Check for updates

OPEN ACCESS

EDITED BY Juei-Tang Cheng, Chang Jung Christian University, Taiwan

REVIEWED BY Oleksandr Serhiyenko, Danylo Halytsky Lviv National Medical University, Ukraine Ali Gohari, Sabzevar University of Medical Sciences, Iran

*CORRESPONDENCE Chunguang Xie, xiecg@cdutcm.edu.cn Hong Gao, cdgh76@163.com

[†]These authors have contributed equally to this work and share first authorship

SPECIALTY SECTION

This article was submitted to Ethnopharmacology, a section of the journal Frontiers in Pharmacology

RECEIVED 12 April 2022 ACCEPTED 27 June 2022 PUBLISHED 08 August 2022

CITATION

Zhang Z, Leng Y, Fu X, Yang C, Xie H, Yuan H, Liang Q, Gao H and Xie C (2022), The efficacy and safety of dachaihu decoction in the treatment of type 2 diabetes mellitus: A systematic review and meta-analysis. *Front. Pharmacol.* 13:918681. doi: 10.3389/fphar.2022.918681

COPYRIGHT

© 2022 Zhang, Leng, Fu, Yang, Xie, Yuan, Liang, Gao and Xie. This is an open-access article distributed under the terms of the Creative Commons Attribution License (CC BY). The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner(s) are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.

The efficacy and safety of dachaihu decoction in the treatment of type 2 diabetes mellitus: A systematic review and meta-analysis

Zehua Zhang^{1†}, Yulin Leng^{1†}, Xiaoxu Fu^{1.2†}, Chan Yang³, Hongyan Xie^{1,2}, Haipo Yuan¹, Qingzhi Liang¹, Hong Gao^{1,2}* and Chunguang Xie^{1.2}*

¹Hospital of Chengdu University of Traditional Chinese Medicine, Chengdu, China, ²TCM Regulating Metabolic Diseases Key Laboratory of Sichuan Province, Hospital of Chengdu University of Traditional Chinese Medicine, Chengdu, China, ³West China Hospital of Sichuan University, Chengdu, China

Background: Type 2 diabetes mellitus (T2DM) is a clinical metabolic syndrome characterized by persistent hyperglycemia, which is caused by defective insulin secretion and decreased function in regulating glucose metabolism. Dachaihu Decoction (DCHD) is a traditional Chinese medicine formula that has been gradually used in T2DM treatment. A comprehensive analysis on the efficacy and safety of DCHD in T2DM treatment is necessary.

Objective: This meta-analysis aimed to systematically assess the clinical efficacy and safety of DCHD in the T2DM treatment and provide a reference for subsequent research and clinical practice.

Methods: Both Chinese and English databases were searched from their inceptions to November 2021. All retrieved studies were screened according to inclusion and exclusion criteria and randomized controlled trials about DCHD on T2DM were enrolled. The quality of the literature was assessed using the bias risk assessment tool in the Cochrane Handbook. Data extraction was performed on the selected studies. Review Manager 5.4 and Stata 16.0 were used for meta-analysis. Sources of heterogeneity were also explored by using meta-regression and subgroup analysis. Funnel plot and Egger's test were used to assess publication bias and the evidence quality was assessed by GRADE.

Results: 17 eligible studies, involving 1,525 patients, were included in this study. Compared with conventional treatment, combined treatment with DCHD was significantly better in improving HbA1c (MD = -0.90%, 95%CI: -1.20 to -0.60, p < 0.01), FBG (MD = -1.08 mmol/L, 95%CI: -1.28 to -0.87, p < 0.01), 2hPG (MD = -1.25 mmol/L, 95%CI: -1.42 to -1.09, p < 0.01), TC (MD = -0.50 mmol/L, 95%CI: -0.70 to -0.30, p < 0.01), TG (MD = -0.44 mmol/L, 95%CI: -0.61 to -0.26, p < 0.01), LDL-C (MD = -0.58 mmol/L, 95%CI: -0.85 to -0.31, p < 0.01), HOMA-IR (SMD = -2.04, 95%CI: -3.09 to -0.99, p < 0.01), HOMA-β (SMD = 2.48, 95%CI: 2.20 to 2.76, p < 0.01) and BMI (MD = -1.52 kg/m², 95%CI: -2.55 to -0.49, p < 0.01). When DCHD used alone, it had a similar efficacy to conventional treatment in HbA1c (MD = -0.04%, 95% CI: -0.17 to 0.09, p = 0.57) and FBG (MD = 0.13 mmol/L, 95%CI: -0.09 to 0.36, p = 0.24). It can also reduce 2hPG, even if not as effective as conventional treatment (MD = 0.54 mmol/L, 95%CI: 0.19 to 0.89, p < 0.01). Due to the small number of included studies, it is unclear whether DCHD used alone has an improving effect on lipid metabolism, BMI, HOMA-IR and HOMA- β . Analysis of adverse events showed DCHD was relatively safe. No obvious publication bias was detected by Funnel plot and Egger's test.

Conclusion: Based on this meta-analysis, we found that the combination with DCHD in the T2DM treatment has more advantages than conventional treatment alone, which can further regulate the glucose and lipid metabolism, reduce insulin resistance, improve islet function and lower BMI. DCHD alone also plays a certain role in regulating glucose. Meanwhile, DCHD is relatively safe. However, limited by the quality and quantity of included studies, the efficacy and safety of DCHD remain uncertain. More high-quality studies are still needed to provide more reliable evidence for the clinical application of DCHD.

Systematic Review Registration: https://www.crd.york.ac.uk/prospero/ display_record.php?ID=CRD42021296718, identifier CRD42021296718.

KEYWORDS

dachaihu decoction, type 2 diabetes mellitus, traditional Chinese medicine, systematic review, meta-analysis

Introduction

Type 2 diabetes mellitus (T2DM) is a clinical metabolic syndrome characterized by persistent hyperglycemia, which is caused by multiple factors such as heredity, environment and immunity. The intrinsic pathological mechanism on T2DM is an impairment in the metabolism and utilization of glucose and lipid due to insulin resistance or insufficient insulin secretion (Roden and Shulman, 2019; Reed et al., 2021). As the disease progresses, it causes damage to target organs such as the kidney, heart, blood vessels and nerves, resulting in dysfunction or failure of tissues and organs and even disability or death (Singh et al., 2022). With the improvement of living standards and changes in diet structure, the prevalence of T2DM has shown an obvious upward trend (Ma, 2018; Lin et al., 2020; Wang et al., 2021). In 2021, there were about 573 million adults with diabetes worldwide, representing approximately 10.5% of the world's population, of whom about 90% had T2DM (IDF, 2021). China has the largest number of diabetics, with around 140 million people suffering from diabetes. T2DM and its complications have become one of the leading causes of death worldwide (Stanaway et al., 2018; Glovaci et al., 2019; Martinez et al., 2020; WHO, 2020). It not only seriously affects the life quality of patients but also brings huge medical and economic burdens to individuals and society (WHO, 2016; Cannon et al., 2018; Zheng et al., 2018). How to effectively prevent the occurrence and development of T2DM and its complications

has always been an important public health issue (Li et al., 2020; Ling et al., 2020; Liu et al., 2020).

The current treatment for T2DM is based on a combination of blood glucose regulation, blood pressure control, lipid microcirculation lowering, improvement and lifestyle intervention (ADA, 2021). However, the existing hypoglycemic drugs have the potential to cause gastrointestinal adverse effects, reduce vitamin B12 concentrations and induce urinary tract infection (Baye et al., 2021). At the same time, even if blood glucose is well controlled, the existence of metabolic memory still makes it difficult to effectively prevent the emergence and progression of T2DM and its complications (Bianchi et al., 2013; Zhang and Wu, 2014; Galicia-Garcia et al., 2020). Therefore, there is an urgent need to find safer and more effective treatments.

In recent years, traditional Chinese medicine (TCM) has gradually shown its unique advantages in treating T2DM. Many studies have shown that TCM treatment can significantly improve the clinical symptoms and life quality of T2DM patients, reduce insulin resistance, decrease the occurrence of adverse effects and consolidate clinical efficacy (Ji et al., 2013; Tong et al., 2013; Lian et al., 2015; Zhang Y. et al., 2019; Pang et al., 2021). Dachaihu Decoction (DCHD) is one of the classical formulas in ancient China. It comes from the *Treatise on Cold Damage and Miscellaneous Diseases* (*Shang Han Za Bing Lun*) by *Zhongjing Zhang*, a famous doctor in the *Eastern Han Dynasty*. It is composed of eight herbs: Chinese Thorowax Root (*Chaihu*,

Bupleurum falcatum L.), Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi), Rhubarb (Dahuang, Rheum palmatum L.), Immature Orange Fruit (Zhishi, Citrus aurantium L.), Pinellia Tuber (Banxia, Pinellia ternata (Thunb.) Makino), White Paeony Root (Baishao, Paeonia lactiflora Pall.), Chinese Date (Dazao, Ziziphus jujuba Mill.) and Fresh Ginger (Shengjiang, Zingiber officinale Roscoe). It has the functions of soothing liver and relieving depression, clearing stomach and purging heat, and is mainly used to treat the syndrome of heat stagnation in liver and stomach. The main active components of DCHD measured by high performance liquid chromatography include paeoniflorin, naringin, hesperidin, neohesperidin, baicalin, baicalein and saikosaponin A (Li et al., 2006; Liu, 2014). Many studies have shown that DCHD has the effect of anti-inflammatory, regulating bile acid metabolism, balancing intestinal flora, protecting liver function and modulating blood lipids (Yoshie et al., 2004; Feng et al., 2019; Yang et al., 2019; Cui H. et al., 2020; Yang et al., 2021; Wang et al., 2022). It can be used in treating diseases such as cholecystitis, acute pancreatitis, bile reflux gastritis, fatty liver and hyperlipidemia (Xu and Sun, 2013; Zou and Wang, 2014; Qian et al., 2016; Dou et al., 2019; Han et al., 2020; Yang et al., 2021). In recent years, DCHD has also been gradually used to treat T2DM. The indications of DCHD for T2DM patients include thirst, bitter taste in mouth, fever, impatience and irascibility, hypochondriac pain, epigastric burning pain, increased eating with rapid hungering, red tongue with yellow coating and wiry or rapid pulse (CDS, 2021). Studies ranging from case reports, retrospective studies to randomized controlled trials (RCTs) suggest that DCHD or its modified may be able to relieve clinical symptoms related to T2DM, improve glucose and lipid metabolism and reduce insulin resistance (Zhang et al., 2015; Zhang, 2016; Li and Yan, 2019; Deng and Huang, 2020; Zou et al., 2021). However, the clinical efficacy of DCHD in T2DM remains uncertain due to limited sample size, inconsistent trial designs, different efficacy indicators and ambiguous methodological quality. Moreover, there is no clinical evidence summarizing the efficacy and safety of DCHD in the T2DM treatment. Therefore, this study comprehensively collected RCTs of DCHD alone or in combination with hypoglycemic drugs in the T2DM treatment and evaluated the clinical efficacy and safety of DCHD, in order to provide a reference for subsequent research and clinical practice.

Materials and methods

Study registration

This systematic review and meta-analysis was conducted and reported under the guidance of the Cochrane Handbook for Systematic Reviews of Interventions version 6.3 (updated 2022) and the Preferred Reporting Items for Systematic Review and Meta-Analysis (PRISMA) 2020 Statement (Moher et al., 2009; Page et al., 2021). The PRISMA 2020 checklist is provided in Supplementary Material S1. Before starting, this study was registered in the International Prospective Register of Systematic Reviews (PROSPERO) (Registration number: CRD42021296718). Data were derived from published clinical studies.

Database and search strategies

We conducted a comprehensive search of three English electronic databases, namely PubMed, EMBASE and the Cochrane Library, and three Chinese electronic databases, including the China National Knowledge Infrastructure (CNKI), Wan Fang Database, and China Science and Technology Journal Database (VIP), from their inceptions to November 2021. The clinical trials related to DCHD, modified DCHD, T2DM were searched using a combination of subject terms and text words. The search terms mainly included: "Dachaihu," "Dachaihu Decoction," "Dachaihu Tang," "Da Chaihu," "Da Chaihu Tang," "Da Chaihu Decoction," "Major Bupleurum Decoction," "Major Bupleurum Tang," "Daisaikoto," "Type 2 Diabetes Mellitus," "Diabetes Mellitus, Type 2," "Type 2 Diabetes," "Diabetes, Type 2" and "Noninsulin-Dependent Diabetes Mellitus". The detailed search strategies containing more search terms are provided in Supplementary Material S2. To understand ongoing studies, the ClinicalTrials.gov database and Chinese Clinical Trial Registry (CHiCTR) were also retrieved. Additionally, references of related reviews and meta-analyses were also screened to discover literature that may be missed in online searches. Only original articles in English and Chinese were included. All literature was selected according to inclusion criteria and exclusion criteria.

Inclusion criteria

Type of studies

RCTs were included without restriction on origins or countries, for they were considered to have high-quality evidence in assessing the effects of interventions. The publication language was limited to English or Chinese.

Type of participants

Adults (at least 18 years old) diagnosed with T2DM were included regardless of complications. Patients had no demographic restrictions such as age, race or gender.

Type of interventions

Interventions involving DCHD or modified DCHD were included, without limit to dosage form (decoction, capsule or granules), frequency or dosage. The experimental group can be DCHD alone or DCHD combined with conventional treatment. The control group can be placebo or conventional treatment. Conventional treatment refers to the classic treatment measures of western medicine, including diabetes health education, diet management, exercise intervention, blood glucose monitoring and hypoglycemic drugs. There is no restriction on the type or dosage form (oral preparation or injection) of hypoglycemic drugs. If the experimental group was combined with conventional treatment, it should be the same as the control group.

Type of comparisons

The following comparisons were made respectively in this study:

DCHD combined with conventional treatment vs. conventional treatment DCHD vs. conventional treatment DCHD vs. placebo

Type of outcome measures

To comprehensively evaluate the efficacy and safety of DCHD in patients with T2DM, the analysis was made from the perspectives of glucose metabolism, lipid metabolism, insulin resistance, pancreatic islet function, body mass index (BMI) and adverse events. RCTs evaluating any of the following outcomes were included:

1) Primary Outcomes

Glucose metabolism index: Glycated hemoglobin (HbA1c), Fasting blood glucose (FBG), 2-h postprandial glucose (2hPG)

2) Secondary Outcomes

Lipid metabolism index: Total cholesterol (TC), Triglyceride (TG), High-density lipoprotein cholesterol (HDL-C), Low-density lipoprotein cholesterol (LDL-C)

Insulin resistance index: Homeostasis model assessment of insulin resistance (HOMA-IR)

Body mass index

If a study reported multiple time points, the result with the longest time point was included in the analysis.

3) Safety Outcomes

Any adverse events that occurred during the study should be recorded, such as the incidence of hypoglycemia, the incidence of adverse events, the incidence of serious adverse events and the incidence of gastrointestinal adverse reactions.

Exclusion criteria

Type of studies

 Studies designed as non-RCTs, such as cohort studies, casecontrol studies, cross-sectional studies, case reports, animal studies and reviews.

- 2) For any replicate studies, the one with more complete data was selected and the other study was excluded.
- 3) Meeting abstracts were excluded if no relevant data were provided.
- Studies were excluded if the full text could not be obtained by searching online or contacting the authors.

Type of participants

- 1) Patients with acute metabolic disorders, such as diabetic ketoacidosis or infections.
- 2) Patients with severe hepatic and renal impairment, severe cardiovascular disease, pregnancy or lactation were excluded.

Type of interventions

The interventions used TCM treatments other than DCHD, such as acupuncture, moxibustion, massage, or acupoint injection.

Type of comparisons

The control group used measures other than conventional treatment.

Type of outcome measures

Studies with obvious data errors, incomplete data, questionable authenticity and lack of required indicators were excluded.

Study selection and data extraction

The search results were imported into EndNote X9 software in the form of bibliography to establish a database. Two researchers independently screened the literature according to the inclusion and exclusion criteria. Firstly, duplicate literature was deleted. Secondly, the literature that did not meet the criteria was preliminarily screened by reading the title and abstract. The literature that was uncertain in the preliminary screening was browsed by reading the full text. After reading the full text, literature that still did not meet the criteria was excluded. If there was any difference, it was determined after discussion or consultation with XF and HX. Two researchers independently extracted data from included studies according to the predesigned data extraction table. If some additional data are needed, we contacted the authors by email. The research data extracted mainly included: first author, publication year, study design, diagnostic criteria, sample size, gender, average age, course of disease, treatment duration, intervention measures, outcome indicators, comorbidity, adverse events and was crosschecked.

Risk of bias assessment

The quality of the literature was assessed using the bias risk assessment tool in the Cochrane Handbook. This part was

embedded in and implemented by Review Manager 5.4 software. This tool assessed seven important bias sources, including random sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessment, incomplete outcome data, selective reporting and other bias. Each included study was assessed for risk of bias from these seven aspects. By evaluating the completeness of research reporting and the correctness of methodological implementation, each aspect was assessed as "high risk," "low risk" or "unclear risk". Two researchers performed independently and examined each other. If there were different opinions on the evaluation results, the third researcher participated in the discussion and made the final decision.

Statistical analysis

Review Manager 5.4 and Stata 16.0 were used for metaanalysis. For binary variables, the relative risk (RR) and 95% confidence interval (CI) were used to express the effect size. For continuous variables, when the same outcome indicator used the same unit, the mean difference (MD) and 95% CI were used to represent the effect size; otherwise, the standardized mean difference (SMD) and 95% CI were used. Heterogeneity was evaluated according to χ^2 test and I² test. If p > 0.1, I² < 50%, it indicated that the heterogeneity between studies was small, and the fixed effect model was used to calculate the pooled effect size. If $p \leq 0.1$, $I^2 \geq 50\%$, it suggested significant statistical heterogeneity among the studies; therefore, the random effect model was used. Subgroup analysis and sensitivity analysis were performed to explore the source of heterogeneity and to judge the stability of the research results. The indicators of HbA1c, FBG and 2hPG included more than 10 studies, and we additionally performed a meta-regression on sample size, publication year and average age to explore the influence of these factors on heterogeneity. Meanwhile, we performed funnel chart and Egger's test to evaluate publication bias on HbA1c, FBG and 2hPG. p > 0.05 indicated no obvious publication bias, and p < 0.05 indicated possible publication bias. Finally, we used the Grading of Recommendations Assessment, Development and Evaluation (GRADE) method to assess the evidence quality.

Subgroup analysis

We prespecified factors that might influence treatment effect and performed subgroup analysis based on these prespecified assumptions to explore sources of heterogeneity. The following subgroup analyses were performed: Course of disease (\leq 5 years or >5 years); Treatment duration (\geq 3 months or <3 months); Average age (\leq 50 years old or >50 years old); Baseline BMI (\leq 24 or >24).

Results

Database search results

A total of 627 studies were retrieved by searching Chinese and English databases. 179 studies were excluded due to duplication. Of the remaining 448 studies, 404 studies were excluded by reading the titles and abstracts. The full texts of the remaining 44 studies were read, and 27 studies were excluded according to the inclusion and exclusion criteria. No additional studies were identified by screening references to relevant reviews and meta-analyses. Finally, 17 eligible studies were included in the quantitative analysis. The preliminary screening of the literature is provided in Supplementary Material S3. Literature excluded after reading the full text and reasons is listed in Supplementary Material S4. A detailed flowchart for screening eligible studies is shown in Figure 1.

Characteristics of included studies

A total of 17 RCTs were included in this study, all completed in China and published between 2002 and 2021 (Liu, 2002; Shen and Wu, 2007; Cui and Chen, 2015; Zhao et al., 2016; Li et al., 2018; Zhang, 2018; Zhang et al., 2018; Gao, 2019; Li, 2019; Zhang H. S. et al., 2019; Bao et al., 2020; Chang, 2020; Ji and Che, 2020; Wang, 2020; Duan, 2021; Zhang Q. J. et al., 2021; Zou et al., 2021). A total of 1,525 patients with T2DM were included in this study, including 771 patients in experimental group and 754 patients in control group. One study (Ji and Che, 2020) evaluated the efficacy of metformin vs. DCHD alone and metformin vs. combination treatment, so we divided this study into part 1 and part 2. In terms of diagnostic criteria, the World Health Organization (WHO) definition was adopted in 10 studies (Shen and Wu, 2007; Cui and Chen, 2015; Zhao et al., 2016; Li et al., 2018; Zhang H. S. et al., 2019; Bao et al., 2020; part1- Ji and Che, 2020; part2- Ji and Che, 2020; Duan, 2021; Zhang Q. J. et al., 2021), the American Diabetes Association (ADA) definition was adopted in 2 studies (Liu, 2002; Zhang, 2018), the Chinese guideline diagnostic criteria were adopted in five studies (Zhang et al., 2018; Gao, 2019; Li, 2019; Wang, 2020; Zou et al., 2021), and the diagnostic criteria were not reported in one study (Chang, 2020). four studies (Zhang, 2018; Zhang H. S. et al., 2019; Bao et al., 2020; Zou et al., 2021) were treated with the original DCHD, and 14 studies (Liu, 2002; Shen and Wu, 2007; Cui and Chen, 2015; Zhao et al., 2016; Li et al., 2018; Zhang et al., 2018; Gao, 2019; Li, 2019; Chang, 2020; part1- Ji and Che, 2020; part2- Ji and Che, 2020; Wang, 2020; Duan, 2021; Zhang Q. J. et al., 2021) were treated with modified DCHD. The composition of DCHD or modified DCHD is shown in Table 1, and none of these studies reported the quality control or chemical analysis of DCHD. There were 13 studies (Shen and Wu, 2007; Zhang, 2018; Zhang et al., 2018; Gao, 2019; Li, 2019;



Zhang H. S. et al., 2019; Bao et al., 2020; Chang, 2020; Ji and Che, 2020; Wang, 2020; Duan, 2021; Zhang Q. J. et al., 2021; Zou et al., 2021) in which the control group was treated with conventional treatment, and the treatment group was treated with DCHD or its modified on the basis of the control group. And there were four studies (Liu, 2002; Zhao et al., 2016; Li et al., 2018; Ji and Che, 2020) in which the control group was treated with conventional treatment, and the treatment group was treated with DCHD or its modified alone. Therefore, none of the above studies were sufficiently blinded. Only one study (Cui and Chen, 2015) claimed to have performed a placebo control. However, the blinding effect could not be achieved due to the difference in dosage form and frequency between placebo and DCHD. The basic characteristics of the included studies are shown in Table 2.

Risk of bias assessment

The risk of bias was assessed using the Cochrane Risk of Bias tool. Of the included studies, eight (Li et al., 2018; Zhang, 2018; Gao, 2019; Zhang H. S. et al., 2019; part 1- Ji and Che, 2020; part 2- Ji and Che, 2020; Zhang Q. J. et al., 2021; Zou et al., 2021) used a random number table and one (Duan, 2021) used a computergenerated random sequence, and these studies were marked as low risk. Other studies claimed to have performed randomization but did not report the specific methods used in the random sequence generation, and these studies were marked as unclear risk. One study (Liu, 2002) used random assignment cards to conceal the random sequence. However, the concealment tightness of this method was not adequately described, and

TABLE 1 Components of DCHD or its	modified used in the included studies.
-----------------------------------	--

References	Formula	Components
Liu (2002)	modified DCHD	Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.), Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi), Golden thread (Huanglian, Coptis chinensis Franch.), Pinellia Tuber [Banxia, Pinellia ternata (Thunb.) Makino], Immature Orange Fruit (Zhishi, Citrus aurantium L.), Rhubarb (Dahuang, Rheum palmatum L.) and Smoked Plum (Wumei, Prunus mume (Siebold) Siebold and Zucc.).
Shen and Wu (2007)	modified DCHD	No specific components are mentioned
Cui and Chen (2015)	modified DCHD	Chinese Thorowax Root (Chaihu, <i>Bupleurum falcatum</i> L.) 15 g, Baical Skullcap Root (Huangqin, <i>Scutellaria baicalensis</i> Georgi) 15 g, Golden thread (Huanglian, <i>Coptis chinensis</i> Franch.) 10 g, Pinellia Tuber [Banxia, <i>Pinellia ternata</i> (Thunb.) Makino) 10 g, White Paeony Root (Baishao, <i>Paeonia lactiflora</i> Pall.) 25 g, Rhubarb (Dahuang, <i>Rheum palmatum</i> L.) 10 g, Immature Orange Fruit (Zhishi, <i>Citrus aurantium</i> L.) 15 g, Thomson Kudzuvine root [Gegen, <i>Pueraria montana</i> var. thomsonii (Benth.) M.R.Almeida] 20 g, Gypsum (Shigao, Gypsum Fibrosum) 30 g, Earthworm (Dilong, Pheretima) 15 g, Peach Seed (Taoren, <i>Prunus persica</i> (L.) Batsch) 15 g, Herba Hedyotidis (Baihuasheshecao, <i>Scleromitrion diffusum</i> (Willd.) RJ.Wang) 15 g, Common Anemarrhena Rhizome (Zhimu, <i>Anemarrhena asphodeloides</i> Bunge) 15 g and Danshen Root (Danshen, <i>Salvia miltiorrhiza</i> Bunge) 30 g. If dry mouth and thirst were identified, Figwort Root (Xuanshen, <i>Scorphularia ningpoensis</i> Hemsl.) 15 g and Snakegourd Root (Tianhuafen, <i>Trichosanthes kirilowii</i> Maxim.) 30 g were added. If shortness of breath and lack of strength were identified, Milkvetch Root (Huangqi, <i>Astragalus mongholicus</i> Bunge) 20 g and Tangshen (Dangshen, Codonopsis pilosula (Franch.) Nanf.) 15 g were added. If dizziness, headache and hypertension were identified, Chrysanthemum Flower (Juhua, <i>Chrysanthemum indicum</i> L.) 15 g, Oyster Shell (Shengmuli, Ostreae Concha) 30 g and Abalone Shell (Shijueming, Haliotidis Concha) 15 g were added. If slimy yellow tongue fur and damp-heat were identified, Largehead Atractylodes Rh (Baizhu, Atractylodes macrocephala Koidz.) 15 g and Chinese Gentian (Longdancao, <i>Gentiana</i> scabra Bunge) 10 g were added
Zhao et al. (2016)	modified DCHD	Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.), Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi), Rhubarb (Dahuang, Rheum palmatum L.), Immature Orange Fruit (Zhishi, Citrus aurantium L.), Pinellia Tuber (Banxia, Pinellia ternata (Thunb.) Makino), White Paeony Root (Baishao, Paeonia lactiflora Pall.), Chinese Gentian (Longdancao, <i>Gentiana</i> scabra Bunge), Hawthorn Fruit (Shanzha, Crataegus pinnatifida Bunge), Virgate Wormwood Herb (Yinchen, Artemisia capillaris Thunb.), Oriental Water Plantain Rhizome (Zexie, Alisma plantago-aquatica L.) and Danshen Root (Danshen, Salvia miltiorrhiza Bunge)
Zhang et al. (2018)	modified DCHD	Chinese Thorowax Root (Chaihu, <i>Bupleurum falcatum</i> L.) 20 g, Baical Skullcap Root (Huangqin, <i>Scutellaria baicalensis</i> Georgi) 20 g, Rhubarb (Dahuang, <i>Rheum palmatum</i> L.) 15 g, Golden thread (Huanglian, <i>Coptis chinensis</i> Franch.) 15 g, Immature Orange Fruit (Zhishi, <i>Citrus aurantium</i> L.) 15 g, Snakegourd Fruit (Gualou, <i>Trichosanthes kirilowii</i> Maxim.) 15 g, Pinellia Tuber [Banxia, <i>Pinellia ternata</i> (Thunb.) Makino] 10 g, White Paeony Root (Baishao, <i>Paeonia lactiflora</i> Pall.) 10 g and Liquorice Root (Gancao, <i>Glycyrrhiza glabra</i> L.) 6 g
Zhang (2018)	DCHD	Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi) 9g, Pinellia Tuber (Banxia, Pinellia ternata (Thunb.) Makino) 8g, White Paeony Root (Baishao, Paeonia lactiflora Pall.) 9g, Immature Orange Fruit (Zhishi, Citrus aurantium L.) 9g, Fresh Ginger (Shengjiang, Zingiber officinale Roscoe) 16g, Rhubarb (Dahuang, Rheum palmatum L.) 7g, Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.) 12 g and Chinese Date (Dazao, Ziziphus jujuba Mill.) 5 pieces
Li et al. (2018)	modified DCHD	Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.) 10 g, Immature Orange Fruit (Zhishi, Citrus aurantium L.) 15 g, Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi) 15 g, Rhubarb (Dahuang, Rheum palmatum L.) 8 g, Golden thread (Huanglian, Coptis chinensis Franch.) 15 g, White Paeony Root (Baishao, Paeonia lactiflora Pall.) 15 g, Common Anemarrhena Rhizome (Zhimu, Anemarrhena asphodeloides Bunge) 15 g and Dried Ginger (Ganjiang, Zingiber officinale Roscoe) 3 g. If dry mouth, thirst or increased eating with rapid hungering were identified, Thomson Kudzuvine root (Gegen, Pueraria montana var. thomsonii (Benth.) M. R. Almeida) and Snakegourd Root (Tianhuafen, Trichosanthes kirilowii Maxim.) were added. If incomplete and irregular bowel movements, slimy yellow tongue fur or slippery pulse were identified, Snakegourd Fruit Kernels (Gualouzi, Trichosanthes kirilowii Maxim.) and Pinellia Tuber (Banxia, Pinellia ternata (Thunb.) Makino) were added
Zhang H. S. et al. (2019)	DCHD	Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.) 12 g. Rhubarb (Dahuang, Rheum palmatum L.) 8 g. Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi) 12 g. Pinellia Tuber [Banxia, Pinellia ternata (Thunb.) Makino] 9 g. Immature Orange Fruit (Zhishi, Citrus aurantium L.) 12 g. Red Peony Root (Chishao, Paeonia lactiflora Pall.) 12 g, Chinese Date (Dazao, Ziziphus jujuba Mill.) 10 g and Fresh Ginger (Shengjiang, Zingiber officinale Roscoe) 10 g
Gao (2019)	modified DCHD	Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.) 20 g, White Paeony Root (Baishao, Paeonia lactiflora Pall.) 10 g, Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi) 20 g, Pinellia Tuber (Banxia, Pinellia ternata (Thunb.) Makino) 10 g, Immature Orange Fruit (Zhishi, Citrus aurantium L.) 15 g, Rhubarb (Dahuang, Rheum palmatum L.) 10 g, Liquorice Root (Gancao, Glycyrrhiza glabra L.) 6 g and Turmeric Root Tuber (Yujin, Curcuma longa L.) 15 g
Li (2019)	modified DCHD	Chinese Thorowax Root (Chaihu, <i>Bupleurum falcatum</i> L.) 20 g, Baical Skullcap Root (Huangqin, <i>Scutellaria baicalensis</i> Georgi) 20 g, Golden thread (Huanglian, <i>Coptis chinensis</i> Franch.) 15 g, Immature Orange Fruit (Zhishi, Citrus aurantium L.) 15 g, Rhubarb (Dahuang, <i>Rheum palmatum</i> L.) 15 g, Snakegourd Fruit (Gualou, <i>Trichosanthes kirilowii</i> Maxim.) 15 g, Pinellia Tuber (Banxia, <i>Pinellia ternata</i> (Thunb.) Makino) 10 g, White Paeony Root (Baishao, <i>Paeonia lactiflora</i> Pall.) 10 g and Liquorice Root (Gancao, <i>Glycyrrhiza glabra</i> L.) 6 g
part1-Ji and Che (2020)	modified DCHD	Chinese Thorowax Root (Chaihu, <i>Bupleurum falcatum</i> L.) 10 g, White Paeony Root (Baishao, <i>Paeonia lactiflora</i> Pall.) 15 g, Baical Skullcap Root (Huangqin, <i>Scutellaria baicalensis</i> Georgi) 10 g, Pinellia Tuber (Banxia, <i>Pinellia ternata</i> (Thunb.) Makino) 10 g, Immature Orange Fruit (Zhishi, <i>Citrus aurantium</i> L.) 10 g, Rhubarb (Dahuang, <i>Rheum palmatum</i> L.) 6 g, sickle senna (Jueming, <i>Senna tora</i> (L.) Roxb.) 15 g, Rhizoma Atractylodis (Cangzhu, <i>Atractylodes lancea</i> (Thunb.) DC.) 15 g, Figwort Root (Xuanshen, <i>Scrophularia ningpoensis</i> Hemsl.) 20 g and Turmeric Root Tuber (Yujin, <i>Curcuma longa</i> L.) 15 g
		(Continued on following pa

(Continued on following page)

TABLE 1 (Continued) Components of DCHD or its modified used in the included studies.

References	Formula	Components
part2-Ji and Che (2020)	modified DCHD	Chinese Thorowax Root (Chaihu, <i>Bupleurum falcatum</i> L.) 10 g, White Paeony Root (Baishao, <i>Paeonia lactiflora</i> Pall.) 15 g, Baical Skullcap Root (Huangqin, <i>Scutellaria baicalensis</i> Georgi) 10 g, Pinellia Tuber (Banxia, <i>Pinellia ternata</i> (Thunb.) Makino) 10 g, Immature Orange Fruit (Zhishi, <i>Citrus aurantium</i> L.) 10 g, Rhubarb (Dahuang, <i>Rheum palmatum</i> L.) 6 g, sickle senna (Jueming, <i>Senna tora</i> (L.) Roxb.) 15 g, Rhizoma Atractylodis (Cangzhu, <i>Atractylodes lancea</i> (Thunb.) DC.) 15 g, Figwort Root (Xuanshen, <i>Scrophularia ningpoensis</i> Hemsl.) 20 g and Turmeric Root Tuber (Yujin, <i>Curcuma longa</i> L.) 15 g
Wang (2020)	modified DCHD	Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.) 20 g, Golden thread (Huanglian, Coptis chinensis Franch.) 15 g, Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi) 20 g, Immature Orange Fruit (Zhishi, Citrus aurantium L.) 15 g, Rhubarb (Dahuang, Rheum palmatum L.) 15 g, Snakegourd Fruit (Gualou, Trichosanthes kirilowii Maxim.) 15 g, White Paeony Root (Baishao, Paeonia lactiflora Pall.) 10 g, Pinellia Tuber [Banxia, Pinellia ternata (Thunb.) Makino] 10 g and Liquorice Root (Gancao, Glycyrrhiza glabra L.) 6 g
Bao et al. (2020)	DCHD	Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.) 15 g, White Paeony Root (Baishao, Paeonia lactiflora Pall.) 10 g, Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi) 10 g, Pinellia Tuber [Banxia, Pinellia ternata (Thunb.) Makino] 10g, Immature Orange Fruit (Zhishi, Citrus aurantium L.) 10 g, Fresh Ginger (Shengjiang, Zingiber officinale Roscoe) 10 g, Rhubarb (Dahuang, Rheum palmatum L.) 6 g and Chinese Date (Dazao, Ziziphus jujuba Mill.) 4 pieces
Chang (2020)	modified DCHD	Chinese Thorowax Root (Chaihu, <i>Bupleurum falcatum</i> L.) 15 g, White Paeony Root (Baishao, <i>Paeonia lactiflora</i> Pall.) 9 g, Baical Skullcap Root (Huangqin, <i>Scutellaria baicalensis</i> Georgi) 9 g, Pinellia Tuber (Banxia, <i>Pinellia ternata</i> (Thunb.) Makino) 9 g, Immature Orange Fruit (Zhishi, <i>Citrus aurantium</i> L.) 9 g, Rhubarb (Dahuang, <i>Rheum palmatum</i> L.) 9 g and Liquorice Root (Gancao, <i>Glycyrrhiza glabra</i> L.) 6 g
Zou et al. (2021)	DCHD	Chinese Thorowax Root (Chaihu, <i>Bupleurum falcatum</i> L.) 15 g, Rhubarb (Dahuang, <i>Rheum palmatum</i> L.) 6 g, Immature Orange Fruit (Zhishi, <i>Citrus aurantium</i> L.) 9 g, Baical Skullcap Root (Huangqin, <i>Scutellaria baicalensis</i> Georgi) 9 g, Pinellia Tuber (Banxia, <i>Pinellia ternata</i> (Thunb.) Makino) 9 g, White Paeony Root (Baishao, <i>Paeonia lactiflora</i> Pall.) 9 g, Chinese Date (Dazao, <i>Ziziphus jujuba</i> Mill.) 4 pieces and Fresh Ginger (Shengjiang, <i>Zingiber officinale</i> Roscoe) 15 g
Duan (2021)	modified DCHD	Chinese Thorowax Root (Chaihu, Bupleurum falcatum L.) 10 g, Rhubarb (Dahuang, Rheum palmatum L.) 10 g, Pinellia Tuber [Banxia, Pinellia ternata (Thunb.) Makino] 15 g, Baical Skullcap Root (Huangqin, Scutellaria baicalensis Georgi) 15 g, White Paeony Root (Baishao, Paeonia lactiflora Pall.) 15 g, Immature Orange Fruit (Zhishi, Citrus aurantium L.) 15 g, Chinese Date (Dazao, Ziziphus jujuba Mill.) 10 g, Fresh Ginger (Shengjiang, Zingiber officinale Roscoe) 6 g, Thomson Kudzuvine root (Gegen, Pueraria montana var. thomsonii (Benth.) M.R.Almeida) 20g, Danshen Root (Danshen, Salvia miltiorrhiza Bunge) 30g, Peach Seed (Taoren, Prunus persica (L.) Batsch) 15g, Common Anemarrhena Rhizome (Zhimu, Anemarrhena asphodeloides Bunge) 15 g and Earthworm (Dilong, Pheretima) 15 g. If dry mouth and thirst were identified, Figwort Root (Xuanshen, Scrophularia ningpoensis Hemsl.) 15 g and Snakegourd Root (Tianhuafen, Trichosanthes kirilowii Maxim.) 30 g were added. If shortness of breath and lack of strength were identified, Nilkvetch Root (Huangqi, Astragalus mongholicus Bunge) 20 g and Tangshen (Dangshen, Codonopsis pilosula (Franch.) Nannf.) 15 g were added. If hypertension was identified, Chrysanthemum Flower (Juhua, Chrysanthemum indicum L.) 15 g, Oyster Shell (Shengmuli, Ostreae Concha) 30 g and Abalone Shell (Shijueming, Haliotidis Concha) 15 g were added. If greasy coating and damp-heat were identified, Largehead Atractylodes Rh (Baizhu, Atractylodes macrocephala Koidz.) 15 g and Chinese Gentian (Longdancao, Gentiana scabra Bunge) 10 g were added
Zhang Q. J. et al., 2021	modified DCHD	Chinese Thorowax Root (Chaihu, <i>Bupleurum falcatum</i> L.) 10 g, Baical Skullcap Root (Huangqin, <i>Scutellaria baicalensis</i> Georgi) 10 g, Immature Orange Fruit (Zhishi, <i>Citrus aurantium</i> L.) 10 g, Pinellia Tuber (Banxia, <i>Pinellia ternata</i> (Thunb.) Makino) 9 g, White Paeony Root (Baishao, <i>Paeonia lactiflora</i> Pall.) 20 g, Rhizoma Atractylodis (Cangzhu, <i>Atractylodes lancea</i> (Thunb.) DC.) 15 g, Chinese Cork-tree (Huangbo, <i>Phellodendron amurense</i> Rupr.) 15 g, Coix Seed (Yiyiren, Coix lacryma-jobi L.) 30 g, Glabrous Greenbrier Rhizome (Tufuling, <i>Smilax glabra</i> Roxb.) 30 g, Silkworm Excrement (Cansha, Faeces Bombycis) 30 g, Appendiculate Cremastra Pseudobulb (Shancigu, <i>Cremastra appendiculata</i> (D.Don) Makino) 10 g and Cyathulae Radix (Chuanniuxi, <i>Cyathula officinalis</i> K.C.Kuan) 15 g

DCHD, dachaihu decoction.

none of the other studies mentioned the assignment concealment method, so all studies were marked as unclear risk. In most studies, the control group was treated with conventional treatment and the experimental group was treated with DCHD alone or in combination with conventional treatment. One study (Cui and Chen, 2015) was placebo-controlled. The frequency and dosage form of the medication varied between experimental group and control group, so participants and researchers were not blinded, and these studies were rated as high risk. All studies did not state whether outcome assessors were blinded and were therefore marked as unclear risk. One study (Liu, 2002) was rated as high risk in outcome data completeness due to an imbalance in the number and reasons on missing patients between groups. One study (Zhao et al., 2016) did not report the number of patients in the two groups at the end of the study, so data integrity could not be judged, and the risk was assessed as unclear. The other studies had no incomplete data. Concerning selective reporting, we were unable to make judgments about the risk because none of the included studies were registered and no study protocols were available. None of the studies had sufficient information to determine whether there was another significant bias risk and were therefore assessed as unclear risk. In general, the

TABLE 2 The characteristics of the included studies.

Disposite transmet fundomized analyzed (EC)1998 MDA1999 WHO1999 WHO1999 WHO2013 CDSISample size (randomized) analyzed (EC)10/07, 52/3540/44, 22/22120/120; 60/60120/120; 60/6086/86; 43/431Gender (MF) (EC)25/27; (16/913/12, 12/1236/24; 20/4032/28; 29/3124/15; 27/1633Gender (MF) (EC)55.81 ± 16.95 \$2.13 ± (1.291.5 (32-61)42.5 ± 11.4; 40.3 ± 12.549.84 ± 428; (53.2 ± 1.45; 50.1 ± 4.833Gourse of disea (pears) (EC)7.9 ± 24; 6.52 ± 7.651.6 months (4 days 3 months)62.2 ± 1.2; 5.8 ± 1.44.25 ± 0.46; 4.25 ± 0.487.2 ± 1.5; 7.0 ± 1.33Course of disea (pears) (EC)6 weels1.year8 weels1.year8 weels3 months3Co-intervention interventionMaintain the original restruent + Dietary intervention1.year8 weels1.yeeks3 months3Tratatment group interventionModified DCHD, 0 to ul/per, bul/tid tid weeksModified DCHD, 1 dose/per day, bid day, bidModified DCHD, 1 dose/per dayModified DCHD, 1 dose/per day, bid day, bidModified DCHD, 1 dose/per day, bid day, bidModified DCHD, 1 dose/per dayModified	First author (year)	Liu (2002)	Shen and Wu (2007)	Cui and Chen (2015)	Zhao et al. (2016)	Zhang et al. (2018)	Zhang et al. (2018)
Normal Supple size (randomized) manycoli (EC)110/87; 52/3549/44; 22/22120/120; 60/60 $120/120; 60/60$ $120/120; 60/60$ $86/86; 43/3$ 1Gender (MP) (EC)5/27, 16/1913/12; 12/1236/24; 20/4032/28; 29/31 $28/15; 27/16$ 3Gender (MP) (EC)55.81 ± 10.54; 52.13 ±15 (32-61)42.5 ± 11.4; 40.3 ± 12.5 83.26 ± 4.15 $50.5 \pm 51; 50.1 \pm 4.8$ 3Corner of Idessee (EVC) $6^{0} \oplus 84.24; 65.2 \pm 7.66$ 1.6 months (4 days-3 months) $6.2 \pm 1.2; 58.4 \pm 1.4$ $4.52 \pm 0.46; 4.25 \pm 0.46$ $72 \pm 1.5; 7.0 \pm 1.3$ 3Corner of Idessee 	Study design	RCT	RCT	RCT	RCT	RCT	RCT
framediantial-disputed periods space periods	Diagnostic criteria	1998 ADA	1999 WHO	1999 WHO	1999 WHO	2013 CDS	ADA
Average age (Yawa) Tig (2) $53.81 \pm 10.54; 52.13 \pm \\ 11.29$ $11.5(32-61)$ $25.5 \pm 1.14; 40.3 \pm 12.5$ $49.84 \pm 4.28; \\ 53.26 \pm 4.15$ $50.5 \pm 5.1; 50.1 \pm 4.8$ 50.5 $50.5 \pm 5.1; 50.1 \pm 4.8$ 50.5 $50.5 \pm 5.1; 50.1 \pm 4.8$ $50.5 \pm $	randomized/	110/87; 52/35	49/44; 22/22	120/120; 60/60	120/120; 60/60	86/86; 43/43	120/120; 60/60
(E/C) 11.29 53.26 ± 4.15 Course of disease (vars) (E/C) 7.69 ± 8.24; 6.52 ± 7.65 1.6 months (4 days-3 months) 6.2 ± 1.2; 5.8 ± 1.4 4.52 ± 0.46; 4.25 ± 0.48 7.2 ± 1.5; 7.0 ± 1.3 3 Treatment duration 6 weeks 1 year 8 weeks 12 weeks 3 months 3 Co-intervention Maintain the original treatment - Dietary intervention Diabetes health education + Dietary discrice intervention Maintain the original treatment / treatment + Dietary intervention NR Modified DCHD, 1 dose/per day,bid + CG I Treatment group interventions Modified DCHD, 6 g, tid Modified DCHD, 100 ml/per, bid/tid Modified DCHD, 1 dose/per day,bid + CG I Itagluide, 0.6 mg, qd, subcutaneous injection Defore mask + Novoline N subcutaneous injection Defore breakfast and dinner Modified DCHD, 1 dose/per day,bid + CG I Control group latelet with education + Distright Modified DCHD, 1 dose/per day,bid + CG I Itagluide, 0.6 mg, qd, subcutaneous injection Defore breakfast and dinner Modified DCHD, 1 dose/per day,bid + CG Itagluide, 0.6 mg, qd, subcutaneous injection Defore breakfast and dinner Modified DCHD, 1 dose/per day,bid + CG Itagluide, 0.6 mg, qd, subcutaneous injection Defore breakfast and dinner Modified DCHD, 1 dose/per day,bid + CG Itagluide, 0.6 mg, qd, subcutaneous injection Defore breakfast and dinner Modified DCH	Gender (M/F) (E/C)	25/27; 16/19	13/12; 12/12	36/24; 20/40	32/28; 29/31	28/15; 27/16	30/30; 35/25
typers) (E/C) Norweich 1 yer 8 weeks 12 weeks 3 months 3 months 3 Co-intervention Maintain the original treatment + Didary intervention Diabetes health education + Diet and exercise intervention + CG Diet and exercise intervention + Diet and exercise intervention + CG Maintain the original treatment + Dietary intervention + CG Maintain the original treatment + Dietary intervention + CG Modified DCHD, 1 dose/per day. NR Testment quoup Modified DCHD, 1 dose/per day. Interventions Modified DCHD, 1 dose/per day. Interventions + Novoline N subcutancous injection before meakfast and dinner Modified DCHD, 1 dose/per day. Interventions + Novoline N subcutancous injection before meakfast and dinner Modified DCHD, 1 dose/per day. Interventions Interventions + Novoline N subcutancous injection before meakfast and dinner Modified DCHD, 1 dose/per day. Interventions + Novoline N subcutancous injection before bedtime, Affer 2 weeks: Novoline N subcutancous injection to SG & O@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@	0 0 7 7		41.5 (32–61)	42.5 ± 11.4; 40.3 ± 12.5		$50.5 \pm 5.1; 50.1 \pm 4.8$	51.1 ± 6.2; 46.1 ± 7.2
Co-intervention Maintain the original intervention Diabetes health education + Diet and exercise intervention intervention Maintain the original intervention <		7.69 ± 8.24; 6.52 ± 7.65	1.6 months (4 days-3 months)	$6.2 \pm 1.2; 5.8 \pm 1.4$	4.52 ± 0.46; 4.25 ± 0.48	$7.2 \pm 1.5; 7.0 \pm 1.3$	3.1 ± 1.4; 3.2 ± 2.1
treatment + Diedary intervention exercise intervention + Medformin, 0.5g, tid Treatment + Medformin, 0.5g, tid treatment + Medformin, 0.5g, tid Modified DCHD, 1 dose/per day,bid + CG I Treatment group interventions Medified DCHD, 6g, tid Modified DCHD, 100 ml/per, bid/tid + CG Modified DCHD, 1 dose/per day Modified DCHD, 1 dose/per day,bid + CG I Control group interventions Metformin, 0.5 g, bid Intensive insulin therapy for 2 weeks: Novoline R subcutaneous injection before bettimes, After 2 weeks: Novoline R subcutaneous injection Before bettimes, After 2 weeks: Novoline R subcutaneous injection 20-30 min before bettimes, After 2 weeks: Novoline R subcutaneous injection 20-30 min before bettimes, After 2 weeks: Novoline R subcutaneous injection 20-30 min before betaffast and dimer Ø@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@	Freatment duration	6 weeks	1 year	8 weeks	12 weeks	3 months	3 months
interventions + CG day, bid 1 dose/per day 1 dose/per day Control group interventions Metformin, 0.5 g, bid Intensive insulin therapy for 2 weeks: Novoline 8 subcutaneous injection before meals + Novoline N subcutaneous injection 20-30 min before bedtime; After 2 weeks: Novoline 30R or 50R subcutaneous injection 20-30 min before breakfast and dinner Placebo, 6 pills, tid Metformin, 0.5 g, tid Liraglutide, 0.6 mg, dl, subcutaneous injection; 1F PBG>7.8 and 2PBG>11.8 mmol/ L, liraglutide was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild was increased to 1.2 mg d. The dose is evaluated every 2 weeks, but the daily dose should not exceed 1.8 mg Flandshild Was increased 1	Co-intervention	treatment + Dietary		Diet and exercise intervention	0	NR	Diabetes health educatior + Diet and exercise intervention
interventions Novoline R subcutaneous injection before meals + Novoline N subcutaneous injection before befakfast and dinner injection; If FBG>7.8 and 2hPG>11.8 mmol/ L, Iraglutide was increased to 1.2 mg qd. The dose is evaluated every 2 weeks; but the daily dose should not exceed 1.8 mg injection before befakfast and dinner Outcome index ©@@@@@@@@@ ©@@@ ©@@@ ©@@@ ©@@@ ©@@@ @@@@@@@@ @@ @@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@@	0 1	Modified DCHD, 6 g, tid	1			Modified DCHD, 1 dose/per day,bid + CG	DCHD, 400 ml, bid + CC
Baseline difference NSD NSD NSD NSD NSD NSD NSD NSD Comorbidity NR	e x	Metformin, 0.5 g, bid	Novoline R subcutaneous injection before meals + Novoline N subcutaneous injection before bedtime; After 2 weeks: Novoline 30R or 50R subcutaneous injection 20–30 min before breakfast and	Placebo, 6 pills, tid	Metformin, 0.5 g, tid	injection; If FBG>7.8 and 2hPG>11.8 mmol/ L, liraglutide was increased to 1.2 mg qd. The dose is evaluated every 2 weeks, but the daily	Hypoglycemic agents
Comorbidity NR NR NR Mperlipemia NR NR M Adverse events NR NR NR NR NR NR NR NR Country CHINA	Dutcome index	2345678	123	123	12356789	123490	234
Adverse events NR NR NR NR No significant change in liver function NR No significant change in liver function NR No significant change in liver function NR Ne	Baseline difference	NSD	NSD	NSD	NSD	NSD	NSD
Country CHINA <	Comorbidity	NR	NR	NR	Hyperlipemia	NR	NR
FundingNRNRNRNRNRFirst author (year)Li et al. (2018)Zhang H. S. et al. (2019)Gao (2019)Li (2019)part1- Ji and Che (2020)Study design Diagnostic criteriaRCTRCTRCTRCTRCTRCT1999 WHO2013 WHO2013 CDSCDS1999 WHO	Adverse events	NR	NR	NR		NR	NR
First author (year)Li et al. (2018)Zhang H. S. et al. (2019)Gao (2019)Li (2019)part1- Ji and Che (2020)Study designRCTRCTRCTRCTRCTRCTDiagnostic criteria1999 WHO2013 WHO2013 CDSCDS1999 WHO	Country	CHINA	CHINA	CHINA	CHINA	CHINA	CHINA
Kyear) RCT RCT RCT RCT RCT Study design RCT 2013 WHO 2013 CDS CDS 1999 WHO	Funding	NR	NR	NR	NR	NR	NR
Diagnostic criteria 1999 WHO 2013 WHO 2013 CDS CDS 1999 WHO		Li et al. (2018)	Zhang H. S. et al. (2019)	Gao (2019)	Li (2019)	part1- Ji and Che (2020)	part2- Ji and Che (2020)
	design	RCT	RCT	RCT	RCT	RCT	RCT
(Conti	Diagnostic criteria	1999 WHO	2013 WHO	2013 CDS	CDS	1999 WHO	1999 WHO
						(C	ontinued on following page)

TABLE 2 (Continued) The characteristics of the included studies.

First author I (year)	Li et al. (2018)	Zhang H. S. et al. (2019)	Gao (2019)		Li (2019)	part1- Ji and	Che (2020)	part2- Ji and Che (2020)
Sample size 1 (randomized/ analyzed) (E/C)	02/102; 51/51	90/90; 45/45	76/76; 38/38		88/88; 44/44	40/40; 20/20		40/40; 20/20
Gender (M/F) 3 (E/C)	31/20; 33/18	25/20; 30/15	21/17; 20/18		24/20; 23/21	11/9; 11/9		10/10; 11/9
Average age 5 (years) (E/C)	52.8; 43.6	53.16 ± 9.37; 52.84 ± 10.26	50.97 ± 6.17; 51.6	4 ± 6.39	51.6 ± 2.5; 52.1 ± 2.3	55.33 ± 9.78; 56.5	50 ± 10.25	55.25 ± 9.93; 56.50 ± 10.25
Course of disease 0 (years) (E/C)).3–5; 0.5–5	6.46 ± 3.73; 5.98 ± 4.09 (months)	NR		7.3 ± 2.3; 7.2 ± 2.1	5.54 ± 4.31; 5.35	± 4.18	$5.62 \pm 4.37;$ 5.35 ± 4.18
Treatment 1 duration	2 weeks	2 weeks	3 months		3 months	1 month		2 months
	Diet and exercise intervention + Blood glucose monitoring	Diabetes health education + Diet a exercise intervention + Blood glucose monitoring	nd NR		NR	NR		NR
	Modified DCHD, 1 dose/per lay, bid	DCHD, 1 dose/per day, 100 ml/p bid + CG	er, Modified DCHD, + CG	1 dose/per day,bid	Modified DCHD, 1 dose/per day, bid + CG	Modified DCHD, 200 ml/per, bid	, 1 dose/per day,	Modified DCHD, 1 dose/per day, 200 ml/per, bid + CG
Control group M interventions	Metformin, 0.25g, tid	Gansulin R, 0.2 U/(kg·d), tid, subcutaneous injection 30 min before meals + Insulin glargine, 0.2U/(kg·d), qd, subcutaneous injection at 22PM + Metformin, 0.5g, bid	increased to 1.2 n	ction. The dose is	Liraglutide, 0.6 mg, qd, subcutaneous injection. The maximum daily dose is no more than 1.8 mg	Metformin, 0.5 g	, tid	Metformin, 0.5 g, tic
Outcome index (1234	23	12391		90	123		023
Baseline difference N	NSD	NSD	NSD		NSD	NSD		NSD
Comorbidity N	NR	NR	NR		NR	NR		NR
Adverse events N	NR	Incidence of hypoglycemia: 3 case (6.64%) in experimental group; 11 cases (24.44%) in control grou			NR	NR		NR
Country C	CHINA	CHINA	CHINA		CHINA	CHINA		CHINA
o S	Hebei Provincial Administration of TCM Li Liping Inheritance Studio Project (No. Hebei TCM 2018-9]	NR	NR		NR	Experts Inheritan	Chinese Medicine ace Studio ject [Department of	
First author V (year)	Wang (2020) E	Bao et al. (2020)	Chang (2020)	Zou e	t al. (2021)	Duan (2021)	Zhang Q. J.	et al. (2021)
Study design R	RCT R	CT R	CT	RCT		RCT	RCT	
Diagnostic criteria 2	2017 CDS 1	999 WHO N	IR	2017 CI	DS .	WHO	1999 WHO	

10.3389/fphar.2022.918681

(Continued on following page)

TABLE 2 (Continued) The characteristics of the included studies.

First author (year)	Wang (2020)	Bao et al. (2020)	Chang (2020)	Zou et al. (2021)	Duan (2021)	Zhang Q. J. et al. (2021)
Sample size (randomized/ analyzed) (E/C)	104/104; 52/52	60/60; 30/30	100/100; 50/50	60/60; 30/30	92/92; 46/46	100/96; 48/48
Gender (M/F) (E/C)	32/20; 31/21	19/11; 18/12	22/28; 24/26	18/12; 19/11	27/19; 29/17	41/9; 39/11
Average age (years) (E/C)	52.12 ± 6.47; 52.44 ± 6.58	44.23 ± 12.80; 40.18 ± 11.79	49.87 ± 7.21; 50.21 ± 8.90	55.51 ± 1.31; 55.47 ± 1.28	62.7 ± 1.4; 63.2 ± 1.8	47.2 ± 9.95; 46.86 ± 10.10
Course of disease (years) (E/C)	5.48 ± 1.77; 5.51 ± 1.82	10.00 ± 14.08; 8.90 ± 14.29 (months)	6.10 ± 1.38; 6.23 ± 1.41	NR	6.2 ± 0.4 ; 6.3 ± 0.7	7.36 ± 4.98; 7.46 ± 4.86
Treatment duration	12 weeks	12 weeks	2 months	14 days	NR	12 weeks
Co-intervention	Diabetes health education + Diet and exercise intervention	Diet and exercise intervention	Conventional treatment	NR	Diet and exercise intervention + Blood glucose monitoring	Diet intervention
Treatment group interventions	Modified DCHD, 1 dose/per day, 300 ml, bid + CG	DCHD, 1 dose/per day, 400 ml, bid + CG	Modified DCHD, 1 dose/per day, bid + CG	DCHD, 1 dose/per day, 150 ml/per, bid + CG	Modified DCHD, 1 dose/per day,tid + CG	Modified DCHD, 1 dose/per day, bid + CG
Control group interventions	Metformin, 0.25 g, tid	Exenatide, 5 µg, bid, subcutaneous injection; Exenatide, 10 µg, bid, subcutaneous injection (Four weeks later)	Insulin glargine, 0.2 IU/(kg·d), qn,subcutaneous injection. Adjust dose according to blood sugar level	Liraglutide, 0.6 mg, qd, subcutaneous injection + Insulin glargine, 0.2 IU/(kg-d), qd, subcutaneous injection	Conventional hypoglycemic drugs	Conventional hypoglycemic drugs
Outcome index	123490	14569	123	1234568	1234	12345689
Baseline difference	NSD	NSD	NSD	NSD	NSD	NSD
Comorbidity	NR	Non-alcoholic fatty liver disease	NR	NR	NR	Hyperuricemia
Adverse events	NR	NR	NR	NR	NR	Treatment group: 4 Loose stool; Control group: 5 abdominal distension; 2 nausea. No special treatment was given
Country	CHINA	CHINA	CHINA	CHINA	CHINA	CHINA
Funding	Key Science and Technology Project of Henan Province, No.092102310010	NR	NR	NR	NR	National Natural Science Foundation of China, No.82004345; Scientific Research Foundation of Guang 'anmen Hospital, No. Y2018-07

Abbreviations: RCT, randomized controlled trial; ADA, american diabetes association; WHO, world health organization; CDS, chinese diabetes society; NR, not reported; DCHD, dachaihu decoction; CG, control group interventions; NSD, no significant difference; TCM, traditional Chinese medicine; Outcome index: ①: HbA1c; ②FBG; ③2hPG; ④BMI; ③TC; ③TG; ③HDL-C; ③HDL-C; ③HOMA-IR; ③HOMA-β.



methodological quality of the included literature was not high. The risk of bias assessment results for included studies is shown in Figure 2.

Primary outcomes

HbA1c

DCHD combined with conventional treatment vs. conventional treatment

Ten studies including 758 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional treatment alone on HbA1c (Shen and Wu, 2007; Zhang et al., 2018; Gao, 2019; Bao et al., 2020; Chang, 2020; Ji and Che, 2020;

Wang, 2020; Duan, 2021; Zhang Q. J. et al., 2021; Zou et al., 2021). According to the heterogeneity test (p < 0.01, $I^2 = 83\%$), a random effect model was selected for statistical analysis. The pooled result showed that compared with conventional treatment, the combination with DCHD could reduce the HbA1c level, and the difference was statistically significant (MD = -0.90%, 95%CI: -1.20 to -0.60, p < 0.01) (Figure 3A). We performed meta-regression on average age, sample size and publication year to identify possible sources of heterogeneity. According to the meta-regression of age, the scatters distribution showed a linear regularity, and the Tau² decreased from 0.18 to 0.10, which suggested that age may be the source of heterogeneity and could explain 53.55% of the variation between studies (p = 0.023, Adj R² = 53.55%) (Figure 4A; Supplementary Material S5.1). We further analyzed from the regression diagram that the

4	Expe	rimen	tal	C	ontrol			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% Cl
Bao et al. 2020	7.35	0.8	30	7.8	0.93	30	10.4%	-0.45 [-0.89, -0.01]	
Chang 2020	5.56	2.41	50	6.5	2.42	50	5.8%	-0.94 [-1.89, 0.01]	
Duan 2021	5.13	0.62	46	6.74	1.06	46	11.2%	-1.61 [-1.96, -1.26]	
Gao 2019	6.74	0.72	38	7.31	0.63	38	11.7%	-0.57 [-0.87, -0.27]	
part-2 Ji and Che 2020	6.03	0.68	20	7.98	0.72	20	10.4%	-1.95 [-2.38, -1.52]	
Shen and Wu 2007	6.1	1.8	22	6.6	2.1	22	4.5%	-0.50 [-1.66, 0.66]	
Wang 2020	6.25	0.85	52	6.97	0.97	52	11.2%	-0.72 [-1.07, -0.37]	
Zhang et al. 2018	7.08	0.81	43	7.71	0.89	43	11.2%	-0.63 [-0.99, -0.27]	
Zhang et al. 2021	7.1	0.64	48	7.83	0.68	48	12.0%	-0.73 [-0.99, -0.47]	
Zou et al. 2021	7.22	0.23	30	7.98	0.85	30	11.6%	-0.76 [-1.08, -0.44]	
Total (95% CI)			379			379	100.0%	-0.90 [-1.20, -0.60]	◆
Total (95% CI) Heterogeneity: Tau ² = 0	0.18; Chi² :	= 52.11		9 (P < 0.	.00001			-0.90 [-1.20, -0.60]	
			l, df = 9	9 (P < 0.	.00001			-0.90 [-1.20, -0.60]	-2 -1 0 1 2 Favours [experimental] Favours [control]
Heterogeneity: Tau ² = 0 Test for overall effect: 2			l, df = 9	9 (P < 0.	.00001			-0.90 [-1.20, -0.60]	-2 -1 0 1 2 Favours [experimental] Favours [control]
Heterogeneity: Tau ² = 0 Test for overall effect: 2	Z = 5.82 (P		l, df = 9 0001)	,	.00001 ontrol); I² = 8		-0.90 [-1.20, -0.60]	
Heterogeneity: Tau ² = 0 Test for overall effect: 2	Z = 5.82 (P	erimen	l, df = 9 0001) tal	,	ontrol); I ² = 8		Mean Difference	Favours [experimental] Favours [control]
Heterogeneity: Tau ² = 0 Test for overall effect: 2	Z = 5.82 (P Expe	rimen SD	l, df = 9 0001) tal	c	ontrol SD); I ² = 8	33%	Mean Difference IV. Fixed, 95% Cl	Favours [experimental] Favours [control] Mean Difference
Heterogeneity: Tau ² = 0 Test for overall effect: 2 Study or Subgroup Li et al. 2018	Z = 5.82 (P Expe <u>Mean</u> 6.54	rimen <u>SD</u> 1.26	l, df = 9 0001) tal <u>Total</u>	C Mean	ontrol SD 1.01); I² = 8 Total	33% Weight	Mean Difference <u>IV. Fixed. 95% Cl</u> 0.16 [-0.28, 0.60]	Favours [experimental] Favours [control] Mean Difference
Heterogeneity: Tau ² = 0 Test for overall effect: 2 Study or Subgroup	Z = 5.82 (P Expe <u>Mean</u> 6.54	rimen 5D 1.26 1.07	l, df = 9 0001) tal <u>Total</u> 51	C <u>Mean</u> 6.38	ontrol SD 1.01 0.72); I ² = 8 <u>Total</u> 51	Weight 8.6% 5.3%	Mean Difference <u>IV. Fixed. 95% Cl</u> 0.16 [-0.28, 0.60] 0.17 [-0.40, 0.74]	Favours [experimental] Favours [control] Mean Difference
Heterogeneity: Tau ² = 0 Test for overall effect: 2 Study or Subgroup Li et al. 2018 part-1Ji and Che 2020	Z = 5.82 (P Expe <u>Mean</u> 6.54 8.15	rimen 5D 1.26 1.07	l, df = 9 0001) tal <u>Total</u> 51 20	C <u>Mean</u> 6.38 7.98	ontrol SD 1.01 0.72); I ² = 8 <u>Total</u> 51 20 60	Weight 8.6% 5.3% 86.2%	Mean Difference <u>IV. Fixed. 95% Cl</u> 0.16 [-0.28, 0.60] 0.17 [-0.40, 0.74]	Favours [experimental] Favours [control] Mean Difference
Heterogeneity: Tau ² = C Test for overall effect: Z Study or Subgroup Li et al. 2018 part-1Ji and Che 2020 Zhao et al. 2016	Z = 5.82 (P Expe <u>Mean</u> 6.54 8.15 6.97	erimen SD 1.26 1.07 0.41	l, df = 9 0001) tal <u>Total</u> 51 20 60 131	Mean 6.38 7.98 7.04	ontrol SD 1.01 0.72); I ² = 8 <u>Total</u> 51 20 60	Weight 8.6% 5.3% 86.2%	Mean Difference <u>IV. Fixed. 95% Cl</u> 0.16 [-0.28, 0.60] 0.17 [-0.40, 0.74] -0.07 [-0.21, 0.07]	Favours [experimental] Favours [control] Mean Difference
Heterogeneity: Tau ² = 0 Test for overall effect: 2 B <u>Study or Subgroup</u> Li et al. 2018 part-1Ji and Che 2020	Z = 5.82 (P Expe <u>Mean</u> 6.54 8.15	rimen 5D 1.26 1.07	l, df = 9 0001) tal <u>Total</u> 51 20	C <u>Mean</u> 6.38 7.98	ontrol SD 1.01 0.72); I ² = 8 <u>Total</u> 51 20	Weight 8.6% 5.3%	Mean Difference <u>IV. Fixed. 95% Cl</u> 0.16 [-0.28, 0.60] 0.17 [-0.40, 0.74]	Favours [experimental] Favours [control] Mean Difference

decrease in HbA1c gradually increased with age. In addition, sample size (p = 0.548, Adj R² = -5.76%) and publication year (p =0.442, Adj R^2 = -0.33%) showed no significant difference on HbA1c (Figure 4B, C; Supplementary Material S5.2, S5.3). Study characteristics such as course of disease, treatment duration and baseline BMI may also contribute to heterogeneity. However, some studies incompletely reported these baseline characteristics, so we could not perform meta-regression on these factors and subgroup analysis was finally used. Subgroup analysis showed that the heterogeneity within each subgroup was not entirely reduced, so these factors cannot be considered as the source of heterogeneity at present (Supplementary Material S6). Sensitivity analysis was also performed, deleting one study at a time, and other studies were analyzed to estimate whether the results might have been significantly affected by a single study. Sensitivity analysis showed that the pooled effect sizes were similar and the result was robust (Figure 5A; Supplementary Material S7.1).

DCHD vs. conventional treatment

Three studies including 262 patients reported the efficacy of DCHD alone on HBA1c compared with conventional treatment (Zhao et al., 2016; Li et al., 2018; Ji and Che, 2020). According to the heterogeneity test (p = 0.48, $I^2 = 0\%$), a fixed effect model was selected for statistical analysis. The result showed that DCHD alone had the same reduction in HBA1c as conventional treatment, and the difference was not statistically significant (MD = -0.04%, 95%CI: -0.17 to 0.09, p = 0.57) (Figure 3B).

Subgroup analysis showed that DCHD can reduce HBA1c in diabetic patients with different course of disease, treatment duration, age and baseline BMI (Supplementary Material S8). After excluding Zhao et al., 2016, the mean difference changed from MD = -0.04%, 95%CI: 0.17 to 0.09 to MD = 0.16%, 95%CI: -0.18 to 0.51. It showed that this study had a great influence on the result due to its relatively large sample size and small standard deviation. However, there was still no heterogeneity, and the result was relatively stable (Figure 5B, Supplementary Material S7.2 and S7.3).

DCHD vs. placebo

One study including 120 patients reported that DCHD could reduce the HbA1c level compared with placebo after 8 weeks treatment (MD = -0.35%, 95%CI: -0.68 to -0.02, p = 0.04) (Cui and Chen, 2015).

FBG

DCHD combined with conventional treatment vs. conventional treatment

Eleven studies including 908 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional treatment alone on FBG (Shen and Wu, 2007; Zhang, 2018; Zhang et al., 2018; Gao, 2019; Zhang H. S. et al., 2019; Chang,



2020; Ji and Che, 2020; Wang, 2020; Duan, 2021; Zhang Q. J. et al., 2021; Zou et al., 2021). According to the heterogeneity test (p = 0.01, $I^2 = 56\%$), a random effect model was selected for statistical analysis. The result showed that compared with conventional treatment, combined with DCHD could reduce the FBG level, and the difference was statistically significant (MD = -1.08 mmol/L, 95% CI: -1.28 to -0.87, p < 0.01)(Figure 6A). Meta-regression was performed to identify possible sources of heterogeneity. Meta-regression according to average age (p = 0.532, Adj R² = -12.30%), sample size $(p = 0.959, \text{Adj } \text{R}^2 = -10.99\%)$ and publication year $(p = 0.959, \text{Adj } \text{R}^2)$ 0.799, Adj $R^2 = -8.55\%$) showed no significant difference on FBG (Figure 7, Supplementary Material S5.4-S5.6). Due to limited information, we performed subgroup analyses on course of disease, treatment duration and baseline BMI to explore possible heterogeneity sources. Subgroup analyses according to course of disease and treatment duration showed that the heterogeneity within these two subgroups was not reduced, so they cannot be considered the source of heterogeneity (Supplementary Material S9A, S9B). The subgroup analysis according to different baseline BMI (≤ 24 or > 24) showed significant subgroup difference (p < 10.01) and the heterogeneity within each group was reduced $(I^2 = 0\%$ and 0%, respectively), which means the baseline BMI may be one of the reasons for the heterogeneity (Supplementary Material S9C). The result showed that compared with conventional treatment, the combination treatment had no statistical difference in reducing FBG in diabetic patients with BMI \leq 24, but there was a significant statistical difference in patients with BMI > 24. Sensitivity analysis showed that the pooled statistics were similar and the result was robust (Figure 5C, Supplementary Material S7.4).

DCHD vs. conventional treatment

Four studies including 349 patients reported the efficacy of DCHD alone on FBG compared with conventional treatment (Liu,

2002; Zhao et al., 2016; Li et al., 2018; Ji and Che, 2020). According to the heterogeneity test (p = 0.48, $I^2 = 0\%$), the fixed effect model was applied. The pooled result did not find the difference between DCHD alone and conventional treatment in FBG (MD = 0.13 mmol/L, 95%CI: -0.09 to 0.36, p = 0.24) (Figure 6B), which means DCHD alone had the same reduction in FBG as conventional treatment. Subgroup analysis showed that DCHD can reduce FBG in diabetic patients with different age, course of disease, treatment duration and baseline BMI (Supplementary Material S10). After excluding Zhao et al., 2016, the mean difference changed from MD = 0.13 mmol/L, 95%CI: -0.09 to 0.36 to MD = -0.07 mmol/L, 95%CI: -0.60 to 0.46. It showed that this study had a great influence on the result. However, there was still no heterogeneity, and the result was still not statistically different, indicating that the result was relatively stable. (Figure 5D, Supplementary Material S7.5 and S7.6).

DCHD vs. placebo

One study including 120 patients reported that DCHD might result in a decrease in FBG compared with placebo after 8 weeks treatment (MD = -0.91 mmol/L, 95%CI: -1.29 to -0.53, p < 0.01) (Cui and Chen, 2015).

2hPG

DCHD combined with conventional treatment vs. conventional treatment

Eleven studies including 908 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional treatment alone on 2hPG (Shen and Wu, 2007; Zhang, 2018; Zhang et al., 2018; Gao, 2019; Zhang H. S. et al., 2019; Chang, 2020; Ji and Che, 2020; Wang, 2020; Duan, 2021; Zhang Q. J. et al., 2021; Zou et al., 2021). According to the heterogeneity test (p = 0.11, $I^2 = 37\%$), a fixed effect model was selected for statistical analysis. The result showed that compared with conventional treatment, combined with



significantly DCHD could reduce the 2hPG level (MD = -1.25 mmol/L, 95% CI: -1.42 to -1.09, p < 0.01)(Figure 8A). Subgroup analysis was used to analyze the effect of the main study characteristics on 2hPG. Within each subgroup of age, course of disease and treatment duration, the effect sizes were statistically significant (p < 0.01), which means that DCHD may have a lowering effect on 2hPG in patients with different ages, course of disease and treatment duration (Supplementary material S11A-C). The subgroup analysis according to different baseline BMI (≤24 or >24) showed that compared with conventional

treatment, the combination treatment had no statistical difference in reducing 2hPG in diabetic patients with BMI \leq 24, but there was a significant statistical difference in patients with BMI > 24 (Supplementary Material S11D). Sensitivity analysis indicated that the results were stable (Figure 5E, Supplementary Material S7.7).

DCHD vs. conventional treatment

Four studies including 349 patients reported the efficacy of DCHD alone on 2hPG compared with conventional treatment (Liu, 2002; Zhao et al., 2016; Li et al., 2018; Ji and Che, 2020).

	Expe	rimen	tal	C	ontrol			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Chang 2020	5.36	0.64	50	6.78	0.61	50	14.8%	-1.42 [-1.67, -1.17]	-
Duan 2021	5.66	1.32	46	6.99	1.14	46	8.8%	-1.33 [-1.83, -0.83]	
Gao 2019	6.18	0.74	38	7.67	0.65	38	13.1%	-1.49 [-1.80, -1.18]	
part-2 Ji and Che 2020	7.42	2.36	20	8.46	3.16	20	1.3%	-1.04 [-2.77, 0.69]	
Shen and Wu 2007	6.2	2.5	22	6.8	2.7	22	1.6%	-0.60 [-2.14, 0.94]	
Wang 2020	6.45	1.12	52	7.12	1.25	52	9.8%	-0.67 [-1.13, -0.21]	
Zhang 2018	5.32	1.86	60	6.47	1.31	60	7.6%	-1.15 [-1.73, -0.57]	
Zhang et al. 2018	6.33	0.74	43	7.35	0.83	43	12.6%	-1.02 [-1.35, -0.69]	- - -
Zhang et al. 2019	7.1	0.85	45	7.96	0.82	45	12.3%	-0.86 [-1.21, -0.51]	
Zhang et al. 2021	7.13	0.78	48	7.86	1.08	48	11.5%	-0.73 [-1.11, -0.35]	
Zou et al. 2021	6.52	1.21	30	7.45	1.36	30	6.5%	-0.93 [-1.58, -0.28]	
Total (95% CI)			454			454	100.0%	-1.08 [-1.28, -0.87]	◆
Heterogeneity: $Tau^2 = 0$.	06: Chi ² :	= 22.89), df = 1	10 (P =)	0.01): I	² = 56%	, D		
Test for overall effect: Z	= 10.24 (P < 0.0	00001)		<i>,</i> .				-2 -1 0 1 2
			,						Favours [experimental] Favours [control]
	Expe	rimen	tal	C	Control	L		Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Tota	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
Li et al. 2018	6.91	1.72	51	6.69	1.85	51	10.4%	0.22 [-0.47, 0.91]	
Liu 2002	7.913	1.887	52	8.453	2.228	35	6.2%	-0.54 [-1.44, 0.36]	
part-1Ji and Che 2020	8.26	3.23	20	8.46	3.16	20	1.3%	-0.20 [-2.18, 1.78]	·
Zhao et al. 2016	6.9	0.66	60	6.72	0.72	60	82.1%	0.18 [-0.07, 0.43]	†■ -
Total (95% CI)			183			166	100.0%	0.13 [-0.09, 0.36]	•
Heterogeneity: Chi ² = 2.4	46, df = 3	(P = 0)	.48); 12	= 0%					
Test for overall effect: Z		•	/.						-2 -1 0 1 2 Favours [experimental] Favours [control]

В С Α 0 0 \bigcirc \bigcirc \bigcirc 0 0 rence C Mean I Mean -1.2 Mean -1.2 0 C 14 1.4 1.4 -1.6 -1.6 80 Sample size 2015 Publication year 100 120 2010 2020 erane an FIGURE 7 Meta-regression of the FBG for DCHD combined with conventional treatment vs. conventional treatment: (A) Average age; (B) Sample size; (C)

Forest plot of the FBG: (A) DCHD combined with conventional treatment vs. conventional treatment; (B) DCHD vs. conventional treatment.

According to the heterogeneity test (p = 0.88, $I^2 = 0\%$), the fixed effect model was applied. The pooled result found statistical significance between DCHD alone and conventional treatment in 2hPG (MD = 0.54 mmol/L, 95%CI: 0.19 to 0.89, p < 0.01) (Figure 8B). Subgroup analyses according to age, course of disease, treatment duration and baseline BMI were performed to explore the effect of these factors on the result (Supplementary Material S12). However, due to the small number of studies within each subgroup, the effect of different intervention levels and different patient baseline characteristics on the result cannot be clearly defined. Sensitivity analysis showed that Zhao et al., 2016 was highly sensitive. After excluding it, the pooled result was reversed and not statistically different (MD = 0.33 mmol/L, 95%CI: 0.50 to 1.16, p = 0.43) (Figure 5F, Supplementary Material S7.8, S7.9), indicating that the result is not robust. Zhao et al., 2016 has a larger weight in the pooled result due to its relatively large sample size, narrow confidence interval, and small standard deviation. Therefore, we are more convinced of the pooled effect size involved in this study, considering that the improvement on 2hPG by DCHD alone is not as good as that of conventional treatment. However, this still needs more high-quality research to verify.

Publication year.

	Expe	eriment	tal	Co	ontrol			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
Chang 2020	6.97	1.13	50	8.3	1.2	50	12.7%	-1.33 [-1.79, -0.87]	
Duan 2021	8.97	2.01	46	10.73	1.43	46	5.2%	-1.76 [-2.47, -1.05]	
Gao 2019	7.76	0.72	38	9.12	0.85	38	21.1%	-1.36 [-1.71, -1.01]	
part-2 Ji and Che 2020	10.83	1.99	20	12.01	4.19	20	0.6%	-1.18 [-3.21, 0.85]	
Shen and Wu 2007	7.4	4.6	22	8.9	3.4	22	0.5%	-1.50 [-3.89, 0.89]	
Vang 2020	13.71	2.32	52	13.77	2.28	52	3.4%	-0.06 [-0.94, 0.82]	
Zhang 2018	8.47	1.82	60	10.37	2.12	60	5.3%	-1.90 [-2.61, -1.19]	
Zhang et al. 2018	7.62	0.87	43	8.85	0.95	43	17.8%	-1.23 [-1.62, -0.84]	
hang et al. 2019	8.67	1.01	45	9.91	1.29	45	11.5%	-1.24 [-1.72, -0.76]	
Zhang et al. 2021	9.67	0.79	48	10.49	1.5	48	11.5%	-0.82 [-1.30, -0.34]	
Zou et al. 2021	8.19	0.82	30	9.46	1.14	30	10.5%	-1.27 [-1.77, -0.77]	
Fotal (95% CI)			454			454	100.0%	-1.25 [-1.42, -1.09]	•
Heterogeneity: Chi ² = 15	5.81, df =	10 (P =	= 0.11):	l ² = 379	6			-	-4 -2 0 2
Test for overall effect: Z	= 15.10 (P < 0.0	00001)						
									Favours [experimental] Favours [control]
	Expe	eriment	al	0	Contro			Mean Difference	Mean Difference
						. T.4	1 14/-		
Study or Subgroup	Mean		Total	Mean	SL	ι οτα	al Weigh	t IV, Fixed, 95% CI	IV, Fixed, 95% CI
Study or Subgroup i et al. 2018				<u>Mean</u> 10.29					IV. Fixed, 95% Cl
i et al. 2018	<u>Mean</u> 10.43	SD 2.83	51	10.29	2.47	7 5	1 11.5%	6 0.14 [-0.89, 1.17]	IV. Fixed. 95% Cl
i et al. 2018 iu 2002	Mean	SD 2.83	51	10.29 12.487	2.47 3.843	75 33	1 11.5% 5 4.6%	6 0.14 [-0.89, 1.17] 6 0.76 [-0.88, 2.39]	IV. Fixed. 95% Cl
i et al. 2018	Mean 10.43 13.243	SD 2.83 3.787	51 52	10.29 12.487	2.47 3.843 4.19	7 5 3 3 9 2	1 11.5% 5 4.6%	6 0.14 [-0.89, 1.17] 6 0.76 [-0.88, 2.39] 6 0.50 [-2.15, 3.15]	IV. Fixed. 95% Cl
i et al. 2018 iu 2002 part-1Ji and Che 2020	Mean 10.43 13.243 12.51	SD 2.83 3.787 4.37	51 52 20	10.29 12.487 12.01	2.47 3.843 4.19	7 5 3 3 9 2 6 6	1 11.5% 5 4.6% 0 1.7%	6 0.14 [-0.89, 1.17] 6 0.76 [-0.88, 2.39] 6 0.50 [-2.15, 3.15] 6 0.58 [0.19, 0.97]	IV. Fixed. 95% CI
i et al. 2018 iu 2002 part-1Ji and Che 2020 Zhao et al. 2016	<u>Mean</u> 10.43 13.243 12.51 8.8	SD 2.83 3.787 4.37 1.26	51 52 20 60 183	10.29 12.487 12.01 8.22	2.47 3.843 4.19	7 5 3 3 9 2 6 6	1 11.5% 5 4.6% 0 1.7% 0 82.2%	6 0.14 [-0.89, 1.17] 6 0.76 [-0.88, 2.39] 6 0.50 [-2.15, 3.15] 6 0.58 [0.19, 0.97]	
i et al. 2018 iu 2002 part-1Ji and Che 2020 Zhao et al. 2016 Fotal (95% CI)	<u>Mean</u> 10.43 13.243 12.51 8.8	SD 2.83 3.787 4.37 1.26 (P = 0.4	51 52 20 60 183 88); I ² =	10.29 12.487 12.01 8.22	2.47 3.843 4.19	7 5 3 3 9 2 6 6	1 11.5% 5 4.6% 0 1.7% 0 82.2%	6 0.14 [-0.89, 1.17] 6 0.76 [-0.88, 2.39] 6 0.50 [-2.15, 3.15] 6 0.58 [0.19, 0.97]	
.i et al. 2018 .iu 2002 part-1Ji and Che 2020 Zhao et al. 2016 Fotal (95% CI) Heterogeneity: Chi ² = 0.6	<u>Mean</u> 10.43 13.243 12.51 8.8	SD 2.83 3.787 4.37 1.26 (P = 0.4	51 52 20 60 183 88); I ² =	10.29 12.487 12.01 8.22	2.47 3.843 4.19	7 5 3 3 9 2 6 6	1 11.5% 5 4.6% 0 1.7% 0 82.2%	6 0.14 [-0.89, 1.17] 6 0.76 [-0.88, 2.39] 6 0.50 [-2.15, 3.15] 6 0.58 [0.19, 0.97]	

DCHD vs. placebo

One study including 120 patients reported that there was no statistically significant difference in 2hPG between DCHD alone and placebo after 8 weeks treatment (MD = -0.53 mmol/L, 95% CI: -1.11 to 0.05, p = 0.07) (Cui and Chen, 2015).

Secondary outcomes

TC

DCHD combined with conventional treatment vs. conventional treatment

Three studies including 216 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional treatment alone on TC (Bao et al., 2020; Zhang Q. J. et al., 2021; Zou et al., 2021). According to the heterogeneity test (p = 0.23, $I^2 = 32\%$), a fixed effect model was used for statistical analysis. The pooled result illustrated that the combination with DCHD was remarkable for lowering TC compared with conventional treatment alone (MD = -0.50 mmol/L, 95%CI: -0.70 to -0.30, p < 0.01) (Figure 9A). Subgroup analyses according to age, course of disease and treatment duration were performed to explore the effect of these factors on the result (Supplementary Material S13). However, due to the small number of studies in each subgroup, the

effect of different ages, course of disease and treatment duration on TC cannot be judged yet. Sensitivity analysis showed the result was robust (Figure 5G, Supplementary Material S7.10).

DCHD vs. conventional treatment

Two studies including 207 patients reported the efficacy of DCHD alone on TC compared with conventional treatment (Liu, 2002; Zhao et al., 2016). According to the heterogeneity test (p = 0.75, $I^2 = 0\%$), the fixed effect model was applied. The pooled effect suggested a greater impact of DCHD than conventional treatment (MD = -0.29 mmol/L, 95%CI: -0.51 to -0.08, p < 0.01) (Figure 9B). Switching to a random effect model did not change the significance of the result, suggesting that the result was robust (Supplementary Material S7.11).

DCHD vs. placebo

No study compared the effect of DCHD with placebo on TC.

ΤG

DCHD combined with conventional treatment vs. conventional treatment

Three studies including 216 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional



FIGURE 9

Forest plot of the lipid metabolism index: (A) TC: DCHD combined with conventional treatment vs. conventional treatment; (B) TC: DCHD vs. conventional treatment; (C) TG: DCHD combined with conventional treatment vs. conventional treatment; (D) TG: DCHD vs. conventional treatment; (F) LDL-C: DCHD vs. conventional treatment; (G) LDL-C: DCHD vs. conventional treatment; (G) LDL-C: DCHD vs. conventional treatment.

treatment alone on TG (Bao et al., 2020; Zhang Q. J. et al., 2021; Zou et al., 2021). According to the heterogeneity test (p = 0.20, $I^2 = 37\%$), a fixed effect model was used for statistical analysis. The pooled result showed that the TG level of the combination treatment group was significantly lower than that of the conventional treatment group (MD = -0.44 mmol/L, 95% CI: -0.61 to -0.26, p < 0.01) (Figure 9C). Subgroup analyses according to age, course of disease and treatment duration were performed to explore the effect of these factors on the result (Supplementary Material S14). However, due to the small number of studies in each subgroup, the effect of different ages, course of disease and treatment duration on TG cannot be judged yet. Sensitivity analysis showed the results were robust (Figure 5H, Supplementary Material S7.12).

DCHD vs. conventional treatment

Two studies including 207 patients reported the efficacy of DCHD alone on TG compared with conventional treatment (Liu, 2002; Zhao et al., 2016). According to the heterogeneity test (p = 0.79, $I^2 = 0\%$), the fixed effect model was applied. The pooled effect indicated that there was no significant difference between the DCHD group and the conventional treatment group (MD = -0.01 mmol/L, 95%CI: -0.21 to 0.20, p = 0.96) (Figure 9D). Switching to a random effect model did not change the result, suggesting that the result was robust (Supplementary Material S7.13).

DCHD vs. placebo

No study compared the effect of DCHD with placebo on TG.

HDL-C

DCHD combined with conventional treatment vs. conventional treatment

No study compared the effect of combination treatment with conventional treatment on HDL-C.

DCHD vs. conventional treatment

Two studies including 207 patients reported the efficacy of DCHD alone on HDL-C compared with conventional treatment (Liu, 2002; Zhao et al., 2016). According to the heterogeneity test (p = 0.81, $I^2 = 0\%$), the fixed effect model was used. The pooled effect indicated that the HDL-C level in the DCHD group was much lower (MD = -0.10 mmol/L, 95% CI: -0.18 to -0.03, p < 0.01) (Figure 9E). Switching to a random effect model did not change the significance of the result, suggesting that the result was robust (Supplementary Material S7.14).

DCHD vs. placebo

No study compared the effect of DCHD with placebo on HDL-C.

LDL-C

DCHD combined with conventional treatment vs. conventional treatment

Two studies including 156 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional treatment alone on LDL-C (Zhang Q. J. et al., 2021; Zou et al., 2021). According to the heterogeneity test (p = 0.92, $I^2 = 0\%$), the fixed effect model was used. The pooled effect indicated that the LDL-C level in the DCHD group was much lower (MD = -0.58 mmol/L, 95%CI: -0.85 to -0.31, p < 0.01) (Figure 9F). Switching to a random effect model did not change the significance of the result, suggesting that the result was robust (Supplementary Material S7.15).

DCHD vs. conventional treatment

Two studies including 207 patients reported the efficacy of DCHD alone on LDL-C compared with conventional treatment (Liu, 2002; Zhao et al., 2016). According to the heterogeneity test (p = 0.02, $I^2 = 81\%$), the random effect model was used. The pooled effect indicated that there was no significant difference between the DCHD group and the conventional treatment group on LDL-C (MD = 0.03 mmol/L, 95%CI: -0.47 to 0.53, p = 0.92) (Figure 9G). Switching to a fixed effect model made the result statistically different, suggesting that the result is not robust (Supplementary Material S7.16). Due to the small number of included studies and the large differences in results between studies, we are not yet able to determine the efficacy of DCHD alone on LDL-C.

DCHD vs. placebo

No study compared the effect of DCHD with placebo on LDL-C.

HOMA-IR

DCHD combined with conventional treatment vs. conventional treatment

Six studies including 510 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional treatment alone on HOMA-IR (Zhang et al., 2018; Gao, 2019; Li, 2019; Bao et al., 2020; Wang, 2020; Zhang Q. J. et al., 2021). According to the heterogeneity test (p < 0.01, $I^2 = 96\%$), a random effect model was used. The pooled result showed that HOMA-IR in the combination treatment group was significantly lower than that in the conventional treatment group (SMD = -2.04, 95%CI: -3.09 to -0.99, p < 0.01) (Figure 10A). Subgroup analyses according to age and course of disease showed that the heterogeneity within these subgroups was not reduced, so they cannot be considered the sources of heterogeneity at present (Supplementary Material S15). We speculated that the heterogeneity might be related to the large individual

differences in HOMA-IR, leading to the large differences in baseline HOMA-IR among study points. It may also be related to the measurement bias caused by the different insulin detection methods. Sensitivity analysis showed that the pooled statistics were similar and the result was robust (Figure 5I, Supplementary Material S7.17).

DCHD vs. conventional treatment

One study including 120 patients reported that there was no significant difference in improving HOMA-IR between DCHD alone and conventional treatment after 12 weeks treatment (MD = -0.20, 95%CI: -0.75 to 0.35, p = 0.47) (Zhao et al., 2016).

DCHD vs. placebo

No study compared the effect of DCHD with placebo on HOMA-IR.

ΗΟΜΑ-β

DCHD combined with conventional treatment vs. conventional treatment

Four studies including 354 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional treatment alone on HOMA- β (Zhang et al., 2018; Gao, 2019; Li, 2019; Wang, 2020). According to the heterogeneity test (p = 0.98, $I^2 = 0\%$), a fixed effect model was used. The result revealed a significant increase in HOMA- β with combination treatment than with conventional treatment alone (SMD = 2.48, 95%CI: 2.20 to 2.76, p < 0.01) (Figure 10B). Subgroup analyses were not performed because the included studies could not be grouped by prespecified factors. Sensitivity analysis showed that the pooled statistics were similar and the result was robust (Figure 5J, Supplementary Material S7.18).

DCHD vs. conventional treatment

No study compared the effect of DCHD with conventional treatment on HOMA- $\beta.$

DCHD vs. placebo

No study compared the effect of DCHD with placebo on $\mbox{HOMA-}\beta.$

BMI

DCHD combined with conventional treatment vs. conventional treatment

Seven studies including 618 patients reported the efficacy of DCHD combined with conventional treatment vs. conventional treatment alone on BMI (Zhang, 2018; Zhang et al., 2018; Bao et al., 2020; Wang, 2020; Duan, 2021; Zhang Q. J. et al., 2021; Zou et al., 2021). According to the heterogeneity test (p < 0.01, $I^2 = 82\%$), a random effect model was used for statistical analysis. The pooled result illustrated that the combination treatment was remarkable for lowering BMI in contrast with conventional treatment alone (MD = -1.52 kg/m^2 , 95%CI: -2.55 to -0.49, p < 0.01) (Figure 10C). Subgroup analysis according to different course of disease showed reduced heterogeneity between subgroups $(I^2 = 59\%$ and 66%, respectively), which means the course of disease may be a source of heterogeneity (Supplementary Material S16A). It also indicated that DCHD might reduce BMI in patients with course of disease ≤5 years, but not in patients with course of disease >5 years. Subgroup analysis according to age showed no reduction in heterogeneity, so age cannot yet be considered a source of heterogeneity (Supplementary Material S16B). Subgroup analysis according to treatment duration was also performed. However, due to limited information, we could not judge whether the treatment duration was the source of heterogeneity (Supplementary Material S16C). Sensitivity analysis indicated that the result was stable (Figure 5K, Supplementary Material S7.19).

DCHD vs. conventional treatment

Two studies including 189 patients reported the efficacy of DCHD alone on BMI compared with conventional treatment (Liu, 2002; Li et al., 2018). According to the heterogeneity test (p = 0.14, $I^2 = 53\%$), the random effect model was applied. The pooled effect showed no significant difference between DCHD alone and conventional treatment (MD = 0.13 kg/m^2 , 95%CI: -0.71 to 0.96, p = 0.77) (Figure 10D), which means that DCHD alone may be as effective as conventional treatment in reducing BMI. Switching to a fixed effect model did not change the result, suggesting that the result was robust (Supplementary Material S7.20).

DCHD vs. placebo

No study compared the effect of DCHD with placebo on BMI.

Adverse events

Of the 17 included studies, only three studies reported adverse events. Zhao et al., 2016 reported that DCHD alone had no significant effect on liver function compared with conventional treatment (ALT: MD = -2.10 U/L, 95%CI: -4.40 to 0.20, p = 0.07; AST: MD = 1.03 U/L, 95%CI: -0.77 to 2.83, p = 0.26) (Figure 11A). Zhang H. S. et al. (2019) reported that there were three cases of hypoglycemia (6.67%) in the combination treatment group, and 11 cases of hypoglycemia (24.44%) in the conventional treatment group. Compared with the conventional treatment, the combination treatment could reduce the incidence of



hypoglycemia, and the difference was statistically significant. Zhang Q. J. et al. (2021) reported four cases of loose stool (8.33%) in the combination treatment group, five cases of abdominal distention and two cases of nausea (14.58%) in the conventional treatment group, all of which were gradually relieved without special treatment. The pooled result showed that the incidence of adverse events in the combination treatment group was significantly lower than that in the conventional treatment group (RR = 0.39, 95%CI: 0.17 to 0.89, p = 0.02) (Figure 11B). None of the other 14 studies mentioned adverse events. The safety indicators assessed in these three studies included liver function, incidence of hypoglycemia and

gastrointestinal adverse reactions, and no serious adverse events were observed. The results indicated that DCHD is relatively safe. However, due to the small sample size, more studies are still needed to further confirm its safety.

Publication bias

Funnel plot and Egger's test were used to assess publication bias on HBA1c, FBG and 2hPG. The funnel plot of HBA1c showed an asymmetric left-right distribution among study points, and two studies deviated far from the combined effect size, suggesting that there may be a large heterogeneity among studies (Figure 12A). Egger's test showed no statistical difference (p = 0.732) (Figure 12B, Supplementary Material S17.1), indicating that there was no obvious publication bias in the studies of HBA1c. The funnel plots of FBG and 2hPG showed roughly symmetrical distribution, which was consistent with Egger's test results (p = 0.266 and 0.886, respectively) (Figure 12C–F, Supplementary Material S17.2, S17.3), suggesting that there was no obvious publication bias in the studies of FBG and 2hPG.

Assessment of evidence quality

The evidence quality was assessed by the GRADE method. The overall evidence quality for each outcome was moderate to very low by assessment. The evidence quality declined mainly due to the high risk of bias, inconsistency between studies and imprecision in results. A summary of the overall evidence for each outcome is presented in Supplementary Material S18.

Discussion

Main results of this research

T2DM is an endocrine disease caused by defective insulin secretion and decreased function in regulating glucose metabolism, manifesting as chronic hyperglycemia and nutrient metabolism disorder (Scheen, 2003). During its development, glucose and lipid metabolism abnormality and insulin resistance form a vicious circle, which leads to the progressive decline of islet β -cell function, and finally promotes the occurrence and development of diabetes and its complications (LeRoith, 2002). DCHD is a classical formula derived from TCM, and several clinical studies have found that its application in the T2DM treatment may achieve better results.

In this study, we analyzed the efficacy and safety of DCHD in the T2DM treatment for the first time, providing the latest systematic evidence for the application of DCHD. We conducted a comprehensive search on both Chinese and English databases and performed a detailed analysis on the outcome indicators from the perspective of glucose and lipid metabolism, insulin resistance, islet function and BMI. Sources of heterogeneity were also explored by using meta-regression and subgroup analysis, and the evidence quality was assessed by GRADE.

Both original DCHD and modified DCHD were included. In TCM, each classical formula is a fixed whole with a definite herbal composition and efficacy. DCHD is no exception. It can be considered for T2DM patients with heat stagnation in liver and stomach syndrome. However, patients' symptoms are varied clinically, so it is necessary to add or subtract herbs based on the original DCHD according to the accompanying symptoms and individual differences. In this way, the formula can be better adapted to the patient's situation and thus exert a better effect. This is syndrome differentiation and treatment, the most important feature of TCM in disease treatment. These modified formulas have similar composition and treatment concepts as the original one and are better adapted to different patient conditions, so they can be considered the same category.

Through comprehensive analysis, some findings were obtained. A total of 627 articles were retrieved and 17 were finally included for meta-analysis. Risk of bias assessment revealed that the methodological quality of the included studies was not high, which was mainly related to the lack of detailed reporting on the specific methods of random sequence generation and allocation concealment, as well as the inadequate implementation of blinding.

Our main finding was that DCHD, in combination or alone, resulted in a statistically significant reduction in blood sugar. Compared with conventional treatment alone, the combination with DCHD can significantly improve FBG (MD = -1.08 mmol/L, 95%CI: -1.28 to -0.87, *p* < 0.01), 2hPG (MD = -1.25 mmol/L, 95% CI: -1.42 to -1.09, p < 0.01) and HbA1c (MD = -0.90%, 95%CI: -1.20 to -0.60, p < 0.01); When used alone, compared with conventional treatment, DCHD had the similar effect in improving FBG (MD = 0.13 mmol/L, 95%CI: -0.09 to 0.36, p =0.24) and HbA1c (MD = -0.04%, 95%CI: -0.17 to 0.09, *p* = 0.57), and can also reduce 2hPG, even if not as effective as conventional treatment (MD = 0.54 mmol/L, 95%CI: 0.19 to 0.89, p < 0.01). It can be seen that combined treatment can achieve a better double hypoglycemic effect, which fully reflects the advantages of integrated traditional Chinese and Western medicine. The improvement of HbA1c by the combined DCHD treatment benefited from the decrease of both FBG and 2hPG. HbA1c reflects the average blood glucose level in the past 2-3 months and is an important criterion for assessing blood glucose control. In 2010, the diabetes management guidelines promulgated by the ADA had included HbA1c \geq 6.5% as one of the diagnostic criteria for T2DM (ADA, 2010). In 2011, WHO also recommended HbA1c as one of the criteria for T2DM diagnosis in areas where conditions are well established (WHO, 2011). As the standardization and consistency of HbA1c testing improved in recent years, HbA1c has also been used as a supplementary diagnostic standard in China (CDS, 2021). The United Kingdom Prospective Diabetes Study (UKPDS) showed that the risk of various complications in T2DM patients was closely related to glycemic control. Each 1% reduction in HbA1c reduced the risk of all diabetes-related endpoints by 21%, the risk of diabetes-related death by 21%, the risk of myocardial infarction by 14%, and the risk of microvascular complications by 37% (Stratton et al., 2000). According to our findings, combining DCHD with conventional treatment could be a beneficial complementary therapy for diabetic patients. In addition, for the combination treatment, we found large heterogeneity in the HbA1c and FBG results, and performed metaregression and subgroup analysis to explore the source of heterogeneity. The results revealed that age may be one of the reasons affecting the efficacy of DCHD on HbA1c (p = 0.023, Adj R^2 = 53.55%), and DCHD may be more effective in lowering HbA1c



in middle-aged and elderly patients. For FBG, different BMI levels may be a source of heterogeneity. DCHD may have a better effect on FBG in obese patients with a BMI greater than 24. Subgroup analysis of 2hPG also found that BMI was one of the factors affecting the efficacy.

In terms of lipid metabolism, the combination treatment had a certain improvement on TC (MD = -0.50 mmol/L, 95%CI: -0.70 to -0.30, p < 0.01), TG (MD = -0.44 mmol/L, 95%CI: -0.61 to -0.26, p < 0.01) and LDL-C (MD = -0.58 mmol/L, 95% CI: -0.85 to -0.31, p < 0.01). This suggested that combined DCHD treatment may be suitable for T2DM patients with abnormal lipid metabolism. Due to the small number of included studies, it is unclear whether DCHD used alone has an improving effect on lipid metabolism.

We also found that the combined DCHD treatment was significantly better than conventional treatment in decreasing HOMA-IR (SMD = -2.04, 95%CI: 3.09 to -0.99, p < 0.01), improving HOMA- β (SMD = 2.48, 95%CI: 2.20 to 2.76, p < 0.01) and reducing BMI (MD = -1.52 kg/m², 95%CI: -2.55 to -0.49, p < 0.01). We found large heterogeneity in HOMA-IR and BMI. Subgroup analysis of HOMA-IR did not reveal a source of heterogeneity. We speculated that the heterogeneity may be related to the large individual differences in HOMA-IR, resulting in the baseline level and improvement degree on HOMA-IR varying widely among study points. It may also be related to measurement bias caused by different insulin detection methods. In addition, the methodological inadequacies

of the included studies, such as the lack of blinding and allocation concealment, may also contribute to the heterogeneity. Subgroup analysis of BMI found that different course of disease may be a reason affecting the efficacy of DCHD (total heterogeneity: $I^2 = 82\%$; subgroup analysis: $I^2 = 59\%$ and 66%, respectively), and DCHD may be more effective in reducing BMI for patients with a course of disease less than 5 years. No significant heterogeneity was found in HOMA- β , suggesting that the results were relatively consistent between individual studies. Due to the small number of included studies, it is unclear whether DCHD used alone has an improving effect on HOMA-IR, HOMA- β and BMI, which remains to be further explored in future research.

Only one study (Cui and Chen 2015) claimed to have performed a placebo control. However, the dosage form and frequency of medications in the experimental and control groups were different, which is likely to break the blindness, so the evidence for DCHD alone vs. placebo is insufficient.

Among the 17 included studies, three studies evaluated the adverse events. One study (Zhao et al., 2016) reported that DCHD had no effect on liver function. One study (Zhang H. S. et al., 2019) reported that combination with DCHD could reduce the incidence of hypoglycemia. One study (Zhang Q. J. et al., 2021) reported four cases of loose stool in the combination treatment group, five cases of abdominal distension and two cases of nausea in the conventional treatment group. The adverse effects were mainly concentrated on the gastrointestinal tract. These adverse



effects resolved spontaneously, and no special treatment was given. No serious adverse events were observed. The results showed that, when used correctly, DCHD is relatively safe. The remaining 14 studies did not report adverse events, suggesting that researchers did not pay enough attention to adverse events. Meanwhile, drug safety needs to be evaluated by multiple indicators, such as blood routine, urine routine, stool routine, liver and kidney function, incidence of hypoglycemia and patient self-reported discomfort, so as to fully reflect the drug's impact on human safety. However, the literature included in this study only reported liver function, the incidence of hypoglycemia and gastrointestinal adverse effects, so the indicators involved were not comprehensive enough. Therefore, more high-quality studies and more comprehensive indicators are needed to further confirm its safety in the future. It is worth noting that although the literature included in this study showed DCHD did not increase the occurrence of adverse effects, the clinical application of DCHD still needs to be considered comprehensively. DCHD may have potential adverse effects or toxicity if improperly applied. DCHD has the functions of soothing liver and relieving depression, clearing stomach and purging heat, and is mainly used to treat the syndrome of heat stagnation in liver and stomach. In DCHD, Chinese Thorowax Root (*Chaihu*, Bupleurum falcatum L.), Baical Skullcap Root (*Huangqin*, Scutellaria baicalensis Georgi) and Rhubarb (*Dahuang*, Rheum palmatum L.) are cold in nature. If the dose is too large or applied to patients with weak constitution, it may injure the yang qi of spleen and stomach, causing abdominal distension, diarrhea, loose stool or epigastric discomfort. Therefore, the clinical application of

10.3389/fphar.2022.918681

DCHD cannot be separated from the principle of syndrome differentiation and treatment, and it should be selected reasonably according to the pathogenesis and constitution. Meanwhile, the dosage of each herb should be adjusted according to the condition to reduce the occurrence of adverse effects. At present, there is no complete report on the toxicity of DCHD. Pharmacological studies have shown that Pinellia Tuber (Banxia, Pinellia ternata (Thunb.) Makino) is a poisonous herb because of its alkaloids, lectins and toxic raphides of calcium oxalate (Ji et al., 2014; Yu H. L. et al., 2015; Mao, 2018). These components may cause mucosal irritation, liver and kidney toxicity, and pregnancy toxicity (Xu et al., 2013; Xie et al., 2016; Xu et al., 2017; Xu et al., 2018; Jin et al., 2019). Therefore, in order to ensure safety, Pinellia Tuber, which has been processed strictly, should be used in DCHD. Through processing, the structure of calcium oxalate raphides can be destroyed, and the lectin protein will also be denatured and inactivated, thus achieving the effect of detoxification (Yu H. et al., 2015; Jin et al., 2019; Jiang and Li, 2021). Studies have also found that the use of processed Pinellia Tuber could greatly reduce the occurrence of poisoning events (Chen et al., 2020). Due to different processing excipients, Pinellia Tuber has different tendencies in efficacy and clinical application (Chen et al., 2020; Jiang and Li, 2021). Clinically, the dose of Pinellia Tuber should be reasonably determined based on the condition and the results of toxicology studies, and attention should be paid to herb concerted application, so as to exert the therapeutic effect and reduce the toxicity. The toxicity of DCHD still needs further pharmacological and toxicological studies to explore.

Funnel plot and Egger's test were performed on HBA1c, FBG and 2hPG, and no publication bias was detected, indicating that the results have certain reliability.

Study on the internal possible mechanism

TCM formulas have played a unique role in preventing and treating T2DM with their multi-component and multi-target advantages. Pharmacological studies have shown that DCHD can improve glucose and lipid metabolism, increase antioxidant enzymes activity, reduce reactive oxygen species (ROS), superoxide dismutase (SOD) and malondialdehyde (MDA) level, and upregulate the expression of pancreatic duodenal homeobox-1 (PDX-1) and MaFA mRNA in pancreatic tissue (Cui Y. R. et al., 2020). That is, it can protect pancreatic β cells by inhibiting oxidative stress. It can also increase glucose transport and improve insulin resistance by modulating the activity of the insulin receptor substrate-1/phosphor inositide-3-kinase/protein kinase B (IRS-1/ PI3K/Akt) pathway in liver tissue (Hou et al., 2020). Another study found that DCHD can significantly reduce blood sugar and cholesterol levels, increase HDL-C levels, improve glucose homeostasis and insulin resistance, reduce hepatic fat deposition and decrease total fat content. The underlying mechanism may be

related to regulating the expression of adiponectin and leptin genes in adipose tissue, inhibiting the proliferation and differentiation of adipose tissue, and balancing intestinal flora (Hussain et al., 2016). Using network pharmacology to explore the mechanism of DCHD in preventing and treating T2DM, it was found that DCHD may be closely related to tumor necrosis factor (TNF) signaling pathway, PI3K/Akt signaling pathway, p53 signaling pathway and apoptotic signaling pathway, and then play a role in regulating inflammatory response, improving insulin resistance and inhibiting pancreatic βcell apoptosis (Ren et al., 2020; Zhao and Xu, 2020; Zhang S. W. et al., 2021). In conclusion, the internal reason for DCHD to regulate glucose and lipid metabolism, reduce insulin resistance and improve islet cell function may be related to inhibiting oxidative stress, regulating inflammatory response, inhibiting islet cell apoptosis, modulating insulin signal transduction, regulating adiponectin and leptin gene expression, and balancing intestinal flora.

Limitation of this study

Although we have tried our best to use standard analytical methods, this study still has some limitations. Firstly, the methodological quality of the included studies was not high, and most studies had unclear randomization methods, lacked blinding, and did not report allocation concealment. Only one study used a placebo control, but there was a risk of unblinding. All of these may lead to a certain risk of bias. Secondly, the included studies were all single-center and small-sample, which may lack representativeness. Thirdly, some studies did not fully report research characteristics such as course of disease, treatment duration and BMI, so the selection of analysis methods for heterogeneity and the exploration of the dominant population were limited. The presence or absence of comorbidities was not reported in most studies, so the subgroup analysis on comorbidities could not be performed. Meanwhile, most of the studies were poorly standardized in reporting dropped cases. Fourthly, the included studies are all Chinese literature, which may have ethnic and regional limitations. None of the included studies were registered and no study protocol was obtained. For positive results in China are more likely to be published, there may be potential publication bias. Finally, adverse events were not reported in most studies, making it difficult to evaluate safety objectively. Therefore, there is still uncertainty about the efficacy of DCHD in the T2DM treatment.

Implications for clinical practice and future research

Based on the above findings and limitations, the following suggestions are provided for future research and practice: Firstly, improve study protocol rigor and strengthen quality control, with

particular attention to the correct implementation of center randomization, allocation concealment and blinding. Placebo control should be used reasonably to eliminate the influence of psychological factors, so as to evaluate the true efficacy and adverse effects of the experimental drug. Secondly, carry out multi-center studies and calculate the sample size reasonably to make the study results more reliable and representative. Thirdly, the report of RCTs should be conducted strictly following the Consolidated Standards of Reporting Trials (CONSORT) statement, with particular emphasis on reporting the age, course of disease, treatment duration, and presence or absence of comorbidities, so as to explore the source of heterogeneity through statistical analysis and further analyze the dominant population. Fourthly, clinical trial registration should be carried out before starting, and both positive and negative results should be reported truthfully to ensure transparency of information and reduce publication bias. Finally, pay attention to observing and monitoring adverse events and establish strict adverse events handling and reporting procedures.

Conclusion

In summary, we found that the combination with DCHD in the T2DM treatment has more advantages than conventional treatment alone, which can further regulate the glucose and lipid metabolism, reduce insulin resistance, improve islet function and lower BMI. DCHD used alone can also play a certain role in regulating blood glucose, but the current evidence is insufficient to clarify the effect of DCHD alone on lipid metabolism, insulin resistance, islet function and BMI. Meanwhile, DCHD is relatively safe. This indicates that DCHD may have a positive effect on T2DM. However, given the limited number of included studies, small sample size and poor methodological quality, the evidence of this study remain uncertain and the results should be interpreted and applied with caution. In the T2DM treatment, clinical decisions still need to be made by considering the patient's overall situation. In the future, more high-quality, large-sample, multi-center, randomized, double-blind and placebo-controlled studies are still needed to provide more reliable evidence for the clinical application of DCHD.

Data availability statement

The original contributions presented in the study are included in the article/Supplementary Material, further inquiries can be directed to the corresponding authors.

References

ADA (2021). 9. Pharmacologic approaches to glycemic treatment: Standards of medical care in diabetes-2021. *Diabetes Care* 44 (Suppl. 1), S111–s124. doi:10.2337/dc21-S009

Author contributions

ZZ and CX conceived the idea and designed the study. ZZ, CY, HY, and QL conducted the literature searching, study selection, data extraction and methodology assessment. ZZ and YL performed this meta-analysis. Disagreement was resolved by discussion with XF and HX. ZZ wrote the original manuscript and HG and CX revised this manuscript. All authors approved the final manuscript before submission.

Funding

This work is supported by the Traditional Chinese Medicine Science and Technology Research Project of Sichuan Province (No. 2020ZD001), the Major Science and Technology Project of Sichuan Province (No. 2022ZDZX0022), National Traditional Chinese Medicine Prevention and Treatment of Diabetic Macrovascular Disease Inheritance and Innovation Team of China (No. ZYYCXTD-C-202209), the National Natural Science Foundation of China (No. 82004341) and Sichuan Science and Technology Department Project (No. 2021JDRC0150).

Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Publisher's note

All claims expressed in this article are solely those of the authors and do not necessarily represent those of their affiliated organizations, or those of the publisher, the editors and the reviewers. Any product that may be evaluated in this article, or claim that may be made by its manufacturer, is not guaranteed or endorsed by the publisher.

Supplementary material

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fphar. 2022.918681/full#supplementary-material

ADA (2010). Standards of medical care in diabetes—2010. Diabetes Care 33 (Suppl. 1), S11-S61. doi:10.2337/dc10-S011

Bao, W. P., Fan, Y. F., and Chu, X. Q. (2020). Efficiency observation of dachaihu decoction and exenatide on newly diagnosed type 2 diabetes mellitus complicated by non-alcoholic fatty liver. *Shanxi J. TCM* 36 (2), 27–29.

Baye, A. M., Fanta, T. G., Siddiqui, M. K., and Dawed, A. Y. (2021). The genetics of adverse drug outcomes in type 2 diabetes: A systematic review. *Front. Genet.* 12, 675053. doi:10.3389/fgene.2021.675053

Bianchi, C., Miccoli, R., and Del Prato, S. (2013). Hyperglycemia and vascular metabolic memory: Truth or fiction? *Curr. Diab Rep.* 13 (3), 403–410. doi:10.1007/s11892-013-0371-2

Cannon, A., Handelsman, Y., Heile, M., and Shannon, M. (2018). Burden of illness in type 2 diabetes mellitus. *J. Manag. Care Spec. Pharm.* 24 (9-a Suppl. l), S5–s13. doi:10.18553/jmcp.2018.24.9-a.s5

CDS (2021). Guideline for the prevention and treatment of type 2 diabetes mellitus in China (2020 edition). *Chin. J. Diabetes Mellit.* 37 (4), 311–398. doi:10. 3760/cma.j.cn115791-20210221-00095

Chang, L. L. (2020). Effect of insulin glargine combined with dachaihu decoction in the treatment of T2DM patients and its effect on blood sugar level. *J. Pract. Diabetol.* 16 (5), 71–72.

Chen, L. J., Li, C. J., Gong, Z. D., Sun, L. J., and Tong, Y. J. (2020). Clinical dosage and toxicity analysis of banxia. *Shandong J. Tradit. Chin. Med.* 39 (3), 298–301. doi:10.16295/j.cnki.0257-358x.2020.03.021

Cui, H., Li, Y., Wang, Y., Jin, L., Yang, L., Wang, L., et al. (2020). Da-chai-hu decoction ameliorates high fat diet-induced nonalcoholic fatty liver disease through remodeling the gut microbiota and modulating the serum metabolism. *Front. Pharmacol.* 11, 584090. doi:10.3389/fphar.2020.584090

Cui, H. Y., and Chen, Y. L. (2015). Clinical observation on the treatment of type 2 diabetes with modified dachaihu decoction. *Hebei J. TCM* 37 (8), 1195–1197. doi:10.3969/j.issn.1002-2619.2015.08.026

Cui, Y. R., Zhou, Q., and Zhu, X. D. (2020). Effect of dachaihu decoction on islet β cell damage induced by oxidative stress in type 2 diabetes model rats. *Mod. Tradit. Chin. Med. Mat. Med-World Sci. Technol.* 22 (5), 1458–1463. doi:10.11842/wst. 20190505003

Deng, Y. X., and Huang, X. Y. (2020). Clinical experience of professor Huang xinyan with the application of dachaihu decoction. *Chin. Med. Mod. Dist. Educ.* 18 (5), 39–41. doi:10.3969/j.issn.1672-2779.2020.05.017

Dou, H. R., Chen, M., Tang, X., and Zhou, L. (2019). Clinical observation of modified dachaihu decoction combined with western medicine in the treatment of bile reflux gastritis after cholecystectomy and its effect on patients' anxiety state. *Yunnan J. Tradit. Chin. Med.* 40 (3), 47–50. doi:10.16254/j.cnki.53-1120/r.2019. 03.021

Duan, J. (2021). Clinical effect analysis of modified dachaihu decoction in the treatment of type 2 diabetes mellitus. *Diabetes New World* 24 (2), 60–62. doi:10. 16658/j.cnki.1672-4062.2021.02.060

Feng, Y., Gao, X. X., and Qin, X. M. (2019). Research progress on lipid-lowering efficacy and mechanism of Chaihu and its related prescriptions. *J. Chin. Med. Mater* 42 (8), 1957–1961. doi:10.13863/j.issn1001-4454.2019.08.049

Galicia-Garcia, U., Benito-Vicente, A., Jebari, S., Larrea-Sebal, A., Siddiqi, H., Uribe, K. B., et al. (2020). Pathophysiology of type 2 diabetes mellitus. *Int. J. Mol. Sci.* 21 (17), 6275. doi:10.3390/ijms21176275

Gao, L. J. (2019). Effects of modified dachaihu decoction on insulin resistance, β -cell function and low inflammatory response in the adjuvant treatment of obesity type 2 diabetes mellitus. *China's Naturop.* 27 (14), 61–62. doi:10.19621/j.cnki.11-3555/r.2019.1431

Glovaci, D., Fan, W., and Wong, N. D. (2019). Epidemiology of diabetes mellitus and cardiovascular disease. *Curr. Cardiol. Rep.* 21 (4), 21. doi:10.1007/s11886-019-1107-y

Han, K., Kwon, O., Park, H. J., Jung, S. Y., Yang, C., and Son, C. G. (2020). Effect of daesiho-tang on obesity with non-alcoholic fatty liver disease: A study protocol for A randomised, double-blind, placebo-controlled pilot trial. *Trials* 21 (1), 128. doi:10.1186/s13063-020-4068-y

Hou, P., Zhou, Q., and Zhu, X. D. (2020). Effects of dachaihu decoction on IRS-1/ PI3K/akt pathway in liver tissue of type 2 diabetic rats. *J. Northwest Minzu Univ. Nat. Sci.* 41 (3), 57–62. doi:10.14084/j.cnki.cn62-1188/n.2020.03.011

Hussain, A., Yadav, M. K., Bose, S., Wang, J. H., Lim, D., Song, Y. K., et al. (2016). Daesiho-tang is an effective herbal formulation in attenuation of obesity in mice through alteration of gene expression and modulation of intestinal microbiota. *PLoS One* 11 (11), e0165483. doi:10.1371/journal.pone.0165483

IDF (2021). *IDF diabetes atlas*. 10th edn. Brussels, Belgium: International Diabetes Federation. [Online]. Available: https://www.diabetesatlas.org (Accessed 0315, 2022).

Ji, J. L., and Che, Z. Y. (2020). 20 cases of type 2 diabetes treated with modified dachaihu decoction combined with metformin. *TCM Res.* 33 (9), 18–21. doi:10. 3969/j.issn.1001-6910.2020.09.06

Ji, L., Tong, X., Wang, H., Tian, H., Zhou, H., Zhang, L., et al. (2013). Efficacy and safety of traditional Chinese medicine for diabetes: A double-blind, randomised, controlled trial. *PLoS One* 8 (2), e56703. doi:10.1371/journal. pone.0056703

Ji, X., Huang, B., Wang, G., and Zhang, C. (2014). The ethnobotanical, phytochemical and pharmacological profile of the genus Pinellia. *Fitoterapia* 93, 1–17. doi:10.1016/j.fitote.2013.12.010

Jiang, Y., and Li, D. X. (2021). Analysis of toxicity and clinical application of Pinellia ternata after processing. *J. Med. Theor. Prac.* 34 (19), 3335–3337. doi:10. 19381/j.issn.1001-7585.2021.19.010

Jin, X. Q., Huang, C. Q., and Zhang, G. (2019). Toxic components and processing mechanism of rhizoma pinelliae. *Lishizhen Med. Mater Med. Res.* 30 (7), 1717–1720. doi:10.3969/j.issn.1008-0805.2019.07.065

LeRoith, D. (2002). Beta-cell dysfunction and insulin resistance in type 2 diabetes: Role of metabolic and genetic abnormalities. *Am. J. Med.* 113 (Suppl. 6A), 3s-11s. doi:10.1016/s0002-9343(02)01276-7

Li, C. Y., Chiu, C. H., Huang, H. S., Lin, C. H., and Wu, T. S. (2006). Highperformance liquid chromatographic method for simultaneous quantification of eight major biologically active ingredients in 'da-chai-hu-tang' preparation. *Biomed. Chromatogr.* 20 (4), 305–308. doi:10.1002/bmc.572

Li, L. P., Zhang, H. Q., Gao, Y. H., Chen, H. W., Wang, J. Y., and Liu, P. S. (2018). Clinical observation on 102 cases of obesity type 2 diabetes treated with dachaihu decoction. *Chin. Med. Mod. Dist. Educ.* 16 (24), 43–45. doi:10.3969/j.issn.1672-2779.2018.24.019

Li, S. S., and Yan, Y. (2019). Cases study of professor yan yong's classic prescription in treating diabetes mellitus. *Chin. J. Ethnomed Ethnopharm* 28 (16), 81–83.

Li, S. W. (2019). Effects of modified dachaihu decoction combined with liraglutide on insulin resistance, β -cell function and low inflammatory response in obese type 2 diabetes mellitus. *Harbin Med. J.* 39 (2), 185–186.

Li, Y., Teng, D., Shi, X., Qin, G., Qin, Y., Quan, H., et al. (2020). Prevalence of diabetes recorded in mainland China using 2018 diagnostic criteria from the American diabetes association: National cross sectional study. *Bmj* 369, m997. doi:10.1136/bmj.m997

Lian, F., Tian, J., Chen, X., Li, Z., Piao, C., Guo, J., et al. (2015). The efficacy and safety of Chinese herbal medicine jinlida as add-on medication in type 2 diabetes patients ineffectively managed by metformin monotherapy: A double-blind, randomized, placebo-controlled, multicenter trial. *PLoS One* 10 (6), e0130550. doi:10.1371/journal.pone.0130550

Lin, X., Xu, Y., Pan, X., Xu, J., Ding, Y., Sun, X., et al. (2020). Global, regional, and national burden and trend of diabetes in 195 countries and territories: An analysis from 1990 to 2025. *Sci. Rep.* 10 (1), 14790. doi:10.1038/s41598-020-71908-9

Ling, W., Huang, Y., Huang, Y. M., Fan, R. R., Sui, Y., and Zhao, H. L. (2020). Global trend of diabetes mortality attributed to vascular complications, 2000-2016. *Cardiovasc Diabetol.* 19 (1), 182. doi:10.1186/s12933-020-01159-5

Liu, H. F. (2002). Clinical and experimental study on therapeutic effect of jiaweidachaihu granula(JG) in improving insulin resistance (Beijing: Beijing University of Chinese Medicine). [Dissertation].

Liu, J., Ren, Z. H., Qiang, H., Wu, J., Shen, M., Zhang, L., et al. (2020). Trends in the incidence of diabetes mellitus: Results from the global burden of disease study 2017 and implications for diabetes mellitus prevention. *BMC Public Health* 20 (1), 1415. doi:10.1186/s12889-020-09502-x

Liu, X. Q. (2014). Studies on HPLC fingerprints of da Chaihu decoction and guifu dihuang prescription (Beijing: Academy of Military Sciences). [Master's thesis].

Ma, R. C. W. (2018). Epidemiology of diabetes and diabetic complications in China. *Diabetologia* 61 (6), 1249–1260. doi:10.1007/s00125-018-4557-7

Mao, S. H. (2018). Study the inflammatory mechanism of Pinellia ternata and pinella pedatisecta and the detoxification mechanism of ginger based on ROS-MAPK/ NLRP3-IL-1 β signaling pathway (Nanjing: Nanjing University of Chinese Medicine). [Master's thesis].

Martinez, R., Lloyd-Sherlock, P., Soliz, P., Ebrahim, S., Vega, E., Ordunez, P., et al. (2020). Trends in premature avertable mortality from non-communicable diseases for 195 countries and territories, 1990-2017: A population-based study. *Lancet Glob. Health* 8 (4), e511–e523. doi:10.1016/s2214-109x(20)30035-8

Moher, D., Liberati, A., Tetzlaff, J., and Altman, D. G. (2009). Preferred reporting Items for systematic reviews and meta-analyses: The PRISMA statement. *PLoS Med.* 6 (7), e1000097. doi:10.1371/journal.pmed.1000097

Page, M. J., McKenzie, J. E., Bossuyt, P. M., Boutron, I., Hoffmann, T. C., Mulrow, C. D., et al. (2021). The PRISMA 2020 statement: An updated guideline for reporting systematic reviews. *Bmj* 372, n71. doi:10.1136/bmj.n71

Pang, G. M., Zhang, F., Wang, K. F., Wu, N., Zhu, P., Fu, Y. X., et al. (2021). Sequential triple therapy of Chinese herbal medicine for type 2 diabetes: A realworld study of 252 patients. J. Tradit. Chin. Med. 62 (18), 1612-1616+1621. doi:10. 13288/j.11-2166/r.2021.18.010

Qian, W., Cai, X., Zhang, X., Wang, Y., Qian, Q., and Hasegawa, J. (2016). Effect of daisaikoto on expressions of SIRT1 and NF-kappaB of diabetic fatty liver rats induced by high-fat diet and streptozotocin. *Yonago Acta Med.* 59 (2), 149–158.

Reed, J., Bain, S., and Kanamarlapudi, V. (2021). A review of current trends with type 2 diabetes epidemiology, aetiology, pathogenesis, treatments and future perspectives. *Diabetes Metab. Syndr. Obes.* 14, 3567–3602. doi:10.2147/DMSO. S319895

Ren, B., Tan, L., Xiong, Y., Ji, W., Mu, J., Pei, Y., et al. (2020). Integrated analysis of the mechanisms of da-chai-hu decoction in type 2 diabetes mellitus by a network pharmacology approach. *Evid. Based Complement. Altern. Med.* 2020, 9768414. doi:10.1155/2020/9768414

Roden, M., and Shulman, G. I. (2019). The integrative biology of type 2 diabetes. Nature 576 (7785), 51–60. doi:10.1038/s41586-019-1797-8

Scheen, A. J. (2003). Pathophysiology of type 2 diabetes. Acta Clin. Belg 58 (6), 335–341. doi:10.1179/acb.2003.58.6.001

Shen, L. X., and Wu, Q. P. (2007). Study on improvement of islet β cell function in patients with type 2 diabetes by integrative Chinese and western medicine. Chongqing Med. 36 (23), 2420–2421. doi:10.3969/j.issn.1671-8348.2007.23.032

Singh, A., Kukreti, R., Saso, L., and Kukreti, S. (2022). Mechanistic insight into oxidative stress-triggered signaling pathways and type 2 diabetes. *Molecules* 27 (3), 950. doi:10.3390/molecules27030950

Stanaway, J. D., Afshin, A., Gakidou, E., Lim, S. S., Abate, D., Abate, K. H., et al. (2018). Global, regional, and national comparative risk assessment of 84 behavioural, environmental and occupational, and metabolic risks or clusters of risks for 195 countries and territories, 1990–2017: A systematic analysis for the global burden of disease study 2017. *Lancet* 392 (10159), 1923–1994. doi:10.1016/s0140-6736(18)32225-6

Stratton, I. M., Adler, A. I., Neil, H. A., Matthews, D. R., Manley, S. E., Cull, C. A., et al. (2000). Association of glycaemia with macrovascular and microvascular complications of type 2 diabetes (UKPDS 35): Prospective observational study. *BMJ* 321 (7258), 405–412. doi:10.1136/bmj.321.7258.405

Tong, X. L., Wu, S. T., Lian, F. M., Zhao, M., Zhou, S. P., Chen, X. Y., et al. (2013). The safety and effectiveness of TM81, a Chinese herbal medicine, in the treatment of type 2 diabetes: A randomized double-blind placebo-controlled trial. *Diabetes Obes. Metab.* 15 (5), 448–454. doi:10.1111/dom.12051

Wang, L., Li, X., Wang, Z., Bancks, M. P., Carnethon, M. R., Greenland, P., et al. (2021). Trends in prevalence of diabetes and control of risk factors in diabetes among US adults, 1999-2018. *Jama* 326 (8), 704–713. doi:10.1001/jama.2021.9883

Wang, S. Y., Xie, Y. C., and Min, L. (2022). Effect of modified da chaihutang on FXR/FGF15/FGFR4 signaling pathway in mice with cholesterol gallstone differentiated into damp-heat syndrome. *Chin. J. Exp. Tradit. Med. Formulae* 28 (3), 15–24. doi:10.13422/j.cnki.syfjx.20220201

Wang, Z. H. (2020). Clinical efficacy of modified dachaihu decoction combined with western medicine in the treatment of obese type 2 diabetes mellitus and its effect on insulin resistance in patients. *Henan Tradit. Chin. Med.* 40 (1), 56–59. doi:10.16367/j.issn.1003-5028.2020.01.0015

WHO (2020). Global health estimates: Life expectancy and leading causes of death and disability. [Online]. Available: https://www.who.int/data/gho/data/themes/mortality-and-global-health-estimates (Accessed 0315, 2022).

WHO (2016). *Global report on diabetes*, Geneva: WHO Press, [Online]. Available: https://www.who.int/publications/i/item/9789241565257 (Accessed 0315, 2022).

WHO (2011). "WHO guidelines approved by the guidelines review committee," in Use of glycated haemoglobin (HbA1c) in the diagnosis of diabetes mellitus: Abbreviated report of a WHO consultation, Geneva: WHO Press, 1–25.

Xu, Z. M., Wang, Y. X., Shan, J. J., Xu, S., and Wang, S. C. (2013). Temperature dependent photoluminescence from polysilsesquioxane xerogels. *Amr* 668 (3), 255–258. doi:10.14148/j.issn.1672-0482.2013.03.01810.4028/www.scientific.net/ amr.668.255

Xie, H. H., Xu, J. Y., Xie, T., Meng, X., Lin, L. L., He, L. L., et al. (2016). Effects of Pinellia ternata (Thunb.) Berit. on the metabolomic profiles of placenta and amniotic fluid in pregnant rats. *J. Ethnopharmacol.* 183, 38–45. doi:10.1016/j.jep. 2016.02.030

Xu, J. Y., Dai, C., Shan, J. J., Xie, T., Xie, H. H., Wang, M. M., et al. (2018). Determination of the effect of Pinellia ternata (Thunb.) Breit. on nervous system development by proteomics. *J. Ethnopharmacol.* 213, 221–229. doi:10.1016/j.jep. 2017.11.014

Xu, J. Y., Zhang, K., Xie, H. H., Shan, J. J., and Xie, T. (2017). Effects of raw Pinellia ternate and ganjiang renshen banxia pill on reproductive related toxicity of mice. *Chin. J. Tradit. Chin. Med. Pharm.* 32 (7), 3154–3157.

Xu, Y. T., and Sun, X. P. (2013). Jiawei dachaihu decoction in treatment of experimental study of cholecystitis. *Chin. J. Exp. Tradit. Med. Formulae* 19 (3), 234–237. doi:10.13422/j.cnki.syfjx.2013.03.076

Yang, C., Wang, T., Chen, J., He, J., Li, Y., Chen, C., et al. (2021). Traditional Chinese medicine formulas alleviate acute pancreatitis: Pharmacological activities and mechanisms. *Pancreas* 50 (10), 1348–1356. doi:10.1097/MPA. 000000000001931

Yang, J. M., Sun, Y., Wang, M., Zhang, X. L., Zhang, S. J., Gao, Y. S., et al. (2019). Regulatory effect of a Chinese herbal medicine formula on non-alcoholic fatty liver disease. *World J. Gastroenterol.* 25 (34), 5105–5119. doi:10.3748/wjg.v25.i34.5105

Yoshie, F., Iizuka, A., Komatsu, Y., Matsumoto, A., Itakura, H., and Kondo, K. (2004). Effects of dai-saiko-to (Da-Chai-Hu-Tang) on plasma lipids and atherosclerotic lesions in female heterozygous heritable kurosawa and kusanagi-hypercholesterolemic (KHC) rabbits. *Pharmacol. Res.* 50 (3), 223–230. doi:10.1016/j.phrs.2004.02.003

Yu, H., Pan, Y., Wu, H., Ge, X., Zhang, Q., Zhu, F., et al. (2015). The alumprocessing mechanism attenuating toxicity of Araceae Pinellia ternata and Pinellia pedatisecta. Arch. Pharm. Res. 38 (10), 1810–1821. doi:10.1007/s12272-015-0556-0

Yu, H. L., Zhao, T. F., Wu, H., Pan, Y. Z., Zhang, Q., Wang, K. L., et al. (2015). Pinellia ternata lectin exerts a pro-inflammatory effect on macrophages by inducing the release of pro-inflammatory cytokines, the activation of the nuclear factor-κB signaling pathway and the overproduction of reactive oxygen species. *Int. J. Mol. Med.* 36 (4), 1127–1135. doi:10.3892/ijmm.2015.2315

Zhang, E., and Wu, Y. (2014). Metabolic memory: Mechanisms and implications for diabetic vasculopathies. *Sci. China Life Sci.* 57 (8), 845–851. doi:10.1007/s11427-014-4710-6

Zhang, H. S., Liu, X. Q., Han, X., Zhang, W., and Xu, J. J. (2019). Effect of dachaihu decoction on hypoglycemic efficacy and blood glucose fluctuation in type 2 diabetes mellitus with initial liver-stomach depression-heat treated by intensive insulin. *Hubei J. TCM* 41 (11), 7–9.

Zhang, H. X., Wang, H. C., Li, S. J., He, L., and Che, Z. Y. (2015). Professor Wang guobin's experience in treating diabetes with dachaihu decoction. *TCM Res.* 28 (8), 42–43. doi:10.3969/j.issn.1001-6910.2015.08.21

Zhang, Q. J., Wu, R., Wang, Z. F., Qin, L., Guo, F. N., An, R., et al. (2021). Clinical study of modified dachaihu decoction on type 2 diabetes mellitus complicated with hyperuricemia. *J. Clin. Exp. Med.* 20 (5), 478–482. doi:10.3969/j.issn.1671-4695. 2021.05.008

Zhang, S. W., Zhang, W. H., Li, H. D., Bai, W., and Gong, Y. B. (2021). Mechanism of dachaihu decoction in the treatment of type 2 diabetes mellitus based on network pharmacology. *World J. Integr. Tradit. West Med.* 16 (7), 1256–1264. doi:10.13935/j.cnki.sjzx.210717

Zhang, W. H., Li, Q. Y., Yang, C. W., Zhou, Y. R., Guo, Z. H., Wang, T. T., et al. (2018). Effect of dachaihu decoction combined with liraglutide on insulin resistance, beta-cell function and low inflammatory response in obese patients with type 2 diabetes mellitus. *Mod. J. Integr. Tradit. Chin. West Med.* 27 (1), 23–264-30. doi:10.3969/j.issn.1008-8849.2018.01.006

Zhang, X. H. (2018). Clinical observation of dachaihu decoction in treating type 2 diabetes mellitus. *Yunnan J. Tradit. Chin. Med.* 39 (7), 44–45. doi:10.16254/j.cnki. 53-1120/r.2018.07.018

Zhang, Y. N. (2016). Dachaihu decoction in the treatment of 32 cases of type 2 diabetes. J. Health Guide (51), 200. doi:10.3969/j.issn.1006-6845.2016.51.192

Zhang, Y., Xie, H. Y., Zhou, L. Y., Tian, Y., and Xie, C. G. (2019). Clinical observation on modified shenqi compound in treatment of 46 cases of newly diagnosed type 2 diabetes mellitus with spleen deficiency. *Chin. Arch. Tradit. Chin. Med.* 37 (12), 2967–2970. doi:10.13193/j.issn.1673-7717.2019.12.035

Zhao, C. F., Li, H., Cheng, L., and Guan, S. (2016). Clinical observation on the treatment of type 2 diabetes mellitus with hyperlipidemia by clearing liver, downbearing turbidity and resolving dampness. *World Chin. Med.* 11 (2), 253–255. doi:10.3969/j.issn.1673-7202.2016.02.018

Zhao, K., and Xu, X. Y. (2020). Study on the therapeutic mechanism of dachaihu decoction on type 2 diabetes based on network pharmacology. *Mod. Tradit. Chin. Med. Mat. Med-World Sci. Technol.* 22 (9), 3225–3234. doi:10.11842/wst.20190628005

Zheng, Y., Ley, S. H., and Hu, F. B. (2018). Global aetiology and epidemiology of type 2 diabetes mellitus and its complications. *Nat. Rev. Endocrinol.* 14 (2), 88–98. doi:10.1038/nrendo.2017.151

Zou, X. M., and Wang, F. R. (2014). Discussion on the theory of "disease caused by overuse" and the prevention and treatment of hyperlipidemia with dachaihu decoction. *Chin. J. Basic Med. Tradit. Chin. Med.* 20 (9), 1199–1200.

Zou, Y. L., Zhang, W. B., He, Y. Y., and Li, H. X. (2021). Effects of dachaihu decoction on blood glucose and blood lipid levels in patients with type 2 diabetes mellitus complicated by obesity. *Mod. Med. Health Res.* 5 (9), 18–20.

Glossary MD mean difference MDA malondialdehyde 2hPG 2-h postprandial glucose NR not reported ADA American Diabetes Association NSD no significant difference BMI body mass index PDX-1 pancreatic duodenal homeobox-1 CDS Chinese Diabetes Society PRISMA the Preferred Reporting Items for Systematic Review CG control group interventions and Meta-Analysis CHiCTR Chinese Clinical Trial Registry PROSPERO International Prospective Register of Systematic CI confidence interval Reviews CNKI China National Knowledge Infrastructure RCTs randomized controlled trials CONSORT Consolidated Standards of Reporting Trials ROS reactive oxygen species DCHD dachaihu decoction **RR** relative risk SMD standardized mean difference FBG fasting blood glucose SOD superoxide dismutase GRADEGrading of Recommendations Assessment Grading of Recommendations AssessmentDevelopment and Evaluation T2DM type 2 diabetes mellitus HbA1c glycated hemoglobin TCM traditional Chinese medicine HDL-C high-density lipoprotein cholesterol TC total cholesterol HOMA-IR homeostasis model assessment of insulin resistance TG triglyceride HOMA-B homeostasis model assessment of beta-cell function TNF tumor necrosis factor IRS-1/PI3K/Akt insulin receptor substrate-1/phosphor UKPDS United Kingdom Prospective Diabetes Study inositide-3-kinase/protein kinase B VIP China Science and Technology Journal Database LDL-C low-density lipoprotein cholesterol WHO World Health Organization.