



# A Meta-Analysis of Survival Outcomes Following Reoperation in Recurrent Glioblastoma: Time to Consider the Timing of Reoperation

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## **Background:** Glioblastoma multiforme (GBM) inevitably recurs, but no standard regimen has been established for recurrent patients. Reoperation at recurrence alleviates mass effects, and the survival benefit has been reported in many studies. However, in most studies, the effect of reoperation timing on survival benefit was ignored. The aim of this meta-analysis was to investigate whether reoperation provided similar survival benefits in recurrent GBM patients when it was analyzed as a fixed or time-dependent covariate.

**Methods:** A systematic literature search of PubMed, EMBASE, and Cochrane databases was performed to identify original articles that evaluated the associations between reoperation and prognosis in recurrent GBM patients.

**Results:** Twenty-one articles involving 8,630 patients were included. When reoperation was considered as a fixed covariate, it was associated with better overall survival (OS) and post-progression survival (PPS) (OS: HR = 0.66, 95% CI 0.61-0.71, p < 0.001,  $l^2 = 0\%$ ; PPS: HR = 0.70, 95% CI 0.57–0.88, p < 0.01,  $l^2 = 70.2\%$ ). However, such a survival benefit was not observed when reoperation was considered as a time-dependent covariate (OS: HR = 2.19, 95% CI 1.47–3.27, p < 0.001; PPS: HR = 0.95, 95% CI 0.82–1.10, p = 0.51,  $l^2 = 0\%$ ). The estimate bias caused by ignoring the time-dependent nature of reoperation was further demonstrated by the re-analysis of survival data in three included studies.

**Conclusions:** The timing of reoperation may have an impact on the survival outcome in recurrent GBM patients, and survival benefits of reoperation in recurrent GBM may be overestimated when analyzed as fixed covariates. Proper analysis methodology should be used in future work to confirm the clinical benefits of reoperation.

Keywords: glioblastoma, recurrence, reoperation, time-dependent covariate, survival

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# INTRODUCTION

Glioblastoma multiforme (GBM) is the most common malignant brain tumor in adults, and the tumor recurs in nearly all patients, with a median survival of 12–15 months from initial diagnosis (1). Although a standard STUPP protocol has been widely used for newly diagnosed GBM (2, 3), no such regimen is fully established for recurrent patients (4–6). Reoperation is an increasingly common treatment option for recurrent GBM patients due to useful symptom relief. In the past decade, some novel intraoperative techniques, such as intraoperative imaging, fluorescence-guided surgery and intraoperative stimulation mapping, have been developed to maximize tumor cytoreduction and minimize surgical morbidity, which significantly improved the surgical management of glioma patients (7).

In addition to improvements in neurologic symptoms and quality of life, survival time is an important indicator in assessing the prognostic value of treatment in tumor diseases. However, the survival benefit of reoperation in recurrent GBM patients remains controversial, owing to a lack of high-level evidence from high-quality randomized trials, prospective studies, and meta-analysis (8, 9). The beneficial effect of repeat resection on survival has been reported in recurrent GBM patients in some studies (10-18), but a reverse effect of reoperation was also shown in other studies (19-27). Of these studies, the survival benefit of reoperation was evaluated by overall survival (OS) or post-progression survival (PPS). Survival outcome in GBM patients is affected by many confounding factors, including age, Karnofsky performance status (KPS), tumor volume, tumor location, treatment schedule, resection extent, and O<sup>6</sup>-methylguanine-DNA methyltransferase (*MGMT*) promoter methylation as well as IDH1 mutation status. These confounding factors are incorporated into survival analyses as fixed covariates or time-dependent covariates. Fixed covariates such as sex remain unchanged during the study. Time-dependent covariates, also called time-varying covariates, may be unknown at the start of observation and may change their value over time. Reoperation is a time-dependent factor in nature but is defined as a fixed covariate in most of the GBM reoperation literature (10-19, 23, 24, 26, 28-30). Improper analysis of a time-dependent covariate as a fixed value always leads to estimate bias, favoring longer survival (21, 31, 32). Recently, a published meta-analysis demonstrated the survival benefit of reoperation in recurrent GBM patients; however, the time-dependent characteristic of reoperation was neglected (33). The aim of this study was to systematically review the survival benefit of reoperation in recurrent GBM patients when analyzed by time-dependent and non-time-dependent methods.

# METHODS

### **Search Strategy**

A comprehensive literature review was conducted to identify the articles covering the association of reoperation with prognosis in GBM from PubMed, EMBASE and the Cochrane Library. The present review was conducted according to the criteria outlined in the PRISMA (Preferred Reporting Items for Systematic reviews and Meta-Analyses) guideline (34). The articles enrolled in this analysis were published between January 1, 2005 and September 1, 2018. The Mesh terms used for literature search included "Glioblastoma," "Recurrenc," and "Reoperation." The following non-Mesh terms were also used: (1) "Glioblastomas," "Astrocytoma, Grade IV," "Astrocytomas, Grade IV," "Grade IV Astrocytoma," "Grade IV Astrocytomas," "Glioblastoma Multiforme," "Giant Cell Glioblastoma," "Giant Cell Glioblastomas," "Glioblastoma, Giant Cell," "Glioblastomas, Giant Cell"; (2) "Recurrences," "Recrudescence," "Recrudescences," "Relapse," "Relapses"; (3) "Surgical Revision," "Surgery, Repeat," "Revision, Surgical," "Revision Surgery," "Revision Surgeries," "Surgery, Revision," "Repeat Surgery," "Revision, Joint," "Joint Revision," "Repeat operation," "Repeat resection," and "Second surgery," Reviews as well as the references of the included studies were also checked to avoid the omission of relevant publications. The eligible studies were restricted to human beings.

## **Inclusion and Exclusion Criteria**

Articles that met the following criteria were included in this study: (1) studies investigating the relationship between reoperation and survival in recurrent GBM patients; (2) multivariate Cox-proportional hazard regression models were used to minimize the bias from baseline confounding factors; and (3) hazard ratios (HR) and 95% confidence intervals (CI) for survival time were available directly. We excluded publications as follows: (1) reviews, laboratory research, and animal experiments; and (2) studies published only in abstract.

# Study Selection and Data Extraction

Study selection was independently performed by two authors, and disagreements were resolved through discussion. The following data were extracted: the first author's name, country of origin, publication year, number of patients, adjustment method, and outcomes (including HRs and 95% CIs).

# **Quality Assessment**

The Newcastle Ottawa Scale (NOS) was used to assess the quality and risk of bias of the included studies (35). Three main aspects of quality were evaluated, including study selection, comparability, and study outcome. The assessment of the included studies was independently conducted by two authors. High scores indicate high quality and low risk of bias.

# **Data Analysis**

The statistical analysis was performed with STATA 12.0 software (StataCorp, College Station, TX, USA). The statistical heterogeneity among studies was assessed by a Q-test and  $I^2$  statistics. If there was significant heterogeneity, the random effect model was used to estimate the pooled HR; otherwise, the fixed effect model was used. Publication bias was assessed by funnel plots and Egger's test. Sensitivity analysis by deleting each included study was conducted to assess the overall robustness of the meta-analysis results.

To illustrate the estimate bias that occurs when a timedependent covariate is treated as known at the start of the study, the survival curves of three included studies (12, 15, 25), in which reoperation was considered as a fixed covariate in the multivariate Cox model, were plotted and reanalyzed by three methods. The traditional Kaplan-Meier method treats the covariate as a fixed value at the start of the study. The landmark method selects a fixed time point after the start of the study as a landmark for conducting the analysis, and it is a valid method to test whether the survival is related to patients' status at a specific time point (36, 37). We selected the 50th percentile and 75th percentile of time between the first and second surgery as the landmark time, as described by Goldman (21). Only patients alive at the landmark time were included in the analysis, and they were separated into reoperation and non-reoperation groups according to their resection status by the landmark time. Then, group comparisons were conducted with the Breslow test. Survival data were also analyzed by the Simon-Makuch method (38), which takes into account the change in covariate status

over time and properly represents the effect of a time-dependent covariate on survival. In the Simon-Makuch method, patients move from non-reoperation group to reoperation group at the time their reoperation occurs. Therefore, the number of patients changes over time.

## RESULTS

#### **Characteristics of the Studies**

The flow chart of the literature selection is presented in **Figure 1**. A total of 738 articles were screened, and 21 articles comprising 8,630 patients were identified (10–28, 39, 40). Of these publications, two articles comprising two independent trials (marked with I and II in the figures and tables) were extracted as four studies (20, 22). Therefore, a total of 23



TABLE 1 | Characteristics of included studies for overall survival analysis.

| Name                          | Study type    | Country     | Patients number | Treatment                                      |  |  |  |
|-------------------------------|---------------|-------------|-----------------|--|--|--|--|
| Chaichana et al. (A) (16)     | Retrospective | America     | 522             | Reoperation+/RT+/chemo vs. /RT+/chemo          |  |  |  |
| Chaichana et al. (B) (16)     | Retrospective | America     | 395             | Multiple reoperation+/RT+/chemo vs. /RT+/chemo |  |  |  |
| Chaichana et al. (C) (16)     | Retrospective | America     | 369             | Third reoperations+/RT+/chemo vs. /RT+/chemo   |  |  |  |
| Chen et al. (A) (11)          | Retrospective | America     | 3882            | Reoperation+/chemo+/RT vs. / RT+/chemo         |  |  |  |
| Chen et al. (B) (11)          | Retrospective | America     | 3575            | Multiple reoperation+/chemo+/RT vs. RT+/chemo  |  |  |  |
| Delgado-Fernandez et al. (12) | Retrospective | Spain       | 121             | Reoperation+ RT+ chemo vs. RT+/chemo           |  |  |  |
| Goldman et al. (21)           | Retrospective | America     | 163             | Reoperation+/RT+/chemo vs. /RT+/chemo          |  |  |  |
| Ortega et al. (A) (23)        | Retrospective | America     | 177             | Reoperation+ RT+/chemo vs. RT+/chemo           |  |  |  |
| Ortega et al. (B) (23)        | Retrospective | America     | 108             | Multiple reoperation+ RT+/chemo vs. RT+/chemo  |  |  |  |
| Skeie et al. (13)             | Retrospective | Norway      | 51              | Reoperation+ RT+/chemo vs. RT+/chemo           |  |  |  |
| Stark et al. (17)             | Retrospective | Germany     | 122             | Reoperation+/chemo +/RT vs. /chemo+/RT         |  |  |  |
| Tugcu et al. (26)             | Retrospective | Turkey      | 50              | Reoperation+/RT vs. /RT                        |  |  |  |
| Tully et al. (40)             | Retrospective | America     | 204             | Reoperation+/chemo+/ RT vs. /chemo+/RT         |  |  |  |
| Woernle et al. (25)           | Retrospective | Switzerland | 98              | Reoperation+/chemo vs. /chemo                  |  |  |  |

RT, radiotherapy; chemo, chemotherapy.

studies were included in this analysis. The characteristics of all included studies are summarized in **Tables 1**, **2**. The quality assessment showed that all studies were of high quality (see **Supplemental Table S1**).

#### **Publication Bias**

Publication bias was evaluated by Egger's test. No publication bias was observed in the analysis of OS (p = 0.69) or PPS (p = 0.15; see **Supplemental Figure S1**).

#### **Sensitivity Analysis**

Sensitivity analysis was conducted by sequentially omitting individual studies to assess whether a single study might significantly affect the overall results. Sensitivity analysis showed no apparent variations in the pooled HR of OS or PPS (**Supplemental Figure S2**), supporting the robustness of the primary findings.

# Impact of Reoperation at Recurrence on OS

In 10 studies (11–13, 16, 17, 21, 23, 25, 26, 40), survival was defined as the time from the first surgery to death or the end of follow-up, and it indicates OS. Our analysis showed that recurrent patients who underwent reoperation had significantly prolonged OS than those who did not (HR = 0.71, 95% CI 0.60–0.85, p < 0.001,  $I^2 = 70\%$ , **Figure 2A**). GBM patients may undergo more than one resection at recurrence. The impact of one reoperation and multiple reoperations on OS was reported in 10 studies (11–13, 16, 17, 21, 23, 25, 26, 40) and 3 studies (11, 16, 23), respectively. Our analysis showed that both one resection and multiple reserves in final studies (11, 16, 23), respectively. Our analysis showed that both one resection and multiple reserves is split. HR = 0.75, 95% CI 0.59–0.95, p = 0.02,  $I^2 = 77.3\%$ ; multiple reoperation: HR = 0.59, 95% CI 0.50–0.71, p < 0.001,  $I^2 = 0\%$ , **Supplemental Figure S3**).

In these studies, reoperation was incorporated into the multivariate Cox model as a fixed covariate in 9 studies (11–13, 16, 17, 23, 25, 26, 40) and as a time-dependent covariate in one

study (21). The difference between the covariates was whether the timing of reoperation was taken into consideration. Then, we evaluated whether reoperation had a similar impact on OS when analyzed with time-dependent and non-time-dependent methods. Reoperation was associated with longer OS when the timing was ignored (HR = 0.66, 95% CI 0.61–0.71, p < 0.001,  $I^2$ = 0%, **Figure 2B**) and was not associated with better OS when timing was taken into account (HR = 2.19, 95% CI 1.47–3.27, p< 0.001, **Figure 2B**).

Given the sparse number of studies with time-dependent analyses, more evidence was needed to demonstrate the estimate bias occurring in survival analysis, where a time-varying covariate is treated as a time-fixed one. The survival data of the three included studies (12, 15, 25) were reanalyzed by the traditional Kaplan-Meier method, landmark method and Simon-Makuch method. Traditional Kaplan-Meier curves showed that patients with reoperation had prolonged OS compared with those without reoperation (Figure 3). However, the landmark analysis showed that survival was not associated with better OS at two specific landmark times (Figure 3). Similarly, survival analysis by the Simon-Makuch method did not shown long OS in patients with reoperation (Woernle et al: HR = 1.45, 95% CI 0.91–2.28, p = 0.12; Zanello et al: HR = 0.99, 95% CI 0.77–1.26, p = 0.91; Delgado-Fernandez et al: HR = 0.79, 95% CI 0.49-1.28, p = 0.91). These results indicated that ignoring the time-dependent nature of reoperation might lead to estimate bias.

# Impact of Reoperation at Recurrence on PPS

When deciding the clinical benefits of treatment at tumor recurrence, to some extent, PPS may be more appropriate for evaluating the survival effect of treatment. In 14 studies (10, 13–15, 18–20, 22, 24, 27, 28, 39), survival was defined as the time from the first recurrence to death or the end of follow-up, and it indicates PPS. There were 11 studies that evaluated the impact of reoperation on PPS (10, 13–15, 18–20, 22, 24, 27). Consistent

 TABLE 2 | Characteristics of included studies for post-progression survival analysis.

| Name                       | Study type    | Country   | Patients number | Treatment                               |  |  |  |
|----------------------------|---------------|-----------|-----------------|---|--|--|--|
| Azoulay et al. (18)        | Retrospective | Canada    | 78              | Reoperation+chemo+/ RT vs. chemo +/ RT  |  |  |  |
| Boiardi et al. (10)        | Retrospective | Italy     | 211             | Reoperation+chemo vs. chemo             |  |  |  |
| Clarke et al. (19)         | Prospective   | America   | 593             | Reoperation+/chemo vs. /chemo           |  |  |  |
| Filippini et al. (I) (20)  | Retrospective | Italy     | 452             | Reoperation+/chemo +/ RT vs. /chemo+/RT |  |  |  |
| Filippini et al. (II) (20) | Retrospective | Italy     | 435             | Reoperation+/chemo +/ RT vs. /chemo+/RT |  |  |  |
| Nava et al. (I) (22)       | Prospective   | Italy     | 203             | Reoperation+/chemo +/ RT vs. /chemo+/RT |  |  |  |
| Nava et al. (II) (22)      | Prospective   | Italy     | 303             | Reoperation+/chemo +/ RT vs. /chemo+/RT |  |  |  |
| Skeie et al. (13)          | Retrospective | Norway    | 51              | Reoperation+/chemo+RT vs. RT+/chemo     |  |  |  |
| Suchorska et al. (24)      | Prospective   | Germany   | 105             | Reoperation+chemo vs. chemo             |  |  |  |
| Vann et al. (14)           | Retrospective | Australia | 117             | Reoperation+/chemo vs. /chemo           |  |  |  |
| Zanello et al. (15)        | Retrospective | France    | 777             | Reoperation+/RT+/chemo vs./chemo+/RT    |  |  |  |
| Sastry et al. (27)         | Retrospective | America   | 368             | Reoperation+chemo+/RT vs. chemo +/RT    |  |  |  |
| Kim et al. (39)            | Retrospective | Korea     | 38              | Reoperation+chemo+RT                    |  |  |  |
| VcGirt et al. (28)         | Retrospective | America   | 285             | Reoperation+chemo+RT                    |  |  |  |

RT, radiotherapy; chemo, chemotherapy.

| Study ID                                    |                | HR for OS (95% CI) | Favors<br>resection | Favors non-<br>resection | Weight% |
|---|----------------|--------------------|---------------------|--------------------------|---------|
| Chaichana et al., 20                        | 13 (A)         | 0.69(0.53-0.90)    |                     |                          | 9.81    |
| Chaichana et al., 20                        | . ,            | 0.61(0.40-0.95)    |                     |                          | 7.16    |
| Chaichana et al., 20                        | . ,            | 0.60(0.32-1.14)    |                     | -                        | 4.79    |
| Chen et al., 2018 (A                        |                | 0.67(0.60-0.74)    |                     |                          | 12.10   |
| Chen et al., 2018 (B                        | )              | 0.57(0.45-0.71)    | -                   |                          | 10.49   |
| Delgado-Fernandez                           | et al., 2017   | 0.48(0.28-0.81)    |                     |                          | 6.03    |
| Goldman et al., 201                         | 3              | 2.19(1.47-3.27)    |                     |                          | 7.67    |
| Ortega et al., 2016 (                       | A)             | 1.00(0.70-1.43)    | -                   | -                        | 8.36    |
| Ortega et al., 2016 (                       | B)             | 0.75(0.43-1.31)    |                     | _                        | 5.62    |
| Skeie et al., 2012 (A                       | .)             | 0.48(0.22-1.01)    |                     |                          | 3.85    |
| Stark et al., 2007                          |                | 0.55(0.30-1.02)    |                     |                          | 4.95    |
| Tugcu et al., 2010                          |                | 0.76(0.34-1.72)    |                     |                          | 3.42    |
| Tully et al., 2016                          |                | 0.65(0.50-0.84)    |                     |                          | 9.87    |
| Woernle et al., 2015                        |                | 0.66(0.39-1.13)    |                     | -                        | 5.87    |
| Overall(I <sup>2</sup> =70.0%, p·           | <0.001)        | 0.71(0.60-0.85)    | $\diamond$          |                          | 100.00  |
| В   |                |                    | .5 1                | 3                        |         |
| Study ID                                    |                | HR for OS (95% CI) | Favors<br>resection | Favors non-<br>resection | Weight% |
| Reoperation as a fix                        | ed covariate   |                    |                     |                          |         |
| Chaichana et al., 20                        | 13 (A)         | 0.69(0.53-0.90)    | -                   |                          | 9.81    |
| Chaichana et al., 20                        | 13 (B)         | 0.61(0.40-0.95)    |                     |                          | 7.16    |
| Chaichana et al., 20                        | 13 (C)         | 0.60(0.32-1.14)    |                     |                          | 4.79    |
| Chen et al., 2018 (A                        | )              | 0.67(0.60-0.74)    | •                   |                          | 12.10   |
| Chen et al., 2018 (B                        | )              | 0.57(0.45-0.71)    | -                   |                          | 10.49   |
| Delgado-Fernandez                           | et al., 2017   | 0.48(0.28-0.81)    |                     |                          | 6.03    |
| Ortega et al., 2016 (                       | A)             | 1.00(0.70-1.43)    | -                   | -                        | 8.36    |
| Ortega et al., 2016 (                       | B)             | 0.75(0.43-1.31)    |                     | -                        | 5.62    |
| Skeie et al., 2012 (A                       | .)             | 0.48(0.22-1.01)    |                     |                          | 3.85    |
| Stark et al., 2007                          |                | 0.55(0.30-1.02)    | -                   |                          | 4.95    |
| Tugcu et al., 2010                          |                | 0.76(0.34-1.72)    |                     |                          | 3.42    |
| Tully et al., 2016                          |                | 0.65(0.50-0.84)    | -                   |                          | 9.87    |
| Woernle et al., 2015                        |                | 0.66(0.39-1.13)    | ~                   |                          | 5.87    |
| Subtotal(I <sup>2</sup> =0%, p=0            | .61)           | 0.66(0.61-0.71)    | v                   |                          | 92.33   |
|   | ne-dependent c | ovariate           |                     |                          |         |
| Reoperation as a tin                        | -              | 2.19(1.47-3.27)    |                     |                          | 7.67    |
| Reoperation as a tin<br>Goldman et al., 201 | 3              | ,                  |                     |                          |         |
|   |                | 2.19(1.47-3.27)    |                     | $\sim$                   | 7.67    |

FIGURE 2 | Forest plots showing pooled results of overall survival analysis in recurrent patients with vs without reoperation (A) and when reoperation was considered as a fixed covariate vs. as a time-dependent covariate (B).

with OS, reoperation at recurrence was associated with improved PPS (HR = 0.78, 95% CI 0.66–0.92, p < 0.01,  $I^2 = 66.5\%$ , **Figure 4A**). In these studies, reoperation was incorporated as a fixed covariate in 8 studies (10, 13–15, 18, 19, 24, 27) and as a time-dependent covariate in 3 studies (20, 22). When the timing of reoperation was ignored, reoperation was also associated with long PPS (HR = 0.70, 95% CI 0.57–0.88, p < 0.01,  $I^2 = 70.2\%$ , **Figure 4B**). However, patients with reoperation had similar PPS as those without reoperation when the timing of reoperation was taken into account (HR = 0.95, 95% CI 0.82–1.10, p = 0.51,  $I^2 = 0\%$ , **Figure 4B**).

Of the 11 studies, 6 studies incorporated the time from the initial diagnosis to first the recurrence into the multivariate Cox

model (10, 13, 20, 22, 27). When the time to first recurrence was not considered, patients who underwent reoperation had a prolonged PPS compared with those who did not (HR = 0.70, 95% CI 0.53–0.93, p = 0.02,  $l^2 = 78.4\%$ , Figure 5). However, comparable PPS was observed between patients with and without reoperation when the time to first recurrence was taken into account (HR = 0.86, 95% CI 0.72–1.03, p = 0.10,  $l^2 = 45.2\%$ , Figure 5).

We further evaluated the impact of time to recurrence on PPS in GBM patients who underwent reoperation. Of the included studies, the cutoff value of the earlier recurrence was defined as 9 months in three studies (22, 39), as 10 months in one study (13) and as 12 months in one study (28). Among



FIGURE 3 | Survival curve plotted by the Kaplan-Meier and landmark methods. The survival data from the studies of Woernle et al. (25) (A), Zanello et al. (15) (B), and Delgado-Fernandez et al. (12) (C) were reanalyzed by the Kaplan-Meier (left) and landmark methods (middle and right). The Kaplan-Meier curve started at the initial diagnosis (left), and the landmark curve started at the 50th percentile (middle) and 75th percentile (right) of time between first and second surgery. In the landmark method, patients were considered to have reoperation if their reoperation occurred by the landmark time and were considered to not have reoperation if they did not undergo reoperation or their reoperation occurred after the landmark time.

GBM patients who underwent reoperation, patients with earlier recurrence had a higher risk of death than those with later recurrence (HR = 1.42, 95% CI 1.22–1.7, p < 0.001,  $l^2 = 0\%$ , **Figure 6**).

# DISCUSSION

Neurosurgeons always pay close attention to OS and progressionfree survival (PFS) when evaluating the clinical benefits of the

| А |   |                                       | Favors               | Favors non-        |  |  |
|---|---|---------------------------------------|----------------------|--------------------|--|--|
| - | Study ID  | HR for PPS (95% CI)                   | resection            | resection          | Weight%                                |  |
|   | Azoulay et al., 2017  | 0.53(0.35-0.83)                       |                      |                    | 7.20                                   |  |
|   | Boiardi et al., 2008  | 0.64(0.46-0.89)                       |                      |                    | 9.29                                   |  |
|   | Clarke et al., 2011   | 1.05(0.87-1.27)                       | 1                    | -                  | 12.57                                  |  |
|   | Filippini et al., 2008 (I)  | 0.98(0.77-1.25)                       | 1                    | ÷                  | 11.32                                  |  |
|   | Filippini et al., 2008 (II)   | 0.86(0.66-1.12)                       |                      | -                  | 10.79                                  |  |
|   | Nava et al., 2014 (II)  | 1.02(0.77-1.35)                       | 1                    | -                  | 10.49                                  |  |
|   | Skeie et al., 2012(B)   | 0.42(0.19-0.92)                       | <b>← *</b> →         |                    | 3.23                                   |  |
|   | Suchorska et al., 2016  | 0.73(0.45-1.19)                       |                      | -                  | 6.31                                   |  |
|   | Wann et al., 2018   | 0.53(0.35-0.80)                       |                      |                    | 7.72                                   |  |
|   | Zanello et al., 2017  | 0.67(0.56-0.81)                       | -                    |                    | 12.68                                  |  |
|   | Sastry et al., 2018   | 0.98(0.67-1.41)                       |                      | -                  | 8.40                                   |  |
|   | Overall (I <sup>2</sup> =66.5%, p=0.001)                                      | 0.78(0.66-0.92)                       | $\Diamond$           |                    | 100.00                                 |  |
|   |   |                                       | .5 1                 | 1.5                |  |  |
| В |   |                                       | Favors               | Favors non-        |  |  |
|   | Study ID  | HR for PPS (95% CI)                   | resection            | resection          | Weight%                                |  |
|   | Reoperation as a fixed covariat   | е                                     |                      |                    |  |  |
|   | Azoulay et al., 2017  | 0.53(0.35-0.83)                       | -                    |                    | 7.20                                   |  |
|   | Boiardi et al., 2008  | 0.64(0.46-0.89)                       | -                    |                    | 9.29                                   |  |
|   | Clarke et al., 2011   | 1.05(0.87-1.27)                       |                      | <b>-</b>           | 12.57                                  |  |
|   | Skeie et al., 2012(B)   | 0.42(0.19-0.92)                       | ← * →                |                    | 3.23                                   |  |
|   | Suchorska et al., 2016  | 0.73(0.45-1.19)                       |                      | -                  | 6.31                                   |  |
|   | Wann et al., 2018   | 0.53(0.35-0.80)                       | -                    |                    | 7.72                                   |  |
|   | Zanello et al., 2017  | 0.67(0.56-0.81)                       |                      |                    | 12.68                                  |  |
|   | Sastry et al., 2018   | 0.98(0.67-1.41)                       |                      | <b>i</b>           | 8.40                                   |  |
|   | Subtotal (I <sup>2</sup> =70.2%, p=0.001)                                     | 0.70(0.57-0.88)                       | $\diamond$           |                    | 67.40                                  |  |
|   | Reoperation as a time-dependent covariate                                     |                                       |                      |                    |  |  |
|   | Filippini et al., 2008 (I)  | 0.98(0.77-1.25)                       | E Contraction (1997) | +                  | 11.32                                  |  |
|   | Filippini et al., 2008 (II)   | 0.86(0.66-1.12)                       |                      | T                  | 10.79                                  |  |
|   | Nava et al., 2014 (II)  | 1.02(0.77-1.35)                       |                      | -                  | 10.49                                  |  |
|   | Subtotal (I <sup>2</sup> =0%, p=0.65)   | 0.95(0.82-1.10)                       |                      | P                  | 32.60                                  |  |
|   | Overall (l <sup>2</sup> =66.5%, p=0.001)                                      | 0.78(0.66-0.92)                       | <b></b>              |                    | 100.00                                 |  |
|   |   | , , , , , , , , , , , , , , , , , , , | .5 1                 | 1.5                |  |  |
|   |   |                                       |                      |                    | •••••••••••••••••••••••••••••••••••••• |  |
|   | wing pooled results of post-progression ate vs. as a time-dependent covariate |                                       | nts with vs. witho   | out reoperation (A | A) and when reoperation was            |  |

primary regimen in tumor patients. Due to the inherent bias arising from prolonged pre-progression survival in patients who undergo repeat resections, the survival benefit of reoperation at tumor progression may be overestimated by measuring OS (41). PPS may be more appropriate to evaluate the survival effect of repeat resections at tumor recurrence. In this study, we conducted a comprehensive analysis to evaluate the association of reoperation with OS and PPS in recurrent GBM patients.

We first analyzed all the included studies and found that reoperation improved both OS and PPS in recurrent patients. In addition, both one resection and multiple resections upon recurrence significantly prolonged OS. These results seem to strongly support surgical management in recurrent GBM patients. However, further analysis showed that better OS and PPS were observed only when reoperation was taken into account as a fixed covariate. We noticed that reoperation was considered a time-dependent covariate in few studies (20-22). To further demonstrate the reliability of our results, survival data in the three included studies, in which the timing of reoperation was not considered, were reanalyzed by treating reoperation as a fixed and a time-dependent covariate, respectively. Consistently, the association of reoperation with improved OS was not found when analyzed by either the landmark or Simon-Makuch method, in which the time-dependent effects of reoperation were considered. In fact, time from initial diagnosis to recurrence or reoperation is unpredictable and variable in GBM patients. Therefore, reoperation is a time-dependent variable in nature. Time-dependent methodology has been well-established in survival analysis (42) and had been used in prognosis analysis of cancers, such as colon cancer (43), breast cancer (44) and leukemic evolution (45). Previous studies reported that ignoring time-dependent bias inevitably overestimated the survival benefit and led to results supporting longer survival (21, 31, 32, 46).

| Study ID                                  | HR for PPS (95% CI) | Favors<br>resection | Favors non-<br>resection | Weight% |
|---|---------------------|---------------------|--------------------------|---------|
| Non-accounting for time to rec            | urrence             | -                   |                          |         |
| Azoulay et al., 2017                      | 0.53(0.35-0.83)     | - <del>* (</del>    | I                        | 7.20    |
| Clarke et al., 2011                       | 1.05(0.87-1.27)     | +                   | ÷                        | 12.57   |
| Suchorska et al., 2016                    | 0.73(0.45-1.19)     |                     | F                        | 6.31    |
| Wann et al., 2018                         | 0.53(0.35-0.80)     |                     |                          | 7.72    |
| Zanello et al., 2017                      | 0.67(0.56-0.81)     |                     |                          | 12.68   |
| Subtotal (I <sup>2</sup> =78.4%, p=0.001) | 0.70(0.53-0.93)     | $\diamond$          |                          | 46.49   |
| Accounting for time to recurrent          | ce                  |                     |                          | 12.68   |
| Boiardi et al., 2008                      | 0.64(0.46-0.89)     | -                   |                          | 9.29    |
| Filippini et al., 2008 (I)                | 0.98(0.77-1.25)     | 1                   | ÷.                       | 11.32   |
| Filippini et al., 2008 (II)               | 0.86(0.66-1.12)     | -                   | t                        | 10.49   |
| Nava et al., 2014 (II)                    | 1.02(0.77-1.35)     |                     | <del>.</del>             | 3.23    |
| Skeie et al., 2012 (B)                    | 0.42(0.19-0.92)     | ( <b>*</b>          | Į.                       | 8.40    |
| Sastry et al., 2018                       | 0.98(0.67-1.41)     |                     | <b></b>                  | 10.49   |
| Subtotal (I <sup>2</sup> =45.2%, p=0.10)  | 0.86(0.72-1.03)     | X                   | 1                        | 53.51   |
| Overall (l <sup>2</sup> =66.5%, p=0.001)  | 0.78(0.66-0.92)     |                     |                          | 100.00  |
|   |                     | 5 1                 | 1.5                      |         |

FIGURE 5 | Forest plots for post-progression survival analysis when accounting for vs. not accounting for the time to first recurrence.



Therefore, the timing of reoperation should be included to evaluate the impact of reoperation on survival in GBM patients.

The importance of the timing of reoperation on survival is also supported by the high risk of death in GBM patients with earlier recurrence. Our results showed that the time to recurrence had a significant impact on the prognostic value of reoperation. The time of recurrence means, to some extent, the time of reoperation. Therefore, the high risk of death in tumor patients with earlier recurrence also underscored the importance of timedependent methodology in survival analysis. GBM patients who experience earlier recurrence usually have more a progressive phenotype of the tumor and/or molecular risk factors such as MGMT promoter unmethylation. These tumor histological and molecular characteristics are independently associated with worse survival. Both tumor-related and treatment-related factors could affect clinical outcomes, but the presence of tumor-related characteristics may have a greater effect on tumor progression and survival than reoperation itself. When the time to recurrence or reoperation was not considered, the impact of some tumorrelated characteristics might be ignored in survival analysis. These understandings can help us to correctly interpret the worse PPS observed in patients with early recurrence. Our results do not mean that reoperation is useless or harmful to patients with early progression. The worse PPS was primarily due to the more progressive characteristics of the tumor rather than reoperation. Therefore, cytotoxic agents in combination with reoperation may be more effective in improving clinical outcomes than reoperation alone. In fact, combining carmustine wafer implantation with surgical resection prolonged OS in newly diagnosed glioblastoma (47, 48). However, its effectiveness remains to be elucidated in recurrent GBM.

# LIMITATIONS

There are several limitations to the present study. First, heterogeneity existed in the pooled OS and PPS analyses. Heterogeneity in the OS analysis resulted from the study incorporating reoperation as a time-dependent covariate. Variation in adjusted confounding factors was the major cause for heterogeneity in PPS analysis. However, sensitive analysis supported the robustness of our primary findings. Second, only a few studies included reoperation as a time-dependent factor. Although all the included studies were of high quality, most of them are retrospective studies and evidence of prospective, randomized control trials was lacking in this meta-analysis. Third, we did not get the entirety of patient data in the three included studies; thus, univariable but not multivariable analyses were conducted. Both the landmark and Simon-Makuch methods have their own limitations, and the results cannot be used to make inferences (36, 49). For instance, the landmark method ignores the patient's status after the specific time and all deaths before that time. Fourth, the molecular marker profile, such as MGMT promoter methylation and IDH1 mutation, is associated with clinical outcomes of GBM patients. Unfortunately, few studies incorporated these molecular markers into the multivariate Cox model (18, 22, 24). Fifth, the first-time treatment regimen and its feasibility may affect the decision of the treatment regimen at progression, which could lead to selection bias. In the case of patients with early recurrence, the chance to have radiotherapy and chemotherapy is low due to the short interval from the previous treatment. In addition, other treatment at recurrence may also influence the survival effects of reoperation. These concerns are not investigated in this study.

# CONCLUSIONS

The timing of reoperation may have an impact on the survival outcome in recurrent GBM patients, and survival benefits of reoperation in recurrent GBM may be overestimated when analyzed as fixed covariates. Proper analysis methodology

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should be used in future work to confirm the clinical benefit of reoperation.

## **AUTHOR CONTRIBUTIONS**

Y-HZ collected data, performed the statistical analysis, and drafted the report. Z-YP participated in data collection analysis. DP, JD-F, and JP participated in the interpretation of the data and critically reviewed the manuscript. Z-FW and Z-QL designed the study and contributed to interpretation and discussion of the results. All authors read and approved the final manuscript.

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### SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fneur. 2019.00286/full#supplementary-material

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**Conflict of Interest Statement:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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