic structure follows an isolation by distance model across the Atlantic and Mediterranean regions (Saavedra et al., 1995; Launey et al., 2002; Diaz-Almela et al., 2004), while the north Atlantic populations clustered into geographical regions associated with oceanic fronts (Vera et al., 2016).

Thus, to explore the genetic population structure and potential footprint of human-mediated spat translocation in European flat oyster populations along the eastern Adriatic coast, we genetically assayed 1178 wild and cultured individuals sampled in two consecutive years, using 13 published microsatellite markers. This study aimed to: (i) examine the possible changes in genetic variation occurring over the spatial and short-term temporal scale of oyster populations, and (ii) investigate the potential impact of spat transfer between different culture sites on the genetic diversity and structure observed in the farm and farm-surrounding populations.

¹<http://www.fao.org/fishery/statistics/global-aquaculture-production/en>

MATERIALS AND METHODS

Oyster Sampling

A total of 1178 *O. edulis* individuals were sampled from 15 sampling sites in two consecutive years, during November 2017 and 2018 (Table 1 and Figure 1). All sampled individuals were within the size of 6–8 cm corresponding to the age of 2 or 3 years. Collection included: (i) wild oysters sampled from eight natural beds covering 600 km of Croatian coastline; (ii) farmed oysters sampled from the three main aquaculture areas along the coastline, and (iii) three farm-associated populations sampled from natural beds located in the vicinity of oyster-farm installations. All farms included in this study were active with annual production of more than 10 tonnes of oysters. A muscle section of each sampled oyster was stored separately in 96% ethanol and later used in genetic analysis. Depending on location, sample sizes range from 12 to 67 individuals (Table 1). Populations were coded according to the sampling year (17, 2017; 18, 2018), regional geographical origin (N, north; M, middle; S, south Adriatic), sampling location, and origin (W, wild; F, farmed; A, farm-associated).

DNA Extraction and Genotyping

Total genomic DNA from muscle was extracted by proteinase K digestion, followed by standard phenol-chloroform extraction protocol. DNA quality and quantity were assessed by spectrophotometry (IMPLEN N50, Germany), following sample dilution to 10 ng μL^{-1} in DNase/RNase free water. A set of 13 microsatellite loci (Launey et al., 2002; Lallias et al., 2009; Vera et al., 2015; see Supplementary Table 1) were split into two PCR multiplex and amplified using the Qiagen multiplex kit with labeled (FAM, NED, VIC and PET, Applied Biosystems) primers following manufacturer recommendations in 12.5 μL reactions. Primer dyes were set up to avoid similar allele size overlapping. PCR multiplex conditions are shown in Supplementary Table 1. Fragments were separated on an ABI3130 automated sequencer (Applied Biosystems) using MacroGen (MacroGen Inc., Seoul, South Korea) services, while peak height values for each microsatellite allele were scored manually by two persons using GeneMapper v.3.5 software (Applied Biosystems).

Hardy-Weinberg Equilibrium, Linkage Disequilibrium and Null Alleles

Software MICROCHECKER 2.2.3 (Van Oosterhout et al., 2004) was used to test for genotyping errors on scored alleles, while the presence and frequency of null alleles were additionally examined by FreeNA (Chapuis and Estoup, 2007). The software computed the F_{ST} statistic, both with exclusion and inclusion of the ENA (Excluding Null Alleles) correction method that efficiently corrects for the positive bias induced by the presence of null alleles on F_{ST} estimation. The bootstrap 95% confidence intervals (CI) for the global F_{ST} values were calculated using 50,000 replicates over loci. Fisher's exact test for deviations from Hardy-Weinberg equilibrium and the linkage disequilibrium (LD) test for all pairs of loci were performed by GENEPOP v.4.0.9 (Rousset, 2008). Exact *P*-values for the individual population or

TABLE 1 | Information of sampling locations, regions, years and codes, along with the number of individual European flat oyster that were genetically assayed with 12 putatively neutral microsatellites.

Location	Region	Year	Pop ID	N	Latitude	Longitude
Wild						
Medulin	North	2017	17N_IW	46	13.9514	44.7592
		2018	18N_IW	54		
Krk	North	2017	17N_KW	37	14.5465	45.1143
		2018	18N_KW	58		
Cres	North	2017	17N_CW	37	14.5185	44.6767
		2018	18N_CW	30		
Ugljan	Middle	2017	17M_UW	49	15.1362	44.0808
		2018	18M_UW	50		
Pakoštane	Middle	2017	17M_PW	48	15.4865	43.8348
		2018	18M_PW	29		
Zečevo	Middle	2017	17M_ZW	67	15.8711	43.6724
Brač	Middle	2017	17M_BW	37	16.4493	43.30313
		2018	18M_BW	18		
Molunat	South	2017	17S_MOW	23	18.2086	42.5431
		2018	18S_MOW	33		
Farmed						
Lim Bay	North	2017	17N_LF	12	13.7338	45.1348
		2018	18N_LF	26		
		2018	18N_LFC	28		
Marina Bay	Middle	2017	17M_MF	45	16.1475	43.5041
		2018	18M_MF	41		
Mali Ston Bay	South	2017	17S_BF	41	17.5251	42.9128
		2018	18S_BF	50		
Farm-associated						
Lim Bay	North	2017	17N_LA	55	13.5903	45.2126
		2018	18N_LA	45		
Marina Bay	Middle	2017	17M_MA	66	16.0855	43.4709
		2018	18M_MA	51		
Mali Ston Bay	South	2017	17S_BA	45	17.702	42.87067
		2018	18S_BA	57		
Overall				1178		

N, sampling size; location Zečevo was investigated in only one consecutive year. In Lim Bay, oysters from two nearby farming concessions were sampled in 2018.

locus tests were estimated using the Markov Chain algorithm (10,000 dememorization steps, 100 batches and 5000 iterations) and the significance of HWE and LD values were adjusted by sequential Bonferroni correction (Rice, 1989).

Genetic Diversity, Test of Demographic Changes and Effective Population Size

Allelic richness (A_r) and the inbreeding coefficient (F_{IS}) were calculated using FSTAT v.2.3 (Goudet, 2002) while the number of alleles (N) and mean effective number of alleles across loci (A_e) were calculated using POPGENE v.1.32 (Yeh et al., 2000). Observed (H_o) and expected (H_e) heterozygosity was calculated in ARLEQUIN v.3.5 (Excoffier et al., 2005) while the contemporary effective population size (N_e) was estimated in the program NeEstimator V2 (Do et al., 2014) using the single-sample linkage disequilibrium method for populations with a

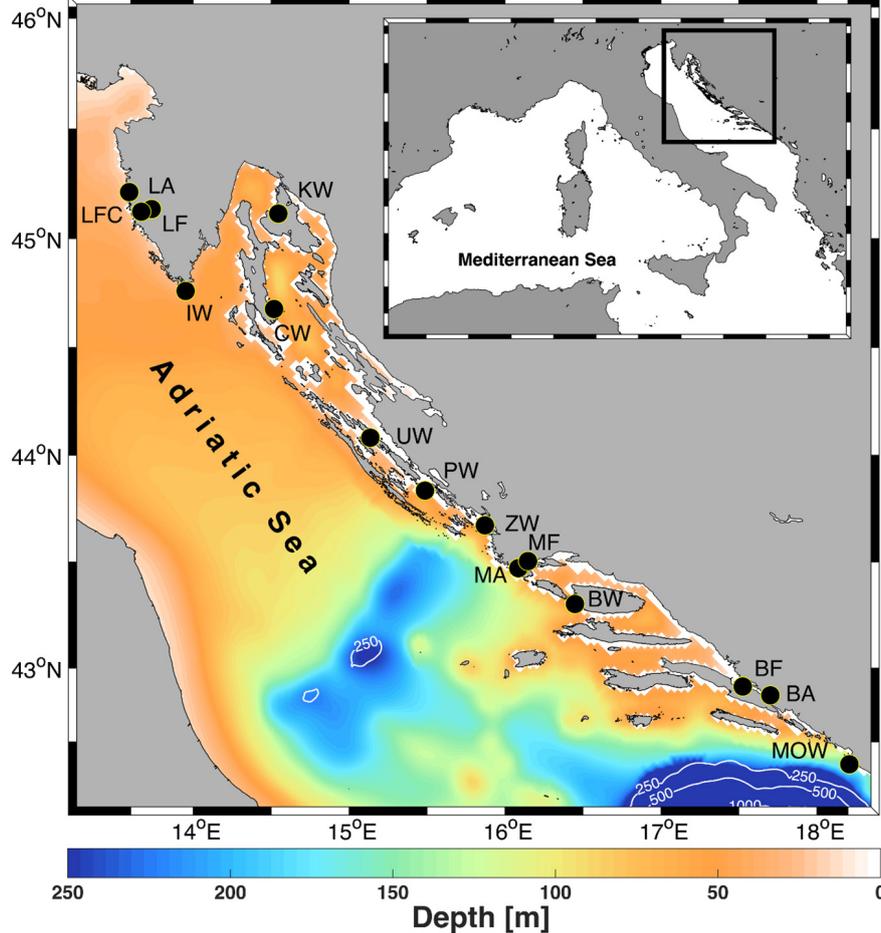


FIGURE 1 | Adriatic Sea bathymetry with sampling locations of wild European flat oyster (IW, Medulin; KW, Krk; CW, Cres; UW, Ugljan; PN, Pakoštane; ZW, Zečevo; BW, Brač; MOW, Molunat), farmed (LF and LFC, Lim Bay; MF, Marina Bay; BF, Mali Ston Bay) and farm-associated adults (LA, Lim Bay; MA, Marina Bay; BA, Mali Ston Bay). Contours are drawn for 250, 500, and 1000 m depths. More information about population abbreviations and sampling years are provided in **Table 1**. The figure has been created using MATLAB 2014a (www.mathworks.com) and GIMP 2.8.16 (www.gimp.org) software.

sample size over 17 individuals. Low frequency alleles ≤ 0.02 were excluded from the analysis.

Evidence for a recent reduction in local population size was tested with the heterozygosity excess method in the BOTTLENECK 1.2.02 software and by the Two-Phased mutation model (TPM), incorporating 90% of single-step mutations and 10% of variance among multiple steps (Piry et al., 1990). Statistical significance was evaluated by Wilcoxon signed-rank test from 10,000 simulation replicates. In addition, Garza and Williamson's *M*-ratio (Garza and Williamson, 2001) for each population was calculated using ARLEQUIN. It is sensitive to population bottlenecks because it measures the proportion of unoccupied allelic states given the range in allele size, and this ratio is reduced as alleles are randomly lost due to drift. *M*-ratio values less than 0.68 are generally indicative of populations that have experienced a recent reduction in size (Garza and Williamson, 2001). Relatedness was calculated with COANCESTRY v1.0 (Wang, 2011), using the triadic likelihood method (Wang, 2007). This estimator was chosen

because it is least biased when data contain many unrelated individuals, as expected in our dataset. Significance of mean differences in relatedness between samples was assessed by 10,000 permutations.

Genetic Differentiation and Population Structuring

Statistical power of tests for genetic homogeneity on the applied data set and sample sizes was assessed by POWSIM software (Ryman and Palm, 2006). Global F_{ST} and pair-wise values were calculated in ARLEQUIN v.3.5 where confidence levels were estimated by 2000 permutations of the dataset. Given that F_{ST} can underestimate population differentiation when highly polymorphic microsatellites are used, the alternative measure *Dest* based on allele identities (Jost, 2008) was also calculated using GENODIVE (Meirmans and Van Tienderen, 2004). Analyses of molecular variance (AMOVA) were carried out with two different analyses of distance, the number of different

alleles (F_{ST}) based on the infinite allele model and the sum of squared size difference (R_{ST}) based on the stepwise mutation model. To investigate the distribution of genetic variability within the eastern Adriatic Sea, spatial Principal Component Analysis (sPCA) were performed on allelic frequencies for each sampling year, by using the R software package *adegenet* (Jombart, 2008; Jombart et al., 2008). Spatial network using a matrix of the inverse Euclidian distance between sampling locations was used for the calculation of Moran's I . sPCA optimizes the product of the variance of individual scores and of Moran's I to summarize genetic variability in a spatial context. The presence of global or local structures was further assessed using the Global and Local random test with 1000 permutations implemented in the *adegenet* package. The lag scores for each of the first two principal components were plotted across geographic space to identify spatial genetic structure. Mantel test with function *mantel.randtest* was used to test spatial structures for each sampling year, by assessing the correlation between genetic distances and geographic distances.

Population variability was further examined with a discriminant analysis of principal components (DAPC) included in the package *adegenet* (Jombart and Ahmed, 2011), using the sampling sites as a prior for each sampling year. The optimal number of clusters was determined based on the Bayesian information criterion (BIC) (Jombart et al., 2010). The function *xvalDapc* was used sequentially with 1000 replicates, to determine the optimal number of principal components ($n = 150$) to retain in the discriminant analysis for both analyses. The *compplot* function was used to calculate posterior membership probability.

RESULTS

Genetic Diversity

A total of 1178 individuals of *O. edulis* were genotyped at 13 microsatellite loci (Table 1 and Figure 1) where the proportion of missing data per locus ranged from 0 to 3.8%, with an average of 1.3%. Locus Oedu12 was immediately excluded from analysis due to the poor amplification result in the dataset. Still, several populations, especially those having farm-associated origin, showed significant deviation from Hardy-Weinberg equilibrium, with tendencies toward heterozygote deficiency at seven of eleven loci (Oedu240, Oedu327, OeduU2, Oedu46 and Oedu212b, Supplementary Tables 2, 3), as revealed by Fisher's exact test. This deficiency in heterozygotes is unlikely to be a technical artifact, as it was observed for the majority of markers. The existence of null alleles can be regarded as the most likely cause, as null alleles are widely observed in other molluscs (Hedgecock et al., 2004; Zhan et al., 2007). MICROCHECKER detected null alleles for the loci Oedu240, Oedu327, OeduU2, Oedu46 and Oedu212b at low frequencies <5% and these loci were retained. The estimation of F_{ST} with and without the ENA correction method gave comparable results; 0.0085 vs. 0.0082 with vs. without the ENA, with overlapping 95% CI. Further, large allele dropout was not detected with MICROCHECKER and no consistent evidence of linkage disequilibrium among pair

of loci was recorded when applying strict Bonferroni correction for multiple tests.

Among the 12 loci examined, all were polymorphic with the number of alleles per locus ranging from 4 to 25 (Supplementary Tables 1, 2). While the expected heterozygosity (He) showed small variations among populations and years (0.87–0.90), the observed heterozygosity (Ho) revealed varying degrees of genetic diversity (0.77–0.87) with reduced values found in 2017 vs. 2018 and in farm-associated groups in comparison to the farm and wild counterparts for both years (Table 2). The indices of effective number of alleles (Ae) and allelic richness (Ar) showed similar diversity levels among populations, origins and years, while the number of alleles per locus was slightly reduced in farmed populations in contrast to wild ones.

The inbreeding coefficient, F_{IS} , ranged from 0.02 to 0.12 in the dataset and was significantly higher than zero in 85% and 50% populations sampled in 2017 and 2018, respectively (Table 2). High disparity of contemporaneous effective population sizes (Ne) in respect to oyster origin was recorded (Table 3). On average, estimates of Ne were 2-fold and 4-fold smaller in the farm-associated group for 2017 (493) and 2018 (547) in comparison to the farmed (1009; 1816) and wild groups (2241; 2332), respectively. In line with temporal sample replicates, both bottleneck tests showed statistical evidence that oyster populations from different origins had undergone a recent reduction in population size (Table 3). Wilcoxon signed-rank tests detected significant heterozygote excess ($p < 0.05$) under the two-phase model in 7/14 and 9/14 populations for 2017 and 2018, respectively. Only 1/14 populations in 2017 and 2018 had a significant probability of heterozygosity excess under the stepwise model (Table 3). In contrast to the first method, M-ratio analyses revealed signatures of genetic bottlenecks in all populations in 2017 and 2018, considering that the observed values (mean M-ratio 0.42 ± 0.05) were lower than the simple threshold criterion (M-ratio < 0.68 ; Garza and Williamson, 2001), which is widely used as a rule of thumb in conservation genetics (Hoban et al., 2013). The average relatedness among farmed (0.027), wild (0.0273) and farm-associated individuals (0.0262) was not statistically different from the 1,000 simulated unrelated individuals (0.032). The average relatedness among farmed and wild populations (0.027) was slightly higher than the average relatedness among farmed and impacted populations (0.025) ($P < 0.05$).

Among-Population Genetic Differentiation

Assessment of the statistical power for both microsatellite data sets in POWSIM revealed that it was possible to detect genetic divergence as low as $F_{ST} = 0.01$ with 100% certainty (χ^2 , Fisher's test) and with 98% certainty for $F_{ST} = 0.001$.

Global genetic differentiation, estimated using F_{ST} and $Dest$, across all 28 populations was 0.0082 ($p < 0.0001$) and 0.064 for the total dataset, supporting a relatively low to moderate level of differentiation among individual populations. The highest global index values of 0.012 and 0.109 were recorded within farmed populations in contrast to the lower levels of global

TABLE 2 | Summary statistics for genetic variation of *Ostrea edulis* in the Adriatic Sea showing the average number of alleles (A), effective number of alleles (Ae), allelic richness (Ar), expected (He) and observed (Ho) heterozygosity, and fixation index (F_{IS}).

	A		Ae		Ar		Ho		He		F _{IS}	
	2017	2018	2017	2018	2017	2018	2017	2018	2017	2018	2017	2018
Wild												
N_IW	16.4 ± 5.9	18.3 ± 6.4	11.2 ± 4.5	12.1 ± 5.2	11.1 ± 3.4	13.5 ± 4.2	0.83 ± 0.2	0.82 ± 0.2	0.87 ± 0.2	0.88 ± 0.1	0.04	0.07*
N_KW	16.7 ± 6.4	18.3 ± 6.4	11.1 ± 4.6	12.1 ± 5.2	11.5 ± 3.6	13.4 ± 4.3	0.80 ± 0.2	0.83 ± 0.1	0.88 ± 0.1	0.87 ± 0.2	0.09*	0.05
N_CW	17.1 ± 5.6	16.2 ± 5.3	11.5 ± 4.6	11.2 ± 4.3	11.6 ± 3.4	13.5 ± 4.0	0.79 ± 0.2	0.82 ± 0.2	0.88 ± 0.2	0.88 ± 0.1	0.08*	0.07*
M_UW	18.3 ± 6.4	17.7 ± 6.6	12.0 ± 4.8	12.1 ± 5.1	11.7 ± 3.3	13.3 ± 4.5	0.82 ± 0.1	0.83 ± 0.2	0.88 ± 0.1	0.87 ± 0.2	0.04	0.05
M_PW	18.0 ± 6.5	15.8 ± 5.2	11.6 ± 4.5	10.7 ± 3.8	11.5 ± 3.4	13.3 ± 4.0	0.78 ± 0.1*	0.87 ± 0.1	0.87 ± 0.2	0.89 ± 0.1	0.12*	0.02
M_ZW	18.8 ± 6.8	–	12.3 ± 5.3	–	11.6 ± 3.6	–	0.81 ± 0.1	–	0.88 ± 0.1	–	0.08*	–
M_BW	17.9 ± 6.1	12.8 ± 4.9	11.6 ± 4.5	9.0 ± 4.0	11.8 ± 3.4	12.5 ± 4.7	0.81 ± 0.1	0.82 ± 0.2	0.89 ± 0.1	0.87 ± 0.2	0.09*	0.03
S_MOW	14.9 ± 4.3	16.1 ± 5.1	10.2 ± 3.8	11.4 ± 4.4	11.6 ± 3.1	13.4 ± 4.0	0.80 ± 0.1*	0.85 ± 0.1	0.88 ± 0.1	0.89 ± 0.1	0.10*	0.05
Overall	23.8 ± 8.0	22.7 ± 7.8	13.7 ± 6.0	13.6 ± 5.7	20.2 ± 7.0	21.1 ± 7.3	0.80 ± 0.1	0.83 ± 0.1	0.89 ± 0.1	0.88 ± 0.1	0.08	0.05
Farmed												
N_LF	10.8 ± 4.0	15.0 ± 5.2	8.1 ± 3.5	10.1 ± 3.8	10.8 ± 4.0	13.0 ± 4.1	0.77 ± 0.2	0.83 ± 0.2	0.87 ± 0.1	0.88 ± 0.2	0.12*	0.06*
N_LFC	–	15.2 ± 4.6	–	10.7 ± 3.7	–	13.2 ± 3.8	–	0.81 ± 0.2	–	0.88 ± 0.2	–	0.08*
M_MF	17.8 ± 6.4	17.5 ± 6.2	11.5 ± 4.8	11.7 ± 4.7	11.3 ± 3.6	13.5 ± 4.4	0.80 ± 0.2*	0.84 ± 0.2	0.86 ± 0.2	0.87 ± 0.2	0.08*	0.04
S_BF	18.1 ± 6.4	18.3 ± 6.6	12.3 ± 5.0	11.8 ± 5.2	12.0 ± 3.4	13.5 ± 4.4	0.84 ± 0.1	0.81 ± 0.1	0.90 ± 0.1	0.87 ± 0.1	0.04	0.09*
Overall	20.4 ± 7.4	21.3 ± 7.8	12.9 ± 5.5	13.0 ± 5.2	20.2 ± 7.2	21.1 ± 7.8	0.81 ± 0.2	0.83 ± 0.1	0.89 ± 0.1	0.88 ± 0.1	0.08	0.07
Farm-associated												
N_LA	18.1 ± 6.7	17.8 ± 6.2	11.6 ± 4.6	11.6 ± 4.7	11.4 ± 3.3	11.4 ± 4.2	0.77 ± 0.1*	0.82 ± 0.1*	0.88 ± 0.2	0.88 ± 0.1	0.12*	0.05
M_MA	19.5 ± 6.7	18.2 ± 6.1	12.8 ± 5.3	11.4 ± 4.9	11.8 ± 3.5	13.3 ± 4.2	0.78 ± 0.1*	0.81 ± 0.2*	0.88 ± 0.2	0.87 ± 0.2	0.11*	0.06*
S_BA	17.3 ± 6.5	17.8 ± 5.7	11.9 ± 5.6	11.6 ± 4.1	11.5 ± 3.6	13.2 ± 3.7	0.79 ± 0.1*	0.80 ± 0.1*	0.88 ± 0.1	0.88 ± 0.1	0.06*	0.09*
Overall	21.8 ± 8.2	21.6 ± 7.1	13.5 ± 5.8	13.0 ± 5.0	20.2 ± 7.3	21.4 ± 7.0	0.78 ± 0.1	0.81 ± 0.1	0.88 ± 0.2	0.88 ± 0.1	0.10	0.07

*p < 0.05.

TABLE 3 | Effective population size (N_E) of *Ostrea edulis* for 12 putatively neutral microsatellite loci and tests for genetic bottlenecks using two models of microsatellite allele mutations (TPM, two phase model and SMM, stepwise model) and the Garza-Williamson index (M ratio, Garza and Williamson, 2001).

	N _E		Wilcoxon test				M ratio	
	2017	2018	TPM 2017	TPM 2018	SMM 2017	SMM 2018	2017	2018
Wild								
N_IW	710 (305, ∞)	573 (288, 9622)	0.01	0.03	0.04	ns	0.42 ± 0.07	0.45 ± 0.04
N_KW	282 (164, 904)	438 (257, 1339)	ns	0.04	ns	ns	0.41 ± 0.06	0.45 ± 0.06
N_CW	221 (137, 524)	∞ (∞, ∞)	0.01	0.03	ns	ns	0.44 ± 0.05	0.42 ± 0.04
M_UW	113 (93, 142)	375 (232, 915)	ns	0.01	ns	0.03	0.42 ± 0.07	0.45 ± 0.04
M_PW	130 (103, 171)	1691 (207, ∞)	ns	ns	ns	ns	0.45 ± 0.03	0.42 ± 0.05
M_ZW	735 (378, 7903)	–	ns	–	ns	–	0.44 ± 0.04	–
M_BW	500 (215, ∞)	∞ (770, ∞)	ns	ns	ns	ns	0.42 ± 0.05	0.42 ± 0.06
S_MOW	273 (126, ∞)	97 (76, 133)	ns	ns	ns	ns	0.4 ± 0.06	0.42 ± 0.05
Overall	2241 (1457, 4650)	2332 (1375, 7042)						
Farmed								
N_LF	–	∞ (230, ∞)	ns	ns	ns	ns	0.36 ± 0.05	0.43 ± 0.07
N_LFC	–	2651 (232, ∞)	–	ns	–	ns	–	0.42 ± 0.04
M_MF	437 (236, 2334)	∞ (913, ∞)	ns	ns	ns	ns	0.41 ± 0.05	0.43 ± 0.04
S_BF	433 (217, 8136)	∞ (673, ∞)	ns	ns	ns	ns	0.44 ± 0.04	0.42 ± 0.04
Overall	1009 (547, 5314)	1816 (608, ∞)						
Farm-associated								
N_LA	484 (258, 2866)	418 (206, ∞)	0.02	0.001	ns	ns	0.43 ± 0.03	0.44 ± 0.04
M_MA	298 (214, 479)	273 (178, 554)	0.01	0.04	ns	ns	0.43 ± 0.05	0.43 ± 0.04
S_BA	433 (403, ∞)	121 (98, 155)	0.03	0.04	ns	ns	0.42 ± 0.05	0.43 ± 0.05
Overall	493 (308, 1133)	547 (413, 794)						

Population codes are explained in Table 1. Populations with sample size smaller than 15 individuals were not included in N_E analysis. For Wilcoxon test, P-values represent one-tailed probabilities for heterozygote excess. Ns, non-significant value.

values observed within the other two population groups, wild (0.008 and 0.062) and farmed-associated (0.005 and 0.040) group, respectively. Pairwise F_{ST} and $Dest$ among populations for both sampling years are outlined in **Supplementary Table 4**. Following Bonferroni correction, 160 of 378 pairwise F_{ST} comparisons were statistically significant when permuted by Fisher's exact test. Pairwise differentiation $Dest$ values displayed the same trend as F_{ST} values after 1000 bootstraps and Bonferroni correction. Mantel test analysis showed strong correlation between two matrices ($r = 0.99$, $p = 0.01$) whereas on average $Dest$ values were 7–8 fold larger than the equivalent statistic.

Farmed samples from the north Adriatic region (17N-LF) showed high and significant pair-wise differentiation (F_{ST} : 0.006 – 0.068, $Dest$: 0.034 – 0.401) in relation to all populations included in the dataset. A similar pattern was observed for other farmed samples from the middle and south region, having less significant interactions in comparison to the north region sample. Although, several wild populations, as 17N-IW, 17N-KW and 18M-BW, showed a break in gene flow toward other populations of different origin, the majority of non-significant comparisons were found within adjacent wild populations sampled throughout the Adriatic. No clear pattern of gene flow reduction was observed among origins, regions or years, which is further supported by the AMOVA results based on F_{ST} and R_{ST} differentiation measures, where less than 0.01% of the total variance was explained by differences among groups and 99.98% was within populations. Only R_{ST} based AMOVA arranged by three geographical regions identified a significant level of sub-structuring ($p < 0.05$), but with evident gene flow among groups and within populations.

The isolation by distance (IBD) analysis for flat European oyster populations from both years revealed a low but statistically significant isolation by distance pattern ($r = 0.285$, $p = 0.04$), where the scatterplot of local densities of distances showed a consistent cloud of points with no clear-cut discontinuation, indicating a gradual cline of genetic differentiation (**Supplementary Figure 1**). However, when this global pattern of genotype distribution based on coordinates was analyzed in greater detail with spatial Analysis of Principal Components (sPCA), a significant global structure appeared (Gtest (2017): obs = 0.005, P -value = 0.01; Gtest (2018): obs = 0.003, P -value = 0.03), differentiating the north from the middle and south genetic pools (**Figure 2**). The first sPCs that were significantly positively autocorrelated in both years ($I_{2017} = 0.0115$, $p = 0.001$; $I_{2018} = 0.009$, $p = 0.02$) separated the north region populations of different origin with a bathymetric threshold of 150 m on one side (black squares) from the middle and south populations on the other side (white squares, **Figure 2**). In addition, connectivity strength of populations from middle and south regions toward the north region varied interannually, where the middle populations in 2017 showed a transition pattern of connectivity between the north-south regions, i.e., with progressive changes instead of sharp boundaries from one patch to another, supporting the isolation by distance process (**Figure 2A**).

The second sPCA scores were less representative than the first in terms of both variance and spatial autocorrelation

($I_{2017} = 0.004$, $p = 0.015$; $I_{2018} = 0.001$, $p = 0.06$), and depicted additional population structuring, i.e., gene flow discontinuity between the middle and southern populations in both years (**Figure 2**). The optimal number of clusters identified using DAPC and successive K-means clustering analyses supported three clusters (**Supplementary Figure 2**). Treating each sampling location as an *a priori* cluster, the DAPC plots showed subtle population structuring following the pattern observed by the sPCA analysis (**Figures 3A,B**). Isolation by distance across the eastern Adriatic with exclusion of two northern populations that formed separated clusters was recorded for 2017. On the other hand, in 2018 northern populations showed a central grouping pattern in PC space with overlapping clusters from both sides, i.e., populations from the middle Adriatic on one side and southern population on the other. Posterior membership probability plots showed slight heterogeneity in cluster stratification (**Figures 3C,D**), where the average assignment score of the individual to its sampling site origin ranged from 75% in 2017 to 77% in 2018. While the majority of farmed individuals showed strong assignment to their actual sampling site origin (82–86%), that was not the case with farm-associated individuals. Namely, for all three regions and for both sampling years, farm-associated individuals had assignment scores below the average (68–72%), where assignment overlapped between sampling locations, with a prevalence toward the farming sites.

DISCUSSION

Human-mediated translocation of wild or hatchery-born *O. edulis* from one location to another, for the purpose of restoration of stocks depleted by intense exploitation and/or outbreak of disease, has been increasingly used in recent decades as a conservation management strategy for endangered species (Seddon et al., 2014; Bromley, 2015). As an ecosystem engineer, *O. edulis* builds biogenic reefs and as such plays a key ecological role for the enhancement of biodiversity and ecosystem services in the marine environment (Gutiérrez et al., 2011; Pogoda et al., 2019). On the other hand, translocation of wild individuals from one location to another for commercial and farming purposes has rarely been studied in detail, even though millions of *O. edulis* have been translocated over the past 200 years (Bromley et al., 2016). Such actions may induce an increase of genetic diversity in recipient populations by mixing genetically divergent populations, and may reduce genetic divergence among geographically distant populations, as already seen in case of the black-lipped pearl oyster in French Polynesia (Lemer and Planes, 2012). In Croatian waters, translocations of juvenile oysters for farming purposes are common. Thus, in the present study, 12 neutral microsatellites loci were used to explore geographically fine-scale population processes of species during two consecutive years, to gain a deeper understanding of the factors shaping genetic connectivity, and to evaluate the impact of seed translocation among regions, by comparing 1178 sampled individuals grouped by origin (wild vs. farmed vs. farmed-associated) and sampling region (north vs. middle vs. south). To the best of our knowledge, this is the first systematic

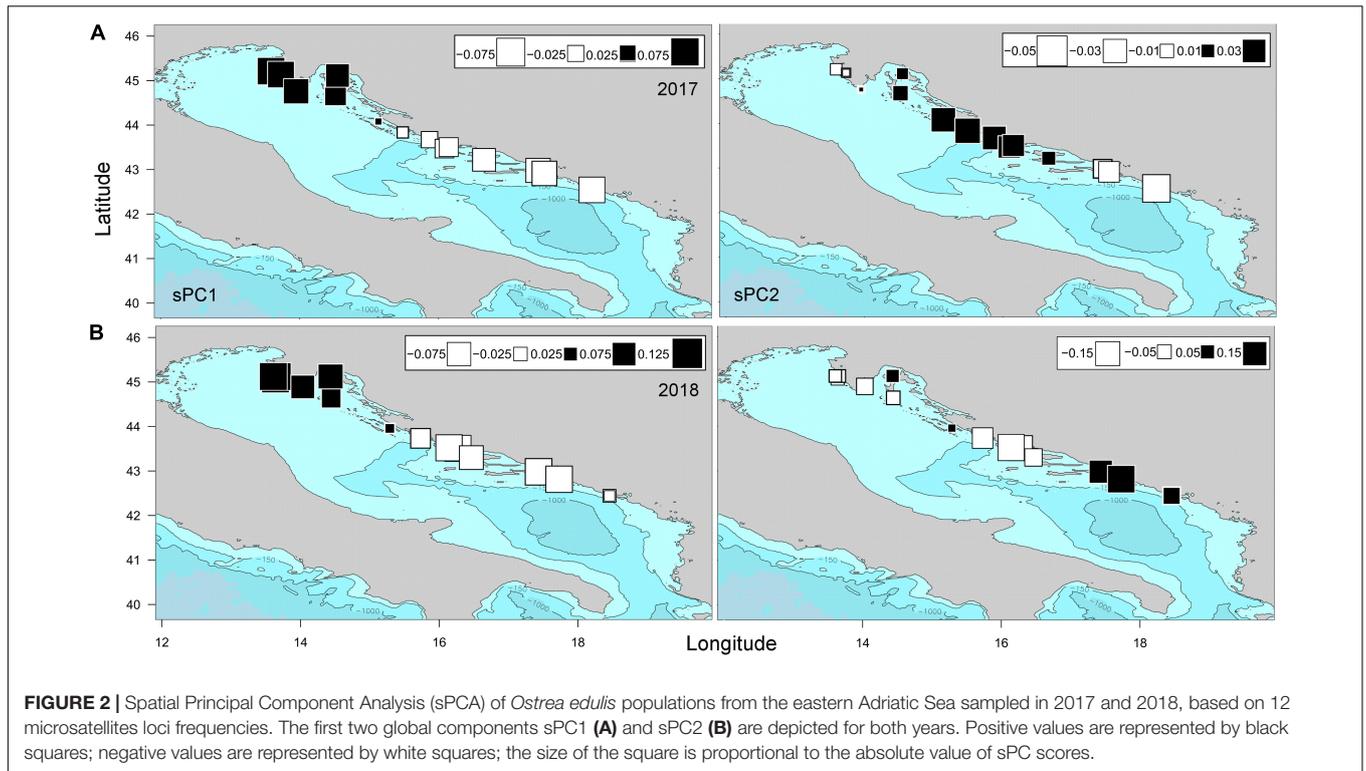


FIGURE 2 | Spatial Principal Component Analysis (sPCA) of *Ostrea edulis* populations from the eastern Adriatic Sea sampled in 2017 and 2018, based on 12 microsatellites loci frequencies. The first two global components sPC1 (A) and sPC2 (B) are depicted for both years. Positive values are represented by black squares; negative values are represented by white squares; the size of the square is proportional to the absolute value of sPC scores.

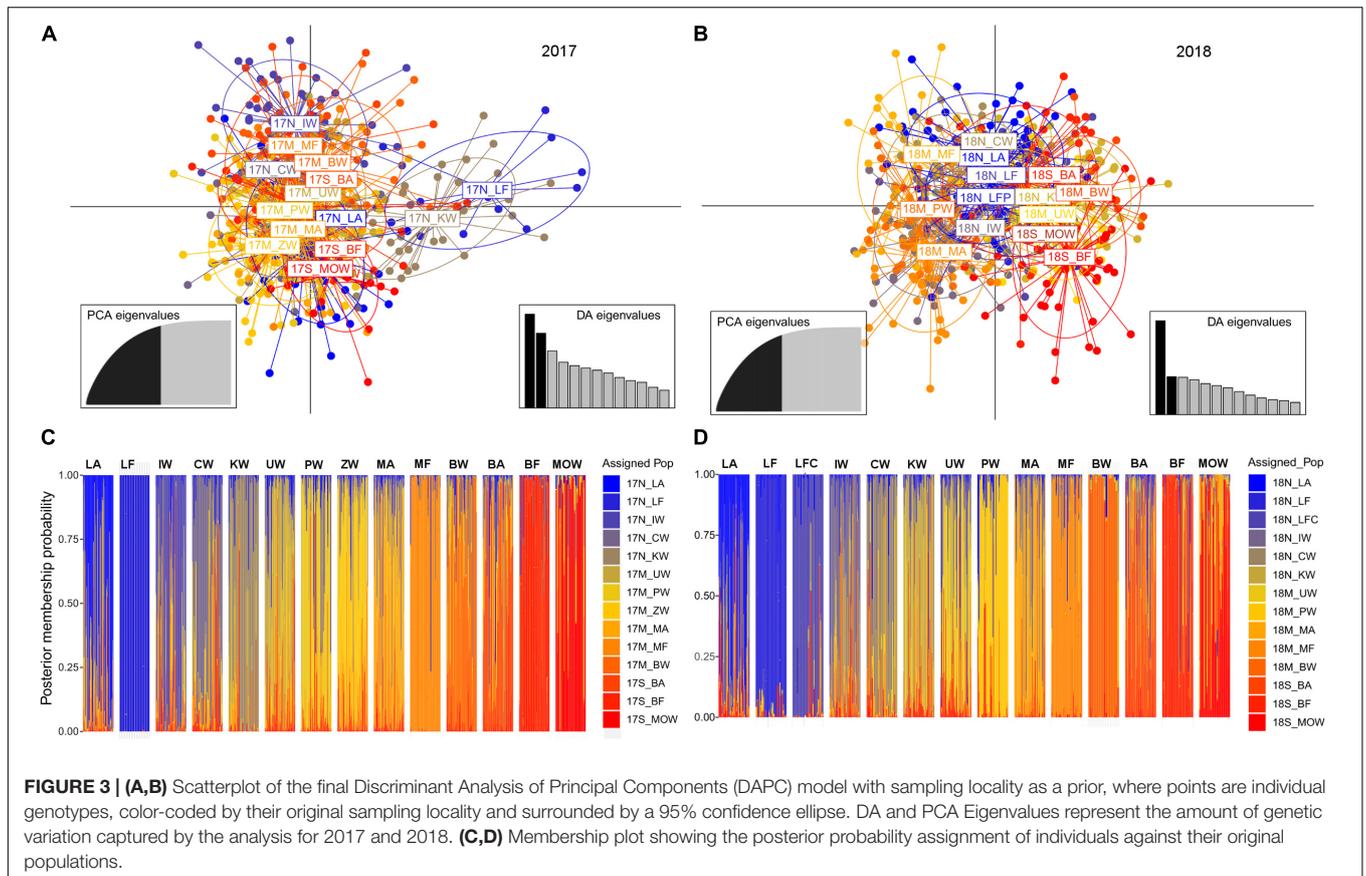


FIGURE 3 | (A,B) Scatterplot of the final Discriminant Analysis of Principal Components (DAPC) model with sampling locality as a prior, where points are individual genotypes, color-coded by their original sampling locality and surrounded by a 95% confidence ellipse. DA and PCA Eigenvalues represent the amount of genetic variation captured by the analysis for 2017 and 2018. **(C,D)** Membership plot showing the posterior probability assignment of individuals against their original populations.

description of genetic diversity and structure of the European flat oyster *O. edulis*, an ecologically and commercially important European bivalve, from the waters of the eastern Adriatic Sea.

Several main findings were outlined in this study. First, we observed a high level of diversity indices ($A_r > 20$, $H_e > 0.88$) in the dataset with the presence of slight inter-annual and inter-populations variation. This corroborated previous reports from the Adriatic and other Mediterranean regions (Launey et al., 2002; Diaz-Almela et al., 2004), even though those studies employed only several of the microsatellite markers used in this study. In addition, observed diversity levels support the findings of significantly higher diversities in flat oyster Mediterranean populations in comparison to Atlantic ones (Diaz-Almela et al., 2004; Vera et al., 2016). This can be linked with: (i) a more favorable temporal window for successful reproduction and consequently a lower variance in effective sizes in the Mediterranean (Launey et al., 2002), and (ii) the absence of oyster parasites in the eastern Adriatic which were responsible for dramatic stock declines in the waters of north Europe in the late 1960s and 1970s (Culloty and Mulcahy, 2007; Berghahn and Ruth, 2015).

While allele richness and expected heterozygosity values showed small temporal and spatial variations within the sampled Adriatic populations, observed heterozygosity was significantly reduced in farm-associated populations in comparison to others, affecting the HWE. Heterozygote deficiencies relative to HWE seem to be a common observation in marine bivalve populations (Huvet et al., 2000; Launey et al., 2002), mainly due to the high frequency of null alleles generated by the extremely high level of polymorphism in the flanking regions targeted by PCR primers (Hedgecock et al., 2004). In this study, several loci were identified as having null alleles, though at relatively low frequencies (<5%). Still, affected loci with heterozygote deficiencies were not consistently recorded in all sampled populations but were more frequently observed in farm-associated populations, explaining the significant multilocus F_{IS} estimates in all abovementioned populations. In addition, all three farm-associated populations, sampled from different geographical regions along the coast (north, middle and south) and in bays where oyster culture has a long tradition and suitable environmental conditions, showed to have contemporaneous N_e estimates significantly smaller than in other wild or farmed groups for both sampling years. The observed differences in N_e between the farmed and farm-associated population at such small scale, within the same bay, were surprising.

Settlement preferences of *O. edulis* larvae and population densities was put forward as a plausible explanation for the observed differences. Namely, the availability of a suitable settlement environment is considered a key driver for the successful recruitment of oyster populations (Möbius, 1877; Korringa, 1946; Low et al., 2007; Smyth et al., 2018) that models the larval connectivity between beds where larvae can delay metamorphosis if suitable settlement cues are absent (Cole and Jones, 1939; Coon et al., 1990). A recent study examined the chemical cues for successful settlement, indicating that most effective settlement cue originates from conspecifics, not the substrate material itself (Rodríguez-Perez et al., 2019). In

addition, increased fertilization success has been recently linked with high population densities, where oysters with a nearest neighbor ≤ 1.5 m were found to brood significantly more larvae than individuals with nearest neighbors ≥ 1.5 m (Guy et al., 2019). Since the concentrated chemical release of adult conspecifics is the driver for dense gregarious localized settlements (Tamburri et al., 2008), it can be argued that during spawning season, oyster farms with a high density of mature oysters per square meter, produce and attract *O. edulis* larvae more successfully than the adjacent wild populations that are scattered at low densities along the bays (< 1 oyster/m², Stagličić et al., 2019). Such an attraction bias was further confirmed by the slightly greater pairwise relatedness observed between farmed and wild populations compared to the farmed and impacted populations (2.7% vs. 2.6%). In the long term, spat settlement disturbance by oyster farms may seriously compromise the viability of surrounding populations, considering that: (i) with increased farming capacities and oyster densities, the fertilization success rate and concertation of chemical cues mediating successful settlement will shift in favour of farming sites, and (ii) with continued illegal harvest of wild oysters from special marine reserves put under protection due to tradition of oyster farming (i.e., Lim and Mali Ston Bays), successful fertilization and settlement may be dramatically reduced without a robust oyster population of sufficient density scale, which may eventually result in population extinction in specific areas. This is contrary to the findings in impacted populations of the black-lipped pearl oyster aquaculture, where the pearl culture promoted transmission of farmed heterogeneity to adjacent wild populations as a result of interbreeding (Lemer and Planes, 2012). In that study, adjacent wild populations tended to have higher genetic diversity values and greater pairwise relatedness coefficient with farmed populations than wild populations, contrary to our findings here. Such differences could arise from the species-specific reproduction and applied culture strategies. While the 2-year culture practice of flat oysters in Croatian waters enables species to spawn only once in the second farming year, the culture cycle of 3–6 years for pearl oysters enables multiple spawning events, starting at an average age of 2 years (Zhu et al., 2019), allowing farmed individuals more opportunity to reproduce with adjacent populations (Lemer and Planes, 2012). Additionally, *O. edulis* as brooding or partial broadcast spawners have a limited dispersal time, with larvae of relatively short planktonic phase (about 2 weeks), and therefore tend to be more aggregated around the parent population (Guy et al., 2019). On the contrary, species with a longer planktonic juvenile stage, such as *Pinctada margaritifera* (up to 4 weeks), are more prone to settling diffusely apart from the parent population due to the effect of broadcast spawning events, larval swimming time, wind and current forcing (Thomas et al., 2014). Still, these farmed-wild oyster interactions in the Adriatic Sea require further attention, since the strong bias in reproductive success, skewed sex ratio and naturally variable recruitment patterns may reduce the effective population sizes and increase inbreeding values (Hedgecock et al., 2007; Lallias et al., 2010b), affecting the abovementioned conclusions.

Molecular signatures of the occurrence of effective size reduction in the Adriatic populations of *O. edulis* were

captured for some populations, with varying results among the heterozygosity-excess and the M -ratio tests. Significant heterozygote excess in the wild populations from the northern Adriatic region and in all farm-associated populations was detected under the TPM mutation model, potentially corresponding to a bottleneck signal. Furthermore, for all populations of *O. edulis*, M ratios were far below the diagnostic value for genetic bottlenecks (0.70; Garza and Williamson, 2001). These results should be interpreted with caution, considering that high reproductive variance is typically observed in bivalves, and the interannual variance of effective population sizes in this dataset increases the rate of false positives in both bottleneck detection tests (Hoban et al., 2013). Still, demographic changes recorded in the northern Adriatic might carry alarming signals that should not be overlooked. Namely, high density flat oyster stock from deeper waters (20–40 m) in the northern Adriatic has been harvested intensively by beam trawls, where the average biomass index per trawl was 780 kg/km² in 2013 and 2014. Such large-scale removal of commercially marketable oysters led to the significant depletion of catches, resulting in a harvest of only 76 kg/km² in 2017 (Ezgeta-Balić et al., 2017). It may be that the continuous harvest of fecund broodstock in recent years increased the inter-individual distance and consequently fertilization success, while also reducing spawner production of the chemical signals that mediate successful larval settlement. The detrimental impact of intense harvesting on *O. edulis* beds throughout Europe is well documented (Hiscock et al., 2013; Thurstan et al., 2013; Smyth et al., 2016), while successful restoration of self-sustained populations is extremely difficult unless sites are closed and properly managed (Beck et al., 2011; Selkoe et al., 2015).

The low global genetic differentiation of flat oyster in the Adriatic Sea ($F_{ST} = 0.0082$) was similar to levels previously documented using polymorphic SNPs ($F_{ST} = 0.0061$; Vera et al., 2019) and 16 microsatellite loci ($F_{ST} = 0.0079$; Vera et al., 2016) on a broad sample collected from the Atlantic area. When populations from the Mediterranean and Atlantic regions were analyzed together using five microsatellite loci, inter-population differentiation was found to be slightly increased ($F_{ST} = 0.019$) and followed the isolation by distance pattern, separating populations according to geographic origin (Launey et al., 2002). The continuous cline of genetic differentiation, with slope varying among sampling years, and a lack of strong genetic structure was observed in the present study. It seems that the present culture practise and ongoing spat translocation promotes genetic heterogeneity in the investigated farms, considering that the main source of genetic differentiation in the dataset was measured in farmed populations ($F_{ST} = 0.012$) in contrast to wild ($F_{ST} = 0.008$) and farmed-associated ($F_{ST} = 0.005$) populations. Posterior membership assignment plots confirmed strong assignment of farmed populations to their sampling site origin (82–86%). Taking into account the reduced diversities and effective sizes of farm-associated populations, it can be hypothesized that unidirectional gene flow from impacted toward farmed populations promoted the absence of genetic differentiation within the farming sites, with the exaction of farmed vs. farm-associated populations from the Lim Bay in

2017. While the level of genetic differentiation greatly depends on spatial variations in population sizes (Prunier et al., 2017) where neutral and selective mechanisms can influence the adaptive potential in farm-associated populations, present study relies on a relatively small number of molecular markers that may limit the statistical probability of detecting heterogeneous patterns of introgression among markers (Putman and Carbone, 2014, and references therein). Namely, selection-driven introgression might affect relatively few markers (Fitzpatrick et al., 2010) and thus use of a dense genome-wide set of single-nucleotide polymorphism (SNP) markers is recommended for allocation of chromosomal regions under selection and detection of introgressed alleles that may be rapidly spread across native populations (Fitzpatrick et al., 2010). These functional genetic markers can reveal important processes of local adaptation among populations that may not be evident based solely on neutral genetic markers, as in case for the Lessepsian migrant *Fistularia commersoni* or for the estuarine fish *Fundulus heteroclitus* where several of the genes identified as having F_{ST} outliers were related to disease resistance, osmoregulation and thermal tolerance (Bernardi et al., 2016; Dayan et al., 2019). Additionally, human-induced environmental shifts tend to be linked with rapid polygenic adaptation that makes it difficult to identify relevant adaptive alleles (Dayan et al., 2019). Still, when population genetic distributions were analyzed in accordance to geographic coordinates by sPCA, two levels of spatial gene flow discontinuity were observed within the basin. The first level, i.e., segregation of the north from the middle and south populations in both years, presents a pattern of genetic differentiation linked with local oceanographic features. The Adriatic is characterized by a large-scale cyclonic meander, with a northerly flow from the Ionian Sea along the eastern coast and a southerly return flow along the western coast (Orlić et al., 1992), with three cyclonic gyres subdividing the basin into three regions (North, Middle and South) in line with our regional population grouping.

While the boundary between the Middle and South Adriatic basins is an area of genetic discontinuity for several marine organisms (Schiavina et al., 2014; Matic-Skoko et al., 2018; Paterno et al., 2019), in the current study the limit between the North and Middle Adriatic basins, located north of the Jabuka Pit, proved to be the main boundary for *O. edulis* populations. The north cyclonic gyre enables high larval retention and promotes homogenization within the region (Artegiani et al., 1997). The second level, which is less representative in both variance and spatial autocorrelation, recognized the weaker gene barrier between the Middle and South Adriatic basins, located north of the South Adriatic Pit. Additionally, weak grouping of the Istria populations in the north Adriatic with southern populations in 2017, and additional grouping with populations from the middle Adriatic in 2018, may represent connectivity associated with spat translocation, which has been occurring between these regions for several years.

In conclusion, human activities associated with oyster culture, stock transfer and overfishing in the eastern Adriatic have significantly mediated the demographic and genetic characteristics of wild populations, especially those associated with aquaculture sites. Several implications were derived from the

spatial robust dataset, i.e., reduced heterozygosity and effective size levels of farmed impacted beds, recent bottleneck signatures and decreased census sizes of wild populations. Thus, the first step toward sustainable shellfish aquaculture is to enable suitable substrate for successful larval recruitment in reinforcement sites where *O. edulis* is still present, but at very low densities, such as farm-impacted sites in marine protected areas in Croatia. Since oyster farming and translocation may pose an environmental risk due to the genetic erosion of wild counterparts and spread of disease, the polluter-pays principle that has been suggested for finfish mariculture should be implemented in the national legislation. A further suggestion is that the annual donation of fertile oysters and their deposition to coastal banks of farm surroundings should be an obligation for each farmer, aimed at retaining high population densities, high fertilization success and settlement of wild populations.

DATA AVAILABILITY STATEMENT

The full dataset of genotypes for this study can be found in the GenoBase of Institute of Oceanography and Fisheries (<http://jadran.izor.hr/~tsegvic/aquapop/GenoBase.html>) and is available from the corresponding authors on reasonable request.

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AUTHOR CONTRIBUTIONS

TŠ-B conceived of the study. IŽ, IT, LG, and NS conducted the sampling. TŠ-B, IŽ, IL, LŽ, and NU conducted the molecular and data analysis. TŠ-B, IT, IL, and LŽ wrote and revised the manuscript. TŠ-B and LG obtained funding for the work.

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SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: <https://www.frontiersin.org/articles/10.3389/fmars.2020.00084/full#supplementary-material>

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