



# Influence of threat and serotonin transporter genotype on interference effects

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Emotion-cognition interactions are critical in goal-directed behavior and may be disrupted in psychopathology. Growing evidence also suggests that emotion-cognition interactions are modulated by genetic variation, including genetic variation in the serotonin system. The goal of the current study was to examine the impact of threat-related distracters and serotonin transporter promoter polymorphism (5-HTTLPR/rs25531) on cognitive task performance in healthy females. Using a novel threat-distracter version of the Multi-Source Interference Task specifically designed to probe emotion-cognition interactions, we demonstrate a robust and temporally dynamic modulation of cognitive interference effects by threat-related distracters relative to other distracter types and relative to no-distracter condition. We further show that threat-related distracters have dissociable and opposite effects on cognitive task performance in easy and difficult task conditions, operationalized as the level of response interference that has to be surmounted to produce a correct response. Finally, we present evidence that the 5-HTTLPR/rs25531 genotype in females modulates susceptibility to cognitive interference in a global fashion, across all distracter conditions, and irrespective of the emotional salience of distracters, rather than specifically in the presence of threat-related distracters. Taken together, these results add to our understanding of the processes through which threat-related distracters affect cognitive processing, and have implications for our understanding of disorders in which threat signals have a detrimental effect on cognition, including depression and anxiety disorders.

**Keywords:** cognition, emotion, interference resolution, threat, serotonin transporter gene, 5-HTTLPR, MSIT

## INTRODUCTION

The ability to successfully carry out a task despite interference from task-irrelevant stimuli is a crucial requirement for goal-directed behavior. According to accepted models of selective attention and cognitive-control, task-irrelevant stimuli interfere with cognitive task performance by competing with task-relevant stimuli for attentional and response-selection resources (Desimone and Duncan, 1995; Miller and Cohen, 2001). However, the impact of distracters on task performance – or conversely, our ability to resist interference from these distracters – can vary considerably, depending on the attributes of the distracters and the attributes of the task itself (Lavie, 2005), as well as on individual differences in susceptibility to various distracters.

Critically, with respect to distracter attributes, such interference can come from both neutral and emotionally salient stimuli, highlighting the fact that emotional and cognitive processes are closely interrelated, giving rise to complex and bidirectional emotion-cognition interactions (Davidson, 2003; Blair et al., 2007). In particular, if neutral distracters impair task performance, threat-related distracters should be even more effective in high-jacking attention and interfering with the task at hand due to the

preferential processing of threat stimuli over non-threat stimuli in the brain. This rapid and automatic processing of threat signals is possible because the amygdala receives threat-related information through a fast subcortical pathway as well as through a slower cortical route (Romanski and LeDoux, 1992; Morris et al., 1999), a finding supported by functional neuroimaging studies showing that the amygdala responds to threat stimuli that are outside of attentional focus or conscious awareness (Whalen et al., 1998; Vuilleumier et al., 2001). From an evolutionary perspective, in humans as in many other species, such preferential processing of potential threat signals serves the adaptive function of facilitating rapid threat detection and fight-or-flight responses essential for survival (Ohman and Mineka, 2001). However, although supported by some studies (Vuilleumier et al., 2001; Dolcos and McCarthy, 2006; Blair et al., 2007; Mitchell et al., 2008), such increased distractibility by threat-related distracters relative to neutral distracters in behavioral measures has not been consistently demonstrated in healthy subjects (Bar-Haim et al., 2007), suggesting that additional modulatory factors may be at play.

Neuroimaging evidence also suggests that the effects of threat distracters on interference processing may dynamically change

over the time-course of the task, because the amygdala response to threat stimuli is temporally dynamic due to both habituation and regulation processes. Salient or novel stimuli initially elicit a strong neural and behavioral response, because they may signal threat or reward, and are thus potentially important to the organism's survival. Habituation refers to a diminished reactivity to a specific stimulus or stimulus class following repeated presentation with no important consequences for the organism, and it is believed to serve an adaptive function of preserving cognitive and behavioral resources and allowing continuous vigilance (Wright et al., 2001). Growing evidence from neuroimaging studies in humans shows that the amygdala habituates to repeatedly presented threat stimuli both in healthy individuals (Breiter et al., 1996; Whalen et al., 1998; Wright et al., 2001) and in patients with anxiety disorders such as post-traumatic stress disorder (Shin et al., 2005). In addition, neuroimaging studies of emotion regulation show a decrease in amygdala response to threat-related stimuli when human subjects actively regulate their emotional response using cognitive-control strategies such as reappraisal, distraction, or suppression (Ochsner et al., 2002; Phan et al., 2005; Eippert et al., 2007; Kim and Hamann, 2007; Wager et al., 2008; McRae et al., 2010), and convergent results have been obtained in animals in the context of fear extinction (Quirk and Beer, 2006; Hartley and Phelps, 2010). This temporally dynamic character of amygdala response to threat stimuli may also be a factor modulating threat-distracter effects on cognitive task performance.

Another important factor that may modulate – or obscure – threat-distracter effects on cognitive task performance is the difficulty level of the task itself. For instance, high perceptual load has been shown to decrease distracter effects relative to low perceptual load for neutral distracters (Rees et al., 1997), although salient distracters such as images of human faces appear to escape this modulation (Lavie et al., 2003). In contrast, high cognitive load increases distracter effects relative to low cognitive load (Lavie, 2005). In particular, a task that is too easy to perform may not allow detection of threat-distracter effects due to ceiling effects in performance, an issue particularly relevant to studies of healthy adults. Ideally, therefore, the impact of threat distracters should be investigated and compared in two different task conditions varying in difficulty, or in the level of cognitive demand required to successfully perform the task.

Finally, growing evidence suggests that common genetic variation in the serotonin system modulates both emotional reactivity and cognitive processing in the human brain, and may also modulate the impact of threat distracters on cognitive task performance. Serotonin, or 5-hydroxytryptamine (5-HT), is known to be involved in a range of behavioral control processes (Cools et al., 2008, 2011; Dayan and Huys, 2009). Serotonergic neurons densely innervate the anterior cingulate cortex (ACC), ventromedial prefrontal cortex (VMPFC), and the amygdala (Hensler, 2006), the key brain circuits involved in resolving interference (Carter et al., 1999) as well as integrating emotional and cognitive influences on behavior (Barbas, 2000; Bechara et al., 2000). Importantly, the serotonin transporter gene (*SLC6A4*) contains a well-studied promoter polymorphism (5-HTT-linked polymorphic region, or 5-HTTLPR; Heils et al., 1996). The short (S) allele, consisting of 14 repeats, has been associated with decreased transporter expression

and decreased 5-HT uptake *in vitro*, compared to the long (L) allele with 16 repeats (Heils et al., 1996; Lesch et al., 1996). In addition, an A → G single nucleotide polymorphism (SNP) within the 5-HTTLPR (rs25531) produces L<sub>A</sub> and L<sub>G</sub> alleles, with the L<sub>G</sub> allele being functionally equivalent to the S allele (Hu et al., 2006). With respect to emotional and stressor reactivity, the S allele has been associated with higher measures of anxiety-related personality traits such as neuroticism (Lesch et al., 1996; Sen et al., 2004) and with an increased attentional bias to negative emotional stimuli such as images of spiders (Osinsky et al., 2008) relative to the L allele. The S allele has also been linked to a greater susceptibility to depression, depressive symptoms and suicide following adverse early-life experiences or stressful life events in adulthood (Caspi et al., 2003; Eley et al., 2004; Kendler et al., 2005; Taylor et al., 2006; Zalsman et al., 2006), findings supported by a recent meta-analysis (Karg et al., 2011, although see Risch et al., 2009). Converging evidence from neuroimaging studies shows that the S or L<sub>G</sub> allele carriers display a heightened amygdala response to threat stimuli (Hariri et al., 2002, 2005; Dannlowski et al., 2007, 2010; Munafo et al., 2008) and an increased functional connectivity between the amygdala and VMPFC during the processing of threat stimuli (Heinz et al., 2005; Pezawas et al., 2005; Friedel et al., 2009), relative to the L/L or L<sub>A</sub>/L<sub>A</sub> group.

Growing evidence also suggests that the 5-HTTLPR/rs25531 modulation extends to cognitive processes (Homberg and Lesch, 2010). Although improved cognitive function in the S or L<sub>G</sub> allele carriers relative to L/L or L<sub>A</sub>/L<sub>A</sub> homozygotes has also been reported (Roiser et al., 2007; Borg et al., 2009), a majority of studies have shown that the S or L<sub>G</sub> allele is associated with a relative impairment in cognitive task performance relative to the L or L<sub>A</sub> allele (da Rocha et al., 2008; Holmes et al., 2010), including dose effects of the L<sub>G</sub> allele on disadvantageous choices in the Iowa Gambling Task (Homberg et al., 2008) and on impulsive responding in the Continuous Performance Task (Walderhaug et al., 2010, although see Lage et al., 2011). Studies of 5-HTTLPR/rs25531 modulation of cognitive interference effects remain few in number. Using a simple flanker interference task, one group (Holmes et al., 2010) reported altered post-error behavioral adjustments in the S or L<sub>G</sub> carriers relative to the L<sub>A</sub>/L<sub>A</sub> group, while another larger study (Olvet et al., 2010) found no effect of 5-HTTLPR/rs25531 genotype on task performance. However, both studies may have been hindered by ceiling effects in task performance, making subtle genetic effects difficult to detect.

In the current study, we employed a novel and demanding threat-distracter version of the Multi-Source Interference Task (MSIT; Bush and Shin, 2006) in healthy females genotyped for the 5-HTTLPR/rs25531 promoter polymorphism, in order to examine the impact of threat-related distracters and 5-HTTLPR/rs25531 genotype on cognitive task performance. Based on previous studies (Vuilleumier et al., 2001; Dolcos and McCarthy, 2006; Blair et al., 2007; Mitchell et al., 2008, although see Bar-Haim et al., 2007), we hypothesized that threat distracters would potentiate interference effects relative to other distracter types and relative to a no-distracter condition. With respect to genetic effects, the simplest model is that functional variants affect gene transcription and protein function in a dose-dependent manner, without dominance, and this model is supported by some evidence

for additive effects of the  $SL_G$  allele on cognitive task performance (Homberg et al., 2008; Walderhaug et al., 2010) as well as on reactivity to environmental adversity (Caspi et al., 2003). Although non-additive effects have also been reported (Kendler et al., 2005), these reports have not been consistent and may be due to ceiling effects in measurement. Therefore, we expected that the  $SL_G$  allele of 5-HTTLPR/rs25531 would increase interference effects in a dose-dependent or additive manner, such that the effect of genotype on interference would follow a specific order:  $L_A/L_A < L_A/SL_G < SL_G/SL_G$ . We further tested two competing hypotheses about the scope of 5-HTTLPR/rs25531 effects on cognitive task performance. Specifically, genetic effects could be present exclusively in the threat-distracter condition, or alternatively, genetic effects could extend to all distracter conditions, irrespective of emotional salience of distracters. We also tested whether the effects of threat distracters change over the time-course of the task, and whether these effects are modulated by task difficulty. We expected that threat distracter effects would decrease over time due to habituation and regulation processes, and that the effects of threat distracters would be greater in the more difficult incongruent task condition compared to the easier congruent task condition.

## MATERIALS AND METHODS

### SUBJECTS

Seventy-one healthy, right-handed Caucasian females aged 18–34 years ( $M = 23.0$  years,  $SD = 4.0$  years) participated in the study. All subjects had normal or corrected-to-normal vision. Exclusion criteria included any serious medical condition, head injury or trauma, lifetime diagnosis of psychiatric illness, current use of a psychoactive medication, and smoking. Only females were studied at this stage, in order to maximize the power to detect genetic modulation of threat-distracter effects in light of prior evidence of interactions between sex hormones and serotonin transporter gene variation on threat reactivity (Josephs et al., 2012), as well as sex differences in the serotonin system (Jovanovic et al., 2008) and in the processing of emotional stimuli in the brain (Klein et al., 2003; Wrase et al., 2003). The study was approved by the University of Michigan Medical School IRB and all subjects provided written informed consent.

### TASK: THREAT-DISTRACTER MSIT

We employed a modified version of the MSIT (Bush et al., 2003; Bush and Shin, 2006). The MSIT is a validated response-interference paradigm which combines the sources of interference from Erikson, Stroop, and Simon tasks, in order to maximally tax the interference processing associated with the ACC (Bush et al., 2003). The MSIT has been shown to produce a robust and temporally stable *interference effect* both in reaction times (RTs) and in accuracy (Bush et al., 2003).

In the MSIT, subjects were presented with a set of three numbers from 0 to 3, one of which was different from the other two (the oddball number). Subjects were instructed to indicate the identity of the oddball number with a corresponding key press: a key press with the index finger if the oddball number was “1,” with the middle finger if the oddball number was “2,” and with the ring finger if the oddball number was “3.” On *congruent* trials, the identity of

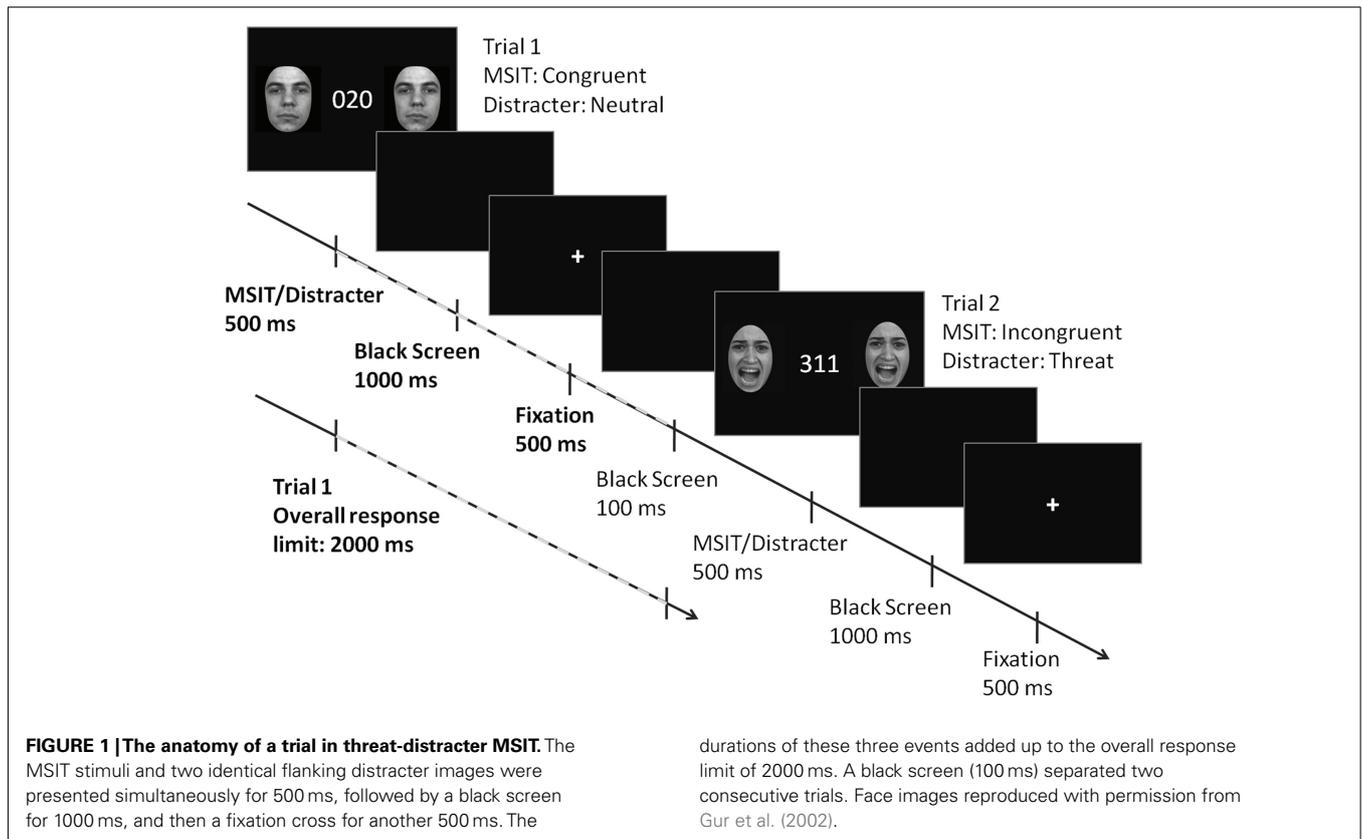
the oddball number corresponds to its location and the other two numbers are 0’s, not related to any valid key press response. On *incongruent* trials, the identity of the oddball number is incongruent with its position and the other two numbers are related to competing key press responses, resulting in stimulus-response incompatibility and response interference. The *incongruent condition* vs. *congruent condition* contrast yields the interference effect in RTs (*Incongruent RT* – *Congruent RT*) and interference effect in accuracy (*Congruent Accuracy* – *Incongruent Accuracy*).

We modified the MSIT to include three categories of task-irrelevant flanker distracters, threat, neutral, and scrambled, in addition to the null distracter condition. Threat distracters were images of human faces signaling the presence of a threat (angry or fearful expression). To isolate the effects specific to emotionally salient stimuli, we included neutral distracters (images of human faces with neutral expression), and scrambled distracters (images retaining the basic oval shape of a face but no facial features). Face stimuli were carefully selected from standardized sets (Ekman and Friesen, 1976; Gur et al., 2002; Tottenham et al., 2009). Angry and fearful faces displayed intense emotion and showed bared teeth and/or open mouth as an additional perceptual homogeneity criterion. In contrast, all neutral faces had closed mouths. All faces were Caucasian, to optimally control for potential sources of variability in emotional responses. All images were presented in grayscale, with hair and background cropped to yield an oval shape. Scrambled distracters were generated from the human face stimuli used in the other two distracter conditions by randomly rearranging the pixels within the oval while preserving the brightness of the image.

### EXPERIMENTAL PROTOCOL

A timeline of events in a single trial is shown in **Figure 1**. The MSIT stimuli and two identical flanking distracter images were presented simultaneously for 500 ms, followed by a black screen for 1000 ms, and then a fixation cross for another 500 ms. The durations of these three events added up to the overall response limit of 2000 ms. A black screen presented for 100 ms separated two consecutive trials. Subjects were instructed to respond as fast and as accurately as they could. The task stimuli were presented and the key press responses collected using E-Prime 2.0.

After a self-timed tutorial in the task and a short practice run, subjects completed a total of 640 trials, divided into 2 runs, four blocks per run, 80 trials per block. A short intermission separated run 1 (blocks 1–4, a total of 320 trials) from run 2 (blocks 5–8, a total of 320 trials). The order of the trials was pseudo-randomized within each block, with the provision that no two consecutive trials (1) had the same correct response or (2) both included threat distracters. Each block lasted approximately 3 min and consisted of 40 congruent and 40 incongruent trials. Within the sets of 40 congruent and 40 incongruent trials, 10 trials included threat distracters (five angry faces, three female, two male or two female, three male; and five fearful faces, three female, two male or two female, three male), 10 trials included neutral distracters (five female, five male), 10 trials included scrambled distracters, and 10 trials were no-distracter trials (i.e., with MSIT stimuli only). The whole experiment lasted approximately 30 min.



**GENOTYPING OF 5-HTTLPR/rs25531**

Genomic DNA was obtained from saliva using the Oragene saliva collection system and extracted using the protocol provided (Genotek, Ontario, Canada). The extracted DNA samples were genotyped for 5-HTTLPR and rs25531 in two steps, according to Wendland et al. (2006). In the first step, the 5-HTTLPR was amplified via polymerase-chain reaction (PCR) using site-specific forward and reverse primers, yielding “short” (14-repeat, 375 bp) and “long” (16-repeat, 419 bp) products. In the second step, the PCR product from the first step was digested with *Hpa* II restriction enzyme to genotype the A → G SNP (rs25531) by identifying L<sub>G</sub> (305 bp) and L<sub>A</sub> alleles. All PCR products were visualized via gel electrophoresis on a 3% agarose gel using ethidium bromide under ultraviolet (UV) light.

**STATISTICAL ANALYSES**

The data were analyzed in a series of steps using repeated-measures Analysis of Variance (ANOVA), correlations, and *t*-tests as implemented in SPSS 19.0. We used two behavioral indices of task performance as dependent variables, RTs on correct trials and accuracy rates. The MSIT interference effects (congruent vs. incongruent) in RTs and in accuracy were used as a global measure of the efficiency of interference processing, with greater interference effects indicating less efficient interference resolution. We conducted two separate 4 × 2 × 3 repeated-measures ANOVAs – one on interference effects in accuracy and one on interference effects in RTs – with distracter type (four levels: threat-related, neutral, scrambled, or null) and run (two levels: pre-intermission

run 1 or post-intermission run 2) as within-subject factors, and 5-HTTLPR/rs25531 genotype (three levels: 0 S<sub>L<sub>G</sub></sub> alleles, 1 S<sub>L<sub>G</sub></sub> alleles, or 2 S<sub>L<sub>G</sub></sub> alleles) as a between-subject factor. Because we conducted two separate ANOVAs, we used a Bonferroni-corrected *p* value of 0.025 as our statistical threshold for the ANOVA results. The *t*-tests and Pearson’s correlations are two-tailed unless stated otherwise.

**RESULTS**

**FINAL SAMPLE**

Out of the 71 healthy female subjects who participated in the study, the data from the final sample of 69 subjects were analyzed and are reported below. The data from two subjects were excluded from analysis due to concerns about task compliance and performance accuracy. One subject did not follow the task instructions and responded to the position of the oddball number rather than to its identity (*M* = 0.05 accuracy on incongruent trials), an occurrence reported in approximately 5% of participants in prior work using the original version of the MSIT (Bush and Shin, 2006). Another subject had a mean accuracy of 0.34 on incongruent trials, corresponding to a chance level of responding in a three-choice task.

**GENOTYPING RESULTS**

We observed the following 5-HTTLPR genotype counts (and frequencies): 25 (0.35) L/L homozygotes, 35 (0.49) L/S heterozygotes, and 11 (0.16) S/S homozygotes (Table 1). The observed genotype frequencies did not deviate from the

**Table 1 | Distribution of 5-HTTLPR and 5-HTTLPR/rs25531 alleles and genotypes.**

5-HTTLPR genotype count (frequency)			5-HTTLPR allele count (frequency)		
L/L	L/S	S/S	L	S	
25 (0.35)	35 (0.49)	11 (0.16)	85 (0.60)	57 (0.40)	

5-HTTLPR/rs25531 genotype count (frequency)						5-HTTLPR/rs25531 allele count (frequency)		
Func L/L	Func L/S		Func S/S			Func L	Func S	
23 (0.32)	36 (0.51)		12 (0.17)			82 (0.58)	60 (0.42)	

L <sub>A</sub> /L <sub>A</sub>	L <sub>A</sub> /L <sub>G</sub>	L <sub>A</sub> /S	L <sub>G</sub> /L <sub>G</sub>	L <sub>G</sub> /S	S/S	L <sub>A</sub>	L <sub>G</sub>	S
23 (0.32)	2 (0.03)	34 (0.48)	0	1 (0.01)	11 (0.16)	82 (0.58)	3 (0.02)	57 (0.40)

S allele and L<sub>G</sub> allele are denoted as functional S alleles.

Hardy–Weinberg Equilibrium ( $\chi^2 = 0.047, p = 0.828$ ). The combined 5-HTTLPR/rs25531 functional genotypes were grouped as follows: 23 (0.32) subjects were L<sub>A</sub>/L<sub>A</sub>, 36 (0.51) subjects were L<sub>A</sub>/L<sub>G</sub>S (2 L<sub>A</sub>/L<sub>G</sub> and 34 L<sub>A</sub>/S<sub>A</sub>), and 12 (0.17) subjects were S/S (1 L<sub>G</sub>/S and 11 S/S). S<sub>L<sub>G</sub></sub> denoted S or L<sub>G</sub> allele (Table 1). Neither the 5-HTTLPR genotype groups nor the 5-HTTLPR/rs25531 genotype groups differed in age, education, or socio-economic status (Table 2).

**BEHAVIORAL RESULTS**

**Robust MSIT interference effects across all distracter conditions**

Consistent with previous reports (Bush et al., 2003; Bush and Shin, 2006), we observed a robust and highly significant MSIT interference effect (i.e., a main effect of congruency) in both measures of task performance. Overall, subjects were significantly less accurate in the incongruent condition compared to the congruent condition (congruent accuracy,  $M = 0.993, SE = 0.001$ ; incongruent accuracy,  $M = 0.838, SE = 0.016$ ; interference effect in accuracy,  $M = 0.158, SE = 0.015; F(1, 66) = 107.290, p < 0.0001$ , partial eta squared = 0.619), and they were also significantly slower to correctly respond in the incongruent condition compared to the congruent condition (congruent RT,  $M = 492\text{ ms}, SE = 11\text{ ms}$ ; incongruent RT,  $M = 710\text{ ms}, SE = 16\text{ ms}$ ; interference effect in RT,  $M = 218\text{ ms}, SE = 9\text{ ms}; F(1, 66) = 579.179, p < 0.0001$ , partial eta squared = 0.898).

The interference effects were robust and highly significant in all four distracter conditions (all  $p$ 's < 0.0001, paired-sample  $t$ -tests). The accuracy results per distracter condition are summarized in Table 3 and the RT results per distracter condition are summarized in Table 4. In addition, the interference effect on accuracy was significant in both runs (run 1,  $M = 0.192, SE = 0.017; t(68) = 11.077, p < 0.0001$ ; run 2,  $M = 0.124, SE = 0.013; t(68) = 9.993, p < 0.0001$ ), although it significantly diminished from run 1 to run 2,  $t(68) = 7.319, p < 0.0001$ , as also indicated by a significant two-way interaction between congruency and run on accuracy,  $F(1, 66) = 72.882, p < 0.0001$ , partial eta squared = 0.525. The interference effect in RTs was also significant in both runs (run 1,  $M = 221\text{ ms}, SE = 9\text{ ms}; t(68) = 26.795, p < 0.0001$ ;

**Table 2 | Demographic profiles of the 5-HTTLPR and 5-HTTLPR/rs25531 genotype groups.**

	S/S (n = 11)	S/L (n = 33)	L/L (n = 25)	$\chi^2$ (p value)
<b>5-HTTLPR GENOTYPE</b>				
Age (years)	22.36 ± 3.50	22.39 ± 4.10	24.08 ± 4.18	19.97 (0.793)
Education (years)	15.64 ± 2.20	15.55 ± 2.60	15.96 ± 1.93	19.51 (0.361)
SES	2.18 ± 0.60	2.30 ± 0.53	2.24 ± 0.44	6.56 (0.363)
	S <sub>L<sub>G</sub></sub> /S <sub>L<sub>G</sub></sub> (n = 12)	S <sub>L<sub>G</sub></sub> /L <sub>A</sub> (n = 34)	L <sub>A</sub> /L <sub>A</sub> (n = 23)	$\chi^2$ (p value)
<b>5-HTTLPR/rs25531 GENOTYPE</b>				
Age (years)	22.17 ± 3.41	22.38 ± 4.02	24.35 ± 4.25	17.67 (0.887)
Education (years)	15.50 ± 2.15	15.56 ± 2.56	16.04 ± 2.00	18.64 (0.415)
SES	2.17 ± 0.58	2.29 ± 0.52	2.26 ± 0.45	5.88 (0.436)

Means and standard deviations are given. No group differences in age, education, or socio-economic status (SES) were found, as assessed with a chi-square ( $\chi^2$ ) test.

run 2,  $M = 216\text{ ms}, SE = 9\text{ ms}; t(68) = 25.463, p < 0.0001$ ), and did not change significantly from run 1 to run 2,  $t(68) = 1.496, p = 0.139$ . These results confirmed that MSIT produced a robust behavioral difference between the easier congruent condition and the more difficult incongruent condition, which persisted across all distracter conditions and across time.

**Threat distracters potentiate MSIT interference effects**

Next, we examined whether threat-related distracters potentiated MSIT interference effects. As hypothesized, the ANOVA on interference effects yielded robust and significant main effects of distracter type on interference effects both in accuracy,  $F(3, 64) = 7.803, p < 0.0001$ , partial eta squared = 0.268, and

**Table 3 | Summary of accuracy data.**

Distracter type	Accuracy (proportion accurate)				
	MSIT condition		MSIT interference effect		
	Congruent	Incongruent	Mean	<i>t</i>	<i>p</i> value
Threat	0.995 (0.013)	0.839 (0.121)	0.156 (0.117)	11.002	<0.0001
Neutral	0.993 (0.014)	0.844 (0.126)	0.149 (0.121)	10.297	<0.0001
Scrambled	0.996 (0.009)	0.834 (0.125)	0.161 (0.120)	11.193	<0.0001
Null	0.990 (0.015)	0.856 (0.117)	0.134 (0.110)	10.177	<0.0001

Means and standard deviations (in parentheses) are given, together with *t* statistics and *p* values for paired-sample *t*-tests (*n* = 69).

**Table 4 | Summary of RT data.**

Distracter type	RT (ms)				
	MSIT condition		MSIT interference effect		
	Congruent	Incongruent	Mean	<i>t</i>	<i>p</i> value
Threat	486 (82)	710 (116)	224 (72)	26.048	<0.0001
Neutral	489 (81)	711 (118)	222 (70)	26.272	<0.0001
Scrambled	489 (87)	714 (117)	225 (71)	26.236	<0.0001
Null	495 (84)	701 (116)	205 (64)	26.781	<0.0001

Means and standard deviations (in parentheses) are given, together with *t* statistics and *p* values for paired-sample *t*-tests (*n* = 69).

in RTs,  $F(3, 64) = 6.309$ ,  $p = 0.001$ , partial eta squared = 0.228. Convergent results were obtained from the ANOVA on accuracy and RTs, which indicated a significant two-way interaction between congruency and distracter type both on accuracy,  $F(3, 64) = 6.465$ ,  $p = 0.001$ , partial eta squared = 0.233, and on RTs,  $F(3, 64) = 8.030$ ,  $p < 0.0001$ , partial eta squared = 0.273. The overall interference effects in accuracy per distracter condition are given in **Table 3** and the overall interference effects in RTs per distracter condition are given in **Table 4**. The interference effects in accuracy in the threat-distracter condition were significantly greater than in the no-distracter condition,  $t(68) = 3.415$ ,  $p = 0.001$ , but not significantly greater than in the neutral-distracter condition,  $t(68) = 0.964$ ,  $p = 0.338$ , or in the scrambled-distracter condition,  $t(68) = 1.017$ ,  $p = 0.313$ . Similarly, the interference effects in RTs were significantly greater with threat distracters present compared to with no distracters present,  $t(68) = 6.308$ ,  $p < 0.0001$ , but not significantly different compared to neutral distracters,  $t(68) = 0.710$ ,  $p = 0.480$ , or scrambled distracters,  $t(68) = 0.211$ ,  $p = 0.833$ . Overall, interference effects in accuracy were significantly greater in the presence of distracters compared to the no-distracter condition (with distracters:  $M = 0.155$ ,  $SE = 0.014$ ; no distracters:  $M = 0.134$ ,  $SE = 0.013$ ;  $t(68) = 4.056$ ,  $p < 0.0001$ ). Similarly, interference effects in RTs were significantly greater in the presence of distracters compared to the no-distracter condition (with distracters:  $M = 220$  ms,  $SE = 8$  ms; no distracters:  $M = 205$  ms,  $SE = 8$  ms;  $t(68) = 5.390$ ,  $p < 0.0001$ ).

#### **Threat-distracter effects on MSIT interference effects are transient**

Overall, there was a robust and highly significant main effect of run both on accuracy [ $F(1, 66) = 68.309$ ,  $p < 0.0001$ , partial eta squared = 0.509] and on RTs [ $F(1, 66) = 104.982$ ,  $p < 0.0001$ , partial eta squared = 0.614]. The overall accuracy in run 1 was  $M = 0.903$ ,  $SE = 0.009$ , whereas in run 2 it significantly increased to  $M = 0.936$ ,  $SE = 0.006$ ,  $t(68) = 7.249$ ,  $p < 0.0001$ . The overall RT in run 1 was  $M = 625$  ms,  $SE = 13$  ms, whereas in run 2 it significantly decreased to  $M = 574$  ms,  $SE = 10$  ms,  $t(68) = 11.708$ ,  $p < 0.0001$ . In addition, there was a significant two-way interaction between distracter type and run on interference effects in accuracy,  $F(3, 64) = 4.290$ ,  $p = 0.008$ , partial eta squared = 0.167, and in RTs,  $F(3, 64) = 11.932$ ,  $p < 0.0001$ , partial eta squared = 0.359. These data are summarized in **Table 5** (accuracy) and **Table 6** (RTs) and graphically shown in **Figure 2A** (accuracy) and **Figure 2B** (RTs).

We also examined how the effects of threat distracters on MSIT interference effects changed over time. In run 1, threat distracters potentiated the interference effects in accuracy relative to neutral distracters,  $t(68) = 3.03$ ,  $p = 0.004$ , scrambled distracters,  $t(68) = 1.74$ ,  $p = 0.09$ , and no distracters,  $t(68) = 3.73$ ,  $p < 0.0001$  (**Figure 2A**). In contrast, in run 2 (following the intermission), the interference effects in accuracy elicited by threat distracters appeared to be lower than those elicited by neutral distracters,  $t(68) = -1.78$ ,  $p = 0.08$ , or scrambled distracters,  $t(68) = -3.24$ ,  $p = 0.002$ , and comparable to the interference effects observed in the no-distracter condition. Interestingly, examining congruent and incongruent trials separately revealed that threat distracters had dissociable and opposite effects on

**Table 5 | Summary of accuracy data (in proportion accurate) in run 1 and run 2.**

Distracter type	Run 1			Run 2		
	MSIT condition		MSIT interference effect	MSIT condition		MSIT interference effect
	Congruent	Incongruent		Congruent	Incongruent	
Threat	0.996 (0.002)	0.788 (0.021)	0.213 (0.020)	0.994 (0.002)	0.884 (0.013)	0.113 (0.013)
Neutral	0.991 (0.002)	0.808 (0.020)	0.184 (0.019)	0.996 (0.002)	0.865 (0.015)	0.136 (0.015)
Scrambled	0.993 (0.002)	0.798 (0.019)	0.196 (0.018)	0.997 (0.001)	0.860 (0.016)	0.141 (0.015)
Null	0.982 (0.004)	0.809 (0.020)	0.173 (0.018)	0.997 (0.001)	0.895 (0.014)	0.103 (0.014)

Means and standard errors (in parentheses) are given.

**Table 6 | Summary of RT data (in ms) in run 1 and run 2.**

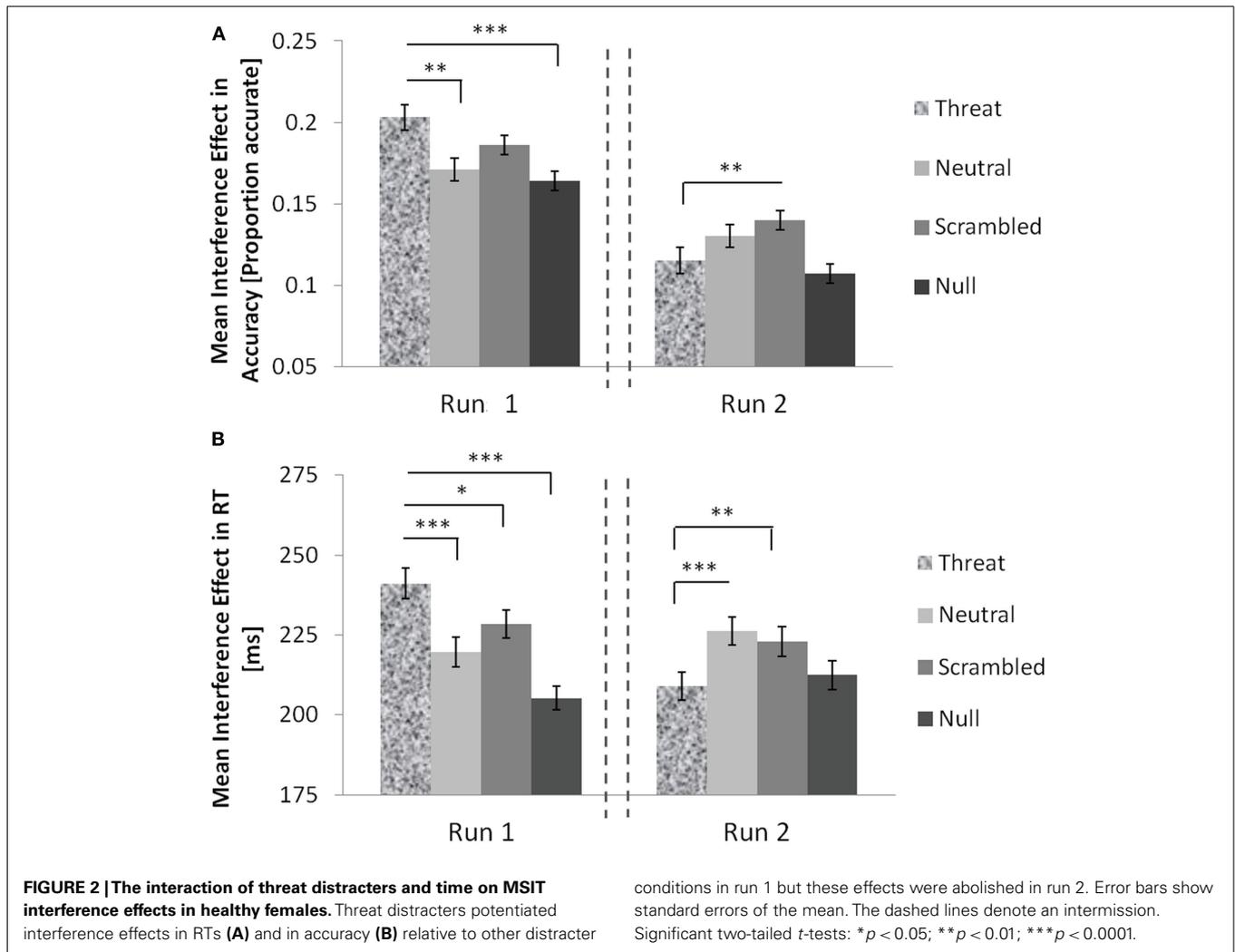
Distracter type	Run 1			Run 2		
	MSIT condition		MSIT interference effect	MSIT condition		MSIT interference effect
	Congruent	Incongruent		Congruent	Incongruent	
Threat	502 (12)	739 (18)	238 (11)	474 (10)	682 (14)	208 (10)
Neutral	516 (13)	734 (17)	217 (10)	465 (9)	689 (15)	224 (10)
Scrambled	513 (13)	737 (17)	224 (10)	471 (10)	689 (15)	219 (10)
Null	528 (13)	733 (17)	204 (8)	682 (14)	674 (15)	211 (10)

Means and standard errors (in parentheses) are given.

accuracy in congruent and incongruent trials across time. As expected, in run 1, subjects were less accurate on the more difficult incongruent trials in the presence of threat distracters than in the presence of neutral distracters,  $t(68) = -2.231$ ,  $p = 0.029$ , or null distracters,  $t(68) = -2.379$ ,  $p = 0.020$ , although not relative to scrambled distracters,  $t(68) = -1.203$ ,  $p = 0.233$ . However, this relationship was reversed in run 2, and subjects appeared more accurate on incongruent trials with threat distracters relative to neutral distracters,  $t(68) = 1.615$ ,  $p = 0.111$ , or scrambled distracters,  $t(68) = 3.010$ ,  $p = 0.004$ , although not different in accuracy compared to incongruent trials with no distracters present,  $t(68) = -0.967$ ,  $p = 0.337$ . In addition, and unexpectedly, in run 1, subjects were actually more accurate on the easy congruent trials in the presence of threat distracters relative to neutral distracters,  $t(68) = 2.013$ ,  $p = 0.048$ , and relative to no distracters,  $t(68) = 3.570$ ,  $p = 0.001$ , although not relative to scrambled distracters,  $t(68) = 0.479$ ,  $p = 0.638$ . In run 2, these apparent performance-enhancing effects of threat distracters were abolished, and subjects' accuracy on congruent trials in the presence of threat distracters did not significantly differ from their accuracy in the presence of neutral distracters,  $t(68) = -0.397$ ,  $p = 0.693$ , scrambled distracters,  $t(68) = -1.413$ ,  $p = 0.162$ , or no distracters,  $t(68) = -1.383$ ,  $p = 0.171$ .

The results were similar for RTs (Figure 2B). In run 1, threat distracters potentiated the interference effects in RTs relative to neutral distracters,  $t(68) = 4.31$ ,  $p < 0.0001$ , scrambled distracters,  $t(68) = 2.38$ ,  $p = 0.020$ , and no distracters,  $t(68) = 7.36$ ,  $p < 0.0001$ . In contrast, in run 2 (following the intermission),

the interference effects in RTs observed in the threat-distracter condition were lower than in the presence of neutral distracters,  $t(68) = -3.87$ ,  $p < 0.0001$ , or scrambled distracters,  $t(68) = -3.28$ ,  $p = 0.002$ , and comparable to the no-distracter condition. As described above for accuracy, threat distracters appeared to have dissociable and opposite effects on the speed of correct responses in congruent and incongruent trials across time. As might be expected, in run 1, subjects were somewhat slower to correctly respond on the more difficult incongruent trials in the presence of threat distracters than in the presence of neutral distracters,  $t(68) = 1.626$ ,  $p = 0.108$ , or no distracters,  $t(68) = 2.595$ ,  $p = 0.012$ , although not relative to scrambled distracters,  $t(68) = 0.407$ ,  $p = 0.685$ . This relationship was reversed in run 2, in which subjects were somewhat faster to correctly respond on incongruent trials with threat distracters relative to neutral distracters,  $t(68) = -1.987$ ,  $p = 0.051$ , or scrambled distracters,  $t(68) = -2.776$ ,  $p = 0.007$ , although still somewhat slower to correctly respond than on incongruent trials with no distracters present,  $t(68) = 1.847$ ,  $p = 0.069$ . In addition, and again unexpectedly, in run 1, subjects were actually faster to accurately respond on the easy congruent trials in the presence of threat distracters relative to neutral distracters,  $t(68) = -5.702$ ,  $p < 0.0001$ , scrambled distracters,  $t(68) = -3.848$ ,  $p < 0.0001$ , or no distracters,  $t(68) = -8.615$ ,  $p < 0.0001$ . This performance-enhancing effect of threat distracters was again transient, as seen above for accuracy. In run 2, the relationship was reversed and subjects were slower to correctly respond on congruent trials with threat distracters relative to neutral distracters,  $t(68) = 4.482$ ,  $p < 0.0001$ ,



scrambled distracters,  $t(68) = 1.613, p = 0.111$ , or no distracters,  $t(68) = 5.925, p < 0.0001$ .

In sum, threat distracters increased the interference effect in accuracy and in RTs compared with neutral or scrambled distracters in the first half of the experiment, but these effects were reversed in the second half, following an intermission. In addition, this transient increase in interference effects in the presence of threat distracters was driven both by a threat-distracter-related impairment in performance on the more difficult incongruent trials, and, unexpectedly, by a threat-distracter-related enhancement in performance on the easy congruent trials.

**5-HTTLPR/rs25531 genotype modulates interference effects irrespective of emotional salience of distracters**

Next, we tested whether the 5-HTTLPR/rs25531 genotype modulated the impact of threat-related distracters on cognitive task performance. Collapsing across both runs and across distracter conditions, genotype did not have a significant effect on interference effects either in accuracy,  $F(2, 66) = 0.983, p = 0.379$ , or in RTs.  $F(2, 66) = 0.399, p = 0.673$ . But there was a significant two-way interaction between genotype and run on interference effects in

accuracy,  $F(2, 66) = 5.111, p = 0.009$ , partial eta squared = 0.134. These results were confirmed by the ANOVA on accuracy, which produced a significant two-way interaction between genotype and run on accuracy,  $F(2, 66) = 4.082, p = 0.021$ , partial eta squared = 0.110.

Specifically, there was an increase in interference effects in accuracy with the number of the SL<sub>G</sub> alleles, which was significant in run 1 (L<sub>A</sub>/L<sub>A</sub>:  $0.156 \pm 0.027$ ; SL<sub>G</sub>/L<sub>A</sub>:  $0.176 \pm 0.021$ ; SL<sub>G</sub>/SL<sub>G</sub>:  $0.243 \pm 0.046$ ;  $r = 0.207, p = 0.044$ , one-tailed correlation) but did not reach significance in run 2 (L<sub>A</sub>/L<sub>A</sub>:  $0.107 \pm 0.021$ ; SL<sub>G</sub>/L<sub>A</sub>:  $0.130 \pm 0.016$ ; SL<sub>G</sub>/SL<sub>G</sub>:  $0.133 \pm 0.036$ ;  $r = 0.103, p = 0.201$ , one-tailed correlation). A comparison of the 5-HTTLPR/rs25531 genotype groups on interference effects in accuracy separately for each distracter condition is given in Figure 3. The increase in interference effects in accuracy with the number of the SL<sub>G</sub> alleles was also significant or marginally significant in all four distracter conditions in run1 (threat:  $r = 0.195, p = 0.054$ ; neutral:  $r = 0.170, p = 0.082$ ; scrambled:  $r = 0.192, p = 0.057$ ; null:  $r = 0.218, p = 0.036$ ; all one-tailed correlations).

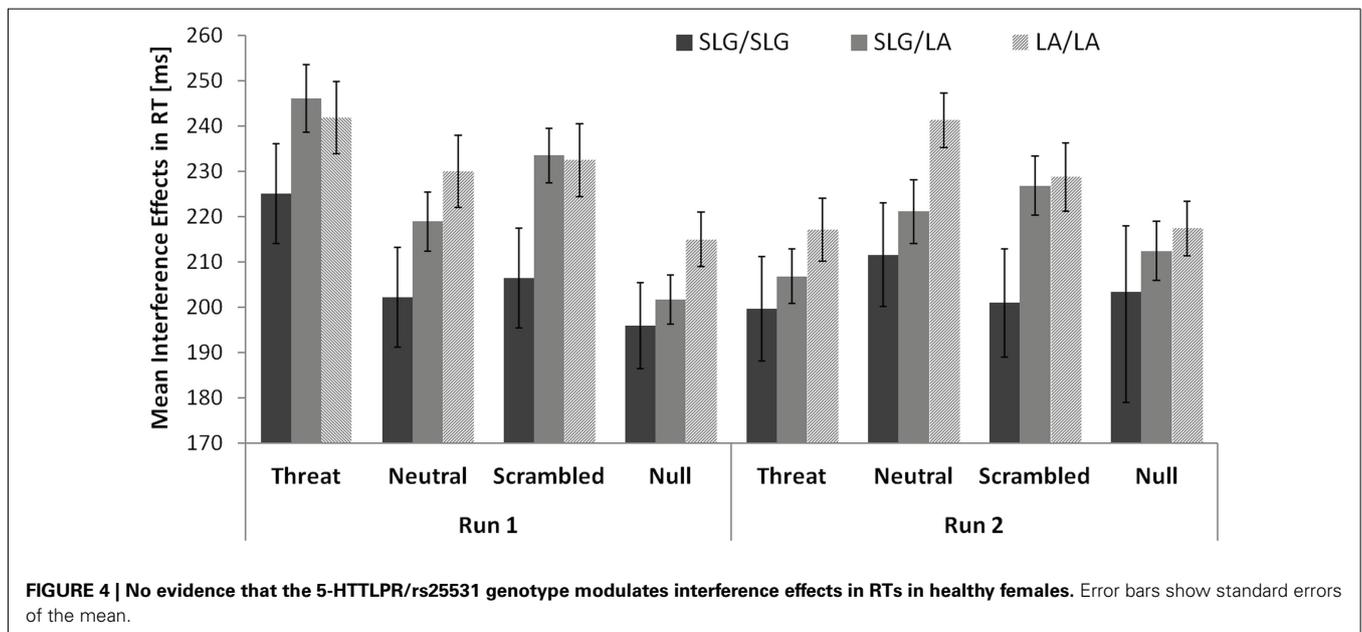
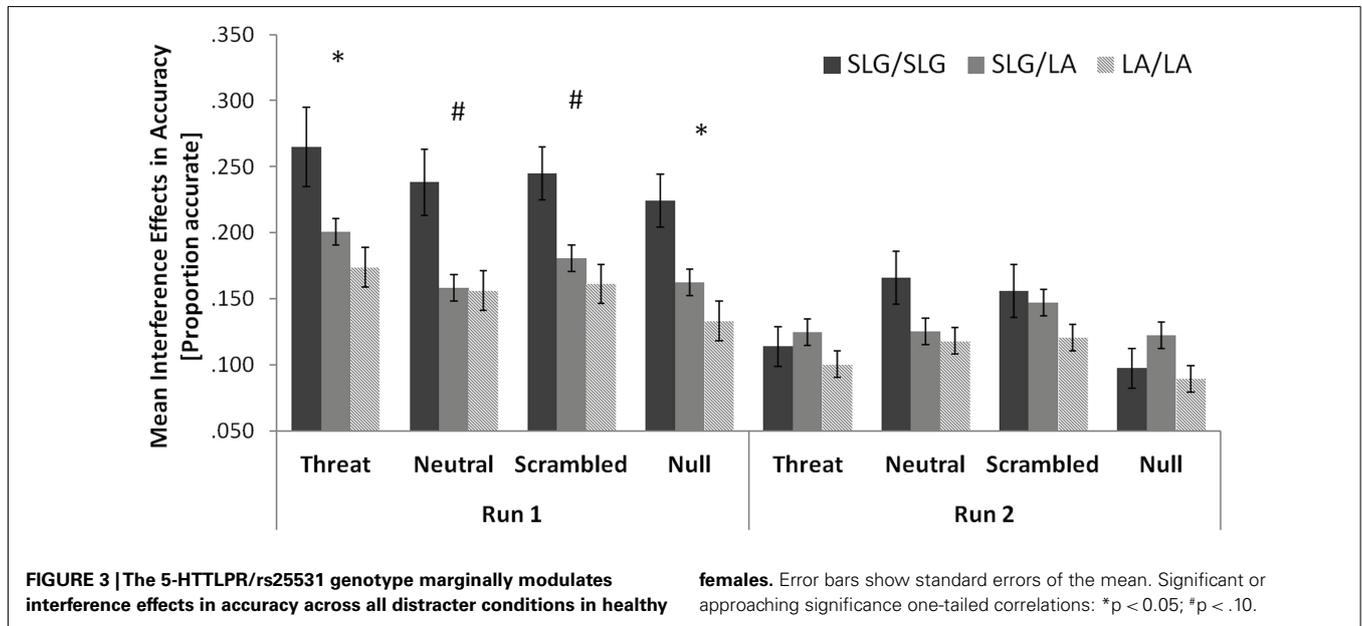
There were no comparable effects of genotype on interference effects in RTs. The magnitude of interference effects in RTs was

not significantly associated with the number of *SLG* alleles either in run 1 ( $L_A/L_A$ :  $230 \pm 14$  ms;  $SL_G/L_A$ :  $225 \pm 12.2$  ms;  $SL_G/SL_G$ :  $207 \pm 20$  ms;  $r = -0.103$ ,  $p = 0.201$ , one-tailed correlation) or in run 2 ( $L_A/L_A$ :  $226 \pm 13$  ms;  $SL_G/L_A$ :  $217 \pm 13$  ms;  $SL_G/SL_G$ :  $204 \pm 24$  ms;  $r = -0.107$ ,  $p = 0.192$ , one-tailed correlation). A comparison of the 5-HTTLPR/rs25531 genotype groups on interference effects in RTs separately for each distracter condition is given in Figure 4.

**DISCUSSION**

Our data demonstrate that threat-related distracters robustly modulate cognitive interference effects but the modulation dynamically changes over time. Threat-related distracters

potentiated interference effects in both accuracy and in RTs relative to non-threat-related distracter types and relative to the no-distracter condition in the first half of the experiment, prior to the intermission. However, these effects were reversed in the second half of the experiment, in which the interference effects in accuracy and in RTs in the presence of threat distracters decreased below the interference effects seen in other distracter conditions, to the level observed when no distracters were present. Furthermore, by examining the congruent and incongruent conditions separately, we were able to show that this transient potentiation of interference effects by threat distracters had a dual source: on the one hand, it was due to a predicted threat-related impairment in task performance in the more difficult incongruent condition



(i.e., subjects were less accurate and slower to correctly respond on incongruent trials in the presence of threat distracters relative to other distracter conditions), but on the other hand, it was also due to an unexpected threat-related enhancement of task performance in the easy congruent condition (i.e., subjects were actually more accurate and faster to correctly respond on congruent trials in the presence of threat distracters compared to other distracter conditions).

We propose that the temporally dynamic character of threat-distracter effects may be due to both habituation and regulation of amygdala response to threat stimuli. Both habituation and regulation would result in diminished amygdala reactivity. Amygdala habituation to threat stimuli has been demonstrated in neuroimaging studies involving both healthy individuals (Breiter et al., 1996; Whalen et al., 1998; Wright et al., 2001) and patients with anxiety disorders such as post-traumatic stress disorder (Shin et al., 2005). A separate line of neuroimaging evidence also shows a decrease in amygdala response to threat-related stimuli when people actively regulate their emotional response using cognitive-control strategies such as reappraisal, distraction, or suppression (Ochsner et al., 2002; Phan et al., 2005; Eippert et al., 2007; Kim and Hamann, 2007; Wager et al., 2008; McRae et al., 2010), with convergent evidence coming from animal studies of fear extinction (Quirk and Beer, 2006; Hartley and Phelps, 2010). We propose that both processes – habituation and regulation of amygdala response to threat stimuli – may be at work in our study. Habituation may be gradually produced by repeated harmless presentation of threat stimuli over the time-course of the task, whereas regulation may be triggered specifically by the intermission separating run 1 from run 2, giving subjects a short reprise from the demands of the task and permitting them to “take stock” and adjust their emotional response to the threat stimuli in run 2. Unfortunately, we are unable to fully dissociate the role of these two processes in the observed decrease in threat-distracter effects on cognitive performance over time using the current study design.

An intriguing finding in our study is the dissociable and opposite character of threat effects on task performance in congruent vs. incongruent task conditions. The transient increase in interference effects in the presence of threat distracters was driven both by threat-distracter-related impairment in performance on the more difficult incongruent trials, and by threat-distracter-related enhancement in performance on the easier congruent trials. Threat-related impairment in task performance has been documented before (Vuilleumier et al., 2001; Dolcos and McCarthy, 2006; Blair et al., 2007; Mitchell et al., 2008), although the findings have been inconsistent (Bar-Haim et al., 2007). Our data suggest that the inconsistencies may come from variable level of task difficulty, with more robust threat-related impairment observed in more difficult task conditions requiring additional time and processing steps to resolve cognitive interference arising from competing stimulus-to-response goal representations, as compared to easier task conditions involving one simple stimulus-to-response mapping.

In this respect, our finding of threat-related enhancement of task performance specific to the easier congruent task condition is informative. We speculate that this threat-related enhancement of both accuracy and speed of correct responding in the easier

task condition may reflect a general priming of the motor system in response to threat signals. Our findings resonate with previous reports of enhanced response speed and force due to exposure to unpleasant stimuli during a preparation of a simple motor response (Coombes et al., 2005, 2009). Consistent with the adaptive function of rapid behavioral response to potential threat signals in the environment, threat-related stimuli may act to prime the motor system for action (Coombes et al., 2005) regardless of their status as task-relevant targets or task-irrelevant distracters. Therefore, both threat-related *enhancement* of task performance in the absence of cognitive interference (easier task condition) and threat-related *impairment* of task performance when the task requires resolution of cognitive interference (more difficult task condition) would reflect the priming of the simple, prepotent motor response – but the primed response itself would be correct in the former case and incorrect in the latter case. We further speculate that the impact of threat distracters on task performance may be mediated primarily through the effects of threat stimuli on the selection and execution of the motor response within broadly defined attentional control processes. Specifically, the detection of a potential threat signal and the subsequent activation of the threat-processing pathway could act either to directly facilitate the execution of the prepotent motor response, or to remove the inhibition of this prepotent response. In either case, performance would be expected to improve when the prepotent response is desired (e.g., in the easier congruent task condition), but suffer when the inhibition of a prepotent response is required for the selection and execution of a correct response (e.g., in the more difficult incongruent task condition). Thus, one possible strategy to reduce threat-related impairment may be to automatize the performance of a given task (i.e., to render the desired task response the prepotent response) through intense practice and habit formation, consistent with the theory of Norman and Shallice (1986).

We also report evidence that the serotonin transporter promoter polymorphism (5-HTTLPR/rs25531) modulates cognitive task performance in healthy female subjects in a global fashion, irrespective of the presence or emotional salience of distracters. Specifically, we observed dose effects of the  $SL_G$  allele on interference effects in accuracy (but not in RTs) in the expected direction:  $L_A/L_A$  interference effects <  $SL_G/L_A$  interference effects <  $SL_G/SL_G$  interference effects. In addition, the modulation of interference effects by 5-HTTLPR/rs25531 genotype was not specific to threat distracters, but instead extended to all four distracter conditions, including threat, neutral, scrambled, and no distracters. Furthermore, the genetic modulation of interference effects was observed exclusively in the first half of the experiment, prior to the intermission, and was abolished in the second half of the experiment.

This pattern of genetic results is particularly intriguing in light of the robust (if transient) potentiation of the interference effects by threat-related distracters observed in the whole sample, collapsing across genotypes. The pattern strongly suggests that the 5-HTTLPR/rs25531 genotype modulates susceptibility to cognitive interference in healthy females in general, rather than to cognitive interference produced specifically by threat-related distracters. In this respect, our results are broadly consistent with the view that the 5-HTTLPR genotype may affect susceptibility

to environmental influences in general rather than modulating specifically the impact of adverse stimuli (Uher, 2008; Belsky and Pluess, 2009), a trait described as *hypervigilance* (Homberg and Lesch, 2010). Thus, the S or L<sub>G</sub> allele is associated with worse behavioral and clinical outcomes in the context of adverse environmental conditions, such as childhood maltreatment or stressful life events, but it can also lead to more favorable outcomes in protective, nurturing environments, relative to the L allele (Caspi et al., 2003; Eley et al., 2004; Taylor et al., 2006). Indeed, Roiser et al. (2009) provided elegant evidence for such increased “framing effects” during decision-making, as well as for the corresponding changes in the amygdala-PFC circuitry, in S/S homozygotes compared to L<sub>A</sub>/L<sub>A</sub> homozygotes. Although the neurobiological mechanisms involved are likely to be highly complex and thus challenging to fully elucidate, we recently proposed one possible molecular mechanism underlying the interaction of stressors and 5-HTTLPR/rs25531 genotype on the amygdala-VMPFC-dorsal raphe nucleus circuitry and the risk of depression (Jasinska et al., 2012).

Some limitations of the current study should be acknowledged. Although our sample size was sufficiently large to give us high statistical power to detect main and interactive effects of the task, it was relatively small to detect genetic effects. The genetic effects in particular should therefore be considered preliminary until replicated in a larger independent sample. It will also be important to replicate the results in both sexes. Furthermore, cognitive function may also be modulated by other functional variants in the serotonin transporter gene (e.g., serotonin transporter intron 2 polymorphism, STin2; Payton et al., 2005; Sarosi et al., 2008), in other serotonergic genes (e.g., *TPH2*; Strobel et al., 2007), or in genes involved in gene-gene interactions with the serotonin transporter gene (e.g., *BDNF*), either in isolation or in interaction with the 5-HTTLPR/rs25531. These

effects were unmeasured in our study. Finally, the level of emotion regulation exerted by subjects while performing the task may also modulate performance on tasks which engage emotion-cognition interactions by altering the activity and functional connectivity within the amygdala-PFC circuitry, consistent with recent reports (Schardt et al., 2010; Enge et al., 2011; Lemogne et al., 2011). Therefore, an important goal of future studies will be to measure and manipulate emotion regulation, particularly with respect to serotonin transporter gene effects, to determine to what degree it alters task performance and can compensate for genetic vulnerability to threat reactivity and to cognitive interference.

In conclusion, using a novel threat-distracter MSIT, we demonstrated that threat distracters robustly but transiently potentiate cognitive interference effects, and that 5-HTTLPR/rs25531 genotype modulation of these cognitive interference effects extends to all distracter conditions, irrespective of emotional salience of distracters, in healthy female subjects. These results add to our understanding of the processes through which threat-related distracters affect cognitive processing, and have implications for our understanding of disorders in which threat signals have a detrimental effect on cognition, including depression and anxiety disorders.

## ACKNOWLEDGMENTS

We thank Ms. Ela Sliwerska for her generous help in carrying out the genetic portion of this study. This research was supported by the Rackham Graduate Student Research Award and the Center for the Education of Women Student Research Award, University of Michigan (Agnes J. Jasinska). When conducting this research, Agnes J. Jasinska was additionally supported by William Orr Dingwall Foundation Fellowship and by Sarah Winans Newman Scholarship from the Center for the Education of Women, University of Michigan.

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**Conflict of Interest Statement:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Received: 26 March 2012; paper pending published: 13 April 2012; accepted: 20 April 2012; published online: 10 May 2012.

Citation: Jasinska AJ, Ho SS, Taylor SF, Burmeister M, Villafuerte S and Polk TA (2012) Influence of threat and serotonin transporter genotype on interference effects. *Front. Psychology* 3:139. doi: 10.3389/fpsyg.2012.00139

This article was submitted to *Frontiers in Emotion Science*, a specialty of *Frontiers in Psychology*.

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