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Exposure to ambient air pollutions and its association with adverse birth outcomes: a systematic review and meta-analysis of epidemiological studies

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Introduction: Air pollution is a significant global public health concern. However, there is a lack of updated and comprehensive evidence regarding the association between exposure to ambient air pollution and adverse birth outcomes (preterm birth, low birth weight, and stillbirth). Furthermore, the existing evidence is highly inconsistent. Therefore, this study aims to estimate the overall association between ambient air pollution and adverse birth outcomes.

Methods: In this study, initially a total of 79,356 articles were identified. Finally, a total of 49 articles were included. We conducted compressive literature searches using various databases, including PubMed, Scientific Direct, *HINARI*, and Google Scholar. Data extraction was performed using Microsoft Excel, and the data were exported to STATA 17 software for analysis. We used the Joanna Briggs Institute's quality appraisal tool to ensure the quality of the included studies. A random effects model was employed to estimate the pooled prevalence. Publication bias was assessed using funnel plots and Egger's regression test.

Results: In this study, the pooled prevalence of at least one adverse birth outcome was 7.69% (95% CI: 6.70–8.69), with high heterogeneity ($l^2 = 100\%$, *p-value* < 0.001). In this meta-analysis, high pooled prevalence was found in preterm birth (6.36%), followed by low birth weights (5.07%) and stillbirth (0.61%). Exposure to PM_{2.5} ($\leq 10 \mu g/m^3$) throughout the entire pregnancy, PM_{2.5} ($\leq 10 \mu g/m^3$) in the first trimester, PM₁₀ (>10 $\mu g/m^3$) during the entire pregnancy, and O₃ ($\leq 10 \mu g/m^3$) during the entire pregnancy increased the risk of preterm birth by 4% (OR = 1.04, 95% CI: 1.03–1.05), 5% (OR = 1.05, 95% CI: 1.04–1.07), respectively. For low birth weight, exposure to PM_{2.5} ($\leq 10 \mu g/m^3$) and PM_{2.5} (>10 $\mu g/m^3$) throughout the entire pregnancy was associated with an increased risk of 13% (OR = 1.13, 95% CI: 1.05–1.21) and 28% (OR = 1.28, 95% CI: 1.23–1.33), respectively.

Conclusion: This study highlighted a significant association between ambient air pollution and adverse birth outcomes. Therefore, it is crucial to implement a compressive public health intervention.

Systematic review registration: The review protocol was registered with the record ID of CRD42024578630.

KEYWORDS

ambient air pollution, outdoor air pollution, adverse birth outcomes, preterm birth, low birth weights, stillbirth

1 Introduction

Air pollution is a major global public health concern, with growing evidence linking exposure during pregnancy to a range of adverse birth outcomes (1, 2). Exposure to ambient air pollutants such as particulate matter (PM), ozone (O₃), nitrogen dioxides (NO₂), and sulfur dioxide (SO₂) has been linked to a range of adverse birth outcomes, including preterm birth, stillbirth, low birth weight, and congenital anomalies (1, 3, 4). There is no evidence for the biological mechanisms underlying these associations, but they are thought to involve disruption of placental function, inflammation, and oxidative stress (5, 6). The World Health Organization (WHO) estimates that ambient air pollution caused 4.2 million premature deaths globally in 2019, highlighting significant implications for maternal and child health (7).

The effects of ambient air pollution on pregnancy can be attributed to both direct biological mechanisms and indirect socioenvironmental factors. For instance, exposure to high levels of PM during critical periods of fetal development can disrupt placental function and fetal growth (8). Socioeconomic disparities often exacerbate the risks associated with air pollution, as marginalized communities frequently reside in areas with higher pollution levels and limited access to healthcare resources (9). Furthermore, the adverse effects of ambient air pollution on fetal development may have long-term consequences, as preterm birth and low birth weight are risk factors for various health problems later in life, including neurological disorders, cardiovascular disease, and diabetes (10).

A review with meta-analysis has revealed that prenatal exposure to ambient $PM_{2.5}$ is associated with an increased risk of stillbirth (11, 12) and decreased birth weights (13). According to Lamichhane et al. (3) and Sun et al. (13), a $10 \mu g/m^3$ increase in PM_{2.5} exposure during pregnancy was associated with a 15 and 13% higher risk of preterm birth, respectively. Zhu et al. (4) reported that a $10 \,\mu\text{g/m}^3$ increase in PM_{2.5} exposure during pregnancy was associated with a 5% increased risk of low birth weight and a 10% increased risk of preterm birth. Similarly, Stieb et al. (1) concluded that there is consistent evidence linking air pollution exposure to preterm birth and low birth weight. Maternal exposure to PM2.5 (per 10 µg/m3 increased) was associated with a 15% increased risk of stillbirth in the entire pregnancy and a 9% increased risk of stillbirth in the third trimester (12). Exposure to major air pollutants throughout pregnancy may increase the risk of low birth weight (14). Several other studies have also indicated possible associations between ambient air pollution and adverse birth outcomes (15-20).

Although several reviews and meta-analyses have explored the association between specific ambient air pollutants and adverse outcomes, their findings have been inconsistent and lack comprehensiveness. Additionally, the conflicting results from previous primary studies underscore the need for a more thorough and integrated analysis of the available evidence. The purpose of this research is to thoroughly estimate the pooled association between ambient air pollutants and adverse birth outcomes, including preterm birth, low birth weight, and stillbirth, while also identifying predictive factors. Up-to-date and comprehensive evidence is crucial for informed decision-making, the development of effective strategies, and support for policymakers and other stakeholders.

2 Methods

This study followed the guidelines of the Preferred Reporting Items for Systematic Reviews and Meta-Analysis (PRISMA) (21) (Figure 1). The review protocol for this study was registered in the International Prospective Register of Systematic Reviews (PROSPERO) with the record ID of CRD42024578630.

2.1 Eligibility criteria

2.1.1 Inclusion criteria

The eligible for this review must meet the following PECOS (Participants/Populations, Exposures, Comparators, Outcomes and Study designs) criteria.

- Participants or populations: the participants are pregnant women at any stage of pregnancy up to birth.
- Exposures: prenatal exposure to ambient air pollution.
- Comparators: pregnant women with lower exposure levels, with or without adverse birth outcomes, as compared to those exposed to higher exposure with adverse birth outcomes.
- Outcomes: the adverse birth outcomes of interest include preterm birth, low birth weight, and stillbirth (reported prevalence [%] and measure of association in adjusted odds ratio [AOR]).
- **Study design**: all observational studies (cross-sectional, cohort, and case control).

Moreover, articles written in English, both published and unpublished, and studies reported from January 1, 2015 to August 30, 2024 were also the inclusion criteria.

2.1.2 Exclusion criteria

Studies that investigated other pregnancy related outcomes besides the specified adverse birth outcomes (preterm birth, low birth weight, and stillbirth). Descriptive epidemiological studies (e.g., descriptive cross-sectional, case reports, and case series), studies without a full report after three personal email contacts with the primary and corresponding authors, conference abstracts, letters to the editors, qualitative studies, systematic reviews, short communications, and commentaries were not considered.

2.1.3 Operational definitions

Low birth weight: "a birth weight of <2,500 g" (22).

Stillbirth: "A baby who dies after 28 weeks of pregnancy, but before or during birth" (23).

Preterm birth: "babies born alive before 37 weeks of pregnancy" (24).



In this study, we categorized the reported concentrations of pollutants into two main groups: $\leq 10 \,\mu\text{g/m}^3$ and $>10 \,\mu\text{g/m}^3$. This categorization was necessary due to the variability in concentration levels reported across different studies.

2.2 Information sources

Both published and grey literature were sources of information for this study. A systematic literature search was undertaken using the following databases: PubMed, Scientific Direct, Google Scholar, and HINAR. The search was conducted for studies published from January 1, 2015, to August 30, 2024. In addition to the electronic database search, further articles were obtained by searching for grey literature through direct Google searches and by reviewing the references of the eligible studies.

2.3 Search strategies

Comprehensive search terms were used to identify relevant studies. These include MeSH terms and key words such as ambient air pollution, outdoor air pollution, adverse birth outcomes, preterm birth, low birth weights, and stillbirth. These search terms were used within PubMed as a template database to finalize an advanced search strategy utilizing the Boolean operators "AND" and "OR." The search strategy was modified as appropriate for other databases and other sources.

2.4 Study screening and selection

All stages of reviewing articles were conducted independently by the two researchers (BD and AKG), with conflict managed by evidence-based discussion with the involvement of the third researcher (GB). From the search, the titles of all identified citations with abstracts were uploaded into Zotero reference manager, and duplicates were removed. Then it was followed by screening the titles and abstracts according to the eligibility criteria. The potential full text of eligible studies was retrieved. All studies that do not meet the inclusion criteria were excluded with reasons and presented in the PRISMA flow chart (21).

2.5 Quality (risk of bias) assessment of the selected studies

The quality of selected eligible studies was evaluated using the Joana Briggs Institute (JBI) critical appraisal checklist for cohort and case-control studies (25). The quality assessment was conducted independently by two reviewers (BD and ABK). In case of any discrepancies encountered during the quality assessment, they were managed through evidence-based discussions with the involvement of a third researcher (GB). Only studies that scored more than 50% on the quality assessment were considered for inclusion in this review (25, 26), as depicted in Table 1. TABLE 1 Summary of the included studies on the association between ambient air pollution and adverse birth outcomes, 2024.

| References | Country/ region | Type of study | Exposure assessment | Pollutants | Outcome | Statistical Methods | Study period | Sample size | Cases | Prevalence (%) | Quality score (%) |
|-------------------------|--------------------|------------------|--|---|----------|---|-----------------|----------------|---------|-------------------|----------------------|
| Yang et al. (42) | China | Cohort | Daily mean concentrations of air pollutants | PM _{2.5} , PM ₁₀ , SO ₂ , NO _{2,} O3, CO | SB | Logistic regressions | 2011–2013 | 95,354 | 859 | 0.9 | 87.5 |
| Quraishi et al. (43) | United States | Cohort | Regulatory and research monitors | PM _{2.5} | PTB, LBW | Linear and Poisson regression | 2006-2012 | 2,099 | 323 | 15.4 | 87.5 |
| Chen et al. (44) | China | Cohort | Tracking Air Pollution | O ₃ | PTB, LBW | Cox proportional-hazards regression model | 2014-2016 | 56,905 | 4,835 | 8.5 | 87.5 |
| Chen et al. (45) | Australia | Birth records | Daily air quality and meteorological data | PM _{2.5} , SO ₂ , NO ₂ , O ₃ | PTB, LBW | Cox-proportional hazards | 2003-2013 | 173,720 | 24,702 | 14.2 | 87.5 |
| Qian et al. (46) | China | Cohort | Daily measurement | PM _{2.5} , CO, PM ₁₀ , SO ₂ , O ₃ | РТВ | Logistic regressions | 2011-2013 | 95,911 | 4,308 | 4.5 | 75 |
| Zhou et al. (47) | China | Cohort | National urban air quality monitoring | PM _{2.5} , PM ₁₀ , O ₃ , CO, NO ₂ , SO ₂ | LBW | Generalized additive model | 2015-2020 | 572,106 | 24,497 | 4.28 | 75 |
| Chen et al. (48) | China | Birth records | Daily measurement | PM ₁₀ , PM _{2.5} , NO ₂ , SO ₂ | РТВ | Cox proportional hazards regression models | 2015-2017 | 13,111 | 614 | 4.7 | 87.5 |
| Zhou et al. (49) | China | Cohort | Monitoring stations | PM _{2.5} , PM ₁₀ , SO ₂ , CO, NO ₂ , O ₃ | РТВ | Generalized additive model | 2015-2020 | 572,116 | 33,669 | 5.88 | 87.5 |
| Melody et al. (50) | Australia | Cohort | Annual estimation | PM _{2.5,} NO ₂ | PTB, LBW | Linear and log-binomial regression models | 2012-2015 | 285,594 | 23, 187 | 8.1 | 75 |
| Li et al. (51) | China | Birth records | Daily measurement | PM _{2.5} | РТВ | Multilevel logistic models | 2014 | 429,865 | 12,810 | 2.98 | 62.5 |
| Zhao et al. (52) | China | Case-control | Monitoring stations | PM ₁₀ | РТВ | Logistic regression modeling | 2010-2012 | 8,969 | 677 | 7.5 | 75 |
| Padula et al. (53) | United States | Birth records | Daily measurement | CO, NO ₂ , PM ₁₀ , PM _{2.5} | PTB | Logistic regression models | 2000-2006 | 252,205 | 28,788 | 11.4 | 62.5 |
| Tapia et al. (17) | Peru | Birth records | Ground measurements, satellite data, and a chemical transport model. | PM _{2.5} | PTB, LBW | Linear and logistic regression model | 2012-2016 | 123,034 | 10, 971 | 8.9 | 62.5 |
| Liu et al. (54) | China | time-series | Ensemble-based models | PM _{2.5} , PM ₁₀ , SO ₂ , NO ₂ , O ₃ , CO | РТВ | General Additive model extend Poisson regression | 2014-2016 | 37,389 | 5,428 | 14.5 | 87.5 |
| Lavigne et al. (55) | Canada | Cohort | 6 Digit-postal code captured | PM _{2.5} , NO ₂ , O ₃ | PTB, LBW | Multivariable mixed-effect logistic regression | 2005-2012 | 818,400 | 90,884 | 11.1 | 87.5 |
| Huang et al. (56) | China | Birth records | Air monitoring data | PM ₁₀ , NO ₂ | РТВ | Multi-pollutant models | 2006-2010 | 50,874 | 3,203 | 6.3 | 87.5 |
| Liu et al. (57) | China | Case-control | National environmental monitoring | PM _{2.5} , PM ₁₀ , SO ₂ , NO ₂ , CO, O ₃ | PTB, LBW | Logistic regression models | 2014-2015 | 86,139 | 1,784 | 2.1 | 87.5 |

(Continued)

TABLE 1 (Continued)

| References | Country/ region | Type of study | Exposure assessment | Pollutants | Outcome | Statistical Methods | Study period | Sample size | Cases | Prevalence (%) | Quality score (%) |
|-------------------------|--------------------|------------------|--|--|----------|--|-----------------|----------------|---------|-------------------|----------------------|
| Yorifuji et al. (58) | Japan | Cohort | Monitoring stations | SO ₂ , NO ₂ | LBW | multilevel logistic regression | 200-2001 | 44,109 | 2,219 | 5 | 87.5 |
| Stieb et al. (59) | Canada | Birth records | Ground-based monitoring data, estimates from remote- sensing, land use variables and, deterministic gradients relative to road traffic | PM2.5, NO2 | PTB, LBW | Generalized estimating equations | 1999–2008 | 3,104,090 | 242,150 | 7.8 | 75 |
| Green et al. (60) | United States | Cohort | Air resources board | PM _{2.5} , SO ₂ , NO ₂ , CO, O ₃ | SB | Logistic regression models | 1999–2009 | 5,788,117 | 26,355 | 0.5 | 75 |
| Mendola et al. (29) | United States | Cohort | Community multiscale air quality | O ₃ | SB | Poisson regression models | 2002-2008 | 223,375 | 992 | 0.44 | 87.5 |
| Ji et al. (61) | China | Case-control | Land use regression | NO ₂ | PTB | Logistic regression | 2014-2015 | 25,493 | 738 | 2.9 | 87.5 |
| Guo et al. (62) | China | Cohort | National environmental monitoring | PM _{2.5} | РТВ | Cox proportional hazards regression | 2014 | 426,246 | 35,261 | 8.3 | 75 |
| Kingsley et al. (63) | United States | Birth records | Hybrid of land-use regression and satellite remote sensing | PM _{2.5} | РТВ | Linear and logistic regression models | 2001-2012 | 61, 640 | 5,007 | 8.1 | 87.5 |
| Ho et al. (64) | Vietnam | Time-series | Fixed monitoring stations | PM _{2.5} | PTB, LBW | Linear and logistic regression model | 2016-2019 | 163,868 | 18, 219 | 11.1 | 87.5 |
| Wang et al. (65) | China | Cohort | Real-time measurement | PM ₁₀ , PM _{2.5} , SO ₂ , NO ₂ , CO, O ₃ | PTB | Generalized additive model | 2018-2019 | 424 | 17 | 4 | 87.5 |
| Xiao et al. (66) | China | Birth records | Satellite-derived estimates or central-site measurements | PM _{2.5} | PTB, LBW | Linear and logistic Regressions | 2011-2014 | 132,783 | 7,117 | 5.36 | 87.5 |
| Zang et al. (67) | China | Cohort | Daily measurement | PM _{2.5} , PM ₁₀ , SO ₂ , NO ₂ , CO, O ₃ | SB | Logistic regression | 2015-2017 | 59,868 | 587 | 0.98 | 75 |
| Arroyo et al. (68) | Spain | Time-series | Daily measurement | PM _{2.5} , NO ₂ , O ₃ | PTB, LBW | Poisson regression models | 2001-2009 | 298,705 | 64,169 | 21.5 | 62.5 |
| DeFranco et al. (69) | United States | Cohort | Monitoring stations | PM _{2.5} | SB | Generalized estimating equation | 2005-2010 | 349,188 | 1,848 | 0.53 | 87.5 |
| Coker et al. (70) | United States | Cohort | Land use regression | PM _{2.5,} NO ₂ , NO | LBW | Bayesian profile regression | 2000-2006 | 804,726 | 16,694 | 2.07 | 62.5 |
| Yuan et al. (71) | China | Cohort | Satellite-based estimates and ground-level measurements | PM _{2.5} | PTB, LBW | multiple linear models | 2013-2016 | 3,692 | 274 | 7.4 | 87.5 |

TABLE 1 (Continued)

| References | Country/ region | Type of study | Exposure assessment | Pollutants | Outcome | Statistical Methods | Study period | Sample size | Cases | Prevalence (%) | Quality score (%) |
|----------------------------|--------------------|-------------------------------------|---|--|----------|--|-----------------|----------------|----------|-------------------|----------------------|
| Li et al. (72) | China | Time-series | Weekly air quality data | PM _{2.5} , PM ₁₀ , O ₃ , SO ₂ , NO _{2, CO} | PTB | Distributed lag non-linear model | 2016-2019 | 120,446 | 5,408 | 4.5 | 87.5 |
| Rammah et al. (73) | United States | Cohort | Daily measurement | O ₃ | SB | Multipollutant models and measure modification | 2008-2013 | 358,366 | 1,599 | 0.45 | 87.5 |
| Nahian et al. (28) | Bangladesh | Birth records | Air quality index | PM _{2.5} , PM ₁₀ , O ₃ , SO ₂ , NO ₂ | PTB, LBW | Logistic regression model | 2014-2017 | 3,206 | 1,287 | 40.1 | 87.5 |
| Mitku et al. (74) | South Africa | Cohort | Land use regression | PM _{2.5} , SO ₂ | PTB, LBW | Generalized Structure Equation | 2013-2017 | 996 | 206 | 20.7 | 62.5 |
| Li et al. (75) | China | Cohort | Satellite remote sensing, meteorological and land use information | PM _{2.5} , PM ₁₀ | РТВ | Cox proportional hazard regression | 2013-2014 | 1,240,978 | 100,433 | 8.1 | 87.5 |
| Bachwenkizi et al. (76) | Africa | Cross- sectional | Global exposure assessment | PM _{2.5} , O ₃ | PTB, LBW | Multivariable logistic regression | 2005–2015 | 131,594 | 17,591 | 13.4 | 87.5 |
| Chu et al. (77) | China | Cohort | Satellite remote sensing | PM _{2.5} | РТВ | Cox proportional hazard models | 2009-2011 | 5,976 | 443 | 7.4 | 87.5 |
| Han et al. (78) | China | Cohort | Inverse distance weighting | PM ₁₀ , O ₃ | РТВ | Logistic and linear regression models | 2014-2016 | 6,693 | 638 | 9.53 | 87.5 |
| Zhang et al. (79) | China | Birth records | Daily measurement | O3 | PTB, LBW | Cox proportional hazard models | 2016-2019 | 34,122 | 2,829 | 8.3 | 75 |
| Kim et al. (80) | Korea | Birth records | Daily measurement | PM ₁₀ | PTB, LBW | Linear and logistic regression | 2010-2013 | 1,742,183 | 148,086 | 8.5 | 62.5 |
| Liang et al. (81) | China | Cohort | Air monitoring stations | PM _{2.5} | PTB, LBW | Cox proportional hazards regressions | 2014-2017 | 1,455,026 | 121, 646 | 8.4 | 62.5 |
| Chen et al. (82) | China | Cohort | Daily measurement | PM _{2.5} , PM ₁₀ , SO ₂ , CO, O ₃ , NO ₂ | PTB, LBW | Cox proportional hazards regression | 2014-2016 | 10,960 | 291 | 2.7 | 87.5 |
| Johnson et al. (83) | United States | Birth records | Air survey and regulatory monitors | PM _{2.5,} NO ₂ | РТВ | Logistic mixed models | 2008-2010 | 132,654 | 10, 271 | 7.7 | 87.5 |
| Siddika et al. (84) | Finland | Cohort | Regional-to-city-scale dispersion modelling and land-use regression | PM _{2.5} , PM ₁₀ , NO ₂ | РТВ | Dispersion modelling and land- use regression | 1984-1990 | 2,568 | 195 | 7.6 | 75 |
| Sun et al. (85) | China | Cohort | Real-time measurement | PM _{2.5} , PM ₁₀ , O ₃ , SO ₂ , NO ₂ | РТВ | logistic regressions model | 2013-2017 | 6,275 | 372 | 5.9 | 75 |
| Hao et al. (86) | United States | Cohort | Ensemble-based models | NO ₂ , PM _{2.5,} O ₃ | РТВ | Logistic regression model | 2000-2015 | 596,926 | 41,936 | 7.03 | 62.5 |
| Fang et al. (87) | China | Longitudinal population study | Daily measurement | PM _{2.5} | PTB, LBW | Generalized additive distributed lag models | 2014-2016 | 10,738 | 303 | 2.8 | 75 |

NB: PTB, Preterm birth; LBW, Low birth weight; SB, Stillbirth; -, not reported.



2.6 Data extraction and management

Data extraction was conducted by two authors (BD and AKG) using a data extraction tool. Any disagreements between the two data extractors were resolved through consensus or with the involvement of a third author (GB). The following information was extracted from the selected studies: author information, study setting, study country, type of study, pollutants, outcomes, statistical methods, study periods, sample size, cases, prevalence, and quality scores. The extracted data was organized in a table format (Table 1).

2.7 Statistical methods and data analysis

The extracted data from Microsoft Excel was transported to STATA version 17 for analysis. The Index of heterogeneity (I^2 statistics) was used to assess variations among the included studies, where values of 25–50%, 50–75%, and >75% indicated low, moderate, and high heterogeneity, respectively (27). The metaprop command in STATA was used to estimate the pooled prevalence. Subgroup analysis were conducted to explore potential variations in the pooled prevalence based on study countries and the nature of the outcomes. Sensitivity

| Authors (Year) | (95% CI) V | Veigh |
|--------------------------------------|----------------------|--------|
| | | |
| Mitku et al. (2023) | 11.35 (9.38, 13.31) | 2.11 |
| Lie et al. (2018) | 8.09 (8.05, 8.14) | 2.54 |
| Bachwenkizi et al. (2022) | 3.33 (3.23, 3.42) | 2.54 |
| Quraishi et al. (2022) | 8.91 (7.69, 10.13) | 2.38 |
| Chen et al. (2018) | 8.03 (7.90, 8.15) | 2.54 |
| Ji et al. (2019) | 2.89 (2.69, 3.10) | 2.54 |
| Qian et al. (2016) | 4.49 (4.36, 4.62) | 2.54 |
| Chen et al. (2021) | 4.68 (4.32, 5.04) | 2.52 |
| Zhou et al. (2022) | 5.88 (5.82, 5.95) | 2.54 |
| Melody et al. (2020) | 6.46 (6.37, 6.55) | 2.54 |
| Zhao et al. (2015) | 7.55 (7.00, 8.09) | 2.50 |
| Padula et al. (2019) | 11.41 (11.29, 11.54) | 2.54 |
| Tapia et al. (2020) | 7.23 (7.09, 7.38) | 2.54 |
| Chen et al. (2023) | 4.80 (4.63, 4.98) | 2.54 |
| Li et al. (2020) | 2.98 (2.93, 3.03) | 2.54 |
| Chen et al. (2021) | 2.20 (1.92, 2.47) | 2.53 |
| Liu et al. (2018) | 14.52 (14.16, 14.87) | 2.52 |
| Lavigne et al. (2018) | 6.25 (6.20, 6.30) | 2.54 |
| Huang et al. (2015) | 6.30 (6.08, 6.51) | 2.54 |
| Liu et al. (2019) | 0.80 (0.74, 0.86) | 2.54 |
| Stieb et al. (2016) | 6.23 (6.20, 6.26) | 2.54 |
| Xiao et al. (2018) | 0.95 (0.90, 1.00) | 2.54 |
| Guo et al. (2018) | 8.27 (8.19, 8.36) | 2.54 |
| Hao et al. (2023) | 7.03 (6.96, 7.09) | 2.54 |
| Arroyo et al. (2016) | 8.23 (8.13, 8.33) | 2.54 |
| Li et al. (2021) | 4.49 (4.37, 4.61) | 2.54 |
| Johnson et al. (2016) | 7.74 (7.60, 7.89) | 2.54 |
| Kingsley et al. (2019) | 8.12 (7.91, 8.34) | 2.54 |
| Ho et al. (2023) | 9.00 (8.86, 9.14) | 2.54 |
| Wang et al. (2022) | 4.01 (2.14, 5.88) | 2.15 |
| Siddika et al. (2020) | 7.59 (6.57, 8.62) | 2.41 |
| Sun et al. (2019) | 5.93 (5.34, 6.51) | 2.50 |
| Chu et al. (2019) | 7.41 (6.75, 8.08) | 2.48 |
| Han et al. (2018) | 9.53 (8.83, 10.24) | 2.48 |
| | | 2.53 |
| Fang et al. (2020) | 2.15 (1.88, 2.43) | |
| Zhang et al. (2024) | 4.79 (4.57, 5.02) | 2.53 |
| Kim et al. (2019) | 4.70 (4.67, 4.73) | 2.54 |
| Yuan et al. (2020) | 4.55 (3.88, 5.22) | 2.48 |
| Liang et al. (2019) | 4.40 (4.37, 4.43) | 2.54 |
| Nahian et al. (2023) | 12.32 (11.18, 13.48) | 2.38 |
| Overall, DL (l' = 100.0%, p = 0.000) | 6.36 (5.66, 7.08) 1 | 100.00 |
| 0 | 25 | |
| v | 20 | |
| | | |

analysis was performed to assess the effect of each individual study on the estimated pooled results. To evaluate publication bias, a funnel plot test and Egger's regression test were used. A meta-regression was employed to identify potential sources of heterogeneity. Finally, the findings of this study are presented using tables, figures, forest plots, and descriptive texts.

3 Results

3.1 Overview of search process

We identified a total of 79,356 studies using a database and through direct Google and citation searching. After duplicate records were removed, 51,592 records were screened for this review. According to the records, only 32,947 studies were sought for retrieval. After being identified for retrieval, 15,063 studies were evaluated for eligibility. Following eligibility, a total of 15,021 studies were excluded due to differences in outcome interest and population differences. Ultimately, a total of 42 studies were included in this review from database sources. In addition to the database sources, seven studies were included in this review from direct Google and citation searching. Finally, a total of 49 articles were included in this study, as presented in the PRISMA flowchart (Figure 1).

3.2 Characteristics of the eligible studies

The majority of the included studies were conducted using birth cohort studies. This meta-analysis included a total of 21,019,317 study participants. The majority of the studies were conducted in China (n=26) and the United States (n=10). This meta-analysis examined

| Authors (Year) | Effect (95% CI) Wei |
|--|--|
| Xiao et al. (2018) | 4.41 (4.35, 4.48) 4 |
| Quraishi et al. (2022) | 6.48 (5.90, 7.06) 4 |
| Chen et al. (2018) | 6.19 (6.13, 6.26) 4 |
| Zhou et al. (2023) | 4.28 (4.25, 4.31) 4 |
| Tapia et al. (2020) | 1.69 (1.65, 1.73) 4 |
| Chen et al. (2023) | 3.69 (3.60, 3.79) 4 |
| Lavigne et al. (2016) | 4.88 (4.83, 4.88) 4 |
| Liu et al. (2019) | 1.27 (1.24, 1.31) 4 |
| Yorifuji et al. (2015) | 5.03 (4.91, 5.15) 4 |
| Stieb et al. (2016) | 1.57 (1.58, 1.58) 4 |
| Arroyo et al. (2016) | 13.25 (13.19, 13.31) |
| Coker et al. (2016) | 2.07 (2.08, 2.09) 4 |
| Ho et al. (2023) | 2.30 (2.26, 2.34) 4 |
| Zhang et al. (2024) | 3.50 (3.38, 3.61) 4 |
| Kim et al. (2019) | 3.80 (3.78, 3.82) 4 |
| Yuan et al. (2020) | 2.87 (2.54, 3.20) 4 |
| Liang et al. (2019) | 3.96 (3.94, 3.98) 4 |
| Nahian et al. (2023) | ÷ 27.82 (27.20, 28.44) 4 |
| Mitku et al. (2023) | 9.34 (8.42, 10.26) 4 |
| Bachwenkizi et al. (2022) | 10.04 (9.96, 10.12) 4 |
| Melody et al. (2020) | 1.66 (1.63, 1.68) 4 |
| Overall, DL (l ² = 100.0%, p = 0.000) | 5.68 (4.84, 6.53) 100 |
| | 25 |
| v | 20 |



the association between ambient air pollutants ($PM_{2.5}$, PM_{10} , SO_2 , NO_2 , O_3 , CO) and adverse birth outcomes (preterm birth, low birth weight, and stillbirths). In this study, the main exposure assessment methods were the daily mean concentration of air pollutants, monitoring stations, land use regression model, and real-time measurement. Among the included studies, Bangladesh had the highest at least one birth outcome (40%) (28), while the United States had the lowest rate (0.44%) (29). The quality score of the included studies was between the ranges of 62.5 and 87.5% (Table 1).

3.3 Meta-analysis

The findings from the random effects model indicated that the pooled prevalence of at least one adverse birth outcome was 7.69% (95% CI: 6.70–8.69), with high heterogeneity (I^2 =100%, *p*-value<0.001) (Figure 2). In this meta-analysis, high pooled prevalence was found in preterm birth (6.36%) (Figure 3), followed by low birth weights (5.07%) (Figure 4) and stillbirth (0.61%) (Figure 5). Subgroup analysis based on study country: the highest pooled prevalence of at least one adverse birth

outcome was observed in Bangladesh at 40.14% (95% CI: 38.45–41.84) and in Spain at 21.5% (95% CI: 21.34–21.63). In contrast, the lowest pooled prevalence of at least one adverse birth outcome was observed in Japan at 5.03% (95% CI: 4.83–5.24) and the United States at 5.12% (95% CI: 4.22–6.02). In addition, subgroup analysis based on the nature of outcomes found that preterm birth and low birth weight were at 11.1% (95% CI: 9.85–12.35), preterm birth at 6.96% (95% CI: 5.86–8.66), low birth weight at 3.79% (95% CI: 2.0–5.59), and stillbirth at 0.615% (95% CI: 0.53–0.699) (Table 2).

In this study, a meta-regression analysis was conducted using the study country and nature of outcomes as factors to identify the source

TABLE 2 Subgroup analysis of the association between ambient air pollution and at least one adverse birth outcome, 2024.

| Variable | Number | | Heter | ogeneity |
|----------------|------------|---------------------|----------------|----------|
| variable | of studies | OR (95%CI) | l ² | p-value |
| Region/country | | | | |
| Africa | 2 | 16.9 (9.5–24.1) | 96.9 | < 0.001 |
| China | 26 | 5.72 (4.6-6.9) | 100 | <0.001 |
| United States | 10 | 5.12 (4.22-6.02) | 100 | <0.001 |
| Australia | 2 | 11.2 (5.2–17.15) | 100 | <0.001 |
| Peru | 1 | 8.92 (8.8–9.08) | - | - |
| Canada | 2 | 9.45 (6.22–12.69) | 100 | <0.001 |
| Japan | 1 | 5.03 (4.83-5.24) | - | - |
| Spain | 1 | 21.5 (21.24–21.63) | - | - |
| Vietnam | 1 | 11.12 (10.97–11.27) | - | - |
| Finland | 1 | 7.59 (6.57–8.62) | - | - |
| Korea | 1 | 8.5 (8.46-8.54) | - | - |
| Bangladesh | 1 | 40.14 (38.45-41.84) | - | - |
| Outcome | | | | |
| PTB and LBW | 20 | 11.1 (9.85–12.35) | 100 | <0.001 |
| РТВ | 20 | 6.96 (5.86-8.66) | 99.9 | <0.001 |
| LBW | 3 | 3.79 (2.0-5.59) | 100 | <0.001 |
| SB | 6 | 0.615 (0.53–0.699) | 98.8 | <0.001 |

NB: PTB, Preterm birth; LBW, Low birth weight; SB, Stillbirth.

of heterogeneity. The finding revealed that the study country was not a statistically significant source of heterogeneity (p = 0.196), but the outcome nature was found to be a statistically significant source of heterogeneity (p < 0.001).

A sensitivity analysis was also performed to evaluate a single study effect on the overall results. The analysis showed for the overall prevalence of at least one adverse birth outcome a slightly broader confidence interval of 7.69% (95% CI: 6.05–8.98) compared to the original pooled prevalence of 7.69% (95% CI: 6.70–8.69), but it does not suggest strong evidence for single study effects. Similarly, sensitivity analysis was also conducted for preterm birth, low birth weight, and stillbirth to examine the effect of a single study on the overall prevalence. The findings suggested that there is no evidence for a single study effect on the overall pooled prevalence (Supplementary material 1).

The funnel plot for the overall analysis showed an asymmetrical distribution as visualized of the included articles, revealing the potential of publication biases (Figure 6). However, the Egger-regression test confirmed that there was no statistically significant presence of publication bias (p-value = 0.1001).

Similarly, funnel plots and Egger-regression tests were conducted to assess publication biases for the specific birth outcomes of preterm birth and low birth weights. For preterm birth, the funnel plots showed an asymmetrical distribution as visualized of the included articles (Figure 7), but the Egger-regression test confirmed that there was no statistically significant (p-value=0.2087) for the presence of publication bias.

In contrast, for low birth weights, the funnel plot also showed an asymmetrical distribution as visualized of the included articles (Figure 8), and the Egger-regression test confirmed the presence of publication bias with statistical significance (p-value=0.0017). To address the publication bias identified for the low birth weight outcome, Duval and Tweedie's "trim and fill" method was conducted (Supplementary material 2).

3.4 Factors associated with adverse birth outcomes

In this meta-analysis, exposure to ambient air pollution, such as $PM_{2.5}$, PM_{10} , and O_3 , was statistically significant for adverse birth outcomes (preterm birth and low birth weight). For preterm birth,



exposure to $PM_{2.5}$ ($\leq 10 \mu g/m^3$) during entire pregnancy, $PM_{2.5}$ ($\leq 10 \mu g/m^3$) in first trimester, PM_{10} (>10 $\mu g/m^3$) during entire pregnancy, and O_3 ($\leq 10 \mu g/m^3$) during entire pregnancy increased the risk by 4% (OR = 1.04, 95% CI: 1.03–1.05), 5% (OR = 1.05, 95% CI: 1.01–1.09), 49% (OR = 1.49, 95% CI: 1.41–1.56), and 5% (OR = 1.05, 95% CI: 1.04–1.07), respectively (Figure 9).

For low birth weight, exposure to $PM_{2.5}$ ($\leq 10 \mu g/m^3$) and $PM_{2.5}$ (>10 $\mu g/m^3$) during entire pregnancy was found to increase the risk by 13% (OR=1.13, 95% CI 1.05–1.21) and 28% (OR=1.28, 95% CI 1.23–1.33), respectively (Figure 10).

4 Discussion

This systematic review and meta-analysis aimed to estimate the pooled association between exposure to ambient air pollution and

adverse birth outcomes (preterm birth, low birth weight, and stillbirth), as well as to identify predictive factors. The pooled prevalence of at least one adverse birth outcome was found to be 7.69% (95% CI: 6.70–8.69), with notable extreme heterogeneity among the included studies ($I^2 = 100$, p < 0.001). Specifically, exposure to ambient air pollution was associated with a 6.36% (95% CI: 5.66–7.06) increased risk of preterm birth, a 5.07% (95% CI: 4.32–5.81) increase in low birth weight, and a 0.65% (95% CI: 0.53–0.7) increase in the risk of stillbirth. These findings suggest that exposure to ambient air pollution during pregnancy negatively affects various birth outcomes (30, 31).

Daba et al. (32) support the present findings, reporting a 15.5% (95% CI: 12.6–18.5) prevalence of adverse pregnancy outcomes linked to indoor air pollution exposure. The WHO reported low birth weight at 15.5% globally, 16.5% in developing countries, and 7% in developed countries in 2015.





| Pollutants and Author | Year | | Effect (95% CI) | 9 Weigh |
|---------------------------------|--------------------------------------|--------------|--------------------|----------------|
| PM2.5 (≤10 μg/π | 13)[Entire pregnancy] | | | |
| Sune et al. | 2019 | | 1.13 (1.03 | , 1.25) 0.8 |
| Xiao et al. | 2018 | • | 1.27 (1.20 | , 1.36) 1.6 |
| Qian et al. | 2016 | i i | | , 1.05) 47.7 |
| Lavigne et al. | 2016 | | 4.00 (2.40 | |
| Subgroup, IV (I ² = | 93.9%, p = 0.000) | il | 1.04 (1.03 | , 1.05) 50.2 |
| PM2.5 (≤10 µg/n | n3)[1st trimester] | | | |
| Liu et al. | 2019 | • | 1.04 (1.01 | , 1.09) 6.71 |
| Ho et al. | 2023 | ۲ | 1.12 (1.02 | , 1.23) 0.97 |
| Subgroup, IV (I ² = | 48.6%, p = 0.163) | _ • | 1.05 (1.01 | ,1.09) 7.68 |
| PM10(≤10 µg/m3 | 3)[3rd trimester] | | | |
| Liu et al. | 2019 | i i i | 1.02 (1.02 | , 1.09) 8.76 |
| Han et al. | 2018 | · · · · | 1.40 (1.08 | , 1.81) 0.08 |
| Subgroup, IV (I ² = | : 75.2%, p = 0.044) | | 1.03 (0.99 | , 1.06) 8.84 |
| PM10 (>10 µg/m | [Entire pregnancy] | | | |
| Qian et al. | 2016 | + | 1.02 (1.00 | , 1.40) 0.27 |
| Zhao et al. | 2015 | - | 1.48 (1.22 | , 1.81) 0.12 |
| Kim et al. | 2019 | | 1.57 (1.49 | , 1.66) 1.52 |
| Subgroup, IV (I ² = | 91.9%, p = 0.000) | 1 | 1.49 (1.41 | , 1.56) 1.91 |
| O3 (≤10 µg/m3) | [Entire pregnancy] | | | |
| Sune et al. | 2019 | | 1.12 (1.05 | , 1.19) 2.19 |
| Hao et al. | 2023 | • | 1.05 (1.02 | , 1.08) 11.9 |
| Qian et al. | 2016 | l 🛉 | 1.05 (1.02 | , 1.07) 17.17 |
| Subgroup, IV (I ² = | 44.0%, p = 0.168) | | 1.05 (1.04 | , 1.07) 31.29 |
| | tween groups: p = 0.000 | | | |
| Overall, IV (I ² = 9 | 94.0%, p = 0.000) | | | , 1.06) 100.00 |
| | 5 | 0 | 5 | |
| | -5 | 0 | 5 | |

The current findings are also supported by numerous studies indicating that exposure to ambient air pollution increases the risk of preterm birth, low birth weight, and stillbirth (30, 31, 33). Variations in findings may be attributed to different factors such as maternal educational level, age differences, socioeconomic conditions, type of pollutants, and the duration and level of exposure during the perinatal period (34, 35). Dzekem et al. (36) highlighted the need to address disparities, such as socioeconomic issues, when examining the relationship between air pollution exposure and pregnancy outcomes. Therefore, it is crucial to ensure that the interaction between pregnant women and their environment is safe and wellmaintained to prevent adverse effects. In this study, the heterogeneity among the included studies was significantly high, a finding that is supported by previous research (14, 31–33, 37, 38). This variability may be attributed to various factors, including differences in study settings, designs, and exposure assessment methods. To identify the potential sources of this heterogeneity, we conducted a subgroup analysis based on the study country and the type of adverse birth outcomes. Subsequently, the meta-regression analysis confirmed that the primary source of heterogeneity was related to the nature of the adverse birth outcomes. This may be linked to the levels and types of pollutants, as well as the conditions of pregnant women.

In this meta-analysis, exposure to $PM_{2.5}$ ($\leq 10 \,\mu g/m^3$) throughout the entire pregnancy and during the first trimester was associated

FI P

| Pollutants and Author | Year | | Effect (95%CI) | % Weight |
|------------------------------------|----------------------|-----------|-----------------------|-------------|
| and Addition | 1 041 | | (337601) | weight |
| PM2.5 (≤10 µg/m3) [| Entire pregnancy] | 1 | | |
| Tapia et al. | 2020 | | 1.11 (1.03, 1.20) | 26.28 |
| Xiao et al. | 2018 | _ | — 1.22 (1.06, 1.41) | 6.20 |
| Subgroup, IV $(I^2 = 18)$ | 8.6% p = 0.268) | | 1.13 (1.05, 1.21) | 32.48 |
| | | | | |
| PM2.5 (>10 µg/m3) [| Entire pregnancy] | 1 | | |
| Mitku et al. | 2023 | - | ··· 1.30 (1.02, 1.42) | 4.75 |
| Bachwenkizi et al. | 2022 | 1 1 | + 1.28 (1.23, 1.34) | 62.77 |
| Subgroup, IV $(I^2 = 0.0)$ | 0% p=0.850) | i i | 1.28 (1.23, 1.33) | 67.52 |
| | | | | |
| Heterogeneity betwe | en groups: p = 0.002 | | | |
| Overail, IV (I ² = 73.5 | % p = 0.010) | | 1.23 (1.19, 1.28) | 100.00 |
| | | · · · · · | | |
| | | 0 1 | | |

with an increased risk of preterm birth, with an OR of 1.04 (95% CI: 1.03–1.05) and 1.05 (95% CI: 1.01–1.04), respectively. Sapkota et al. (37) reported that exposure to $PM_{2.5}$ at levels of 10 µg/m³ during pregnancy increased the risk of preterm birth with an OR of 1.15 (95% CI: 1.14–1.16). Liu et al. (39) found a similar positive association, reporting an OR of 1.15 (95% CI: 1.07–1.23) for $PM_{2.5}$ exposure during pregnancy. Additionally, Lamichhane et al. (3) estimated an OR of 1.14 (95% CI=1.06–1.22) for preterm birth per 10µg/m³ increase in $PM_{2.5}$ exposure during the entire pregnancy. Therefore, these findings highlighted the importance of protecting pregnant mothers from $PM_{2.5}$ exposure to reduce the risk of adverse outcomes of preterm birth (39–41).

In this study, we found a 49% increase in the risk of preterm birth for each >10 μ g/m³ increase in PM₁₀ exposure during the entire pregnancy. Stieb et al. (1) reported a high risk of preterm birth associated with a 20 μ g/m³ increase in PM₁₀ over the same period. Similarly, Lamichhane et al. (3) noted that exposure to PM₁₀ increased the risk of preterm birth by 23% for each 10 μ g/m³ increment. The slight difference in findings may be attributed to variation in exposure levels and duration. Therefore, addressing ambient air pollution is essential for reducing the occurrence of preterm birth.

In the present study, exposure of O_3 at levels of $\leq 10 \,\mu\text{g/m}^3$ during pregnancy was positively associated with a 5% increased risk of preterm birth, with (OR = 1.05, 95% CI: 1.04–1.07). This finding is consistent with previous research (31, 33), which suggests that

exposure to ozone throughout pregnancy may significantly elevate the risk of preterm birth. These findings underscoring the need for protective measures to minimize pregnant women's exposure to ozone.

In this meta-analysis, maternal exposure to $PM_{2.5}$ during the entire pregnancy at levels of >10 µg/m³ was associated with a 28% increase in the risk of low birth weight (OR = 1.28, 95% CI: 1.23–1.33). Additionally, exposure to $PM_{2.5}$ (\leq 10 µg/m³) also showed a 13% increase in risk (OR = 1.13, 95% CI: 1.05–1.21). These findings suggest that as $PM_{2.5}$ exposure increases, the risk of low birth weight also increases. Supporting this, Zhu et al. (4) reported a 5% increase in low birth weight per 10 µg/m³ increment in $PM_{2.5}$ exposure during the entire pregnancy (OR = 1.05, 95% CI: 1.02–1.07). The difference in findings may be attributed to variations in exposure assessment methods. Thus, exposure to $PM_{2.5}$ throughout pregnancy could significantly impact the final birth weight.

4.1 Limitation of the study

This study focusses exclusively on studies conducted in the English language. Additionally, it does not explore the underlying mechanisms that link ambient air pollution to adverse birth outcomes. Furthermore, this study also focused on selected adverse birth outcomes, but other birth outcomes like congenital anomalies or birth defects and others might be important.

5 Conclusion

This meta-analysis highlighted a significant association between ambient air pollution and adverse birth outcomes. In this study, PM_{2.5}, PM₁₀, and O₃ were found to be positively associated with these adverse birth outcomes (preterm birth and low birth weight). Given these findings, it is essential for healthcare professionals, the Ministries of Health, non-governmental organizations, and other relevant stakeholders to implement compressive public health interventions aimed at reducing the incidence of adverse birth outcomes related to ambient air pollution. Such interventions could include policies to improve air quality, strict regulations, public awareness campaigns, and targeted support for vulnerable populations of pregnant women. Furthermore, to gain a deeper understanding of the mechanisms underlying the association between ambient air pollution and adverse birth outcomes, future research is highly recommended. Investigating these mechanisms will provide valuable insights that can help inform more effective strategies for mitigating the risks associated with air pollution during pregnancy. In addition, future researchers are encouraged to investigate the impact of ambient air pollution on congenital anomalies and other significant adverse birth outcomes.

Data availability statement

The original contributions presented in the study are included in the article/Supplementary material, further inquiries can be directed to the corresponding author.

Author contributions

BD: Methodology, Software, Validation, Writing – original draft, Writing – review & editing. GB: Conceptualization, Data curation, Visualization, Writing – review & editing. AG: Data

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Supplementary material

The Supplementary material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fpubh.2024.1488028/ full#supplementary-material

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